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Article Effect of data scaling on color device model fitting

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Effect of data scaling on color device model fitting

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Abstract: Output devices in print production can be characterized by different characterization methods. One commonly used method of color device characterization is least squares fitting. In essence, the least squares fitting is used to determine the coefficients of a predetermined polynomial, such that the sum of squared differences between the values predicted by the model and the empirical data is minimal. The choice of the polynomial order and the cross product terms which best describe the behavior of a certain device is not obvious. This paper is a part of a larger study which investigates the criteria in the measurement data which can be used for optimal model selection. The part of the study covered in this paper addresses the data over fitting problem. It is investigated by comparing the performance of models of different polynomial orders on two different domains.

Keywords: regression model, characterization data, printer characterization, color reproduction

1 Introduction

The color management systems are used to ensure accurate color reproduction in print production processes. Current industrial standard are the ICC (International Color Consortium) based systems. In the second half of the 1980s, with the growing use of computers in print production, different software and equipment vendors developed proprietary color management solutions which were not interoperable. The ICC is a consortium of those vendors founded in the year 1993

with the aim of developing a universal color management solution. The ICC soon devised a device profile format which has the aim of describing the color reproduction behavior of a given device used in the color reproduction process. The ICC profile format, defined by the ICC Profile Specification, consists of various data structures which provide a mechanism for color transforms. The general color transformation models and the data structures which map to those models are specified. However, the procedures and techniques used to populate the profile data structures are not specified (ICC, 2004). Various techniques, such as global and local regression methods, distance weighted techniques and neural networks can be used for this purpose, and the results depend on their choice (Sharma, 2003). The global regression methods are very commonly used and are investigated in this paper. Analytical expression used as regression models are polynomials. Since color spaces are three dimensional spaces, the polynomials used to model the transforms from one to another color space are multivariate. The choice of the polynomial order and the cross products (individual polynomial terms) affects the model accuracy. Low order polynomials may not be sufficient when modeling devices with non-linear relationships between inputs and outputs. High order polynomials, on the other hand, may exaggerate the minor fluctuations in the data and introduce local extremes i.e. over fit the data (Green, 2002).

2 Theoretical

When fitting a model to a dataset, the model form is predetermined. The least squares solution when determining the model parameters may give satisfactory prediction accuracy with one model, while some other model may not be as accurate (Harris et al., 2001). The key to successful modeling, when using polynomial models, is to determine the optimal model order and, in multivariate models its cross products. High order models can over fit the data, causing the model to be very well adapted to the dataset used to determine the model parameters, but to be inaccurate in other regions of its domain, i.e. to have poor generalization properties. Low order models may not even adapt to the dataset at satisfactory accuracy. Another interesting case is that high order models may not adapt well even to the dataset.

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3 Experimental

This study was carried out on a desktop thermal inkjet printer with commercial inks supplied with the printer. The printer driver takes RGB values as inputs. To obtain device data, a standard TC9.18 RGB test chart was printed on plain 80g/m² paper, (spectral reflectances) measured and the device responses using a spectrophotometer with $45^{\circ}/0^{\circ}$ measurement geometry and type A illuminant. Although the choice of training set samples affects the model accuracy and depends on the substrate used, and sample selection methods were developed (Chou et al., 2010), a standard chart suits the purpose of this study. The obtained data consisted of values of RGB inputs and their corresponding spectral reflectances under the D50 illuminant. The spectral reflectances were converted to L*a*b* values using the 2° observer color matching functions, which were used as the model domain. Two domains were tested, one with L*a*b* values expressed in native L*a*b* units, and one where ranges of L* [0,100] and a*,b* [-128, 127] values were mapped to the range [0,1]. In the case of the first L*a*b* domain, the corresponding RGB digital counts were left unchanged, i.e. the [0,255] range was used. In the case of the scaled L*a*b* domain, the RGB values were also scaled on the [0,1] range.

The abovementioned input and output values were fitted with four different models, and the significance of terms was tested by backward elimination procedure (partial F-test). Special software was written for this purpose, allowing blocks of terms of a given order to be tested. The evaluation of model accuracies was carried out by comparing the values predicted by the models to the measured values for each of the 918 test chart patches. The obtained ΔE errors are Euclidean distances in the L*a*b* space, and the minimum, maximum, mean and median of errors were calculated. The confidence intervals at 95% confidence level were also calculated as they are useful when estimating whether the mean values of the observed cases can be considered to differ significantly. Many studies on the acceptability of ΔE results yielded different results due to imperfect uniformity of L*a*b* space (Thomas et al., 2010). The evaluation was carried out with two data sets, one being the one to which the models were fitted, and one containing different RGB, L*a*b* value pairs from those to which the models were fitted. The evaluation of models with the independent set is important because complex models can give excellent least-square fits, but on the other hand, exaggerate the

small fluctuations in the data and produce local extremes. Instead of modeling the underlying relationship in the data, the model can turn out to be very imprecise for values different from those it was fitted to. In this case, the results from both data sets were similar and therefore only one dataset evaluation results are presented in this paper.

3.1 Fourth order model with 23 terms

 $\begin{bmatrix} 1 \ R \ G \ B \ RG \ GB \ RB \ R^2 \ G^2 \ B^2 \ RGB \ R^2G \ G^2B \ B^2R \ R^2B \ G^2R \ B^2G \ R^3 \ G^3 \ B^3 \ R^2GB \ RG^2B \ RG^2B \ RGB^2 \end{bmatrix}$

| | Ν | $\overline{\Delta E}$ | Min | Median | Max | C.I. 95% |
|--------------|-----|-----------------------|------|--------|-------|----------|
| RGB->Lab, D1 | 918 | 4,51 | 0,21 | 3,98 | 14,65 | 0,37 |
| RGB->Lab, D2 | 918 | 5,05 | 0,27 | 4,80 | 14,93 | 0,30 |

Table 1. "Statistics of the 23-4 model performance".

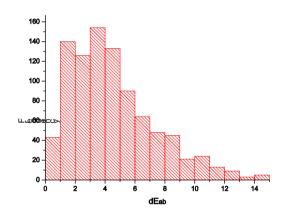
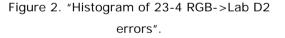


Figure 1. "Histogram of 23-4 RGB->Lab D1 errors".



3.2 Fifth order model with 26 terms

 $\begin{bmatrix} 1 \ R \ G \ B \ RG \ GB \ RB \ R^2 \ G^2 \ B^2 \ RGB \ R^2 \ G^2 \ B \ B^2 \ R \ R^2 \ B \ G^2 \ R \ B^2 \ G^3 \ B^3 \ R^2 \ GB \ RG^2 \ B \ RG^2 \ B^3 \ RG^2 \ B^3 \ RG^3 \ B \ RG^3 \ RG^3 \ B \ RG^3 \ B \ RG^3 \ B \ RG^3 \ B \ RG^3 \ RG^3 \ B \ RG^3 \ B \ RG^3 \ B \ RG^3 \ B \ RG^3 \ RG^3 \ B \ RG^3 \ B \ RG^3 \ B \ RG^3 \ RG^3$

| | N | $\overline{\Delta E}$ | Min | Median | Max | C.I. 95% |
|--------------|-----|-----------------------|------|--------|-------|----------|
| RGB->Lab, D1 | 918 | 4,91 | 0,72 | 4,59 | 15,64 | 0,30 |
| RGB->Lab, D2 | 918 | 4,91 | 0,72 | 4,59 | 15,64 | 0,30 |

Table 2. "Statistics of the 26-5 model performance".

180 -

160

140

120

100

80 -

40

20

0

<u>0</u> 2

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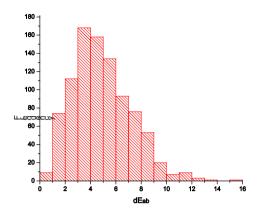
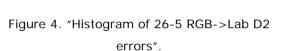


Figure 3. "Histogram of 26-5 RGB->Lab D1 errors".



dEab

14

12

3.3 Seventh order model with 26 terms

 $\begin{bmatrix} 1 \ R \ G \ B \ RG \ GB \ RB \ R^{2} \ G^{2} \ B^{2} \ RGB \ R^{2}G \ G^{2}B \ B^{2}R \ R^{2}B \ G^{2}R \ B^{2}G \ R^{3} \ G^{3} \ B^{3} \ R^{2}GB \ RG^{2}B \ RG^{2}B \ RG^{2}B \ RG^{2}B^{2} \ R^{3}G^{3}B^{2} \ R^{2}G^{3}B^{2} \ R^{2}G^{2}B^{3} \end{bmatrix}$

| | N | $\overline{\Delta E}$ | Min | Median | Max | C.I. 95% |
|--------------|-----|-----------------------|------|--------|-------|----------|
| RGB->Lab, D1 | 918 | 19,63 | 0,97 | 16,77 | 55,56 | 1,49 |
| RGB->Lab, D2 | 918 | 4,94 | 0,31 | 4,68 | 15,88 | 0,30 |

Table 3. "Statistics of the 26-7 model performance".

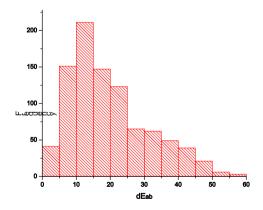


Figure 5. "Histogram of 26-7 RGB->Lab D1 errors".

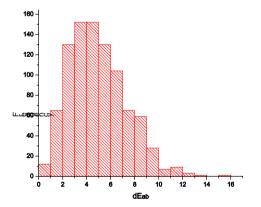


Figure 6. "Histogram of 26-7 RGB->Lab D2 errors".

3.4 Sixth order model with 29 terms

 $\begin{bmatrix} 1 \ R \ G \ B \ RG \ GB \ RB \ R^2 \ G^2 \ B^2 \ RGB \ R^2 \ G^2 \ B \ B^2 \ R \ R^2 \ B \ G^2 \ R \ B^2 \ G^2 \ R \ B^2 \ G^3 \ B^3 \ R^2 \ GB \ RG^2 \ B \ RG^2 \ B^3 \ R^3 \ G^3 \ B^3 \ R^2 \ GB \ RG^2 \ RG^2 \ B^3 \ RG^2 \ B^3 \ R^3 \ G^3 \ B^3 \ R^2 \ GB \ RG^2 \ RG^2 \ B^3 \ RG^2 \ B^3 \ RG^2 \ B^3 \ RG^2 \ RG$

| | N | $\overline{\Delta E}$ | Min | Median | Max | C.I. 95% |
|--------------|-----|-----------------------|------|--------|-------|----------|
| RGB->Lab, D1 | 918 | 7,91 | 0,67 | 6,56 | 30,19 | 0,68 |
| RGB->Lab, D2 | 918 | 4,85 | 0,36 | 4,54 | 15,97 | 0,30 |

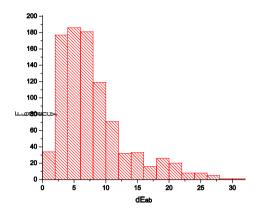


Table 4. "Statistics of the 29-6 model performance".

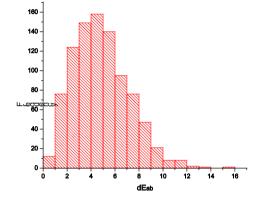


Figure 7. "Histogram of 29-6 RGB->Lab D1 errors".

Figure 8. "Histogram of 29-6 RGB->Lab D2 errors".

3.5 Seventh order model with 35 terms

 $\begin{bmatrix} 1 \ R \ G \ B \ RG \ GB \ RB \ R^{2} \ G^{2} \ B^{2} \ RGB \ R^{2}G \ G^{2}B \ B^{2}R \ R^{2}B \ G^{2}R \ B^{2}G \ R^{3} \ G^{3} \ B^{3} \ R^{2}GB \ RG^{2}B \ RG^{2}B \ RG^{2}B^{3} \ R^{3}G^{2}B^{2} \ R^{2}G^{3}B^{2} \ R^{2}G^{3}B^{2} \ R^{2}G^{3}B^{2} \ R^{2}G^{3}B \ R^{2}GB^{3} \ R^{3}GB^{2} \ RG^{3}B^{2} \ RG^{3} \ RG^{3}$

| | N | $\overline{\Delta E}$ | Min | Median | Max | C.I. 95% |
|--------------|-----|-----------------------|------|--------|-------|----------|
| RGB->Lab, D2 | 918 | 4,41 | 0,37 | 4,18 | 13,77 | 0,30 |

| Table 5 | "Statistics | of the | 35-7 | model | performance". |
|---------|-------------|--------|------|-------|---------------|
|---------|-------------|--------|------|-------|---------------|

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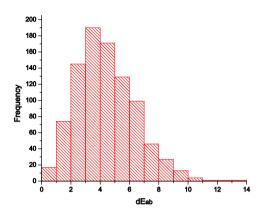


Figure 9. "Histogram of 35-7 RGB->Lab D2 errors".

As can be seen from Table 1 and Figure 1 and Figure 2, the 23-4 model performed better on the native L*a*b* domain. Adding additional cross product terms and forming more complex models resulted in poorer performance of those models on the native L*a*b* domain. The 26-5 model's median and mean of errors were increased by ΔE 0,6 and ΔE 0,4 respectively. The maximum error was increased by ΔE 1. The 26-7 model had the worst performance. With the median and the mean of errors of around ΔE 17 and ΔE 20, and the maximum error of around ΔE 56, it is unsuitable for use as the errors are extreme. The 29-6 model also performed worse than simpler models. The median error of ΔE 7, the mean error of ΔE 8, and the maximum error of ΔE 30 are considered to be quite large. On the scaled L*a*b* domain, the performance of the initial 23-6 model was worse than on the native L*a*b* domain. However, more complex models formed by adding additional cross product terms showed better performance, as the median and the mean of errors were reduced. The maximum error was increased. The 26-5 model's mean and median of errors were reduced by the amount that is not considered to be significant as it falls within the limits of the confidence interval. The maximum error was increased by approximately ΔE 0,7. The 26 -7 model performed similarly. Its mean and median of errors were reduced by insignificant amount, and the maximum error was increased by almos AE 1. The 29 -6 model's mean and the median of errors were reduced by arothed0,2. The maximum error was increased by around ΔE 1.

As all of the tested additional cross products improved the model performance, they were all gathered to form a new 35-7 model with all of them included. The 35 term model performed better than the basic 23-4 model and better than all of the more complex models with additional cross products. Its mean and median of

errors were reduced by approximat $\Delta E = 0,6$, and the maximum error was reduced by around $\Delta E = 1,2$. Although it was not tested in this study, it is likely that even more complex models could further improve the accuracy on the scaled L*a*b* domain.

| | Native L* | a*b* | | | Scaled L*a*b* | | | |
|-------|-----------|-------|--------|--------|---------------|--------|----------|----------|
| | R,G,B | L* | a* | b* | R,G,B | L* | а* | b* |
| | 0 | 18,38 | -47,63 | -44,68 | 0 | 0,1838 | 0,315176 | 0,326745 |
| Range | 255 | 92,11 | 58,45 | 69,98 | 1 | 0,9211 | 0,731176 | 0,776392 |
| Ratio | | 0,29 | 0,416 | 0,45 | | 0,7373 | 0,416 | 0,45 |

Table 6. "Domain - co domain ratios".

The polynomial models higher than order fourth over fitted the data on the native L*a*b* domain. On the scaled L*a*b* domain the seventh order model performed well. The data displayed in Table 2 presents the ranges of the data on the two domains, and the ratios of the data ranges on the two domains with respect to their co domains. The purpose of calculating the ratios was to describe the data spread and explain why the over fitting occurs on one domain, and does not occur on the other one. However, the ratios are the same for a* and b* values, and addition test with L* values scaled to produce a ratio of 0,29 gave even better results. The only apparent difference between these two domain lies in the fact that the maximum input values are, in the case of the native L*a*b* domain 255, and in the case of the scaled L*a*b* domain 1. Putting those values under power, x^n , in the first case 255^n produces large values, while in the second case 1^n leaves the value of 1 unchanged. Another consideration that has to be taken into account is that putting values greater than 1 under power produces larger values, while putting values smaller than one produces smaller values than the value under the power. Keeping in mind that there is a positive relationship between the RGB and L* values the results become confusing.

4 Conclusion

The ratio indicators used in this study did not provide the complete explanation of the differences between the two domains. They do indicate differences, as it was noted by different scaling of L* values. The differences that they indicate come

from producing different gradients between data points by scaling the data. If the data is fairly linear, the gradient values between pairs of points are similar. Scaling the data by different factors can affect those gradients and make the data less linear. However, in this case, both domain and co domain values were scaled by the same factor, leaving the gradients unchanged. The fact that putting value smaller than 1 under the power produces a value smaller than the value under the power produced better results when modeling variables with positive relationship can be explained by simple adaptation of the estimated coefficients to model this relationship. Aside from the obvious difference of the behavior of maximum values 255 or 1 under the power, it was noted that by scaling the L* value by 255 produced better results than scaling it by 100. This effect should be further studied and tested on other, a* and b* values as well as it appears to have a potential for further improvements.

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