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**Working Paper**

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Working Paper, No. 866

**Provided in Cooperation with:**

School of Economics and Finance, Queen Mary University of London

*Suggested Citation:* Mumtazy, Haroon; Mussoz, Alberto (2018) : The evolving impact of global, region-specific and country-specific uncertainty, Working Paper, No. 866, Queen Mary University of London, School of Economics and Finance, London

This Version is available at:

<https://hdl.handle.net/10419/210423>

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# The evolving impact of global, region-specific and country-specific uncertainty

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Working paper No. 866

September 2018

ISSN1473-0278

## School of Economics and Finance



# The evolving impact of global, region-specific and country-specific uncertainty\*

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## Abstract

We develop a dynamic factor model with time-varying parameters and stochastic volatility, estimate it with several variables for a large number of countries and decompose the variance of each variable in terms of contributions from uncertainty common to all countries ('global uncertainty'), region-specific uncertainty and country-specific uncertainty. Among other findings, the estimates suggest that global uncertainty plays a primary role in explaining the volatility of inflation, interest rates and stock prices, although to a varying extent over time, while all uncertainty components are found to play a non-negligible role for real economic activity, credit and money for most countries.

JEL Codes: C15, C32, E32

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\*The authors would like to thank for useful comments Nicholas Bloom, Oliver De Groot, Jesus Fernandez-Villaverde, Fabio Fornari, Geoff Kenny, Francesco Mongelli, an anonymous referee of the ECB Working Paper Series and participants to presentations of previous versions of this paper at the ECB Workshop on "The Economics of Uncertainty" (Frankfurt, Germany, November 2017), the Stanford Institute for Theoretical Economics (SITE) workshop on "Macroeconomics of Uncertainty and Volatility" (Stanford, US, August 2018) and the 33rd Annual Congress of the European Economic Association (Cologne, Germany, August 2018). The views expressed in this paper are those of the authors, and not necessarily those of the European Central Bank or the Eurosystem.

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Key Words: Dynamic Factor Model, Time-Varying Parameters, Stochastic Volatility,  
Uncertainty Shocks, Global Uncertainty

## 1 Introduction

After a period characterised by increasing macroeconomic stability observed in several advanced economies from the mid- to late 1980s onwards, known as the Great Moderation, the world economy has experienced a substantial increase in financial and macroeconomic volatility as a result of the global financial crisis starting in the summer of 2007, a major global recession between 2008 and 2009 and regional crises such as the sovereign debt crisis in Europe starting in 2010. The prolonged adverse effects of these developments, despite the implementation of several unconventional monetary policy measures by the major central banks around the world, have led to a marked increase in the degree of uncertainty prevailing in several countries, which is likely to be a significant factor behind the slow pace of the recent recoveries observed in several advanced economies.

As a result, the analysis of the role of volatility and uncertainty in the macroeconomy has regained a prominent role in recent years. This is reflected in the publication of several studies on the role of uncertainty shocks during the course of the past decade (see Bloom, 2014, for a recent overview of this literature). Several of these studies conclude that unexpected large changes in uncertainty (or the closely related concepts of risk and volatility) represent an important source of macroeconomic fluctuations (see for example Bloom, 2009, and Christiano et al., 2014) and also explain a significant fraction of the contraction in real GDP observed during the latest global recession of 2008-2009, known as the Great Recession (see for example Bloom et al., 2012, and Stock and Watson, 2012).

From a policy perspective, measuring, monitoring and analysing the impact of uncertainty is very important for various reasons. First, uncertainty can affect the macroeconomy through several channels, whether it reflects exogenous factors such as natural disasters or geopolitical turmoil, thus representing a source of macroeconomic fluctuations, or whether it arises as an endogenous response to other macroeconomic forces, such as specific aggregate demand shocks or aggregate supply shocks, thus contributing to amplify their effects. Indeed, heightened uncertainty can transmit through the macroeconomy by affecting spending decisions of households and firms, for example inducing them to postpone consumption and investment, as well as financial markets, for example if expected asset price volatility leads to increased risk premia which are then transmitted to higher cost of credit to families and companies. Second, the degree of macroeconomic uncertainty and volatility may affect the effectiveness of economic policies. For example, slowdowns characterised by a high degree of uncertainty might require a more substantial monetary policy stimulation package to support the economy and achieve a desired increase in aggregate demand compared to recessions coinciding with a more muted degree of uncertainty. As a result, the assessment of macroeconomic uncertainty is very much at the centre of attention of policymakers, as discussed in speeches of central bankers (Bernanke, 2007; Carney, 2016; Praet, 2015) and in policy articles (ECB, 2016; Haddow et al., 2013; Kose and Terrones, 2012).

Against this background, the purpose of this paper is to provide estimates of common global, common regional and country-specific macroeconomic uncertainty and to assess the economic impact of the associated uncertainty shocks. More precisely, we address the following two questions. Do fluctuations in uncertainty that are common among advanced economies or common to specific regions such as the Euro Area, North-America or Asia matter more for macroeconomic volatility than country-specific uncertainty shocks? Has the relative importance of these different sources of uncertainty changed over time? In order to carry out this investigation, we build a dynamic factor

model with time-varying factor loadings and stochastic volatility allowing for the estimation of uncertainty that is common across a large set of advanced economies, uncertainty that is common to specific regions (Euro Area, other European countries, North-America, Asia and Oceania) and country-specific uncertainty. We then calculate the contribution of each of these components to the volatility of a large range of macroeconomic and financial series for each country in the panel. The time-varying factor loadings imply that we can assess if the relative importance of these components has changed over time. This represents an advantage compared to the alternative approach adopted often in the literature which consists first in deriving one estimate of macroeconomic uncertainty and then using it as if it were an observable time series within an econometric model such as a recursive VAR to derive inference on the effects of uncertainty shocks on the economy (see for example Caggiano et al., 2014, Basu and Bundick, 2017, Bachman et al., 2013, and Gilchrist et al., 2014, in addition to several of the above-mentioned papers). As noted by Carriero et al. (2016), such a two-step approach has several limitations, including possible omitted variable bias and non-fundamentalness of the errors, linked to the fact that the second step is typically based on small scale VAR models. Our approach allows to overcome such limitations, as the derivation of the uncertainty measures and the inference on the associated uncertainty shocks are derived within a coherent econometric framework including several variables, thereby increasing the reliability of the estimates. Moreover, to the best of our knowledge, this paper is the first to provide a comprehensive estimate of common macroeconomic uncertainty and its economic impact at regional level, along with corresponding estimates for global common macroeconomic uncertainty and country-specific macroeconomic uncertainty. The analysis is based on a large set of quarterly financial and macroeconomic variables spanning from 1960 to 2016 for 22 OECD countries, including eleven Euro Area economies, five other European countries and six other countries. For each of the 22 countries we consider 20 variables and the sample is completed with

20 additional international variables, some referring to prices of commodities such as oil, gas and gold, while other ones representing a small sample of long time series for selected indicators for a number of emerging economies. Overall, 460 time series are included in the sample.

The main results of the empirical analysis are the following. First, all the uncertainty measures display significant recurrent fluctuations, with evidence of alternating periods of high and low persistent uncertainty found for most cases. A historical perspective appears to be very informative, showing for example that the most recent temporary but marked increase in macroeconomic uncertainty associated to the global economic and financial crisis of 2008/2009, which can be observed in most estimates of uncertainty (global, for most regions and for most countries), is not unprecedented and indeed often comparable to uncertainty increases emerging during the first half of the 1970s and early 1980s. Second, we find that all uncertainty measures appear to be strongly countercyclical, with periods of marked increased uncertainty often emerging just before or during the vast majority of recessions, and a strong positive correlation of these measures with inflation. Third, the relative importance of the various uncertainty measures in explaining the volatility of the variables considered appears to differ both over time, geographically (across country and region) and for different variables, but all of them - global uncertainty, region-specific uncertainty and country-specific uncertainty - play a non-negligible role in most cases. Specifically, for real economic activity, credit and money all components appear to be important in most countries, while the volatility of inflation, interest rates and stock prices seems to be driven primarily by the global common uncertainty component in most countries, although to a varying extent over time. By contrast, region-specific uncertainty drives most of the exchange rate volatility especially for all Euro Area countries and for the countries in North-America and Oceania, while for the other countries either country-specific or idiosyncratic uncertainty prevail in importance.

This paper is closely related to various recent developments in the empirical macroeconomic

literature. The aim of the paper is similar in spirit to the work on international business cycles (see for example Kose et al., 2003) and the research on inflation co-movements (see Mumtaz and Surico, 2012) that has sought to establish the importance of a common factor in explaining the movements in these variables. We focus on comovement in the second moment and show that this feature is important. Our analysis is also closely related to the recent literature that has focused on estimating proxies for economic uncertainty for the purposes of monitoring its evolution and deriving estimates of their impact on the economy. Much of this literature has focused on deriving uncertainty measures for the US economy (see for example Carriero et al., 2015, Carriero et al., 2016, and Jurado et al., 2015, which also include an overview of this literature), although a number of recent studies have also provided estimates of global uncertainty along with related country-specific uncertainty measures (Cesa-Bianchi et al, 2014; Berger et al., 2016; Ozturk and Sheng, 2017; Mumtaz and Theodoridis, 2017). However, there is a lack of estimates of common region-specific uncertainty, despite the obvious fact that some sources of uncertainty for several economies in specific regions are common, including the Euro Area as a result of the process of European economic and monetary integration and of the monetary policy changes of the European Central Bank. One exception is represented by Baker et al. (2016), which however derive a measure of European economic policy uncertainty, based on newspaper articles regarding policy-related economic uncertainty, instead of European or Euro Area macroeconomic uncertainty.<sup>1</sup> In contrast to these studies, we investigate the role of alternative sources of uncertainty with special attention to sources of *common* movements in uncertainty, by explicitly accounting for uncertainty common to various regions as well as uncertainty that is common across the entire set of countries. Moreover, we focus on common *macroeconomic* uncertainty reflected in real and nominal aggregate variables

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<sup>1</sup>See in particular the European Economic Policy Uncertainty Index, which runs from January 1997 onwards and is based on data for Germany, France, the UK, Italy, Spain and the Netherlands. The index can be found in the website associated to the paper Baker et al., (2016): [http://www.policyuncertainty.com/europe\\_monthly.html](http://www.policyuncertainty.com/europe_monthly.html). For other Euro Area uncertainty measures under development see ECB (2016).



as well as in financial variables. Finally, our analysis is a generalisation of the investigation by Muntaz and Theodoridis (2017) as we use a substantially more comprehensive data set and allow for time-varying parameters in the factor model.

Our results have potentially important policy implications. Indeed, accounting for different sources of uncertainty can inform the assessment of the macroeconomic landscape and the optimal policy response. For example, if increased macroeconomic uncertainty is predominantly driven by the country-specific uncertainty component then a set of domestic policy measures might represent the most appropriate response to mitigate its potentially adverse effects. By contrast, if the regional common uncertainty component is the main driver of a specific macroeconomic uncertainty spike observed in several countries of that region, then a set of coordinated policy measures by national authorities of that region might be warranted. Finally, heightened macroeconomic uncertainty driven mainly by the global uncertainty component in specific periods might be beyond the control of national or even regional policy authorities (such as the European Central Bank for the Euro Area countries) if acting in isolation and might require, under specific circumstances, coordinated policy responses at global level. By showing the changing role of the different components of uncertainty in explaining the volatility of several core macroeconomic variables, we suggest that it is important to monitor all three sources of uncertainty, global, region-specific and country-specific, in order to understand developments in macroeconomic fluctuations as well as inflation and financial cycles, and inform the economic policy process.

The paper is organised as follows: Section 2 introduces the empirical model and provides details on the estimation method, the model specification and the dataset used. The results from the empirical model are presented in Section 3, including uncertainty estimates and the role of uncertainty shocks via a variance decomposition analysis. Section 4 provides conclusions. A supplemental appendix includes various annexes that provide further details on the model, the

data and results. More specifically, Annex I includes a detailed description of the technical aspects of the model estimation, Annex II describes in detail the data set, while Annex III reports a more comprehensive set of results.

## 2 Global, region-specific and country-specific uncertainty

In this section we describe the econometric model used and provide some details on the estimation and the dataset. Annex I provides more details on the technical aspects of the estimation.

### 2.1 The model

In order to estimate country-specific, region-specific and global (also referred to as ‘world’) measures of uncertainty, we use a dynamic factor model with time-varying volatility and time-varying factor loadings. The factor model is defined as

$$X_{it} = B_{i,t}^W F_t^W + B_{i,t}^R F_t^R + B_{i,t}^C F_t^C + v_{it} \quad (1)$$

where  $X_{it}$  is a panel of macroeconomic and financial data for the set of industrialised countries described below. This panel of data is summarised by four components: a set of  $K$  factors common to all countries  $F_t^W$ ,  $K$  region-specific factors  $F_t^R$  for each region, a set  $K$  country-specific factors  $F_t^C$  for each country and idiosyncratic components  $v_{it}$ . The region and the country-specific factors are distinguished from the world factors by placing zero restrictions on the factor loadings. For example, series belonging to country  $i$  load on the regional factor specific to the region where the country is assigned and have a zero factor loading associated with all other regional factors. The

world, regional and the country factors follow VAR processes:

$$F_t^W = c^W + \sum_{j=1}^P \beta_j^W F_{t-j}^W + (\Omega_t^W)^{1/2} e_t^W \quad (2)$$

$$F_t^R = c^R + \sum_{j=1}^P \beta_j^R F_{t-j}^R + (\Omega_t^R)^{1/2} e_t^R \quad (3)$$

$$F_t^C = c^C + \sum_{j=1}^P \beta_j^C F_{t-j}^C + (\Omega_t^C)^{1/2} e_t^C \quad (4)$$

Note that equations 2, 3 and 4 allow the world, country and regional factors to have a dynamic relationship. The idiosyncratic components have an AR transition equation

$$v_{it} = \sum_{j=1}^J \rho_j v_{it-j} + h_{it}^{1/2} \varepsilon_{it} \quad (5)$$

with  $e_t^W, e_t^R, e_t^C, \varepsilon_{it} \sim N(0, 1)$ . Following Del Negro and Otrok (2008), we allow for time-varying factor loadings. Collecting the factor loadings at time  $t$  in a matrix  $B_{i,t} = [B_{i,t}^W; B_{i,t}^R; B_{i,t}^C]$ , the law of motion describing their time-variation is given by:

$$B_{i,t} = B_{i,t-1} + (Q_i^B)^{1/2} U_t \quad (6)$$

Note that the error terms in equations 2, 3, 4 and 5 are heteroscedastic. The error covariance matrices in the VAR models 2, 3 and 4 are defined as:

$$\Omega_t^J = (A^J)^{-1} H_t^J (A^J)^{-1'} \quad (7)$$

where  $J = W, R, C$ .  $A^J$  are lower triangular and  $H_t^J$  are diagonal matrices defined as

$$H_t^J = \text{diag}(S_k^J \lambda_t^J) \quad (8)$$

The time-varying volatility is captured by  $\lambda_t^J$  with  $S_k$  representing scaling factors for  $k = 1, 2, \dots, K$ . The overall volatilities evolve as AR(1) processes:

$$\ln \lambda_t^J = \alpha^J + \beta^J \ln \lambda_{t-1}^J + (Q^J)^{1/2} \eta_t^J \quad (9)$$

The structure defined by equation 7 suggests that the volatility specification captures the *overall* volatility in the orthogonalized residuals of the VAR models. As explained in Carriero et al. (2015), the common volatilities can be interpreted as the average of the variance of the shocks with equal weight given to each individual volatility. Note that the errors to these equations represent the shocks to ‘world’, region and country factors. Thus  $\lambda_t^W, \lambda_t^R, \lambda_t^C$  capture the average volatility of the unpredictable part of the common component, the region-specific and the country-specific component. We interpret these volatilities as measures of uncertainty associated with global economic conditions, region-wide economic conditions and country-specific economic conditions.

The variance of the shocks to the idiosyncratic components are also assumed to be heteroscedastic with  $h_{it}$  evolving as a stochastic volatility process

$$\ln h_{it} = a_i + b_i \ln h_{it-1} + q_i^{1/2} n_{it} \quad (10)$$

The structure of the model implies that the unconditional variance of each series can be written

as a function of  $\Omega_t^J$  ( $J = W, R, C$ ) and  $h_t$ . In particular

$$\text{var}(X_{it}) = (B_{i,t}^W)^2 \text{var}(F_t^W) + (B_{i,t}^R)^2 \text{var}(F_t^R) + (B_{i,t}^C)^2 \text{var}(F_t^C) + \text{var}(e_{it}) \quad (11)$$

where the variance terms on the RHS of equation 11 can be calculated using the standard VAR formula for the unconditional variance. Note that these variance terms are time-varying as they are functions of  $\lambda_t^W, \lambda_t^R, \lambda_t^C$  and  $h_{it}$  respectively. The volatility of each series in our panel is thus driven by uncertainty that is common to all countries, uncertainty that is common to specific regions (Euro Area, other Europe, North-America, Asia and Oceania), uncertainty that is country-specific and a residual term that captures sectoral volatility and data uncertainty. Our framework, therefore, allows us to calculate how volatility of key series (such as GDP growth, CPI inflation, interest rates, credit and stock market prices) is driven by uncertainty that is common to all countries and uncertainty that is region-, country- and series-specific. As we allow for time-varying factor loadings, the contribution of each of these components is time-varying.

The underlying intuition of this empirical model is related to the procedure used in Jurado et al. (2015) to estimate US economic uncertainty. The uncertainty measure in that study is the average time-varying variance in the unpredictable component of a large set of real and financial time-series. The volatility specification in our factor model has a similar interpretation – it attempts to capture the average volatility in the *shocks* to the factors that summarise real and financial conditions. In contrast to Jurado et al. (2015), however, our model allows the estimation of uncertainty at the country and at the ‘world’ and regional level.<sup>2</sup>

The model proposed is more general than those employed in Mumtaz and Surico (2012) and Berger et al. (2016) along a number of dimensions. These studies focus on the volatility of

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<sup>2</sup>Mumtaz and Theodoridis (2017) use a factor model with common stochastic volatility to estimate the time-varying impact of aggregate uncertainty shocks on the US economy.

the unpredictable component of output growth or inflation factors. In contrast, our analysis focuses on average volatility associated with shocks to factors that span a range of macroeconomic and financial variables. In other words, our focus is on *aggregate economic* uncertainty rather than a narrower measure focussing on a particular variable. Moreover, our analysis considers the possibility of uncertainty at the regional level. As mentioned above, recent events such as the sovereign debt crisis of 2010 in the Euro Area, the Asian financial crisis starting in 1997 as well as specific economic and geopolitical events affecting in particular countries in regions with close trade and financial linkages have highlighted the importance of this economic block. Finally, the proposed model also generalises the work in Mumtaz and Theodoridis (2017), by considering a more comprehensive data set and allowing for time-varying factor loadings.

## 2.2 Estimation and model specification

The factor model described in equations 1 to 10 is estimated via Gibbs sampling. Annex I provides details on the priors and the conditional posterior distributions. In short, the algorithm exploits the fact that, given the factors, the model consists of a sequence of regressions with time-varying parameters and VARs with stochastic volatility, where the conditional posteriors are well known and easily sampled from.

In the benchmark specifications, we use 20,000 replications and base our inference on the last 1,000 replications. The recursive means of the retained draws (see technical appendix) show little fluctuation providing support for convergence of the algorithm.<sup>3</sup>

In order to maintain parsimony, the lag lengths in the VARs ( $L$ ) are fixed at 2. In addition, we allow for first order serial correlation in the idiosyncratic errors  $v_{it}$ . The number of factors is an

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<sup>3</sup>Annex I presents results from a small Monte-Carlo experiment that shows that this MCMC algorithm performs well.

important specification choice. In the benchmark model, we fix the number of common, region-specific and country-specific factors to 3.<sup>4</sup> While in theory it is possible to use (likelihood-based) model selection criteria to select the number of factors, the large number of state-variables in the model make an accurate calculation of the likelihood infeasible.<sup>5</sup>

## 2.3 Data

The data includes a large set of quarterly financial and macroeconomic variables spanning from the first quarter of 1960 to the fourth quarter of 2016 for 22 OECD countries, including eleven Euro Area economies (Germany, France, Italy, Spain, the Netherlands, Belgium, Austria, Finland, Greece, Ireland and Portugal), five other European countries (the UK, Sweden, Denmark, Switzerland and Norway), two North-American countries (the US and Canada), two Asian countries (Japan and South Korea) and two Oceanian countries (Australia and New Zealand). For each of the 22 countries we consider 20 variables, ranging from real economic activity variables (real GDP, real private consumption expenditure, real gross fixed capital formation, industrial production, retail sales), consumer prices (CPI), labour market variables (employment and the unemployment rate), asset prices (stock market prices and house prices), interest rates (short-term interest rates and long-term interest rates), credit market variables (total credit to the private sector and bank loans to the non-financial private sector), money (narrow money and broad money), international trade variables (real exports and real imports) and exchange rates (the nominal effective exchange rate and the US dollar exchange rate). The sample is completed with 20 more international variables, including 8 time series referring to international prices of commodities such as crude oil, natural gas, agricultural products (food, beverages and raw materials), fertilizers and metals (pre-

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<sup>4</sup>This implies that each series for each country loads on 9 factors. Given that we only have 20 series per country, we consider this as the upper limit on the number of factors.

<sup>5</sup>The benchmark model contains 3836 state-variables. Our attempts to calculate the likelihood via a (Rao Blackwellized) particle filter suggest that the estimate is extremely sensitive to initial conditions.

cious metals and other metals and minerals), and 12 time series for selected indicators available over the time span mentioned above for a number of emerging economies (China, India, Turkey, Mexico and South Africa). Overall, 460 time series are included in the sample. Annex II provides details on the data definitions and sources.

### 3 Empirical results

#### 3.1 Estimates of uncertainty components

The measures of macroeconomic uncertainty based on the dynamic factor model are represented by the posterior estimates of the common standard deviation of the shocks to the global factors  $(\lambda_t^W)^{1/2}$ , the posterior estimates of the common standard deviation of the shocks to the regional factors  $(\lambda_t^R)^{1/2}$  and the posterior estimates of the common standard deviation of the shocks to the country-specific factors  $(\lambda_t^C)^{1/2}$ , for  $C = 1, 2, \dots, 22$ . They are displayed in Figure 1 (global), Figure 2 (Euro Area), Figure A in Annex III (other regions) and Figures B and C in Annex III (countries).

Figure 1 displays the global uncertainty measure, along with global recessions as dated by the IMF and several selected events which arguably have a global nature or relevance, either relating to major economic events (dashed vertical lines) or associated to major geopolitical events with significant economic implications such as turmoil in the Middle East with implications for global oil prices (dotted vertical lines). A visual inspection of the figure suggests that global uncertainty spikes are often associated to recessions, as most global recessions are preceded (mid-1970s and early 1980s) or accompanied (2009) by marked increases in global macroeconomic uncertainty. The early 1990s global recession appears to represent an exception, possibly explained by the fact that in some countries such as the US the recession and associated increased uncertainty took place earlier (around 1991) than in most European countries, where the expansionary effect of



German re-unification more than offset the decline in Euro Area foreign demand and implied that a recession was experienced only later (1992 or 1993). As expected, the largest increase in global macroeconomic uncertainty can be observed in 2008, as most countries in the sample experienced increased financial volatility, banking crises and a major recession, although the spike does not seem to be significantly higher than those observed in the first half of the 1970s or early 1980s. Major geopolitical events leading to marked adverse oil price shocks in the mid-1970s and 1979-1980 appear to be factors which can be associated to significant increases in global macroeconomic uncertainty, but in more recent decades similar events seem to have a more limited effect, possibly due to the increased resilience of advanced economies to oil price shocks. Indeed, as discussed in Blanchard and Gali (2007) and Blanchard and Riggi (2103), although in recent years the global economy has witnessed various oil shocks of sign and magnitude comparable to those of the 1970s, their macroeconomic impact has been much more limited. Among the major economic events, while several appear to have had a limited impact on global macroeconomic uncertainty, including the Asian financial crisis starting in 1997 and the start of the Dotcom bubble crash around 2000, other ones such as financial turbulence in housing markets and interbank money markets leading to the start of the financial crisis in the summer of 2007 appear to coincide with a marked increase in global macroeconomic uncertainty.

<FIGURE 1 AROUND HERE>

A comparison of the dynamics of the common global uncertainty measure with an alternative global uncertainty measure is presented in the top panel of Figure D of Annex III (with all indices reported in standardised form, i.e. demeaned and divided by their respective standard deviation, to enhance the comparison). More precisely, the chart shows the estimated common global uncertainty measure along with the Global News Index (GNI) of Baker, Bloom and Davis (2016).

The correlation of our measure with the GNI measure from 1997 (i.e. the starting point of the latter) onwards is close to zero (-0.06), which is not surprising as the GNI measure is based on news references to specific uncertainty aspects, in particular relating to economic policies. Overall, the differences across these two indicators can be associated to the fact that they aim at capturing different aspects of global uncertainty: macroeconomic versus economic policy, such that they could be seen as providing complementary, rather than substitute, information.

Among the region-specific common uncertainty measures, a particularly interesting one to analyse is that for the Euro Area, given the multiple steps toward economic and monetary integration that the countries that adopted the Euro have implemented over the past decades. The common Euro Area uncertainty measure is shown in Figure 2, along with Euro Area recessions as identified by the CEPR Euro Area Business Cycle Dating Committee (grey shaded areas) and several selected events which arguably have a Euro Area nature, either relating to the process of European economic and monetary integration (dashed vertical lines) or associated to changes in the ECB monetary policy (dotted vertical lines). Also in this case it appears that heightened uncertainty is often associated to recessions, as it can be found in coincidence with all of the recessionary periods reported, with the exception of the very latest recession (2011-2013), which arguably was experienced by most Euro Area countries in somewhat different periods and with different intensity (severity and duration), a fact reflected in the different dynamics of the country-specific uncertainty measures (Figure B in Annex III). In contrast to the case of the global uncertainty measure, the increase in Euro Area uncertainty which can be observed during the 2008-2009 recession is not the highest by historical standards, being clearly more limited than that observed in the mid-1970s. Overall, the increased uncertainty in 2008-2009 appears to include a stronger global component than a regional common component, suggesting that it can be associated to multiple causes and channels of transmission with a marked international component. At the same time, significant

increases in 2008-2009 can be observed not only for the Euro Area common uncertainty measure, but also for other region-specific measures such as those for North-America and Asia (Figure A in Annex III), indicating that the impact of the Great Recession was felt globally but to a different degree in different areas. As regards the effects of specific events relating to European integration, it can be observed that the ERM crises of late 1992 and mid-1993 coincide with increased Euro Area uncertainty, but these events also overlap with the early 1990s Euro Area recession. While for several years the ECB operated in an environment characterised by low uncertainty, this is less the case since 2007. At the same time, it appears that the inception of the European sovereign debt crisis in 2010 and the re-intensification of the crisis in 2012 are not associated with increased common Euro Area uncertainty. This could be explained by the fact that not only such episodes had distinctively heterogeneous effects across Euro Area countries, but it can also be argued that the impact on uncertainty was mitigated by some timely policy measures, such as the introduction of the Securities Markets Programme (SMP), the joint EC/ECB/IMF programme of financial assistance to Greece in the summer of 2010 and then the ECB's announcement of the Outright Monetary Transactions (OMT) in the summer of 2012.

<FIGURE 2 AROUND HERE>

The estimated common Euro Area uncertainty measure appears to display markedly different dynamics than other Euro Area or European uncertainty measures (second top panel of Figure D in Annex III). This applies to both the VSTOXX index and a weighted average of the Economic Policy Uncertainty (EPU) index of Baker, Bloom and Davis (2016) for the largest Euro Area economies (Germany, France, Italy and Spain). The former is a stock market implied volatility (of the EURO STOXX 50<sup>®</sup> Index) measure and can be characterised as a financial market uncertainty measure. Its correlation with the common Euro Area uncertainty measure based on our model

from 2000 (i.e. the starting period of the VSTOXX) onwards is highly significant (0.63), but also suggests that about one third of the time they move in different direction. This is not surprising, as our measure includes some financial variables but also a majority of macroeconomic variables, such that we characterise our measure as a macroeconomic uncertainty measure. The correlation of our measure with the Euro Area EPU index (available from 2001 onwards) is low (0.22), as in the case of the global uncertainty measures. Similar differences can also be detected between the US-specific uncertainty estimate and alternative uncertainty measures proposed for the US (see the lower panels of Figure D in Annex III).<sup>6</sup>

The country-specific estimates of macroeconomic uncertainty confirm that most recessionary episodes are accompanied by a rise in country-specific macroeconomic uncertainty, unless they coincide with a rise in either global or region-specific macroeconomic uncertainty. However, several episodes of heightened uncertainty can also be detected coinciding with other events which are not classified as recessions. An example of such episode is represented by the German re-unification of 1990, which gave rise to a significant increase in the German-specific uncertainty measure, not surprisingly given the unique nature of such event. For several country-specific uncertainty measures, as is the case for the global and regions-specific uncertainty measures, it is noticeable that the 1970s was the decade characterised by the highest degree of volatility, often more marked than during the period of the recent economic and financial crisis, which highlights the importance to undertake such analysis with a historical sample spanning several decades to assess recent developments in a broader perspective.

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<sup>6</sup>For instance, the correlation between the US-specific uncertainty measure and the corresponding one of Jurado et al. (2015) is 0.37, that with the VIX is 0.15 and those with the economic policy uncertainty indices for the US of Baker, Bloom and Davis (2016) are -0.25 for the historical index (available from 1971 to 2014) and 0.21 for the baseline index (available from 1985 onwards) (all standardised indices).

## 3.2 Co-movement of uncertainty

In order to assess to which extent the uncertainty measures relate to each other, it can be instructive to look at the cross-correlations between pairs of uncertainty measures. These are reported in Table A of Annex III.

The correlation between the global common uncertainty measure and the region-specific uncertainty measures is on average around 0.50, a similar figure found for the average correlation between pairs of region-specific measures. While these numbers indicate a significant degree of co-movement among international uncertainty measures, they also suggests that they often capture different components of overall macroeconomic uncertainty, as about half of the time they move in different direction. The global measure and the region-specific measures are also correlated with the country-specific measures (on average about 0.30 for the global and about 0.40 for the region-specific), confirming however that more than half of the time they move in different direction, thereby indicating that they often capture different components of overall uncertainty.

As regards the country-specific macroeconomic uncertainty measures, the average cross-correlation for all pairs of measures is 0.32. Among Euro Area countries the average cross-correlation is only marginally higher (0.35), while it is lower among the countries grouped under the "other Europe" region (0.26) and the Asian region (0.10), in contrast to the strong pairs of correlations between US and Canada and Australia and New Zealand (around 0.60 for both pairs).

The estimates of uncertainty clearly point to a negative correlation with real economic activity growth and a positive co-movement with inflation. The countercyclical nature of uncertainty is confirmed by the negative correlations found for most measures with real GDP growth, both at country level and at aggregate global or Euro Area level (for which aggregate real GDP are readily available, in contrast to the other regions). Indeed, as shown in Table B of Annex III, the

contemporaneous correlations between the uncertainty measures and real GDP quarterly growth of the corresponding country or area (first column) is in most cases negative and clearly significant.

Another feature which emerges from these estimates is the strong positive correlation with inflation, which is in most cases of a magnitude (in absolute value) even greater than that between uncertainty and real GDP growth (second column in Table B of Annex III).

### 3.3 Variance decomposition

In order to assess the extent to which shocks to the different uncertainty components drive the overall volatility of key macroeconomic and financial variables, forecast error variance decompositions are considered. More precisely, using equation 11 the unconditional variance of each variable is decomposed into the contributions of the various uncertainty components (global:  $\lambda_t^W$ , region-specific:  $\lambda_t^R$  and country-specific:  $\lambda_t^C$ , for  $C = 1, 2, \dots, 22$ ) with the residual capturing the idiosyncratic, or variable-specific, volatility. Since the variances in the model are time-varying, the implied decomposition changes over time as well, and it is instructive to assess both the average contributions over the whole sample period as well as the evolution over time of these contributions.

Starting with real economic activity, Table 1 reports the average variance decomposition for a set of real economic activity variables, namely nine variables ranging from real GDP and its components real consumption and real investment to employment and industrial production (detailed results for most of the specific variables in this set are reported in Tables C to G in Annex III). Specifically, the table reports the average contributions (averages of median, 16th percentile and 84th percentile) over the whole sample period of each uncertainty component to real economic activity for each country, region and for the whole world. Looking at overall averages (last row), for all countries and over the whole sample period it appears that idiosyncratic uncertainty ex-

plains 56% of total volatility of real economic activity, much more than country-specific uncertainty (16%), region-specific uncertainty (12%) and global uncertainty (16%). The relative importance of the various uncertainty components does not seem to differ much across countries, except that for the groups of countries in the North-American and Asian regions the region-specific uncertainty component seems to be more important (just above 40% for both regions). For most countries idiosyncratic uncertainty is clearly the most important source of volatility of real economic activity, but also the other three uncertainty components play a significant role in most cases. Looking at contributions over time, it appears that global uncertainty (especially for European countries) and country-specific uncertainty (except for the Asian countries) have gradually become less important on average, while idiosyncratic uncertainty (except for the North-American countries) seems to play a gradually more important role (Figure 3).

<TABLE 1 AND FIGURE 3 AROUND HERE>

As regards consumer price inflation, on average for the majority of countries idiosyncratic uncertainty is also the most important driver of volatility, with global uncertainty representing the second most important component in most cases (Table 2). By contrast, country-specific uncertainty and region-specific uncertainty seem to explain minor fractions of volatility, with few exceptions (notably region-specific uncertainty for the Asian countries appears also important). The importance of global uncertainty for consumer price inflation volatility is in line with the findings of Ciccarelli and Mojon (2010) and Mumtaz and Surico (2012), who provide empirical evidence on the importance of the common international component of inflation, suggesting that inflation in industrialized countries is largely a global phenomenon. From a historical perspective, it is noticeable that the importance of global uncertainty in explaining inflation increased during the 1980s and 1990s, but since then it has become less important on average, although remaining

clearly significant (Figure 4).

<TABLE 2 AND FIGURE 4 AROUND HERE>

For short-term interest rate volatility, global uncertainty is the most important driver on average as well as for most countries and regions (Table 3). A similar picture emerges for long-term interest rate volatility (Table H in Annex III). By contrast, region-specific and country-specific uncertainty appear to be of negligible importance. For most countries, the role of global uncertainty in explaining interest rate volatility has even increased over time (Figure 5). This evidence on the importance of the global uncertainty component for interest rate volatility is in line with the evidence reported in some studies on the existence of a global yield curve (Diebold et al., 2008), along with the declining path of interest rates observed in most countries over the past four decades.

<TABLE 3 AND FIGURE 5 AROUND HERE>

Stock price volatility also appears to be driven first and foremost by global uncertainty, followed in importance by idiosyncratic uncertainty, for most countries on average (Table 4). At the same time, the contribution of region-specific uncertainty seems to be non-negligible for several countries. By contrast, country-specific uncertainty seems to play a negligible role in stock price volatility in all countries. Over time, global uncertainty seems to have been gaining importance in driving stock price volatility on average, with signs of slight diminishing importance only over the past decade (Figure 6). The relevance of the global uncertainty component for stock price volatility supports the view on the presence of a global financial cycle discussed in some recent studies (Miranda-Agrippino and Rey, 2015). This evidence is somewhat in contrast to the case of house price volatility, for which the evidence points to the overwhelming importance of the idiosyncratic



uncertainty component for most countries (see detailed results in Table I and Figure K in Annex III).

<TABLE 4 AND FIGURE 6 AROUND HERE>

As regards other variables, for credit volatility all four uncertainty components seem to play a non-negligible role on average for most countries and regions (Tables J and K and Figures L and M in Annex III). By contrast, for monetary aggregates the idiosyncratic uncertainty component appears to be the most important driver of volatility, with however a non-negligible role also for all the other components in most cases (Tables L and M and Figures N and O in Annex III). Finally, in contrast to the other variables, the evidence for exchange rate volatility differs markedly across groups of countries. Indeed, region-specific uncertainty is clearly the most important source of exchange rate volatility for all Euro Area countries, as well as for the countries in the North-America and Oceania groups (Table N in Annex III), with its relevance strongly increasing over the past three decades (Figure P in Annex III). For the other European countries, the country-specific uncertainty component is the main driver of exchange rate fluctuations, while for Asian countries it is the idiosyncratic uncertainty component to play a major role in explaining exchange rate volatility.

## 4 Conclusions

In this paper we build a dynamic factor model with time-varying factor loadings and stochastic volatility allowing for the estimation of uncertainty that is common across a large set of advanced economies, uncertainty that is common at regional level and country-specific uncertainty. On the basis of a large sample of data comprising 460 quarterly time series for financial and macroeconomic variables for 22 OECD countries spanning from 1960 to 2016, we provide estimates of these

three different components of macroeconomic uncertainty, quantify their impact in explaining the volatility of aggregate real and nominal variables and assess their changing role over time.

Overall, we find that all uncertainty estimates display significant recurrent fluctuations and that the marked increase in macroeconomic uncertainty associated to the global economic and financial crisis of 2008/2009, which can be observed in the global common uncertainty measure, some of the region-specific uncertainty measures as well as in most country-specific uncertainty measures, is not unprecedented and indeed often comparable to uncertainty increases emerging during the first half of the 1970s and early 1980s. Moreover, we find that all uncertainty measures appear to be strongly countercyclical, with periods of marked increased uncertainty often emerging just before or during most recessions, and a strong positive correlation of these measures with inflation. Finally, the relative importance of the various uncertainty measures in explaining the volatility of the variables considered appears to differ somewhat over time and across country and region, but all of them - global uncertainty, region-specific uncertainty and country-specific uncertainty - play a non-negligible role in most cases, including for real economic activity, credit and money. Global common uncertainty appears to play a primary role in explaining the volatility of inflation, interest rates and stock prices in most countries, although to a varying extent over time. Region-specific uncertainty drives most of the exchange rate volatility for all Euro Area countries as well as countries in North-America and Oceania, while for the other countries either country-specific or idiosyncratic uncertainty prevail in importance.

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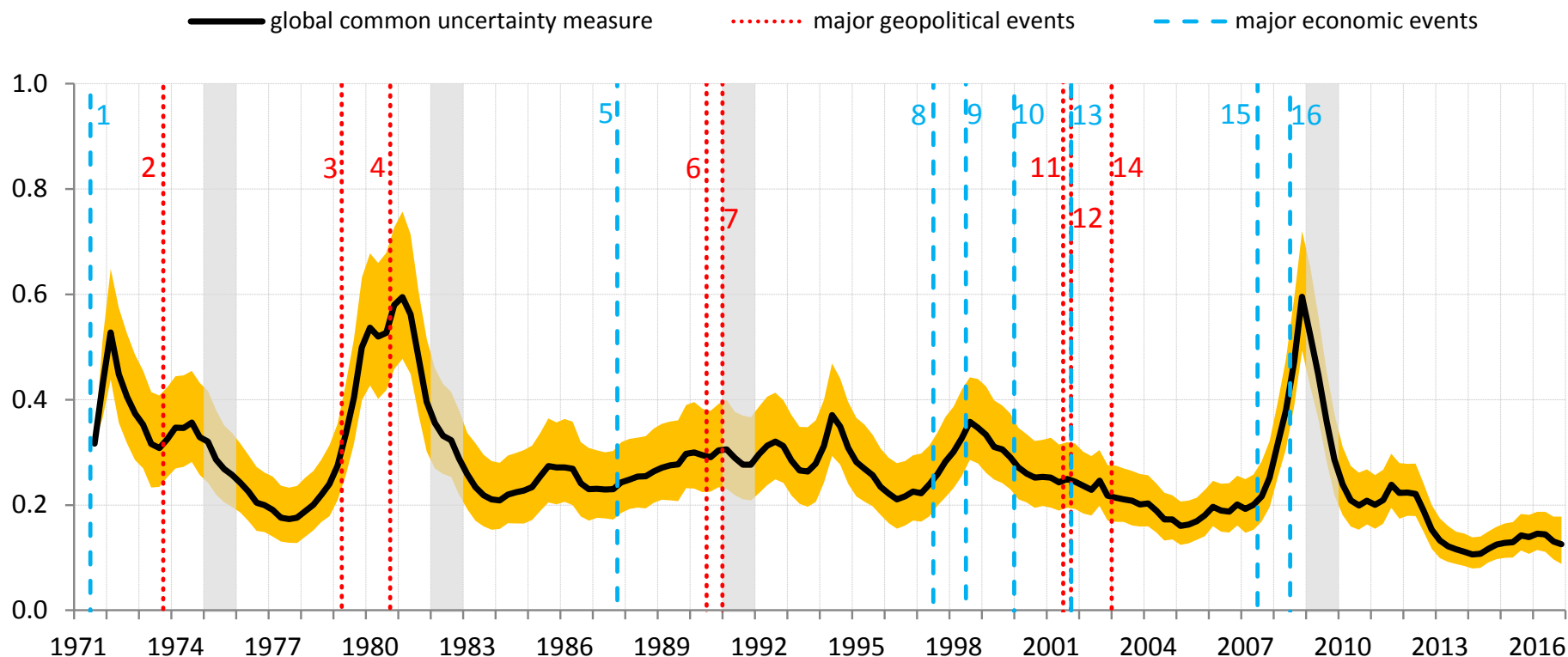
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Chart 1 – Global uncertainty

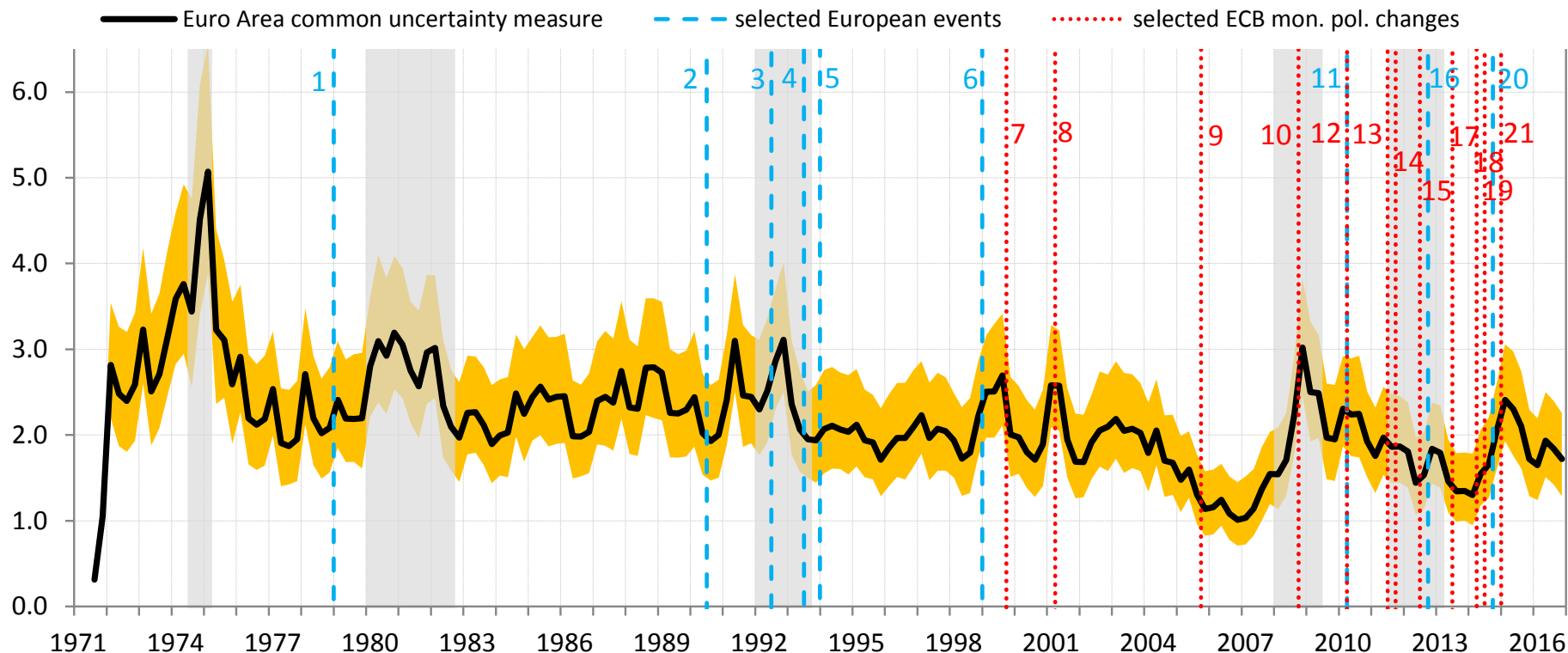


Source: Own estimates, IMF.

Notes: Estimate of the common standard deviation of shocks to the world factors (median and 68 percentile band). Grey areas delimit global recessions as dated by the IMF (April 2009 World Economic Outlook, Box 1.1. on Global Business Cycles). Events associated to vertical dashed and dotted lines are:

1	August 1971: End of Bretton Woods system	9	August 1998: Russian financial crisis
2	October 1973: Arab-Israel conflict and OPEC oil embargo	10	March 2000: Start of Dotcom bubble crash
3	April 1979: Iranian revolution and end of monarchy in Iran	11	September 2001: US terrorist attacks
4	September 1980: Invasion of Iran by Iraq, leading to the Iran-Iraq war	12	October 2001: Start of war in Afghanistan
5	October 1987: Black Monday	13	December 2001: Argentina's debt default
6	August 1990: Invasion of Kuwait by Iraq	14	March 2003: Start of Second Persian Gulf War
7	January-February 1991: First Persian Gulf war	15	August 2007: Start of financial crisis
8	July 1997: Start of Asian financial crisis	16	September 2008: Lehman Brothers bankruptcy

Chart 2 – Euro Area uncertainty



Source: CEPR and own calculations.

Notes: Estimate of the common standard deviation of shocks to the Euro Area factors (median and 68 percentile band). Grey areas delimit Euro Area recessions as dated by the CEPR Euro Area Business Cycle Dating Committee. Events associated to vertical dashed and dotted lines are:

1	March 1979: Creation of the European Monetary System (EMS)	12	May 2010: Sovereign debt crisis (Greece IMF/ECB/EC programme)
2	July 1990: Start of Stage One of EMU	13	August 2011: ECB reactivates SMP programme
3	September 1992: Exchange Rate Mechanism (ERM) crisis	14	December 2011: ECB announces two 3yLTROs
4	August 1993: ERM currency fluctuation bands increased to 15%	15	August 2012: OMTs announcement
5	January 1994: Start of Stage Two of EMU	16	October 2012: Inauguration of European Stability Mechanism (ESM)
6	January 1999: Start of Stage III of EMU	17	July 2013: ECB introduces forward guidance
7	November 1999: Start of ECB interest rate tightening cycle	18	June 2014: ECB announces TLTROs and cuts DFR to negative levels
8	May 2001: Start of ECB interest rate loosening cycle	19	September 2014: ECB announces CBPP3 and ABSPP and cuts DFR
9	December 2005: Start of ECB interest rate tightening cycle	20	November 2014: SSM enters into force
10	October 2008: Start of ECB interest rate loosening cycle	21	January 2015: ECB announces expanded Asset Purchase Programme
11	May 2010: ECB introduces SMP programme		

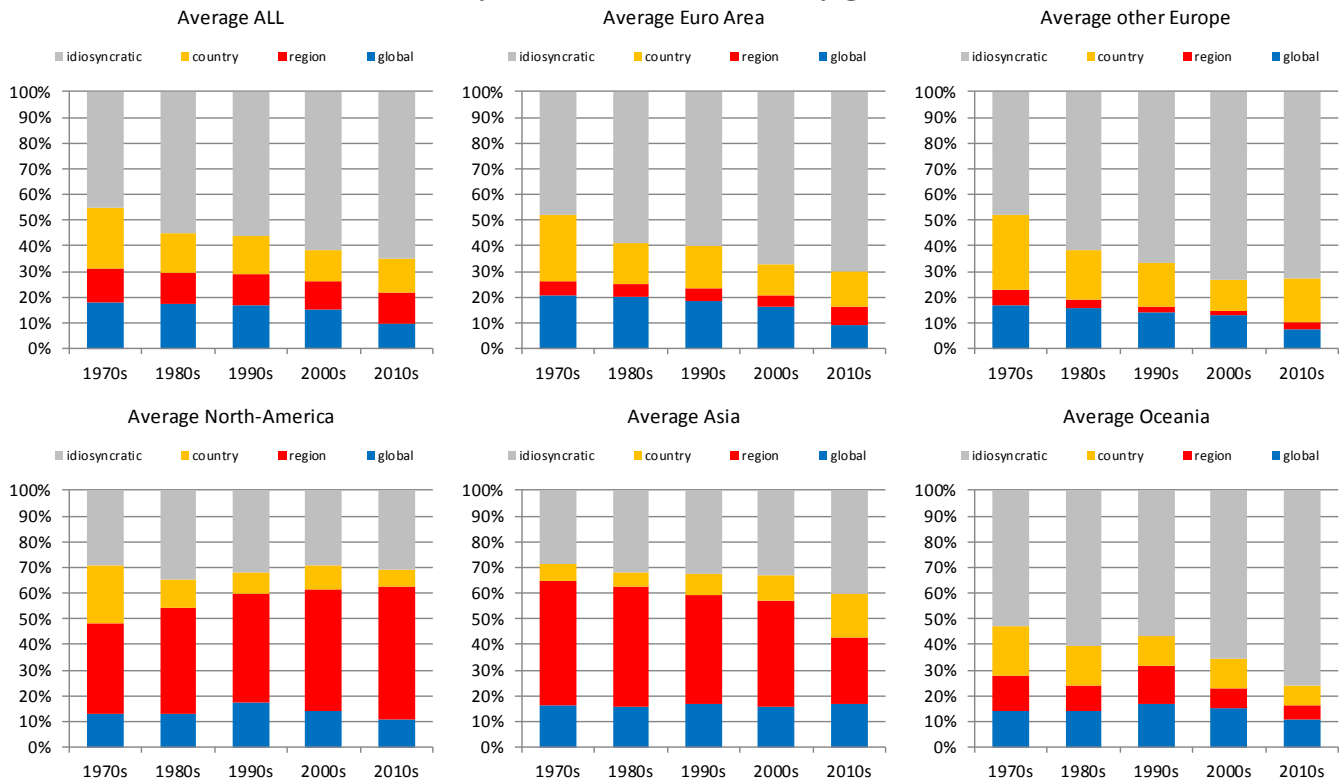
**Table 1 – Variance decompositions: contributions of uncertainty components to the volatility of real economic activity growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	8%	16%	28%	5%	9%	15%	10%	20%	33%	56%
France	10%	18%	32%	5%	9%	16%	8%	16%	28%	57%
Italy	10%	18%	32%	3%	6%	11%	6%	13%	23%	64%
Spain	10%	18%	31%	2%	4%	7%	11%	19%	30%	59%
Netherlands	11%	19%	32%	4%	7%	13%	9%	18%	31%	56%
Belgium	17%	27%	42%	2%	4%	7%	10%	18%	29%	51%
Austria	13%	20%	32%	3%	7%	12%	3%	8%	15%	65%
Finland	9%	16%	28%	2%	3%	6%	8%	14%	25%	67%
Greece	7%	13%	23%	1%	3%	5%	13%	23%	40%	62%
Ireland	6%	13%	24%	1%	3%	6%	10%	19%	32%	65%
Portugal	9%	16%	29%	1%	3%	7%	10%	17%	28%	63%
UK	7%	12%	23%	1%	3%	5%	11%	20%	32%	65%
Sweden	8%	15%	27%	2%	3%	7%	12%	23%	38%	59%
Denmark	9%	15%	25%	1%	3%	6%	14%	22%	32%	60%
Switzerland	7%	14%	25%	2%	3%	7%	9%	15%	23%	68%
Norway	7%	13%	24%	2%	4%	7%	5%	14%	25%	70%
US	7%	15%	29%	23%	41%	60%	6%	15%	30%	30%
Canada	7%	13%	24%	28%	45%	61%	3%	9%	18%	33%
Japan	8%	15%	27%	19%	34%	51%	6%	14%	27%	37%
Australia	10%	16%	27%	4%	9%	17%	8%	16%	29%	59%
New Zealand	7%	13%	24%	6%	12%	22%	5%	11%	19%	64%
Korea	10%	17%	29%	35%	51%	65%	1%	4%	9%	28%
Av. Euro Area	10%	18%	30%	3%	5%	10%	9%	17%	28%	60%
Av. other Europe	8%	14%	25%	2%	3%	6%	10%	19%	30%	64%
Av. North-America	7%	14%	27%	25%	43%	61%	5%	12%	24%	31%
Av. Asia	9%	16%	28%	27%	42%	58%	4%	9%	18%	33%
Av. Oceania	8%	15%	26%	5%	11%	20%	6%	13%	24%	61%
Average ALL	9%	16%	28%	7%	12%	19%	8%	16%	27%	56%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of real economic activity growth (average of contributions to real GDP growth, real private consumption growth, real gross fixed capital formation, employment growth, unemployment rate, industrial production growth, retail sales growth, real export growth and real import growth) over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.



**Chart 3 – Variance decompositions: contributions of uncertainty components to the volatility of real economic activity growth over time**



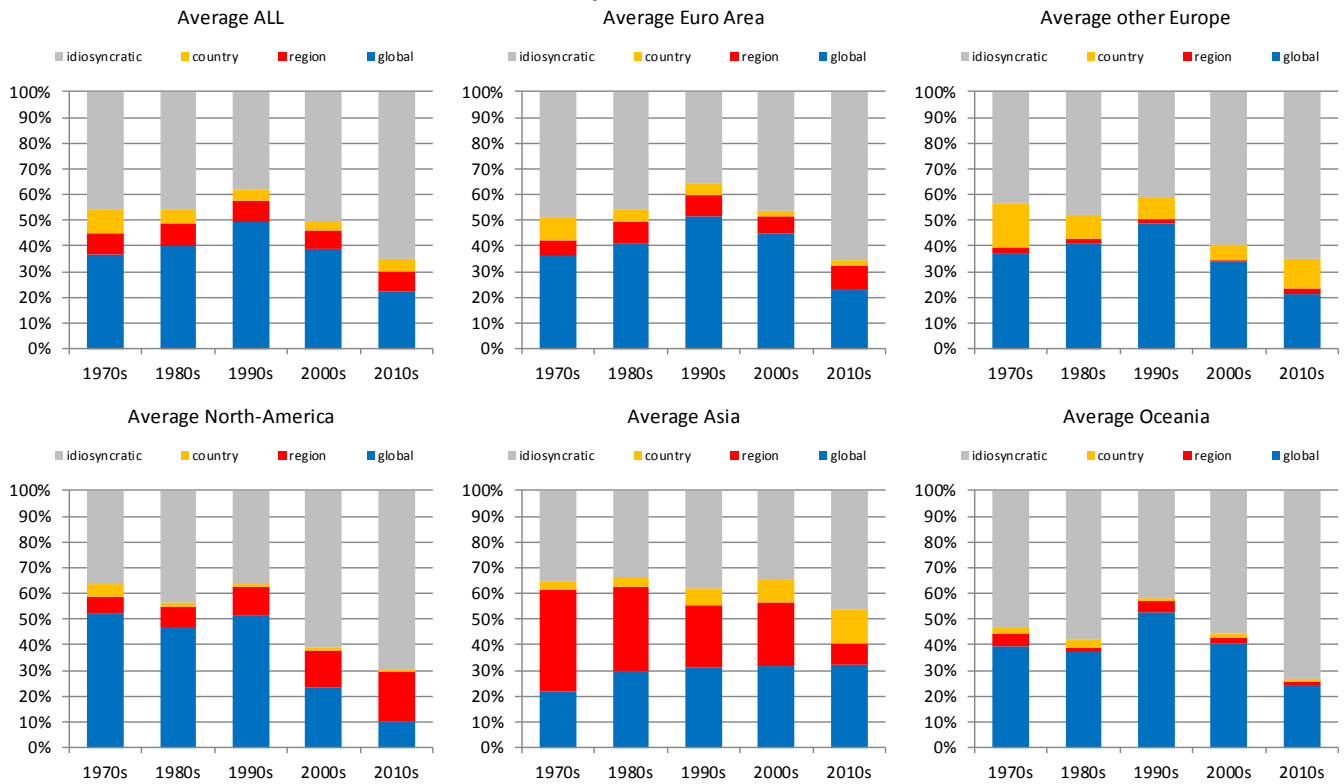
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of real economic activity growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table 2 – Variance decompositions: contributions of uncertainty components to the volatility of CPI inflation**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	6%	18%	41%	5%	12%	26%	2%	7%	19%	63%
France	38%	61%	81%	2%	7%	17%	0%	0%	1%	32%
Italy	43%	63%	81%	1%	4%	10%	0%	0%	2%	33%
Spain	25%	43%	65%	1%	4%	9%	1%	3%	9%	50%
Netherlands	22%	40%	64%	6%	15%	30%	1%	4%	11%	41%
Belgium	13%	30%	56%	2%	6%	16%	0%	2%	7%	62%
Austria	9%	24%	49%	5%	15%	36%	8%	21%	43%	39%
Finland	31%	53%	76%	2%	6%	14%	1%	3%	10%	38%
Greece	14%	31%	57%	1%	5%	12%	0%	3%	17%	62%
Ireland	21%	43%	70%	3%	7%	16%	1%	3%	10%	47%
Portugal	27%	45%	65%	1%	4%	10%	1%	4%	13%	48%
UK	16%	35%	63%	0%	1%	3%	3%	11%	27%	53%
Sweden	21%	43%	69%	0%	1%	4%	1%	5%	19%	51%
Denmark	25%	46%	72%	0%	1%	5%	6%	16%	31%	37%
Switzerland	17%	33%	55%	0%	1%	3%	2%	6%	14%	60%
Norway	14%	32%	57%	1%	3%	8%	3%	12%	32%	53%
US	21%	41%	64%	4%	14%	38%	0%	2%	8%	43%
Canada	19%	37%	60%	2%	8%	29%	1%	2%	7%	53%
Japan	7%	21%	50%	10%	28%	56%	3%	11%	30%	39%
Australia	15%	32%	58%	1%	5%	14%	1%	3%	9%	61%
New Zealand	27%	48%	71%	0%	2%	6%	0%	1%	4%	49%
Korea	14%	37%	67%	9%	26%	55%	1%	2%	7%	34%
Av. Euro Area	23%	41%	64%	3%	8%	18%	1%	5%	13%	47%
Av. other Europe	19%	38%	63%	0%	2%	4%	3%	10%	25%	51%
Av. North-America	20%	39%	62%	3%	11%	34%	0%	2%	7%	48%
Av. Asia	11%	29%	58%	9%	27%	56%	2%	7%	18%	37%
Av. Oceania	21%	40%	64%	1%	3%	10%	0%	2%	7%	55%
Average ALL	20%	39%	63%	3%	8%	19%	2%	6%	15%	48%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of CPI inflation over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart 4 – Variance decompositions: contributions of uncertainty components to the volatility of CPI inflation over time**



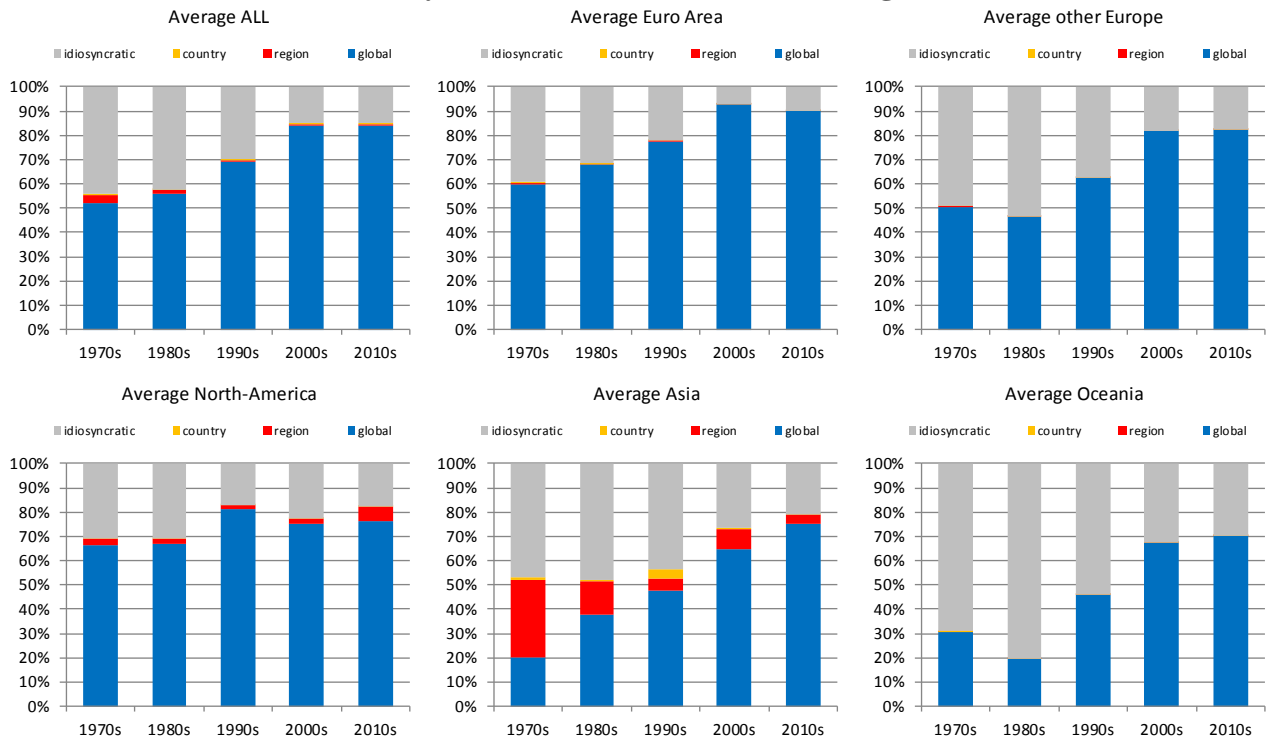
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of CPI inflation. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table 3 – Variance decompositions: contributions of uncertainty components to the volatility of short-term interest rate changes**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	61%	84%	95%	0%	0%	0%	0%	0%	1%	16%
France	79%	91%	97%	0%	0%	0%	0%	0%	0%	8%
Italy	82%	91%	97%	0%	0%	1%	0%	0%	0%	8%
Spain	67%	84%	92%	0%	0%	1%	0%	0%	0%	16%
Netherlands	75%	88%	96%	0%	0%	1%	0%	0%	0%	11%
Belgium	75%	89%	96%	0%	0%	0%	0%	0%	0%	11%
Austria	54%	73%	89%	0%	0%	1%	0%	0%	0%	26%
Finland	9%	28%	52%	0%	0%	0%	0%	0%	0%	72%
Greece	52%	75%	92%	0%	0%	1%	0%	1%	11%	23%
Ireland	69%	83%	93%	0%	0%	1%	0%	0%	1%	17%
Portugal	37%	59%	80%	0%	0%	0%	0%	0%	3%	40%
UK	39%	65%	85%	0%	0%	1%	0%	0%	1%	34%
Sweden	53%	72%	87%	0%	0%	1%	0%	0%	1%	28%
Denmark	77%	89%	96%	0%	0%	0%	0%	0%	0%	11%
Switzerland	38%	64%	85%	0%	0%	0%	0%	0%	0%	36%
Norway	7%	27%	60%	0%	0%	0%	0%	0%	0%	72%
US	27%	62%	86%	1%	3%	9%	0%	0%	1%	35%
Canada	67%	85%	95%	1%	3%	9%	0%	0%	0%	13%
Japan	3%	27%	65%	0%	6%	37%	0%	0%	1%	68%
Australia	28%	50%	70%	0%	0%	0%	0%	0%	0%	50%
New Zealand	19%	41%	65%	0%	0%	1%	0%	0%	0%	59%
Korea	46%	69%	86%	7%	20%	42%	1%	2%	7%	8%
Av. Euro Area	60%	77%	89%	0%	0%	1%	0%	0%	2%	23%
Av. other Europe	43%	64%	82%	0%	0%	0%	0%	0%	1%	36%
Av. North-America	47%	73%	90%	1%	3%	9%	0%	0%	1%	24%
Av. Asia	25%	48%	76%	4%	13%	40%	0%	1%	4%	38%
Av. Oceania	23%	45%	68%	0%	0%	1%	0%	0%	0%	55%
Average ALL	48%	68%	84%	0%	2%	5%	0%	0%	1%	30%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of short-term interest rate changes over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart 5 – Variance decompositions: contributions of uncertainty components to the volatility of short-term interest rate changes over time**



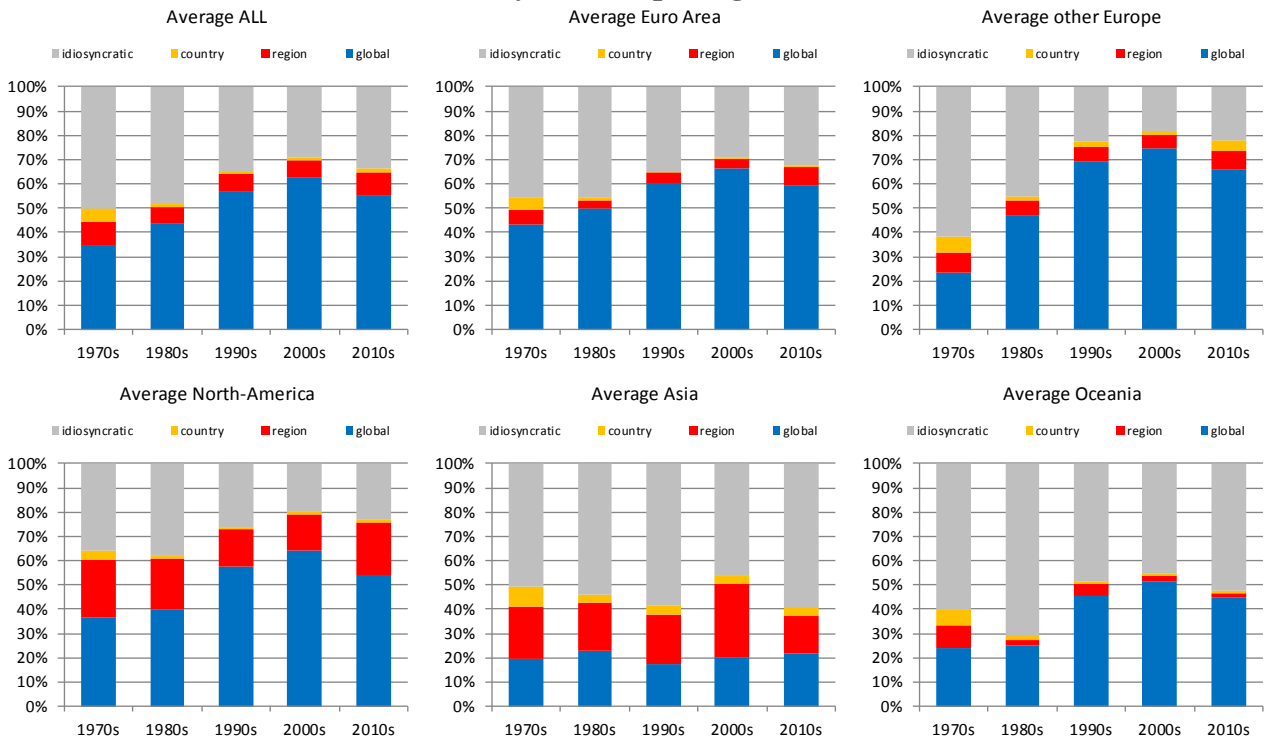
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of short-term interest rate changes. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table 4 – Variance decompositions: contributions of uncertainty components to the volatility of stock price growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	62%	79%	91%	2%	6%	13%	0%	1%	3%	14%
France	66%	83%	93%	2%	6%	15%	0%	0%	1%	11%
Italy	42%	62%	80%	2%	4%	11%	0%	0%	1%	34%
Spain	10%	22%	43%	0%	1%	2%	0%	1%	3%	77%
Netherlands	65%	81%	92%	3%	8%	18%	0%	0%	1%	11%
Belgium	51%	69%	84%	2%	5%	11%	0%	0%	2%	26%
Austria	46%	66%	84%	2%	6%	13%	0%	1%	4%	27%
Finland	31%	53%	75%	2%	7%	16%	1%	2%	7%	38%
Greece	5%	13%	26%	0%	0%	1%	0%	1%	7%	85%
Ireland	36%	58%	78%	2%	6%	14%	1%	4%	13%	31%
Portugal	16%	30%	50%	2%	4%	9%	1%	4%	13%	62%
UK	39%	58%	77%	3%	7%	16%	1%	3%	11%	31%
Sweden	40%	58%	76%	4%	9%	19%	1%	4%	14%	29%
Denmark	37%	52%	69%	2%	5%	11%	1%	4%	9%	39%
Switzerland	38%	59%	79%	3%	8%	18%	1%	3%	8%	30%
Norway	38%	55%	74%	1%	3%	8%	0%	1%	6%	41%
US	30%	53%	76%	6%	20%	47%	0%	1%	9%	25%
Canada	24%	48%	74%	5%	17%	43%	0%	2%	6%	33%
Japan	11%	25%	48%	2%	11%	33%	1%	3%	10%	61%
Australia	36%	56%	76%	0%	2%	7%	1%	3%	9%	39%
New Zealand	9%	20%	41%	2%	6%	15%	1%	2%	6%	72%
Korea	6%	16%	36%	14%	33%	60%	2%	6%	15%	45%
Av. Euro Area	39%	56%	73%	2%	5%	11%	0%	1%	5%	38%
Av. other Europe	38%	57%	75%	2%	6%	14%	1%	3%	9%	34%
Av. North-America	27%	51%	75%	5%	19%	45%	0%	1%	7%	29%
Av. Asia	9%	20%	42%	8%	22%	47%	1%	4%	13%	53%
Av. Oceania	23%	38%	58%	1%	4%	11%	1%	2%	7%	55%
Average ALL	34%	51%	69%	3%	8%	18%	1%	2%	7%	39%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of stock price growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart 6 – Variance decompositions: contributions of uncertainty components to the volatility of stock price growth over time**



Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of stock price growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

The evolving impact of global, region-specific and country-specific  
uncertainty: Online supplementary appendix

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# Annex I - Model specification and estimation

## 1 Model

The dynamic factor model is defined as

$$X_{it} = B_{i,t}^W F_t^W + B_{i,t}^R F_t^R + B_{i,t}^C F_t^C + v_{it} \quad (1)$$

$$B_{i,t} = B_{i,t-1} + (Q_i^B)^{1/2} U_t \text{ where } B_{i,t} = [B_{i,t}^W; B_{i,t}^R; B_{i,t}^C] \quad (2)$$

$$F_t^W = c^W + \sum_{j=1}^P \beta_j^W F_{t-j}^W + (\Omega_t^W)^{1/2} e_t^W \quad (3)$$

$$F_t^R = c^R + \sum_{j=1}^P \beta_j^R F_{t-j}^R + (\Omega_t^R)^{1/2} e_t^R \quad (4)$$

$$F_t^C = c^C + \sum_{j=1}^P \beta_j^C F_{t-j}^C + (\Omega_t^C)^{1/2} e_t^C \quad (5)$$

$$v_{it} = \sum_{j=1}^J \rho_j v_{it-j} + h_{it}^{1/2} \varepsilon_{it} \quad (6)$$

$$R_t = \text{diag}(h_{1t}, \dots, h_{Nt}) \quad (7)$$

$$\Omega_t^W = (A^W)^{-1} H_t^W (A^W)^{-1'}, H_t^W = \text{diag}(S_k^W \lambda_t^W), k = 1, 2, \dots, N \quad (8)$$

$$\Omega_t^R = (A^R)^{-1} H_t^R (A^R)^{-1'}, H_t^R = \text{diag}(S_k^R \lambda_t^R), k = 1, 2, \dots, N \quad (9)$$

$$\Omega_t^C = (A^C)^{-1} H_t^C (A^C)^{-1'}, H_t^C = \text{diag}(S_k^C \lambda_t^C), k = 1, 2, \dots, N \quad (10)$$

$$\ln \lambda_t^W = \alpha^W + \beta^W \ln \lambda_{t-1}^W + (Q^W)^{1/2} \eta_t^W \quad (11)$$

$$\ln \lambda_t^R = \alpha^R + \beta^R \ln \lambda_{t-1}^R + (Q^R)^{1/2} \eta_t^R \quad (12)$$

$$\ln \lambda_t^C = \alpha^C + \beta^C \ln \lambda_{t-1}^C + (Q^C)^{1/2} \eta_t^C \quad (13)$$

$$\ln h_{it} = a_i + b_i \ln h_{it-1} + q_i^{1/2} n_{it} \quad (14)$$

$$U_t, \varepsilon_{it}, e_t^W, e_t^R, e_t^C, \eta_t^W, \eta_t^R, \eta_t^C, n_{it} \sim N(0, 1) \quad (15)$$

## 2 Estimation

### 2.1 Priors and starting values

#### 2.1.1 Factor loadings and factors

The initial values for  $B_{i,t}^j$  is normal and is assumed to be  $N(B_{i,0}, V_B)$  where  $B_{i,0}$  is set equal to the loadings obtained using a principal component estimate of  $F_t = [F_t^W, F_t^R, F_t^C]$  over  $T_0 = 40$  observations. The variance  $V_{i,B}$  is assumed to be equal to the OLS estimate of the coefficient covariance. The prior for  $Q_i^B$  is inverse Wishart with scale matrix  $Q_{i,0}^B = V_{i,B} \times T_0 \times \kappa$  where  $\kappa = 3.5 \times 10^{-4}$  as in Cogley and Sargent (2005) and prior degrees of freedom  $TT_0 = \dim Q_{i,0}^B + 1$ .

The initial estimate of the factors  $F_t^{PC}$  provides the initial value of the factors  $F_{0 \setminus 0}$  with the initial variance set equal to the identity matrix.

#### 2.1.2 VAR Coefficients

Following Banbura et al. (2010) we introduce a natural conjugate prior for the VAR parameters  $\tilde{b}^j = \{c^j, \beta^j\}$  via dummy observations for  $j = W, R, C$ . In our application, the prior means are chosen as the OLS estimates of the coefficients of an AR(1) regression estimated for each endogenous variable using a training sample. The overall prior tightness of this prior  $\tau = 0.1$ .

A similar procedure is used to set the prior for  $\rho$  with prior tightness parameter  $\tau_\rho = 1$

### 2.1.3 Elements of $S, A$ and the parameters of the common volatility transition equation

The elements of  $S^j, j = W, R, C$  have an inverse Gamma prior:  $P(s_i^j) \sim IG(S_{0,i}^j, V_0^j)$ . The degrees of freedom  $V_0$  are set equal to 1. The prior scale parameters are set by estimating the following regression:  $\bar{\lambda}_{it}^j = S_{0,i}^j \bar{\lambda}_{it}^j + \varepsilon_t^j$  where  $\bar{\lambda}_{it}^j$  is the first principal component of the stochastic volatilities  $\tilde{\lambda}_{it}^j$  obtained using a univariate stochastic volatility model for the residuals of each equation of the VAR in equation 3 estimated via OLS using the principal components  $F_t^{PC}$ .

The prior for the off-diagonal elements  $A^j, j = W, R, C$  is  $A_0 \sim N(\hat{a}^{ols}, V(\hat{a}^{ols}))$  where  $\hat{a}^{ols}$  are the off-diagonal elements of the inverse of the Cholesky decomposition of  $\hat{v}^{ols}$ , with each row scaled by the corresponding element on the diagonal. These OLS estimates are obtained using the initial VAR model described above.  $V(\hat{a}^{ols})$  is assumed to be diagonal with the elements set equal to 10 times the absolute value of the corresponding element of  $\hat{a}^{ols}$ .

We set a normal prior for the unconditional mean  $\mu^j = \frac{\alpha^j}{1-\beta^j}$  for  $j = W, R, C$ . This prior is  $N(\mu_0, Z_0)$  where  $\mu_0 = 0$  and  $Z_0 = 10$ . The prior for  $Q^j$  is  $IG(Q_0, V_{Q0})$  where  $Q_0$  is the average of the variances of the transition equations of the initial univariate stochastic volatility estimates and  $V_{Q0} = 5$ . The prior for  $\beta^j$  is  $N(F_0, L_0)$  where  $F_0 = 0.8$  and  $L_0 = 1$ .

### 2.1.4 Parameters of the idiosyncratic shock volatility transition equation

We set a normal prior for the unconditional mean  $\tilde{\mu} = \frac{a}{1-b}$ . This prior is  $N(\mu_0, Z_0)$  where  $\mu_0 = 0$  and  $Z_0 = 10$ . The prior for  $q_i$  is  $IG(q_0, V_{q0})$  where  $q_0 = 0.01$  and  $V_{q0} = 5$ . The prior for  $b$  is  $N(F_0, L_0)$  where  $F_0 = 0.8$  and  $L_0 = 1$ .

## 2.2 Gibbs algorithm

Following Del Negro and Otrok (2005) we fix the initial conditions for the the stochastic volatilities to fix the scale of the factors. As discussed in Del Negro and Otrok (2005) the sign of the factors and factor loadings is not identified separately. Notice, however, that our interest does not focus on recovering these two objects separately in this exercise. We are instead interested in the volatility of the shocks to the factors and this is unaffected by switch in sign of the factors. In addition, as the product of the factors and the factor loadings is unaffected by the sign indeterminacy, we can recover the contribution of each variance component to the variance of  $X_{it}$ .

The Gibbs algorithm cycles through the steps described below. Note that the superscript  $j = W, R, C$ . Note also that  $F_t = [F_t^W, F_t^R, F_t^C]$  and  $B_{i,t} = [B_{i,t}^W; B_{i,t}^R; B_{i,t}^C]$ . The coefficients of the transition equations are given by  $\tilde{b}^j = \{c^j, \beta^j\}$ .

1.  $G(F_t \setminus \Xi)$ : Given a draw for all other parameters (denoted by  $\Xi$ ), the algorithm of Carter and Kohn (2004) is used to sample from the conditional posterior distribution of  $F_t$ . The state-space of the model is:

$$\begin{aligned} X_{it}^{**} &= B_{i,t} F_t^{**} + R_t^{1/2} \varepsilon_{it} \\ F_t &= \mu + f F_{t-1} + \check{Q}_t^{1/2} E_t \end{aligned}$$

where  $X_{it}^{**} = X_{it} - \sum_{j=1}^J \rho_j X_{it-j}$ ,  $F_t^{**} = (F_t - \sum_{j=1}^J \rho_j F_{t-j})$ ,  $E_t = [e_t^W; e_t^R; e_t^C]$  and  $\check{Q}_t$  is block diagonal matrix with  $\Omega_t^W, \Omega_t^R, \Omega_t^C$  on the main diagonal. The conditional posterior is:  $F_t \setminus X_{it}, \Xi \sim N(F_{T \setminus T}, P_{T \setminus T})$  and  $F_t \setminus F_{t+1}, X_{it}, \Xi \sim N(F_{t \setminus t+1, F_{t+1}}, P_{t \setminus t+1, B_{t+1}})$  where  $t = T-1, \dots, 1$ . As shown by Carter and Kohn (2004) the simulation proceeds as follows. First we use the Kalman filter to draw  $F_{T \setminus T}$  and  $P_{T \setminus T}$  and then proceed backwards in time using  $F_{t|t+1} = F_{t|t} + P_{t|t} f' P_{t+1|t}^{-1} (F_{t+1} - f F_{t|t} - \mu)$  and  $P_{t|t+1} = P_{t|t} - P_{t|t} f' P_{t+1|t}^{-1} f P_{t|t}$ . Here  $f$  denotes the autoregressive coefficients of the transition equations 3, 4, 5 in companion form, while  $\mu$  denotes the pre-determined regressors in the transition equations in companion form.

2.  $G(B_{i,t} \setminus \Xi)$ : Given a draw for the factors and the variance of the idiosyncratic component and the serial correlation coefficients  $\rho_j$ , a separate linear time-varying parameter regression model with heteroscedasticity and serial correlation applies to each  $X_{it}$ . In particular, the model for each  $i$  is

$$\begin{aligned} X_{it} &= B_{i,t} F_t + v_{it} \\ v_{it} &= \sum_{j=1}^J \rho_j v_{it-j} + h_{it}^{1/2} \varepsilon_{it} \\ B_{i,t} &= B_{i,t-1} + (Q_i^B)^{1/2} U_t \end{aligned}$$

The model can be transformed to remove heteroscedasticity and serial correlation by creating  $X_{it}^* = \frac{(X_{it} - \sum_{j=1}^J \rho_j X_{it-j})}{\sqrt{h_{it}}}$ ,  $\tilde{F}_t^* = \frac{(F_t - \sum_{j=1}^J \rho_j F_{t-j})}{\sqrt{h_{it}}}$ . This is then a linear state-space model for each  $i$  with iid disturbances with a unit variance and given  $Q^B$  the Carter and Kohn (2004) algorithm is used to draw from the conditional posterior of  $B_{i,t}$ .

3.  $G(Q_i^B \setminus \Xi)$ : Given  $B_{i,t}$ , this conditional posterior is inverse Wishart with scale matrix  $(B_{i,t} - B_{i,t-1}) + Q_{i,0}^B$  and degrees of freedom  $T + TT_0$
4.  $G(\rho \setminus \Xi)$ : Given a draw for the factors, the factor loadings and the variances  $h_{it}$ , a heteroscedastic AR(j) regression applies to each  $i$ :

$$v_{it} = \sum_{j=1}^J \rho_j v_{it-j} + h_{it}^{1/2} \varepsilon_{it}$$

The heteroscedasticity can be removed by dividing both sides by  $\sqrt{h_{it}}$ . Letting,  $y_{it} = \frac{v_{it}}{\sqrt{h_{it}}}$  and  $x_{it} = \frac{[v_{it-1}, v_{it-2}, \dots, v_{it-j}]}{\sqrt{h_{it}}}$  the conditional posterior for  $\rho = [\rho_1, \rho_2, \dots, \rho_j]$  is normal  $N(M^*, V^*)$ :

$$\begin{aligned} M^* &= (V_\rho^{-1} + x'_{it} x_{it})^{-1} (V_\rho^{-1} \rho_0 + x'_{it} y_{it}) \\ V^* &= (V_\rho^{-1} + x'_{it} x_{it})^{-1} \end{aligned}$$

where  $\rho_0$  and  $V_\rho$  are the prior mean and variance for  $\rho$ .

5.  $G(h_{it} \setminus \Xi)$ : Given a draw for the factors, the parameters of the transition equation 14, the serial correlation coefficients  $\rho_j$  and the factor loadings  $B_{i,t}$ , a univariate stochastic volatility model applies for each  $i$ :

$$\begin{aligned} \tilde{v}_{it} &= h_{it}^{1/2} \varepsilon_{it} \\ \ln h_{it} &= a_i + b_i \ln h_{it-1} + q_i^{1/2} n_{it} \end{aligned}$$

where  $\tilde{v}_{it} = v_{it} - \sum_{j=1}^J \rho_j v_{it-j}$ . A particle Gibbs step (described below) is used to draw  $h_{it}$ .

6.  $G(\tilde{b}^j \setminus \Xi)$ . Given a draw of  $\lambda_t^j$ , the left and the right hand side variables of the VAR:  $y_t = F_t$  and  $x_t = [c, F_{t-1}, F_{t-2}, \dots, F_{t-j}]$  can be transformed to remove the heteroscedasticity in the following manner

$$\tilde{y}_t = \frac{y_t}{\lambda_t^{1/2}}, \tilde{x}_t = \frac{x_t}{\lambda_t^{1/2}}$$

Then the conditional posterior distribution for the VAR coefficients is standard and given by

$$N(\tilde{b}^*, \tilde{\Omega} \otimes (X^{*'} X^*)^{-1})$$

where  $\tilde{b}^* = (X^{*'} X^*)^{-1} (X^{*'} Y^*)$ ,  $\tilde{\Omega} = A^{-1} \text{diag}(S) A^{-1'}$  and  $Y^*$  and  $X^*$  denote the transformed data appended with the dummy observations.

7.  $G(A^j \setminus \Xi)$ . Given a draw for the VAR parameters (equations 3, 4 and 5 respectively) the model can be written as  $A^{j'} (v_t^j) = \tilde{e}_t^j$  where  $v_t^j = F_t^j - (c^j + \sum_{p=1}^P \beta_p^j F_{t-p}^j)$  and  $\text{VAR}(\tilde{e}_t^j) = H_t^j$ . This is a system of linear equations with a known form of heteroscedasticity. The conditional distributions for a linear regression apply to each equation of this system after a simple GLS transformation to make the errors homoscedastic. The  $k$ th equation of this system is given as  $v_{kt}^j = -\alpha v_{-kt}^j + \tilde{e}_{kt}^j$  where the subscript  $k$  denotes the  $k$ th column while  $-k$  denotes columns 1 to  $k-1$ . Note that the variance of  $\tilde{e}_{kt}^j$  is time-varying and given by  $\lambda_t^j S_k^j$ . A GLS transformation involves dividing both sides of the equation by  $\sqrt{\lambda_t^j S_k^j}$  to produce  $v_{kt}^{j*} = -\alpha v_{-kt}^{j*} + \tilde{e}_{kt}^{j*}$  where  $*$  denotes the transformed variables and  $\text{var}(\tilde{e}_{kt}^{j*}) = 1$ . The conditional posterior for  $\alpha^j$  is normal with mean and variance given by  $M^*$  and  $V^*$ :

$$\begin{aligned} M^* &= \left( V(\hat{a}^{ols})^{-1} + v_{-kt}^{j*'} v_{-kt}^{j*} \right)^{-1} \left( V(\hat{a}^{ols})^{-1} \hat{a}^{ols} + v_{-jt}^{j*'} v_{jt}^{j*} \right) \\ V^* &= \left( V(\hat{a}^{ols})^{-1} + v_{-jt}^{j*'} v_{-jt}^{j*} \right)^{-1} \end{aligned}$$

8.  $G(S^j \setminus \Xi)$ . Given a draw for the VAR parameters (equations 3, 4 and 5 respectively),  $A^{j'}(v_t^j) = \tilde{e}_t^j$ . The  $k$ th equation of this system is given by  $v_{kt}^j = -\alpha v_{-kt}^k + \tilde{e}_{kt}^j$  where the variance of  $e_{kt}^j$  is time-varying and given by  $\lambda_t^j S_k^j$ . Given a draw for  $\lambda_t^j$  this equation can be re-written as  $\tilde{v}_{kt}^j = -\alpha \tilde{v}_{-kt}^j + \tilde{e}_{kt}^j$  where  $\tilde{v}_{kt}^j = \frac{v_{kt}^j}{\lambda_t^{j,1/2}}$  and the variance of  $\tilde{e}_{kt}^j$  is  $S_k^j$ . The conditional posterior is for this variance is inverse Gamma with scale parameter  $\tilde{e}_{kt}^{j'} \tilde{e}_{kt}^j + S_{0,j}$  and degrees of freedom  $V_0 + T$ .
9. Elements of  $\lambda_t^j$ . Conditional on the VAR coefficients, and the parameters of the volatility transition equation, the model has a multivariate non-linear state-space representation. Following recent developments in the seminal paper by Andrieu et al. (2010), we employ a particle Gibbs step to sample from the conditional posterior of  $\tilde{h}_t^j = \ln \lambda_t^j$ . Andrieu et al. (2010) show how a version of the particle filter, conditioned on a fixed trajectory for one of the particles can be used to produce draws that result in a Markov Kernel with a target distribution that is invariant. However, the usual problem of path degeneracy in the particle filter can result in poor mixing in the original version of particle Gibbs. Recent developments, however, suggest that small modifications of this algorithm can largely alleviate this problem. In particular, Lindsten et al. (2014) propose the addition of a step that involves sampling the ‘ancestors’ or indices associated with the particle that is being conditioned on. They show that this results in a substantial improvement in the mixing of the algorithm even with a few particles.<sup>1</sup> As explained in Lindsten et al. (2014), ancestor sampling breaks the reference path into pieces and this causes the particle system to collapse towards something different than the reference path. In the absence of this step, the particle system tends to collapse to the conditioning path. We employ particle Gibbs with ancestor sampling in this step.

Let  $\tilde{h}_t^{(g-1)}$  denote the fixed trajectory, for  $t = 1, 2, \dots, T$  obtained in the previous draw of the Gibbs algorithm  $g - 1$ . Here we suppress the superscript  $j = W, R, C$  for notational simplicity. The algorithm is applied the three non-linear state space systems defined by the observation and transition equations:

$$\begin{aligned} F_t^j &= c^j + \sum_{p=1}^P \beta_p^j F_{t-p}^j + (\Omega_t^j)^{1/2} e_t^W \\ \Omega_t^j &= (A^j)^{-1} H_t^j (A^j)^{-1'} , H_t^j = \text{diag}(\lambda_t^j S^j) \\ \ln \lambda_t^j &= \alpha^j + \beta^j \ln \lambda_{t-1}^j + (Q^j)^{1/2} \eta_t^j \end{aligned}$$

We denote the remaining parameters of the model by  $\Xi$ , and  $m = 1, 2, \dots, M$  represents the particles. The conditional particle filter with ancestor sampling proceeds in the following steps:

1. (a) For  $t = 1$ 
  - i. Draw  $\tilde{h}_1^{(m)} \setminus \tilde{h}_0^{(m)}, \Xi$  for  $m = 1, 2, \dots, M - 1$ . Fix  $\tilde{h}_1^{(M)} = \tilde{h}_1^{(g-1)}$
  - ii. Compute the normalised weights  $p_1^{(m)} = \frac{w_1^{(m)}}{\sum_{j=1}^M w_1^{(j)}}$  where  $w_1^{(m)}$  denotes the conditional likelihood:  $\left| \Omega_1^{(m)} \right|^{-0.5} - 0.5 \exp \left( e_1 \left( \Omega_1^{(m)} \right)^{-1} e_1' \right)$  where  $e_1 = F_1 - \left( c + \sum_{j=1}^P \beta_j F_{1-j} \right)$  and  $\Omega_1^{(m)} = A^{-1} H_1^{(m)} A^{-1'}$  with  $H_1^{(m)} = \text{diag} \left( \exp \left( \tilde{h}_1^{(m)} \right) S \right)$ .
- (b) For  $t = 2$  to  $T$ 
  - i. Resample  $\tilde{h}_{t-1}^{(m)}$  for  $m = 1, 2, \dots, M - 1$  using indices  $a_t^{(m)}$  with  $\Pr \left( a_t^{(m)} = m \right) \propto p_{t-1}^{(m)}$
  - ii. Draw  $\tilde{h}_t^{(m)} \setminus \tilde{h}_{t-1}^{(a_t^{(m)})}, \Xi$  for  $m = 1, 2, \dots, M - 1$  using the transition equation of the model. Note that  $\tilde{h}_{t-1}^{(a_t^{(m)})}$  denotes the resampled particles in step (a) above.
  - iii. Fix  $\tilde{h}_t^{(M)} = \tilde{h}_t^{(g-1)}$
  - iv. Sample  $a_t^{(M)}$  with  $\Pr \left( a_t^{(M)} = m \right) \propto p_{t-1}^{(j)} \Pr \left( \tilde{h}_t^{(g-1)} \setminus \tilde{h}_{t-1}^{(m)}, \alpha^j, \beta^j, Q^j \right)$  where  $\Pr \left( \tilde{h}_t^{(g-1)} \setminus \tilde{h}_{t-1}^{(j)}, \alpha^j, \beta^j, Q^j \right)$  is computed as  $\left| Q^j \right|^{-0.5} - 0.5 \exp \left( \tilde{\eta}_t^{(m)} (Q)^{-1} \tilde{\eta}_t^{(m)'} \right)$  where  $\tilde{\eta}_t = \tilde{h}_t^{(g-1)} - \left( \alpha^j + \beta^j \tilde{h}_{t-1}^{(m)} \right)$ . This constitutes the ancestor sampling step. If  $a_t^{(M)} = M$  then the algorithm collapses to the simple particle Gibbs.

<sup>1</sup>See Nonejad (2015) for a recent application of this algorithm.

- v. Update the weights  $p_t^{(m)} = \frac{w_t^{(m)}}{\sum_{j=1}^M w_t^{(j)}}$  where  $w_1^{(m)}$  denotes the conditional likelihood:  $\left| \Omega_t^{(m)} \right|^{-0.5} - 0.5 \exp \left( e_t \left( \Omega_t^{(m)} \right)^{-1} e_t' \right)$  where  $e_t = F_t - \left( c + \sum_{j=1}^P \beta_j F_{t-j} \right)$  and  $\Omega_t^{(m)} = A^{-1} H_t^{(m)} A^{-1'}$  with  $H_t^{(m)} = \text{diag} \left( \exp \left( \tilde{h}_t^{(m)} \right), S^j \right)$ .

vi. End

- (c) Sample  $\tilde{h}_t^{(g)}$  with  $\Pr \left( \tilde{h}_t^{(g)} = \tilde{h}_t^{(m)} \right) \propto p_T^{(m)}$  to obtain a draw from the conditional posterior distribution

We use  $M = 50$  particles in our application. The initial values  $\mu_0$  defined above are used to initialise step 1 of the filter.

8.  $G(\alpha^j, \beta^j, Q^j \setminus \Xi)$ . We re-write the transition equation in deviations from the mean (the superscript  $j = W, R, C$  is suppressed below for simplicity)

$$\tilde{h}_t - \mu = \beta \left( \tilde{h}_{t-1} - \mu \right) + \eta_t \quad (16)$$

where the elements of the mean vector  $\mu$  are defined as  $\frac{\alpha}{1-\beta}$ . Conditional on a draw for  $\tilde{h}_t$  and  $\mu$  the transition equation 16 is a simply a linear regression and the standard normal and inverse Gamma conditional posteriors apply. Consider  $\tilde{h}_t^* = \beta \tilde{h}_{t-1}^* + \eta_t$ ,  $\text{VAR}(\eta_t) = Q$  and  $\tilde{h}_t^* = \tilde{h}_t - \mu$ ,  $\tilde{h}_{t-1}^* = \tilde{h}_{t-1} - \mu$ . The conditional posterior of  $\beta$  is  $N(\theta^*, L^*)$  where

$$\begin{aligned} \theta^* &= \left( L_0^{-1} + \frac{1}{Q} \tilde{h}_{t-1}^* \tilde{h}_{t-1}^{*'} \right)^{-1} \left( L_0^{-1} F_0 + \frac{1}{Q} \tilde{h}_{t-1}^* \tilde{h}_t^* \right) \\ L^* &= \left( L_0^{-1} + \frac{1}{Q} \tilde{h}_{t-1}^* \tilde{h}_{t-1}^{*'} \right)^{-1} \end{aligned}$$

The conditional posterior of  $Q$  is inverse Gamma with scale parameter  $\eta_t' \eta_t + Q_0$  and degrees of freedom  $T + V_{Q0}$ .

Given a draw for  $\beta$ , equation 16 can be expressed as  $\tilde{\Delta} \tilde{h}_t = C \mu + \eta_t$  where  $\tilde{\Delta} \tilde{h}_t = \tilde{h}_t - \beta \tilde{h}_{t-1}$  and  $C = 1 - \beta$ . The conditional posterior of  $\mu$  is  $N(\mu^*, Z^*)$  where

$$\begin{aligned} \mu^* &= \left( Z_0^{-1} + \frac{1}{Q} C' C \right)^{-1} \left( Z_0^{-1} \mu_0 + \frac{1}{Q} C' \tilde{\Delta} \tilde{h}_t \right) \\ Z^* &= \left( Z_0^{-1} + \frac{1}{Q} C' C \right)^{-1} \end{aligned}$$

Note that  $\alpha$  can be recovered as  $\mu(1 - \beta)$

9.  $G(a_i, b_i, q_i \setminus \Xi)$ . Given a draw for  $h_{it}$ , the conditional posterior distributions for the parameters of the transition equations 14 are as described in step 8.

## 2.3 A Monte-Carlo experiment

In order to examine the performance of this algorithm, we consider a small Monte-Carlo experiment

### 2.3.1 Data Generating Process

We generate data from the following dynamic factor model with 2 factors:

$$X_{it} = B_i^W F_t^W + B_i^R F_t^R + B_i^C F_t^C + R_t^{1/2} \varepsilon_{it}$$

where the factor loadings  $B_i$  are drawn from  $N(0, 0.5)$  and  $i = 1, 2, \dots, 80$ . We assume that there are four countries and the first two load on the factor  $F_t^R$ .

The dynamics of the country factors are defined as

$$\begin{pmatrix} F_{1t}^C \\ F_{2t}^C \end{pmatrix} = \begin{pmatrix} 0.7 & -0.05 \\ 0.05 & 0.7 \end{pmatrix} \begin{pmatrix} F_{1t-1}^C \\ F_{2t-1}^C \end{pmatrix} + \begin{pmatrix} 0.2 & -0.05 \\ 0.05 & 0.2 \end{pmatrix} \begin{pmatrix} F_{1t-2}^C \\ F_{2t-2}^C \end{pmatrix} + \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix}, \text{var} \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix} = \Omega_t^C$$

The variance process is defined as

$$\begin{aligned}\Omega_t^C &= A^{-1} (S\lambda_t) A^{-1'} \\ A &= \begin{pmatrix} 1 & 0 \\ -1 & 1 \end{pmatrix} \\ S &= \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix} \\ \ln \lambda_t^c &= 0.9 \ln \lambda_{t-1}^c + (0.5)^{\frac{1}{2}} v_t\end{aligned}$$

The dynamics of the world and regional factors are defined as:

$$\begin{pmatrix} F_{1t}^j \\ F_{2t}^j \end{pmatrix} = \begin{pmatrix} 0.7 & -0.05 \\ 0.05 & 0.7 \end{pmatrix} \begin{pmatrix} F_{1t-1}^j \\ F_{2t-1}^j \end{pmatrix} + \begin{pmatrix} 0.05 & -0.05 \\ 0.05 & 0.05 \end{pmatrix} \begin{pmatrix} F_{1t-2}^j \\ F_{2t-2}^j \end{pmatrix} + \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix}, \text{var} \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix} = \Omega_t^j$$

for  $j = W, R$ . Note that  $\Omega_t^j$  is generated in exactly the same manner as described for the country factors above.

We generate 300 observations for  $X_{it}$  and drop the first 100 observations to reduce the influence of initial conditions. The experiment is repeated 100 times. At each iteration, the factor model is estimated using the MCMC algorithm described above using 5000 iterations with a burn-in of 4000 observations. The retained draws are used to calculate the contribution of  $\lambda_t^j$  to the unconditional variance of each variable. In figures 1 and 2 we present the estimated time-varying contribution of world uncertainty. It is clear from the figures that the estimated contribution closely tracks the true value. A similar conclusion can be discerned for the contribution of regional uncertainty shown in figure 3. This provides evidence that the MCMC algorithm described above displays a satisfactory performance.

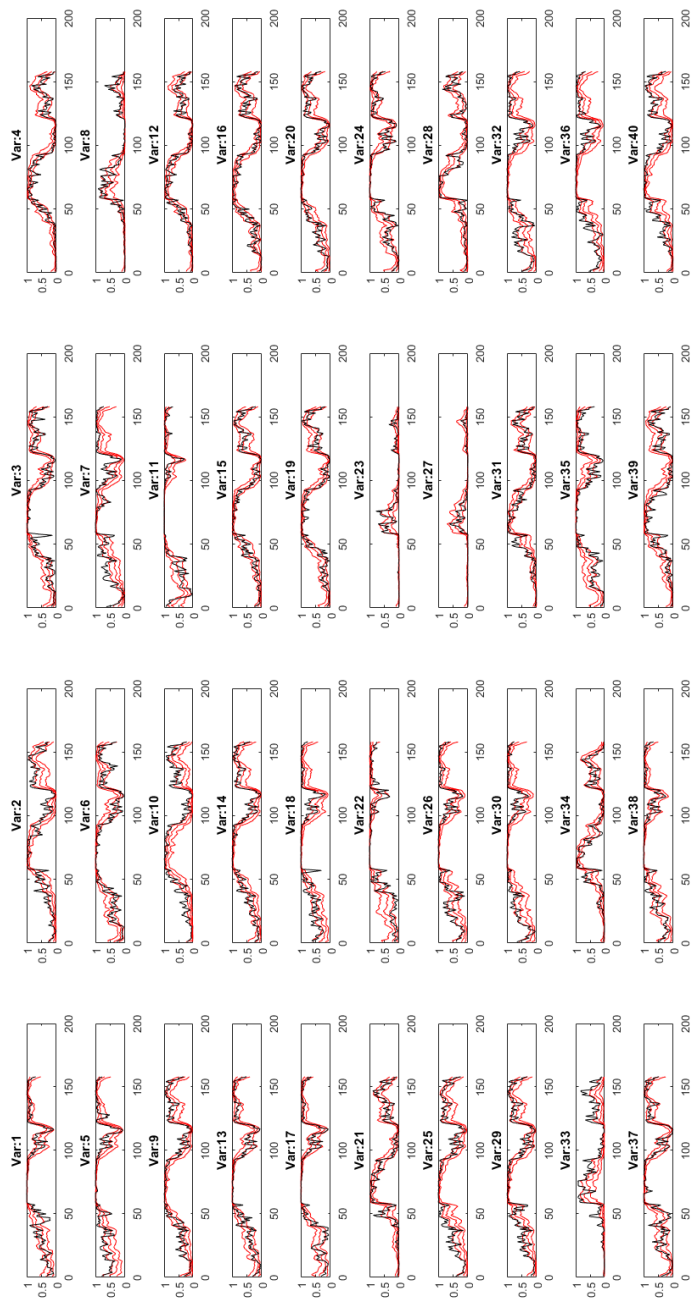


Figure 1: Contribution of World Uncertainty. The black line is the true value. The red lines represent the distribution of the estimate across the Monte-Carlo replications (Median and 1 SD band).

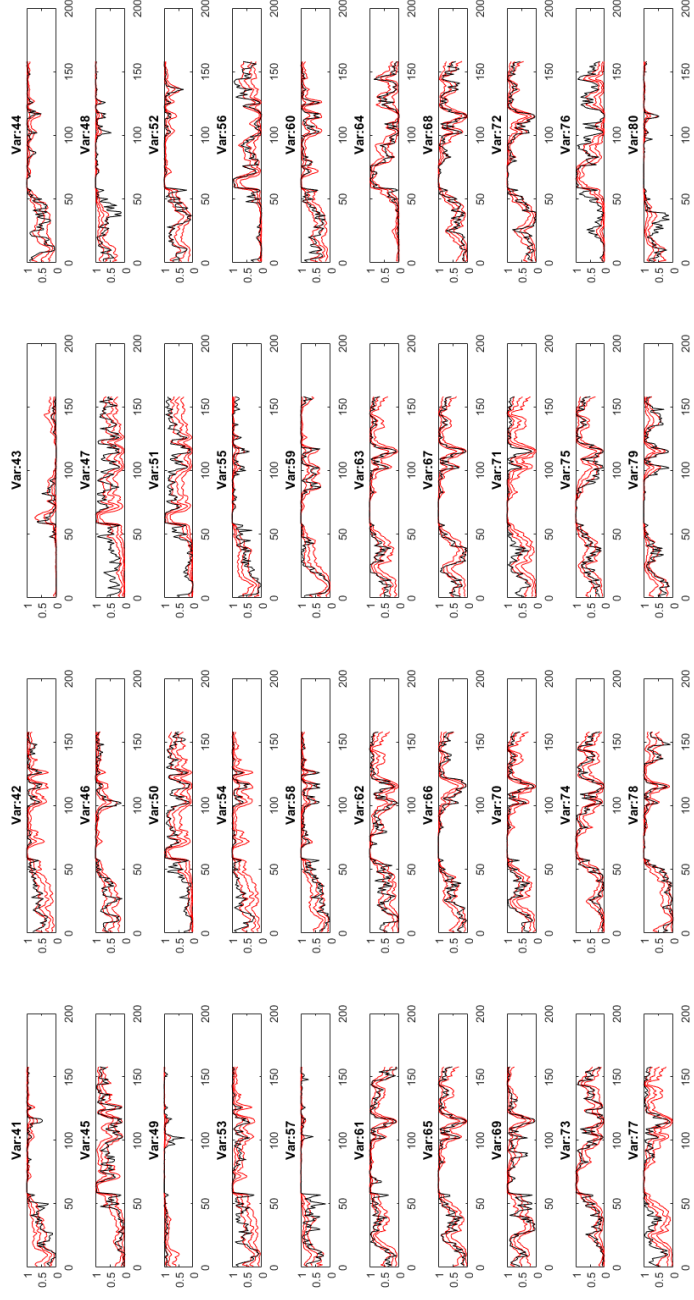


Figure 2: Contribution of World Uncertainty (continued). The black line is the true value. The red lines represent the distribution of the estimate across the Monte-Carlo replications (Median and 1 SD band).



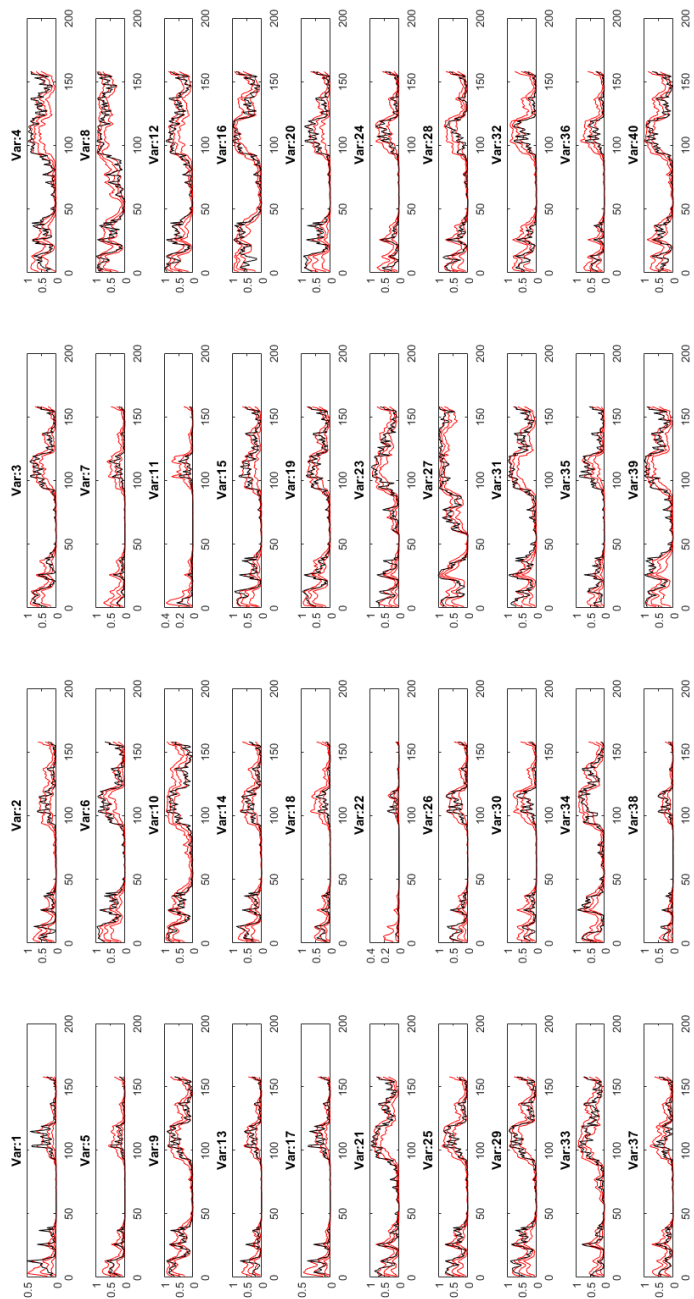


Figure 3: Contribution of Regional Uncertainty. The black line is the true value. The red lines represent the distribution of the estimate across the Monte-Carlo replications (Median and 1 SD band).

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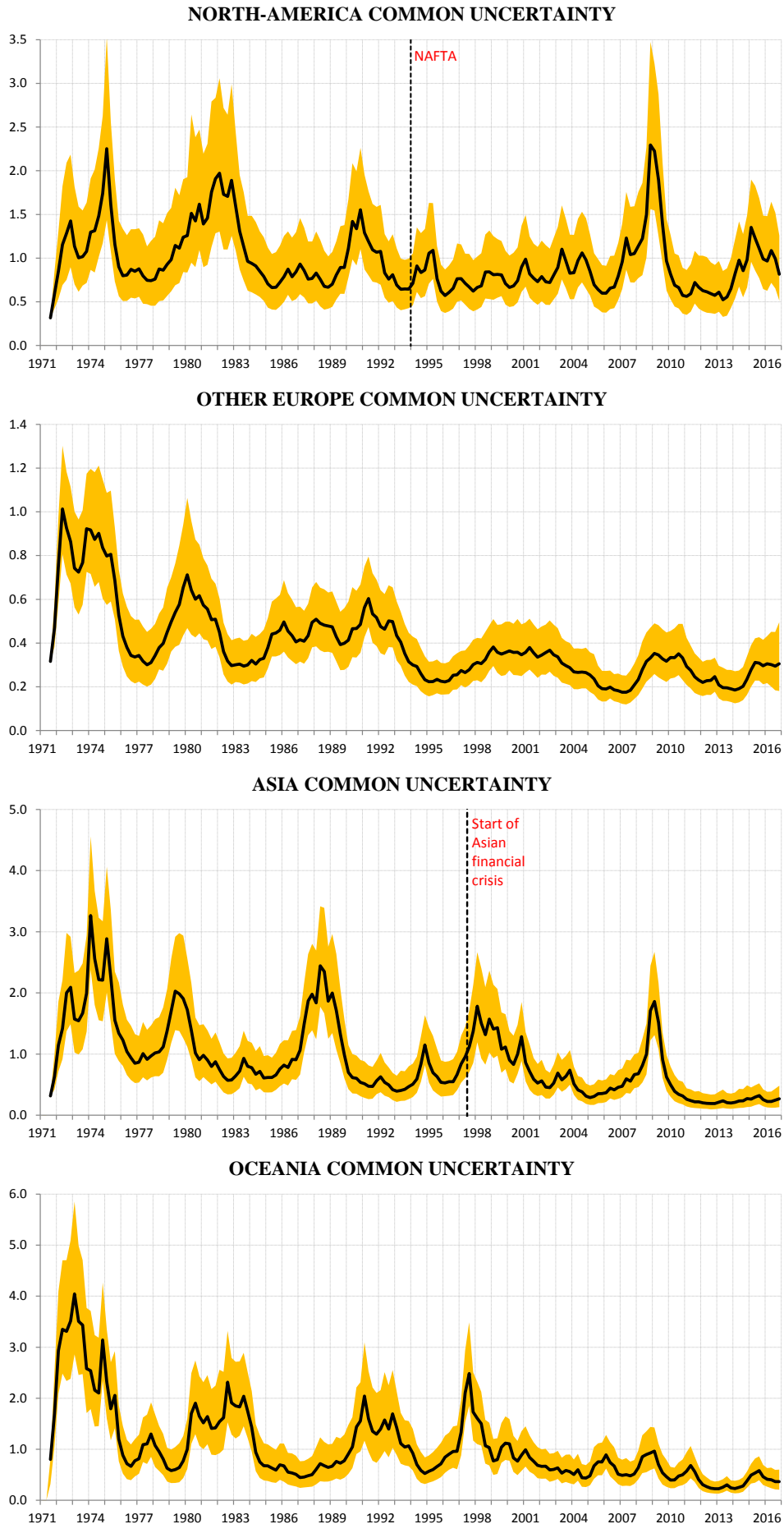
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## Annex II – Data: definitions and sources

Variable	Definition	Source
real GDP	Gross Domestic Product (GDP), volumes	BIS, Eurostat, IMF, OECD
real private consumption	Private final consumption expenditure, volumes	BIS, Eurostat, IMF, OECD
real gross fixed capital formation	Gross fixed capital formation, total, volume	BIS, Eurostat, IMF, OECD
real exports	Exports of goods and services, volume	BIS, Eurostat, IMF, OECD
real imports	Imports of goods and services, volume	BIS, Eurostat, IMF, OECD
employment	Total employment, number of people	BIS, ECB, IMF, OECD
unemployment rate	Unemployment rate, percetn of the labour force	BIS, ECB, IMF, OECD
industrial production	Industrial production, total industry excluding construction, index	ECB, IMF, OECD
retail sales	Sales, total retail trade, volume index	ECB, Fed, IMF, OECD
consumer prices	Consumer prices, index	ECB, OECD
stock prices	Stock prices, index	BIS, ECB, IMF, OECD
house prices	House prices, index	BIS, ECB, IMF, OECD
short-term interest rates	Three-month interest rates (Treasury bonds or 3-month Euribor), percent	BIS, ECB, IMF, OECD
long-term interest rates	Ten-year interest rate (government bond yield)	BIS, ECB, IMF, OECD
private sector credit	Total credit to the private sector, outstanding amounts	BIS, ECB, IMF, OECD
bank loans	Bank loans to the non-financial private sector, outstanding amounts	BIS, ECB, IMF, OECD
narrow money	M1	BIS, ECB, IMF, OECD
broad money	M3 (or M2, or M4)	BIS, ECB, IMF, OECD
nominal effective exchange rate	Nominal effective exchange rate	ECB, IMF
US dollar exchange rate	US dollar exchange rate (or SDRs per US dollar for the US ), average of daily rates	ECB, IMF, OECD
Crude oil, average	Crude oil price, monthly average, nominal US dollars	World Bank
Natural gas	Natural gas price index, monthly average, nominal US dollars	World Bank
Agriculture: Beverages	Agriculture: Beverages, price index, monthly average, nominal US dollars	World Bank
Agriculture: Food	Agriculture: Food, price index, monthly average, nominal US dollars	World Bank
Agriculture: Raw Materials	Agriculture: Raw Materials, price index, monthly average, nominal US dollars	World Bank
Fertilizers	Fertilizers, price index, monthly average, nominal US dollars	World Bank
Metals & Minerals	Metals & Minerals, price index, monthly average, nominal US dollars	World Bank
Precious Metals	Precious Metals, price index, monthly average, nominal US dollars	World Bank

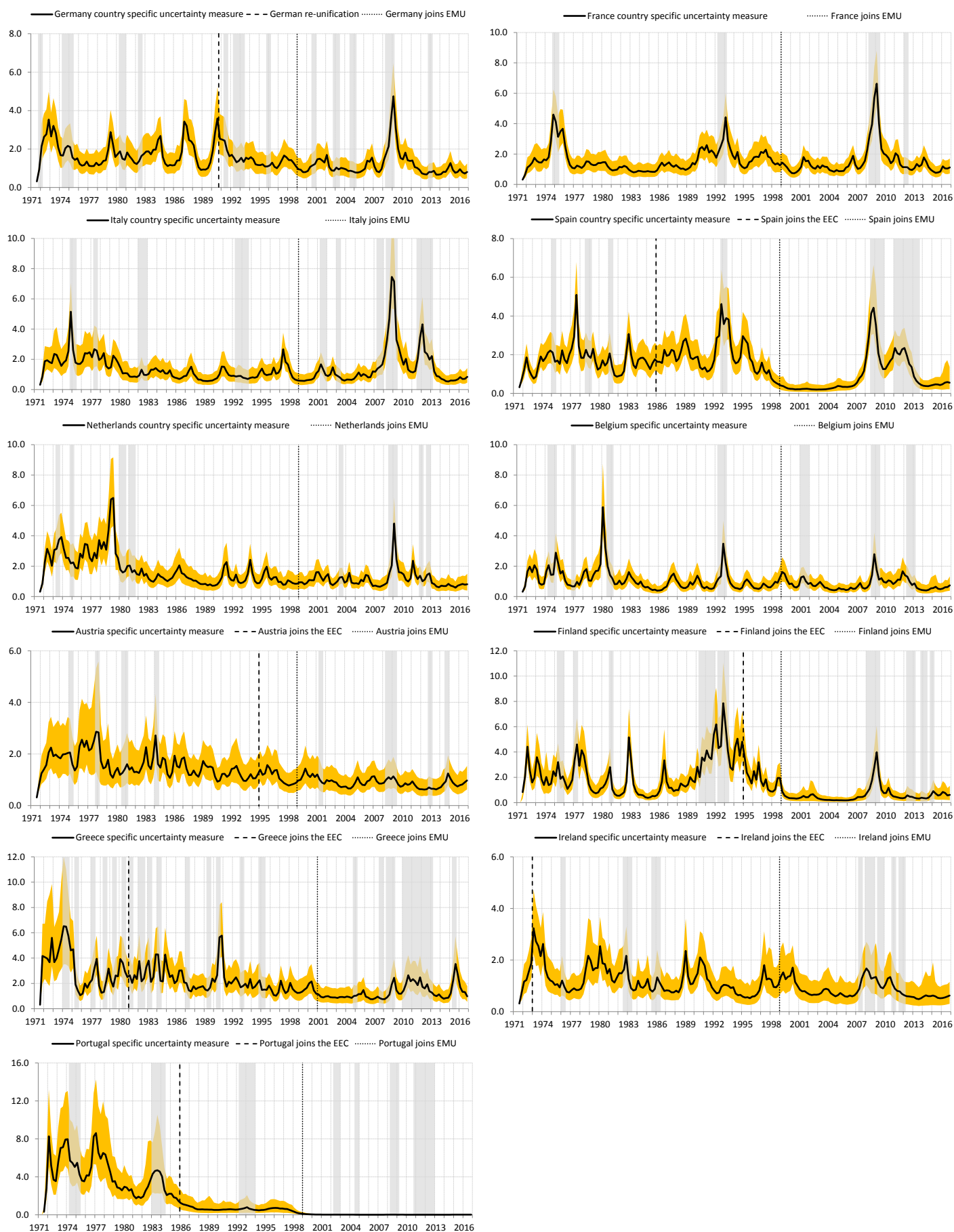
## Annex III – Additional estimates

### Chart A – Other region-specific uncertainty estimates



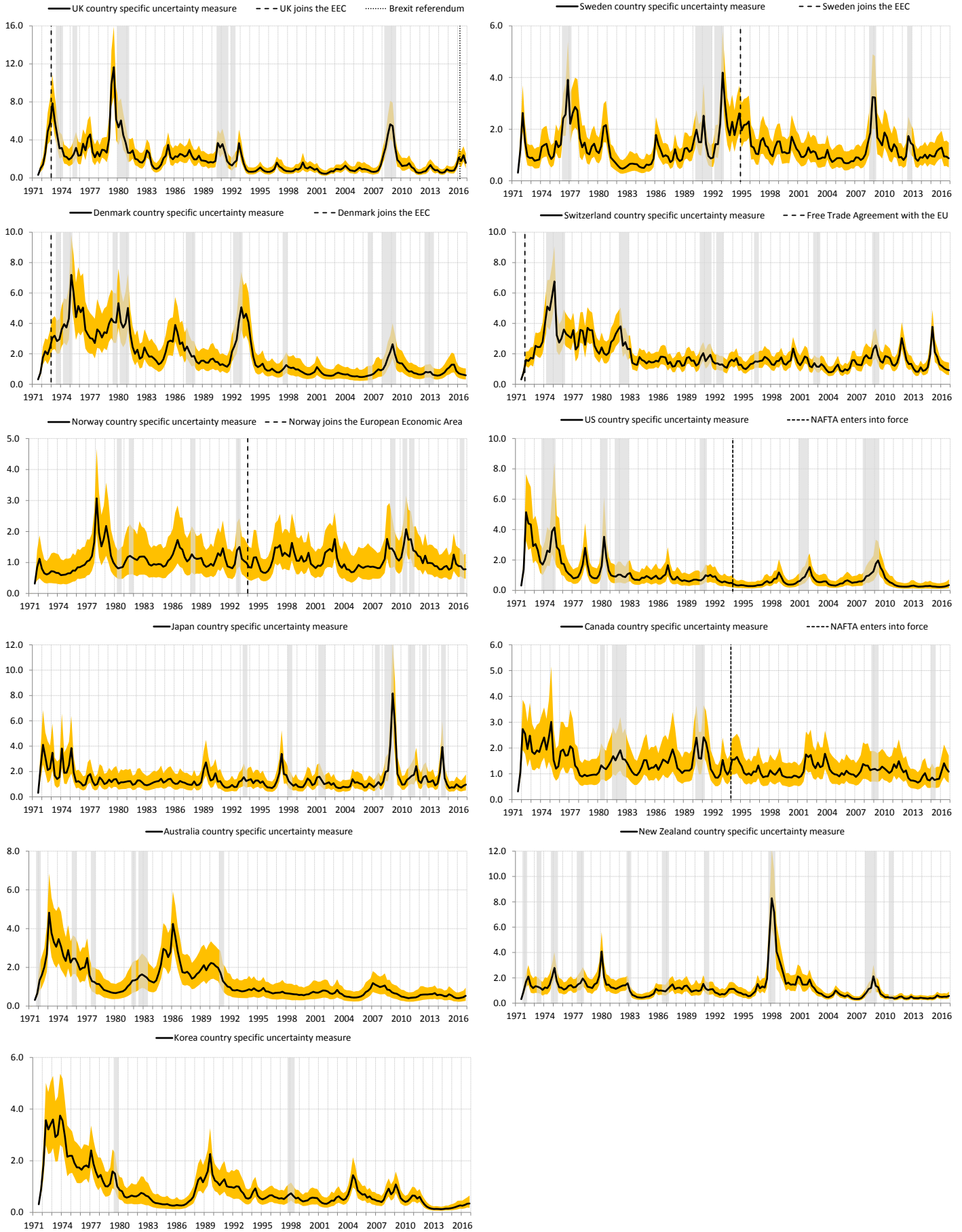
Notes: Estimate of the common standard deviation of shocks to region-specific factors (median and 68 percentile band).

## Chart B – Euro Area country-specific uncertainty estimates



Notes: Estimate of the common standard deviation of shocks to the country factors (median and 68 percentile band). Grey areas delimit recessions as dated according to a “two or more consecutive quarters of negative quarterly real GDP growth” rule.

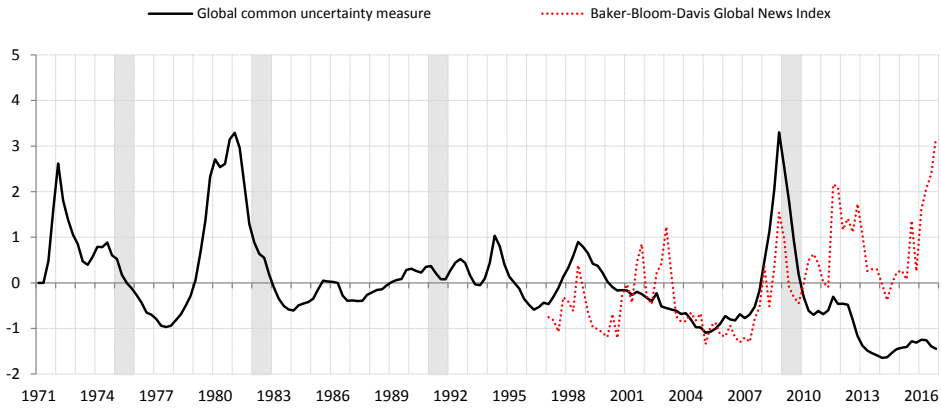
## Chart C – Other country-specific uncertainty estimates



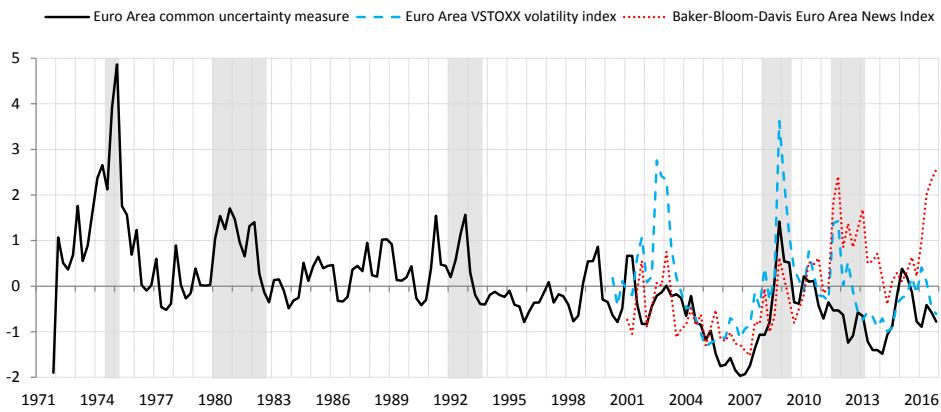
Notes: Estimate of the common standard deviation of shocks to the country factors (median and 68 percentile band). Grey areas delimit recessions as dated according to a “two or more consecutive quarters of negative quarterly real GDP growth” rule, except for the US, for which they are based on the NBER’s Business Cycle Dating Committee.

## Chart D – Alternative uncertainty indicators

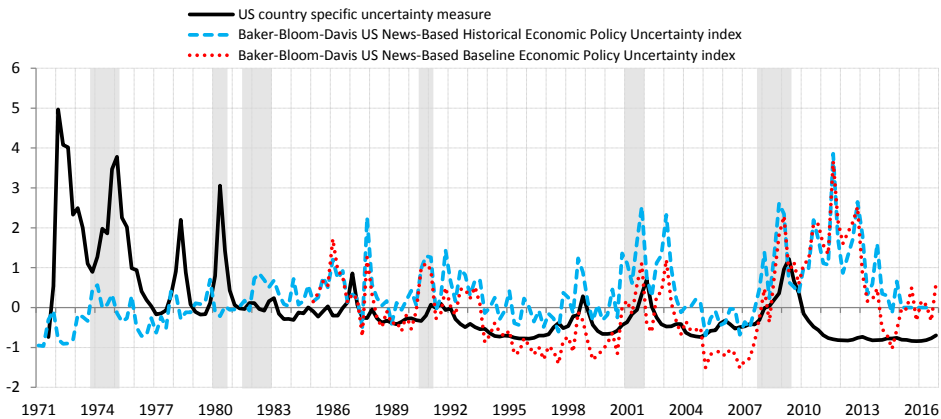
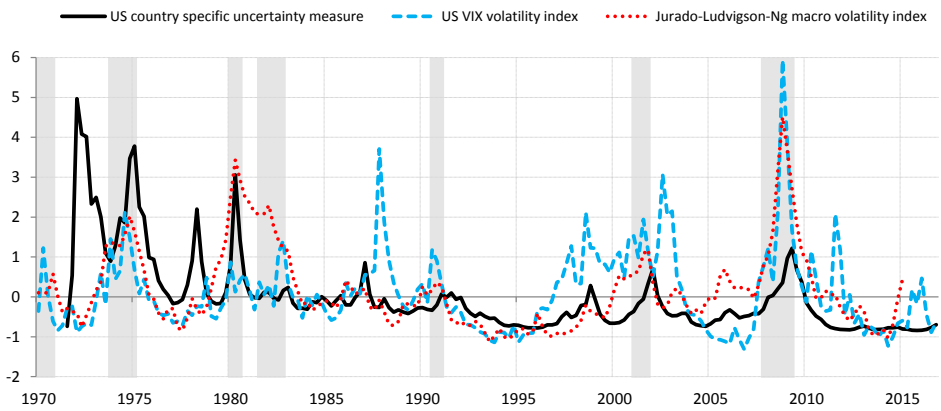
### ALTERNATIVE GLOBAL UNCERTAINTY MEASURES



### ALTERNATIVE EURO AREA COMMON UNCERTAINTY MEASURES



### ALTERNATIVE US UNCERTAINTY MEASURES



Source: Baker, Bloom and Davies (2016), CEPR, ECB, Fed of St Louis FRED-QD, Jurado, Ludvigson, Ng (2015), IMF, NBER and own calculations.

Notes: All indicators normalised. Grey areas delimit global recessions as dated by the IMF (April 2009 World Economic Outlook, Box 1.1. on Global Business Cycles), Euro Area recessions as dated by the CEPR Euro Area Business Cycle Dating Committee, and US recessions as dated by the NBER's Business Cycle Dating Committee.





**Table B – Correlations between uncertainty measures, real GDP growth and CPI inflation**

	correlation uncertainty - real GDP quarterly growth	correlation uncertainty - CPI inflation
Global	<b>-0.23</b>	<b>0.46</b>
Euro Area	<b>-0.14</b>	<b>0.56</b>
Germany	<b>-0.06</b>	<b>0.26</b>
France	<b>-0.47</b>	<b>0.03</b>
Italy	<b>-0.36</b>	<b>0.19</b>
Spain	<b>-0.33</b>	<b>0.37</b>
Netherlands	<b>-0.03</b>	<b>0.52</b>
Belgium	<b>-0.04</b>	<b>0.36</b>
Austria	0.09	<b>0.61</b>
Finland	<b>-0.19</b>	<b>0.20</b>
Greece	<b>-0.03</b>	<b>0.58</b>
Ireland	<b>-0.11</b>	<b>0.39</b>
Portugal	0.19	<b>0.76</b>
UK	<b>-0.10</b>	<b>0.53</b>
Sweden	<b>-0.37</b>	<b>0.12</b>
Denmark	<b>-0.05</b>	<b>0.68</b>
Switzerland	<b>-0.43</b>	<b>0.47</b>
Norway	<b>-0.16</b>	-0.01
US	<b>-0.02</b>	<b>0.45</b>
Canada	<b>-0.06</b>	<b>0.45</b>
Japan	<b>-0.23</b>	<b>0.23</b>
Australia	<b>-0.06</b>	<b>0.63</b>
New Zealand	0.06	<b>0.32</b>
Korea	<b>-0.24</b>	<b>0.28</b>

Source: ECB, Eurostat, OECD, own estimates.

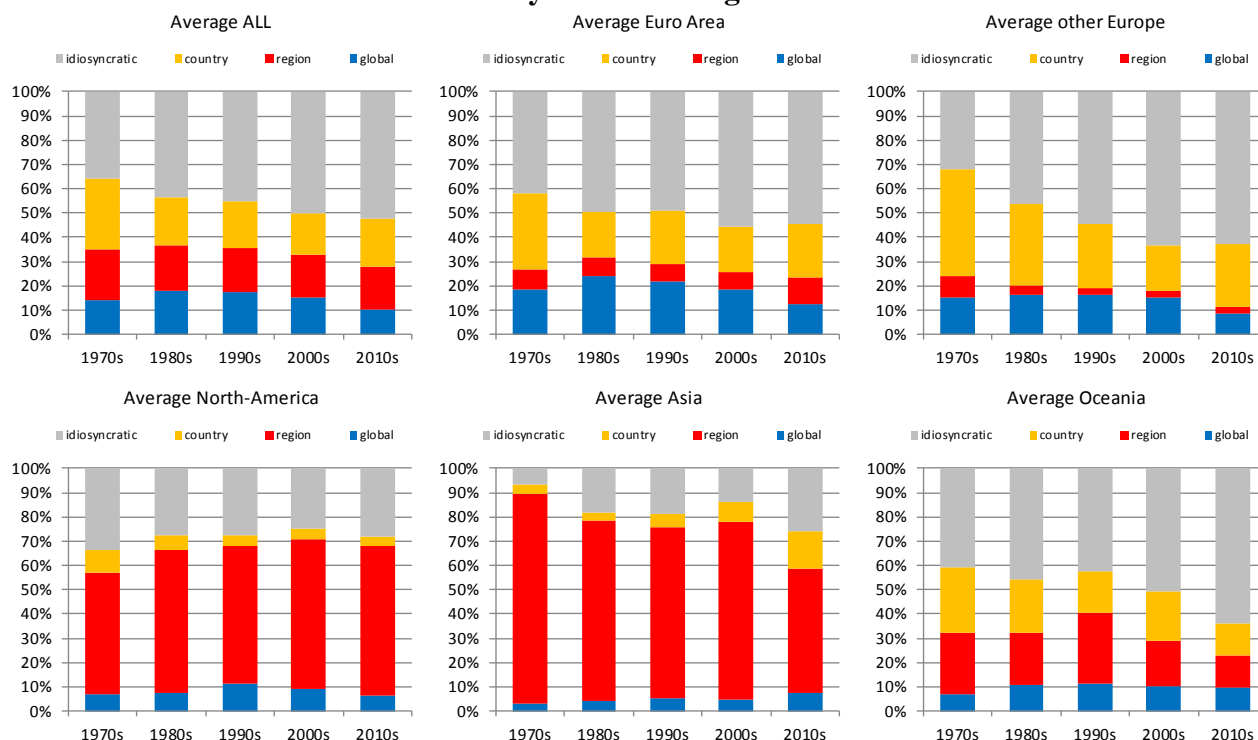
Notes: The table reports the contemporaneous correlation between uncertainty measures and real GDP quarterly growth or CPI inflation over 1971Q3-2016Q4. Global real GDP (global CPI) is represented by the aggregate OECD real GDP (OECD CPI) as computed by the OECD and Euro Area real GDP (Euro Area CPI) is represented by the aggregate Euro Area real GDP (HICP) as reported in the Area Wide Model database. Negative values for the correlations with real GDP growth and positive values for the correlations with CPI inflation are highlighted.

**Table C – Variance decompositions: contributions of uncertainty components to the volatility of real GDP growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	6%	18%	42%	6%	14%	28%	18%	37%	60%	31%
France	9%	22%	46%	10%	21%	38%	11%	25%	43%	32%
Italy	12%	27%	52%	2%	6%	14%	8%	18%	36%	49%
Spain	12%	24%	45%	1%	4%	11%	17%	33%	51%	39%
Netherlands	5%	14%	33%	5%	14%	28%	14%	32%	56%	40%
Belgium	18%	38%	63%	1%	3%	8%	12%	31%	58%	29%
Austria	7%	22%	48%	6%	16%	34%	2%	10%	24%	52%
Finland	10%	21%	42%	1%	2%	6%	7%	14%	26%	63%
Greece	3%	7%	19%	1%	3%	9%	9%	23%	52%	66%
Ireland	1%	7%	22%	0%	2%	7%	3%	12%	32%	79%
Portugal	6%	18%	41%	0%	1%	4%	2%	11%	28%	71%
UK	5%	12%	29%	0%	1%	3%	17%	33%	54%	54%
Sweden	6%	14%	31%	0%	1%	4%	19%	42%	69%	43%
Denmark	4%	10%	26%	1%	3%	9%	24%	37%	50%	49%
Switzerland	10%	25%	51%	4%	11%	26%	8%	19%	36%	45%
Norway	3%	11%	31%	2%	5%	10%	4%	16%	41%	67%
US	3%	10%	27%	30%	60%	84%	1%	3%	14%	27%
Canada	2%	7%	21%	28%	55%	79%	2%	8%	22%	30%
Japan	2%	6%	17%	34%	58%	79%	4%	13%	30%	23%
Australia	2%	6%	20%	8%	21%	42%	3%	10%	27%	63%
New Zealand	4%	14%	36%	7%	24%	53%	12%	30%	56%	32%
Korea	1%	4%	13%	64%	86%	96%	0%	1%	5%	9%
Av. Euro Area	8%	20%	41%	3%	8%	17%	10%	22%	42%	50%
Av. other Europe	5%	15%	34%	1%	4%	10%	14%	29%	50%	52%
Av. North-America	3%	8%	24%	29%	57%	81%	1%	6%	18%	29%
Av. Asia	1%	5%	15%	49%	72%	88%	2%	7%	18%	16%
Av. Oceania	3%	10%	28%	7%	22%	47%	7%	20%	41%	47%
Average ALL	6%	15%	34%	10%	19%	30%	9%	21%	40%	45%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of real GDP growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart E – Variance decompositions: contributions of uncertainty components to the volatility of real GDP growth over time**



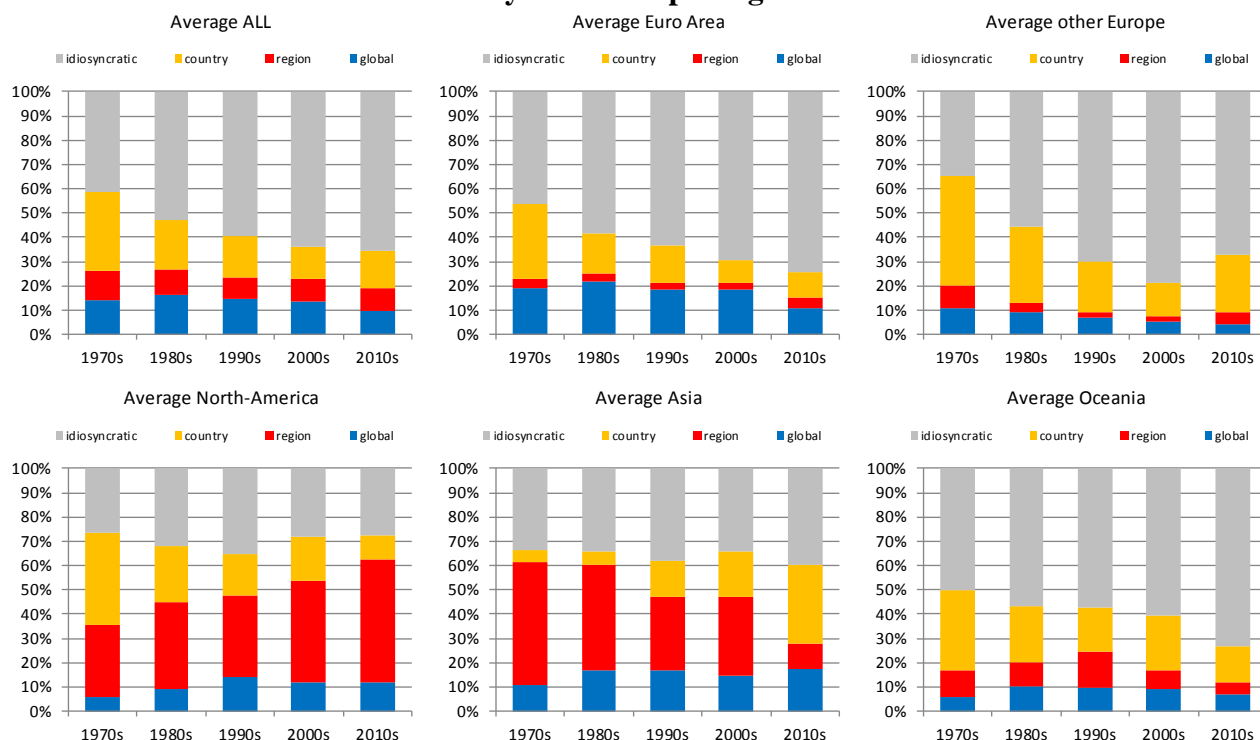
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of real GDP growth. Idiosyncratic contribution derived as residual. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table D – Variance decompositions: contributions of uncertainty components to the volatility of consumption growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	3%	10%	28%	2%	5%	12%	7%	18%	36%	67%
France	6%	16%	39%	2%	7%	18%	6%	16%	33%	61%
Italy	9%	23%	48%	1%	4%	10%	2%	5%	13%	69%
Spain	9%	19%	37%	1%	2%	5%	18%	33%	52%	46%
Netherlands	8%	18%	38%	1%	2%	6%	10%	22%	41%	58%
Belgium	15%	33%	59%	0%	1%	5%	4%	12%	27%	54%
Austria	5%	14%	32%	1%	3%	6%	3%	10%	23%	73%
Finland	6%	17%	38%	1%	4%	13%	10%	21%	39%	58%
Greece	6%	19%	44%	0%	2%	6%	1%	3%	16%	76%
Ireland	3%	11%	30%	1%	3%	10%	6%	19%	44%	67%
Portugal	6%	20%	46%	1%	3%	12%	11%	23%	40%	53%
UK	1%	5%	14%	1%	4%	10%	24%	43%	64%	49%
Sweden	4%	10%	22%	1%	3%	7%	10%	25%	49%	62%
Denmark	1%	5%	15%	1%	4%	11%	18%	31%	45%	61%
Switzerland	5%	13%	33%	1%	4%	12%	11%	23%	40%	59%
Norway	1%	4%	13%	2%	6%	12%	4%	12%	29%	79%
US	3%	12%	31%	11%	31%	61%	9%	25%	52%	32%
Canada	3%	9%	24%	18%	44%	73%	6%	18%	41%	28%
Japan	1%	4%	14%	27%	52%	75%	8%	23%	46%	22%
Australia	2%	6%	17%	3%	8%	18%	9%	20%	37%	66%
New Zealand	3%	11%	29%	4%	12%	28%	10%	26%	50%	51%
Korea	10%	26%	54%	5%	18%	47%	2%	6%	17%	49%
Av. Euro Area	7%	18%	40%	1%	3%	9%	7%	17%	33%	62%
Av. other Europe	3%	7%	19%	1%	4%	10%	14%	27%	45%	62%
Av. North-America	3%	11%	28%	15%	37%	67%	8%	22%	46%	30%
Av. Asia	5%	15%	34%	16%	35%	61%	5%	14%	32%	35%
Av. Oceania	3%	8%	23%	4%	10%	23%	9%	23%	44%	59%
Average ALL	5%	14%	32%	4%	10%	21%	9%	20%	38%	56%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of consumption growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart F – Variance decompositions: contributions of uncertainty components to the volatility of consumption growth over time**



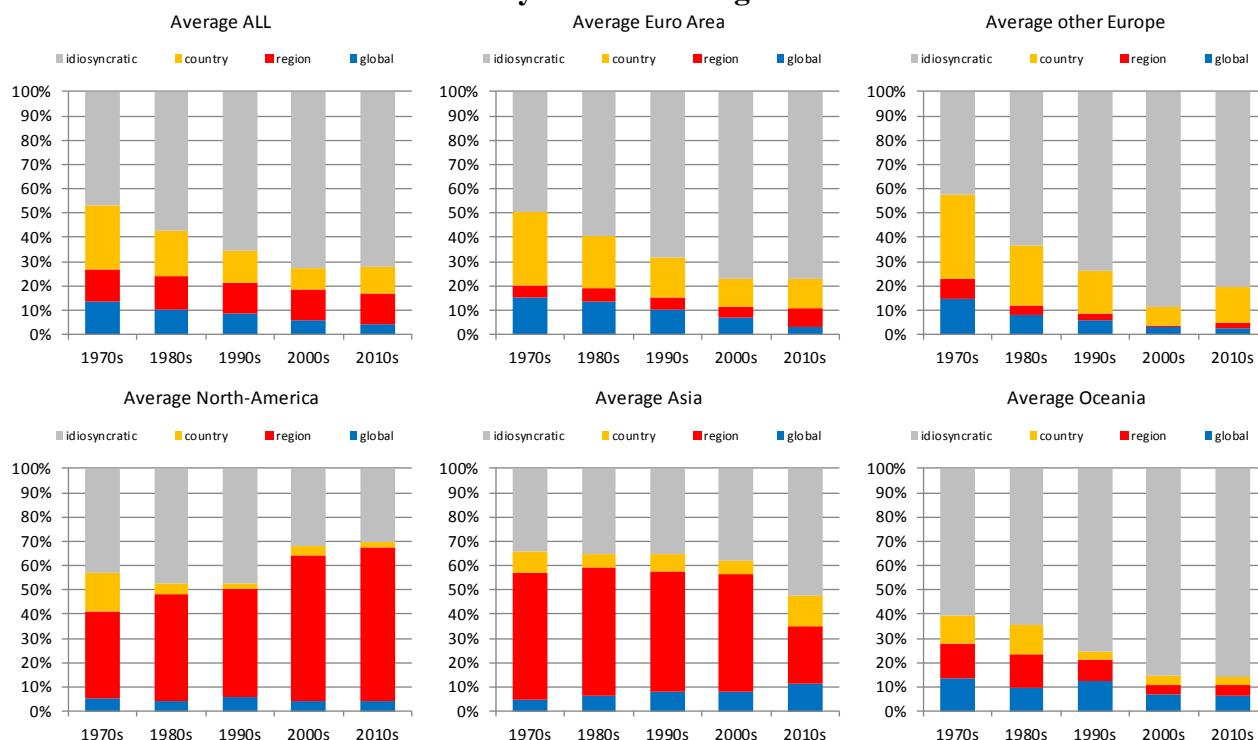
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of consumption growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table E – Variance decompositions: contributions of uncertainty components to the volatility of investment growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	1%	5%	17%	6%	14%	28%	11%	23%	43%	57%
France	7%	17%	37%	5%	13%	28%	4%	9%	20%	61%
Italy	2%	7%	20%	2%	5%	12%	10%	22%	41%	66%
Spain	7%	16%	33%	1%	3%	9%	15%	29%	48%	53%
Netherlands	1%	4%	13%	4%	9%	19%	13%	29%	50%	58%
Belgium	4%	14%	36%	1%	4%	12%	5%	19%	47%	63%
Austria	3%	9%	22%	1%	2%	5%	3%	10%	22%	79%
Finland	2%	7%	21%	0%	1%	5%	5%	11%	25%	80%
Greece	4%	12%	29%	0%	2%	6%	2%	10%	37%	76%
Ireland	2%	7%	22%	0%	2%	5%	13%	29%	46%	62%
Portugal	5%	14%	32%	0%	1%	4%	3%	13%	29%	72%
UK	2%	8%	22%	1%	3%	8%	4%	11%	27%	79%
Sweden	3%	8%	19%	2%	5%	14%	13%	31%	59%	55%
Denmark	2%	6%	20%	1%	3%	7%	16%	28%	43%	63%
Switzerland	3%	9%	24%	1%	3%	7%	6%	14%	28%	74%
Norway	1%	3%	12%	1%	3%	7%	4%	13%	31%	80%
US	1%	5%	15%	28%	56%	80%	3%	9%	25%	30%
Canada	1%	5%	15%	15%	41%	74%	1%	2%	9%	52%
Japan	2%	6%	18%	27%	50%	74%	2%	8%	21%	36%
Australia	4%	15%	43%	2%	6%	16%	5%	13%	29%	66%
New Zealand	1%	5%	15%	4%	12%	27%	0%	1%	3%	82%
Korea	3%	9%	27%	20%	44%	71%	2%	7%	20%	40%
Av. Euro Area	3%	10%	26%	2%	5%	12%	8%	19%	37%	66%
Av. other Europe	2%	7%	19%	1%	3%	9%	9%	20%	37%	70%
Av. North-America	1%	5%	15%	22%	48%	77%	2%	6%	17%	41%
Av. Asia	2%	7%	22%	24%	47%	72%	2%	7%	20%	38%
Av. Oceania	3%	10%	29%	3%	9%	22%	2%	7%	16%	74%
Average ALL	3%	9%	23%	6%	13%	24%	6%	16%	32%	63%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of investment growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart G – Variance decompositions: contributions of uncertainty components to the volatility of investment growth over time**



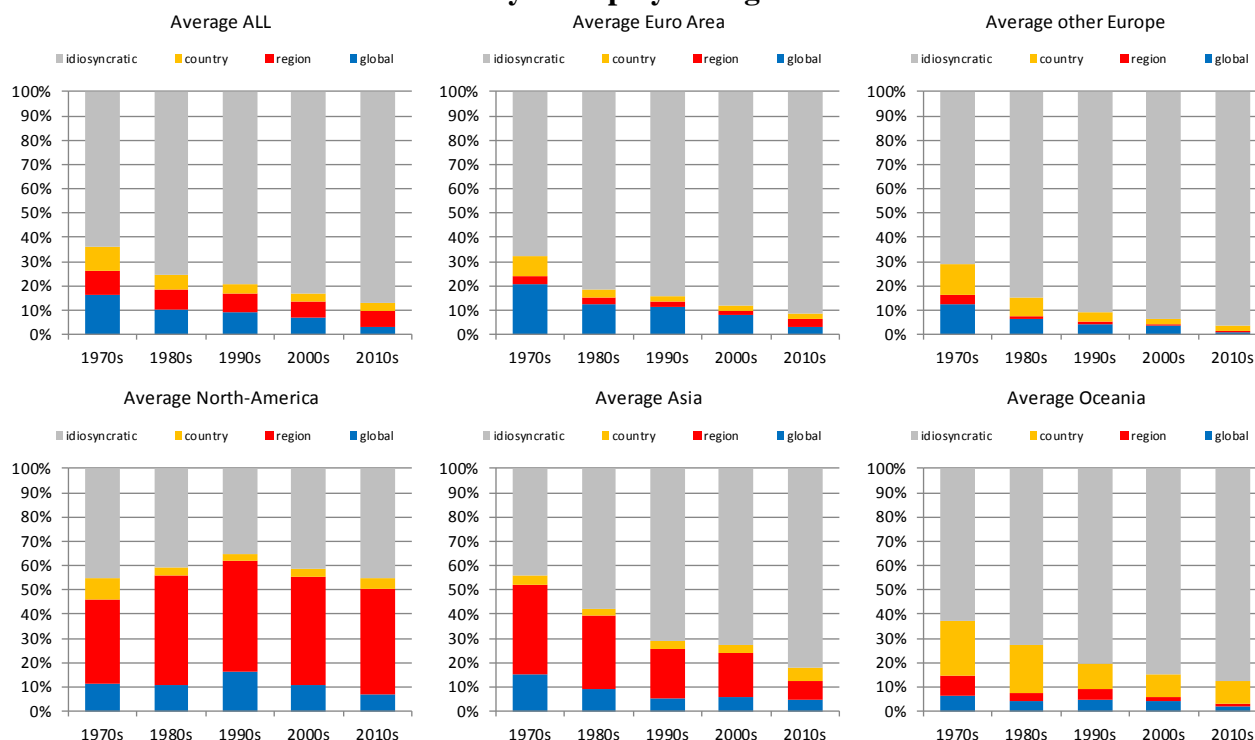
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of investment growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table F – Variance decompositions: contributions of uncertainty components to the volatility of employment growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	6%	20%	45%	0%	2%	7%	1%	4%	13%	74%
France	4%	12%	31%	1%	2%	8%	0%	1%	3%	85%
Italy	4%	12%	31%	1%	5%	16%	2%	6%	16%	77%
Spain	2%	5%	13%	0%	0%	1%	0%	0%	0%	95%
Netherlands	2%	7%	20%	0%	2%	5%	0%	1%	4%	90%
Belgium	6%	20%	49%	1%	5%	15%	1%	2%	8%	73%
Austria	2%	5%	13%	2%	5%	11%	0%	1%	5%	89%
Finland	1%	3%	12%	1%	3%	9%	1%	5%	13%	89%
Greece	5%	13%	29%	0%	1%	3%	1%	3%	14%	83%
Ireland	6%	19%	44%	0%	1%	5%	2%	10%	37%	69%
Portugal	5%	12%	28%	1%	2%	5%	2%	5%	13%	81%
UK	4%	12%	30%	0%	1%	4%	0%	1%	6%	86%
Sweden	1%	4%	12%	1%	2%	6%	2%	8%	25%	86%
Denmark	3%	8%	21%	1%	3%	7%	8%	19%	33%	71%
Switzerland	1%	3%	10%	0%	0%	1%	0%	0%	1%	96%
Norway	1%	2%	8%	0%	0%	1%	0%	1%	5%	96%
US	2%	9%	26%	23%	47%	72%	1%	4%	13%	41%
Canada	5%	15%	37%	16%	38%	66%	1%	5%	16%	42%
Japan	1%	5%	20%	17%	38%	65%	1%	4%	14%	53%
Australia	1%	5%	18%	1%	4%	13%	11%	29%	56%	62%
New Zealand	1%	3%	10%	1%	3%	9%	0%	0%	1%	93%
Korea	3%	11%	31%	2%	9%	27%	1%	3%	12%	77%
Av. Euro Area	4%	12%	29%	1%	3%	8%	1%	3%	11%	82%
Av. other Europe	2%	6%	17%	0%	1%	4%	2%	6%	14%	87%
Av. North-America	4%	12%	32%	19%	42%	69%	1%	4%	14%	41%
Av. Asia	2%	8%	25%	9%	24%	46%	1%	4%	13%	65%
Av. Oceania	1%	4%	14%	1%	4%	11%	5%	15%	29%	77%
Average ALL	3%	9%	25%	3%	8%	16%	2%	5%	14%	78%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of employment growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart H – Variance decompositions: contributions of uncertainty components to the volatility of employment growth over time**



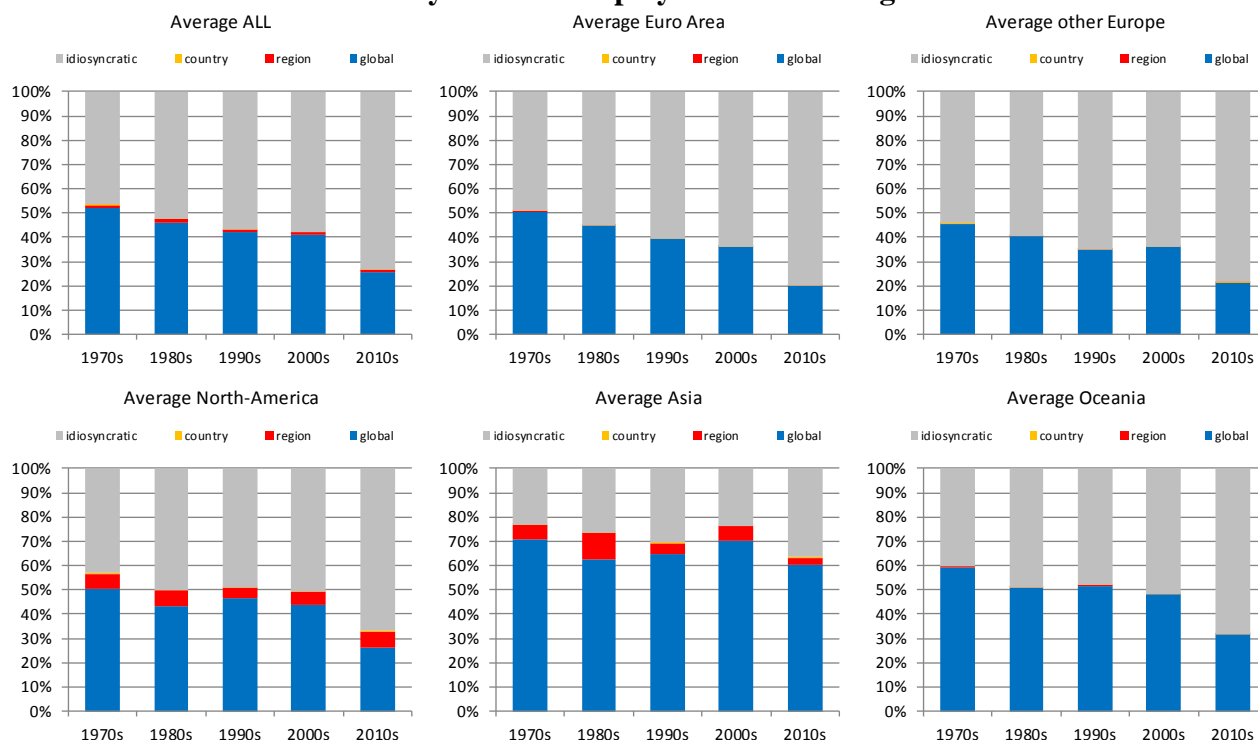
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of employment growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table G – Variance decompositions: contributions of uncertainty components to the volatility of the unemployment rate changes**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	11%	30%	55%	0%	0%	0%	0%	0%	0%	70%
France	12%	33%	60%	0%	0%	0%	0%	0%	0%	67%
Italy	8%	22%	47%	0%	0%	0%	0%	0%	0%	78%
Spain	15%	38%	64%	0%	0%	0%	0%	0%	0%	62%
Netherlands	39%	62%	83%	0%	0%	0%	0%	0%	0%	37%
Belgium	42%	63%	82%	0%	0%	0%	0%	0%	0%	36%
Austria	60%	77%	90%	0%	0%	0%	0%	0%	0%	22%
Finland	22%	41%	64%	0%	0%	0%	0%	0%	0%	59%
Greece	12%	26%	45%	0%	0%	0%	0%	0%	1%	73%
Ireland	8%	22%	45%	0%	0%	0%	0%	0%	0%	78%
Portugal	8%	20%	39%	0%	0%	1%	0%	1%	3%	79%
UK	12%	29%	53%	0%	0%	0%	0%	0%	0%	71%
Sweden	13%	32%	61%	0%	0%	0%	0%	0%	2%	67%
Denmark	39%	61%	81%	0%	0%	1%	0%	0%	1%	39%
Switzerland	10%	26%	52%	0%	0%	0%	0%	0%	0%	74%
Norway	15%	35%	62%	0%	0%	0%	0%	0%	1%	65%
US	13%	32%	61%	2%	9%	26%	0%	0%	2%	59%
Canada	29%	54%	78%	1%	3%	10%	0%	0%	1%	43%
Japan	33%	58%	81%	0%	1%	5%	0%	0%	0%	41%
Australia	36%	61%	82%	0%	0%	0%	0%	0%	1%	38%
New Zealand	18%	38%	63%	0%	0%	0%	0%	0%	0%	62%
Korea	45%	74%	92%	2%	11%	34%	0%	0%	2%	14%
Av. Euro Area	22%	40%	61%	0%	0%	0%	0%	0%	1%	60%
Av. other Europe	18%	37%	62%	0%	0%	0%	0%	0%	1%	63%
Av. North-America	21%	43%	70%	1%	6%	18%	0%	0%	1%	51%
Av. Asia	39%	66%	86%	1%	6%	20%	0%	0%	1%	27%
Av. Oceania	27%	50%	72%	0%	0%	0%	0%	0%	0%	50%
Average ALL	23%	43%	65%	0%	1%	4%	0%	0%	1%	56%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of the unemployment rate changes over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart I – Variance decompositions: contributions of uncertainty components to the volatility of the unemployment rate changes over time**



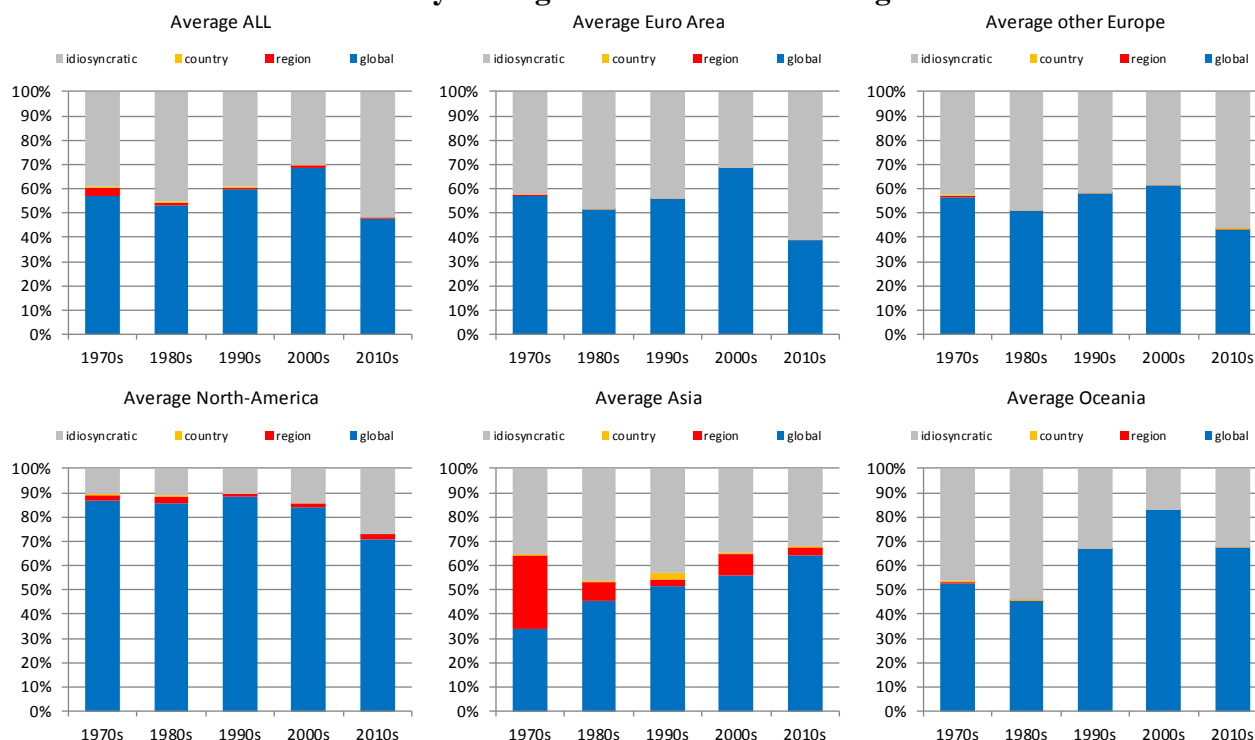
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of the unemployment rate changes. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table H – Variance decompositions: contributions of uncertainty components to the volatility of long-term interest rate changes**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	60%	79%	91%	0%	0%	0%	0%	0%	0%	20%
France	77%	89%	96%	0%	0%	0%	0%	0%	0%	11%
Italy	24%	45%	70%	0%	0%	0%	0%	0%	0%	54%
Spain	16%	38%	64%	0%	0%	0%	0%	0%	0%	62%
Netherlands	74%	87%	95%	0%	0%	0%	0%	0%	0%	13%
Belgium	59%	79%	92%	0%	0%	0%	0%	0%	0%	21%
Austria	3%	10%	26%	0%	0%	0%	0%	0%	0%	90%
Finland	4%	16%	44%	0%	0%	0%	0%	0%	0%	84%
Greece	21%	45%	72%	0%	0%	1%	0%	1%	8%	54%
Ireland	68%	84%	94%	0%	0%	0%	0%	0%	0%	16%
Portugal	19%	38%	60%	0%	0%	0%	0%	0%	2%	62%
UK	73%	87%	95%	0%	0%	1%	0%	0%	1%	12%
Sweden	34%	62%	84%	0%	0%	1%	0%	0%	1%	38%
Denmark	11%	33%	62%	0%	0%	0%	0%	0%	0%	67%
Switzerland	50%	72%	88%	0%	0%	1%	0%	0%	0%	28%
Norway	7%	21%	47%	0%	0%	0%	0%	0%	0%	79%
US	67%	83%	93%	1%	2%	8%	0%	0%	1%	14%
Canada	68%	85%	95%	0%	2%	7%	0%	0%	0%	13%
Japan	9%	28%	56%	1%	3%	12%	0%	0%	0%	69%
Australia	44%	72%	92%	0%	0%	1%	0%	0%	1%	28%
New Zealand	32%	54%	76%	0%	0%	1%	0%	0%	0%	45%
Korea	42%	71%	90%	6%	18%	44%	0%	2%	6%	9%
Av. Euro Area	39%	56%	73%	0%	0%	0%	0%	0%	1%	44%
Av. other Europe	35%	55%	75%	0%	0%	1%	0%	0%	0%	45%
Av. North-America	68%	84%	94%	0%	2%	8%	0%	0%	1%	14%
Av. Asia	26%	50%	73%	3%	11%	28%	0%	1%	3%	39%
Av. Oceania	38%	63%	84%	0%	0%	1%	0%	0%	0%	37%
Average ALL	39%	58%	76%	0%	1%	4%	0%	0%	1%	40%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of long-term interest rate changes over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart J – Variance decompositions: contributions of uncertainty components to the volatility of long-term interest rate changes over time**



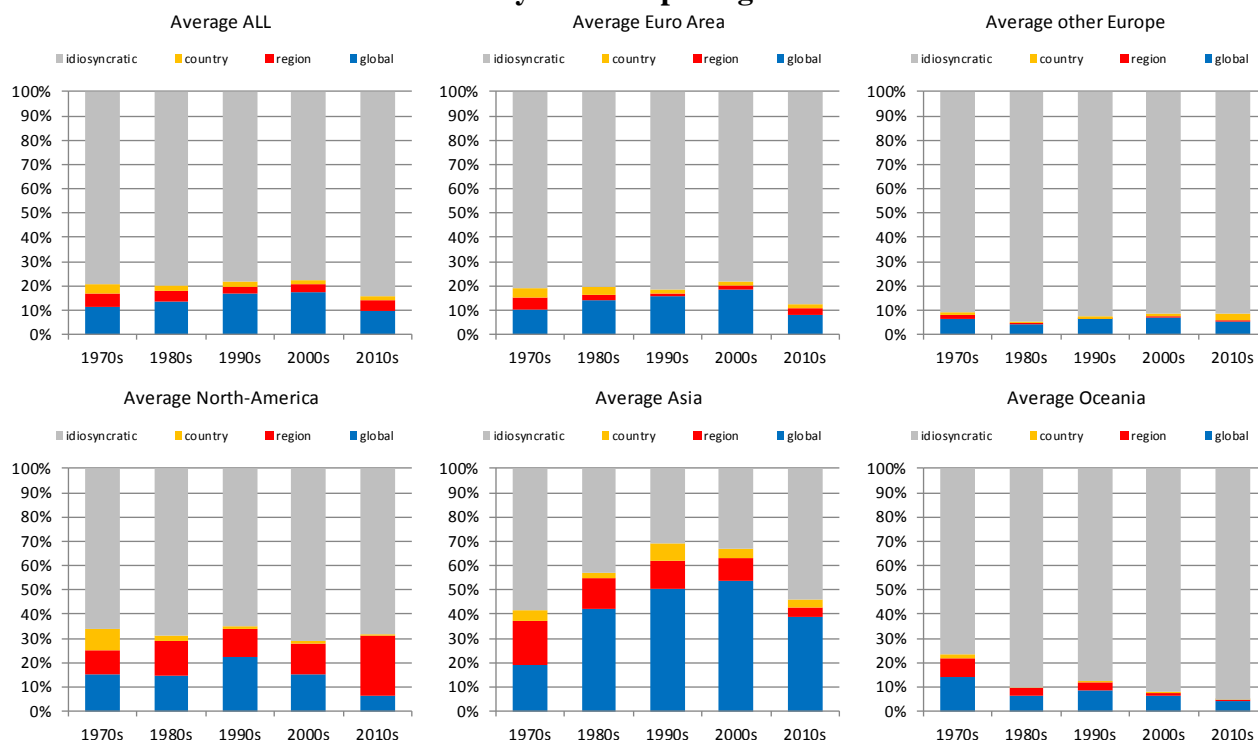
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of long-term interest rate changes. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table I – Variance decompositions: contributions of uncertainty components to the volatility of house price growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	15%	32%	53%	0%	0%	2%	0%	1%	4%	67%
France	3%	9%	22%	0%	0%	1%	0%	0%	0%	91%
Italy	6%	21%	47%	0%	2%	5%	0%	0%	2%	77%
Spain	3%	9%	21%	0%	1%	3%	0%	0%	1%	90%
Netherlands	4%	13%	32%	1%	3%	9%	0%	0%	1%	84%
Belgium	3%	9%	24%	1%	4%	10%	0%	1%	3%	86%
Austria	2%	6%	16%	4%	10%	20%	1%	4%	11%	80%
Finland	3%	8%	20%	0%	1%	4%	0%	0%	1%	90%
Greece	5%	15%	37%	0%	2%	6%	0%	2%	18%	81%
Ireland	2%	6%	16%	0%	1%	3%	2%	8%	24%	85%
Portugal	9%	26%	52%	1%	4%	12%	2%	8%	24%	63%
UK	3%	8%	20%	0%	0%	1%	0%	1%	4%	91%
Sweden	2%	5%	12%	1%	1%	4%	0%	1%	5%	92%
Denmark	2%	6%	14%	0%	0%	1%	0%	0%	1%	94%
Switzerland	2%	6%	15%	0%	1%	2%	1%	3%	7%	91%
Norway	2%	5%	13%	0%	1%	2%	0%	1%	2%	94%
US	6%	28%	68%	3%	14%	41%	1%	5%	20%	52%
Canada	1%	3%	8%	4%	13%	31%	0%	0%	2%	84%
Japan	22%	49%	72%	1%	6%	21%	0%	0%	1%	46%
Australia	3%	8%	21%	0%	1%	4%	0%	0%	2%	90%
New Zealand	3%	8%	22%	2%	5%	14%	0%	1%	2%	86%
Korea	15%	35%	61%	5%	18%	45%	2%	8%	23%	39%
Av. Euro Area	5%	14%	31%	1%	2%	7%	1%	2%	8%	81%
Av. other Europe	2%	6%	15%	0%	1%	2%	0%	1%	4%	92%
Av. North-America	3%	15%	38%	4%	14%	36%	1%	3%	11%	68%
Av. Asia	19%	42%	67%	3%	12%	33%	1%	4%	12%	42%
Av. Oceania	3%	8%	21%	1%	3%	9%	0%	1%	2%	88%
Average ALL	5%	14%	30%	1%	4%	11%	1%	2%	7%	80%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of house price growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart K – Variance decompositions: contributions of uncertainty components to the volatility of house price growth over time**



Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of house price growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

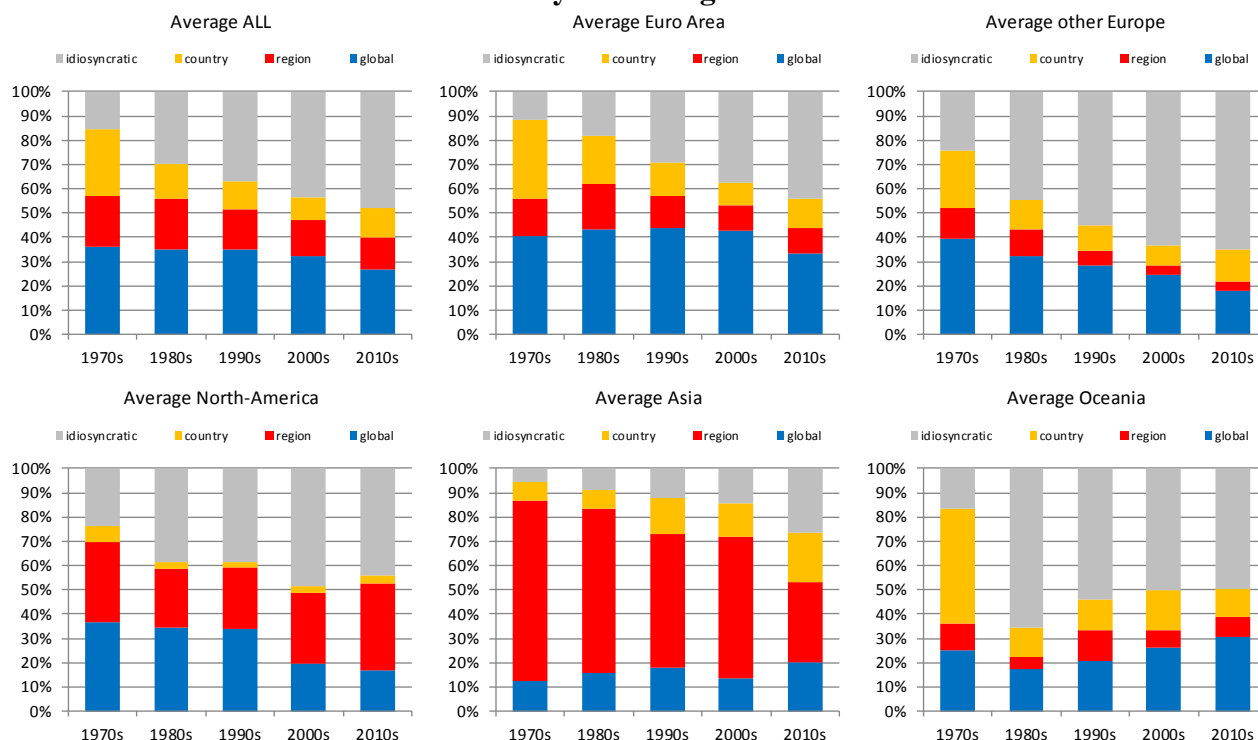


**Table J – Variance decompositions: contributions of uncertainty components to the volatility of credit growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	24%	45%	69%	10%	22%	39%	3%	9%	23%	24%
France	41%	60%	79%	2%	5%	13%	1%	2%	7%	32%
Italy	27%	49%	74%	17%	35%	57%	2%	6%	15%	10%
Spain	32%	54%	76%	14%	32%	52%	1%	4%	11%	11%
Netherlands	20%	37%	62%	1%	3%	9%	7%	15%	26%	45%
Belgium	18%	36%	60%	5%	13%	28%	7%	19%	38%	32%
Austria	32%	54%	76%	3%	6%	14%	9%	24%	45%	16%
Finland	26%	46%	70%	3%	8%	20%	1%	3%	10%	43%
Greece	5%	22%	51%	0%	3%	11%	24%	59%	92%	16%
Ireland	8%	21%	47%	9%	19%	35%	8%	23%	47%	37%
Portugal	17%	33%	56%	2%	5%	11%	19%	30%	44%	32%
UK	6%	19%	46%	7%	18%	36%	3%	9%	24%	54%
Sweden	7%	20%	43%	8%	19%	35%	6%	16%	36%	45%
Denmark	17%	30%	47%	0%	1%	4%	0%	0%	2%	68%
Switzerland	35%	57%	79%	0%	0%	1%	19%	39%	61%	3%
Norway	11%	20%	37%	0%	0%	1%	0%	1%	3%	79%
US	7%	21%	51%	13%	40%	73%	1%	4%	19%	34%
Canada	17%	37%	63%	5%	17%	44%	1%	2%	10%	43%
Japan	7%	21%	50%	24%	54%	80%	3%	10%	27%	15%
Australia	12%	25%	48%	1%	3%	8%	4%	11%	24%	61%
New Zealand	10%	22%	42%	6%	15%	31%	14%	28%	47%	35%
Korea	3%	11%	32%	36%	64%	86%	5%	15%	35%	10%
Av. Euro Area	23%	41%	65%	6%	14%	26%	8%	18%	33%	27%
Av. other Europe	15%	29%	50%	3%	8%	15%	6%	13%	25%	50%
Av. North-America	12%	29%	57%	9%	29%	59%	1%	3%	15%	39%
Av. Asia	5%	16%	41%	30%	59%	83%	4%	12%	31%	13%
Av. Oceania	11%	23%	45%	3%	9%	20%	9%	20%	35%	48%
Average ALL	17%	34%	57%	7%	17%	31%	6%	15%	29%	34%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of credit growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart L – Variance decompositions: contributions of uncertainty components to the volatility of credit growth over time**



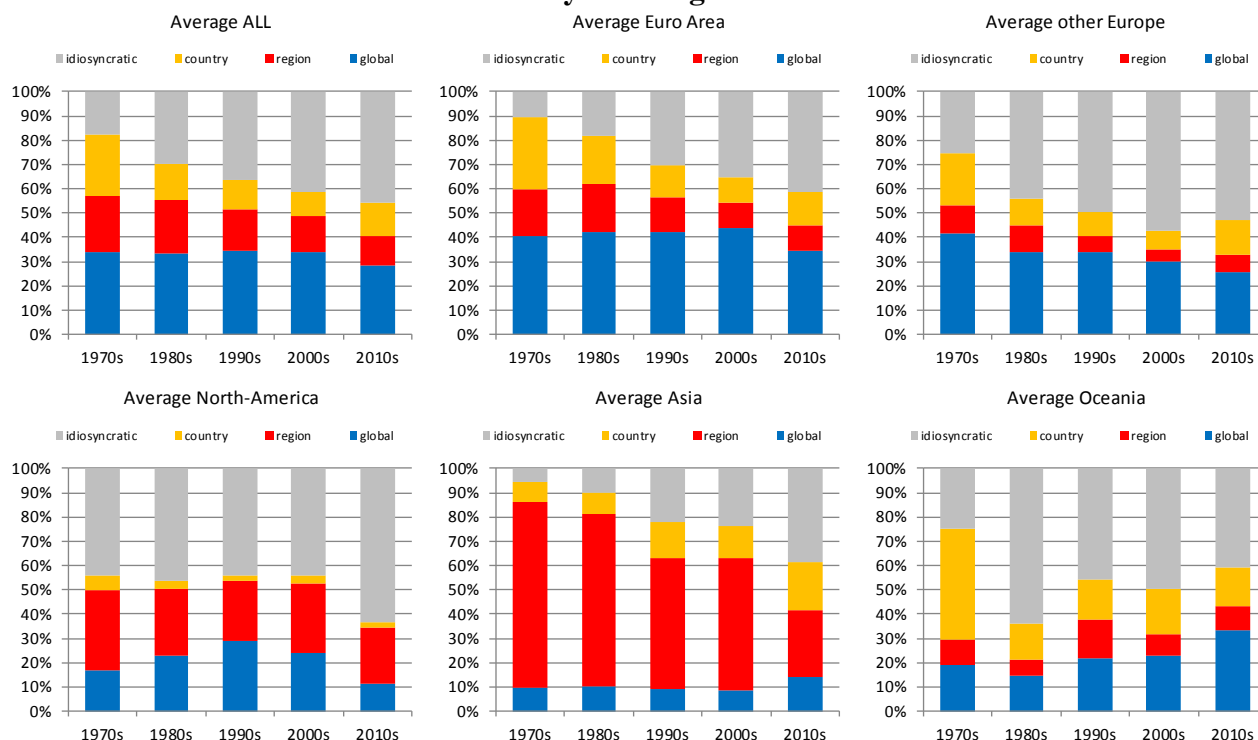
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of credit growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table K – Variance decompositions: contributions of uncertainty components to the volatility of loan growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	29%	51%	75%	10%	23%	40%	4%	12%	26%	15%
France	32%	54%	76%	9%	20%	37%	1%	3%	8%	24%
Italy	28%	50%	74%	17%	36%	58%	1%	5%	15%	9%
Spain	32%	53%	75%	13%	29%	48%	1%	4%	12%	14%
Netherlands	15%	31%	55%	1%	2%	7%	7%	16%	28%	51%
Belgium	22%	40%	64%	3%	8%	18%	3%	9%	22%	43%
Austria	35%	56%	78%	2%	6%	14%	9%	24%	46%	13%
Finland	16%	34%	60%	2%	7%	19%	1%	3%	8%	56%
Greece	5%	22%	52%	0%	3%	11%	25%	61%	92%	14%
Ireland	10%	26%	53%	11%	23%	41%	9%	24%	49%	27%
Portugal	20%	37%	59%	2%	6%	14%	20%	31%	46%	26%
UK	13%	35%	66%	7%	20%	41%	1%	4%	14%	40%
Sweden	9%	25%	53%	8%	20%	37%	7%	18%	39%	37%
Denmark	18%	32%	53%	0%	1%	4%	0%	0%	2%	66%
Switzerland	35%	57%	78%	0%	0%	1%	19%	39%	61%	4%
Norway	9%	16%	31%	0%	0%	1%	0%	0%	2%	83%
US	4%	12%	32%	13%	34%	64%	1%	4%	17%	50%
Canada	12%	32%	60%	6%	21%	50%	1%	3%	11%	45%
Japan	3%	9%	29%	28%	57%	81%	3%	12%	33%	22%
Australia	10%	22%	44%	1%	2%	6%	5%	13%	28%	62%
New Zealand	9%	21%	42%	7%	19%	37%	15%	31%	51%	30%
Korea	3%	11%	32%	33%	61%	84%	5%	14%	32%	15%
Av. Euro Area	22%	41%	66%	6%	15%	28%	7%	17%	32%	26%
Av. other Europe	17%	33%	56%	3%	8%	17%	5%	12%	24%	46%
Av. North-America	8%	22%	46%	9%	27%	57%	1%	3%	14%	47%
Av. Asia	3%	10%	30%	30%	59%	82%	4%	13%	33%	19%
Av. Oceania	10%	22%	43%	4%	10%	22%	10%	22%	40%	46%
Average ALL	17%	33%	56%	8%	18%	32%	6%	15%	29%	34%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of loan growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart M – Variance decompositions: contributions of uncertainty components to the volatility of loan growth over time**



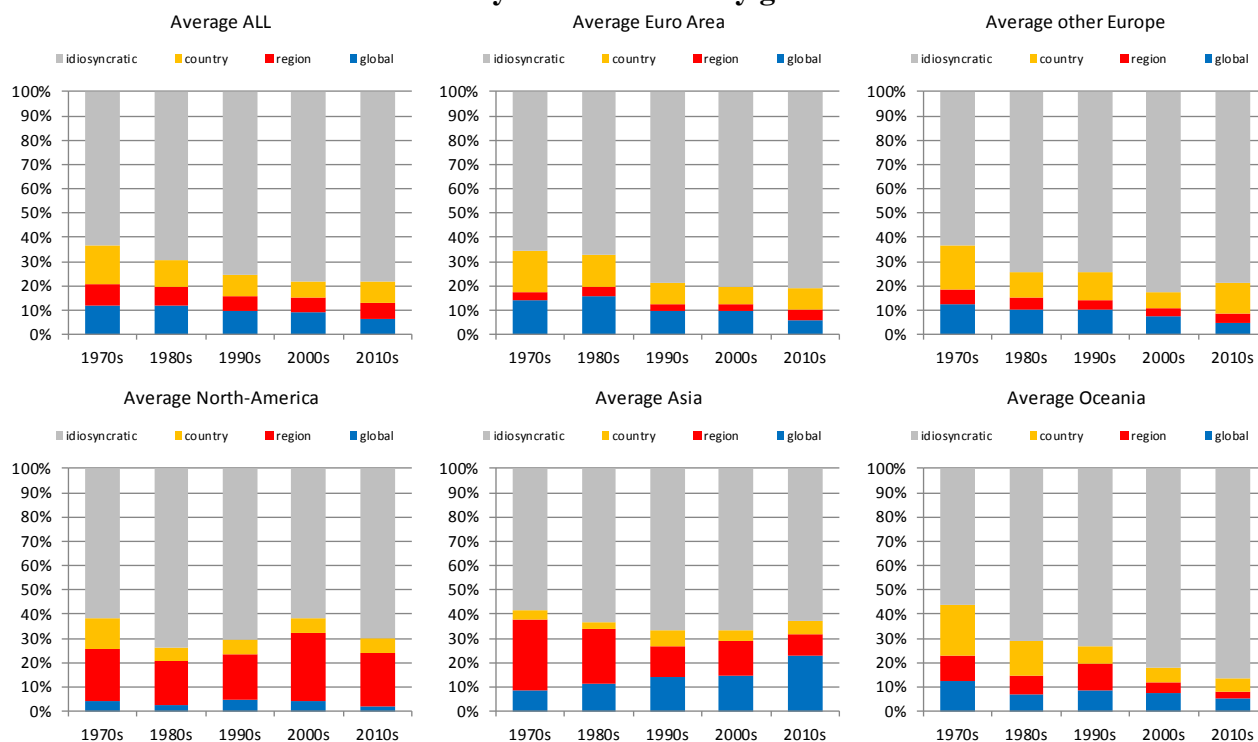
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of loan growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table L – Variance decompositions: contributions of uncertainty components to the volatility of narrow money growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	2%	6%	17%	0%	1%	2%	4%	9%	19%	84%
France	5%	14%	32%	2%	5%	13%	5%	14%	32%	67%
Italy	9%	23%	48%	1%	3%	9%	3%	9%	21%	65%
Spain	12%	27%	52%	1%	3%	10%	3%	8%	19%	61%
Netherlands	3%	8%	24%	1%	4%	11%	0%	1%	5%	87%
Belgium	2%	5%	13%	2%	6%	13%	6%	16%	35%	73%
Austria	3%	9%	23%	0%	1%	3%	0%	1%	5%	89%
Finland	2%	7%	21%	1%	4%	11%	2%	6%	15%	84%
Greece	2%	8%	25%	0%	2%	6%	3%	13%	50%	77%
Ireland	2%	7%	21%	1%	5%	13%	9%	25%	51%	63%
Portugal	4%	12%	31%	1%	2%	6%	7%	20%	39%	66%
UK	2%	7%	22%	3%	9%	22%	1%	3%	11%	80%
Sweden	1%	2%	7%	2%	4%	9%	21%	43%	68%	51%
Denmark	4%	11%	26%	3%	6%	14%	3%	8%	17%	75%
Switzerland	4%	14%	34%	0%	1%	3%	1%	2%	5%	83%
Norway	4%	13%	31%	0%	1%	5%	0%	3%	11%	84%
US	1%	4%	17%	5%	17%	44%	1%	3%	11%	76%
Canada	1%	3%	10%	9%	26%	53%	3%	11%	28%	60%
Japan	7%	18%	40%	3%	13%	37%	0%	2%	9%	67%
Australia	4%	12%	31%	2%	5%	15%	6%	18%	39%	64%
New Zealand	2%	4%	12%	4%	9%	20%	1%	3%	8%	83%
Korea	3%	10%	29%	7%	23%	50%	2%	7%	21%	60%
Av. Euro Area	4%	11%	28%	1%	3%	9%	4%	11%	27%	74%
Av. other Europe	3%	9%	24%	2%	4%	11%	5%	12%	23%	75%
Av. North-America	1%	4%	13%	7%	21%	49%	2%	7%	20%	68%
Av. Asia	5%	14%	35%	5%	18%	44%	1%	5%	15%	64%
Av. Oceania	3%	8%	22%	3%	7%	17%	4%	11%	24%	74%
Average ALL	4%	10%	26%	2%	7%	17%	4%	10%	24%	73%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of narrow money growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart N – Variance decompositions: contributions of uncertainty components to the volatility of narrow money growth over time**



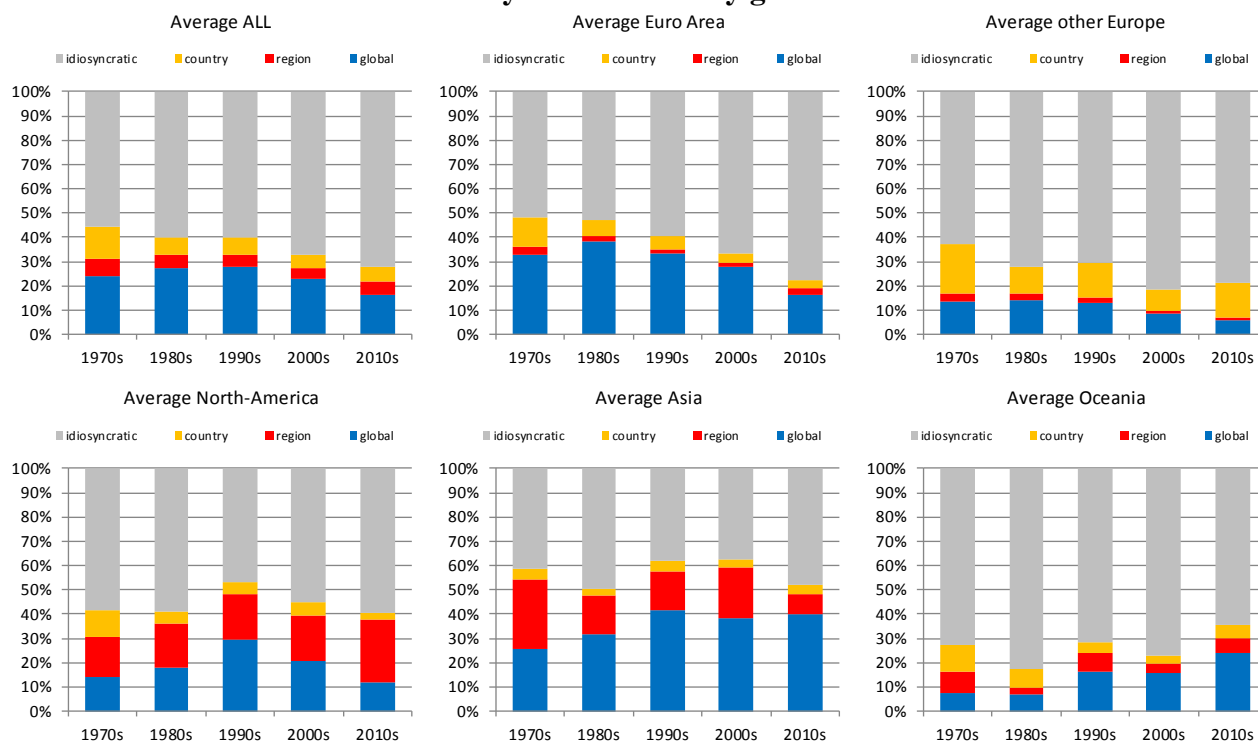
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of narrow money growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table M – Variance decompositions: contributions of uncertainty components to the volatility of broad money growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	10%	23%	46%	0%	2%	6%	2%	7%	17%	68%
France	24%	42%	65%	1%	2%	5%	3%	9%	19%	48%
Italy	18%	37%	63%	0%	1%	3%	0%	1%	3%	61%
Spain	36%	59%	79%	1%	2%	5%	0%	1%	4%	38%
Netherlands	7%	19%	40%	1%	3%	7%	1%	2%	7%	77%
Belgium	7%	18%	38%	1%	4%	8%	6%	14%	30%	65%
Austria	15%	30%	54%	1%	2%	6%	0%	1%	4%	67%
Finland	13%	27%	50%	0%	1%	4%	1%	3%	8%	69%
Greece	7%	26%	56%	1%	3%	9%	2%	9%	38%	63%
Ireland	3%	10%	26%	1%	3%	8%	6%	20%	45%	67%
Portugal	26%	49%	74%	0%	1%	4%	1%	3%	10%	47%
UK	7%	19%	42%	1%	4%	9%	1%	3%	10%	75%
Sweden	2%	6%	17%	0%	0%	2%	31%	55%	78%	38%
Denmark	2%	5%	16%	1%	3%	7%	1%	3%	8%	89%
Switzerland	4%	12%	33%	1%	3%	9%	1%	3%	8%	81%
Norway	5%	15%	35%	0%	1%	2%	1%	3%	10%	82%
US	1%	5%	20%	5%	19%	50%	2%	10%	29%	66%
Canada	15%	34%	61%	5%	19%	48%	0%	2%	7%	45%
Japan	22%	48%	77%	4%	13%	35%	0%	1%	4%	38%
Australia	4%	11%	28%	0%	2%	6%	2%	6%	17%	82%
New Zealand	6%	16%	37%	3%	10%	23%	2%	7%	15%	68%
Korea	7%	23%	52%	7%	24%	54%	2%	6%	19%	47%
Av. Euro Area	15%	31%	54%	1%	2%	6%	2%	6%	17%	61%
Av. other Europe	4%	11%	29%	1%	2%	6%	7%	13%	23%	73%
Av. North-America	8%	20%	41%	5%	19%	49%	1%	6%	18%	56%
Av. Asia	14%	35%	64%	5%	19%	45%	1%	3%	11%	43%
Av. Oceania	5%	13%	32%	2%	6%	14%	2%	6%	16%	75%
Average ALL	11%	24%	46%	2%	5%	14%	3%	8%	18%	63%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of broad money growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart O – Variance decompositions: contributions of uncertainty components to the volatility of broad money growth over time**



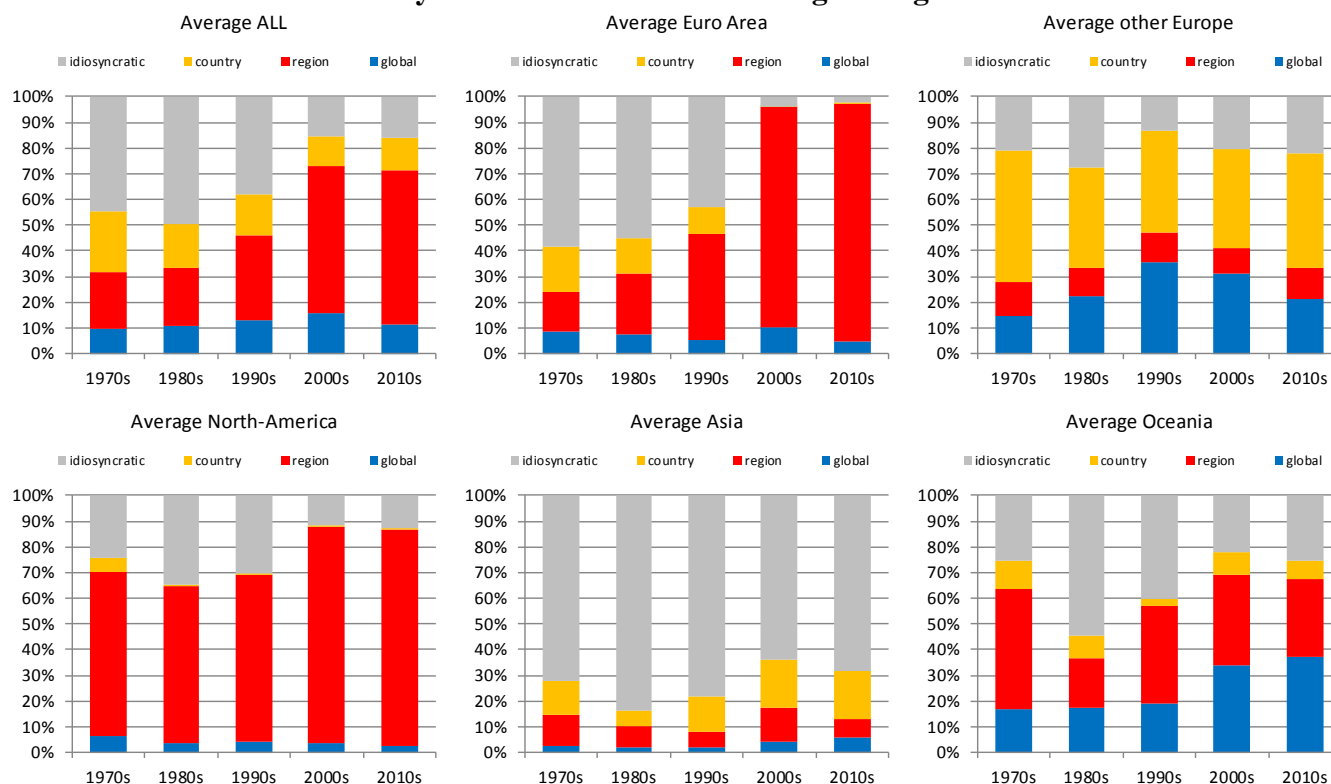
Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of broad money growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

**Table N – Variance decompositions: contributions of uncertainty components to the volatility of nominal effective exchange rate growth**

	global			region-specific			country-specific			idiosyncratic
	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	16 <sup>th</sup> p.	median	84 <sup>th</sup> p.	residual*
Germany	2%	7%	19%	48%	62%	74%	0%	1%	3%	31%
France	4%	8%	20%	43%	53%	64%	0%	0%	1%	38%
Italy	3%	8%	18%	28%	38%	47%	0%	0%	0%	55%
Spain	2%	7%	19%	34%	42%	50%	28%	41%	50%	10%
Netherlands	2%	5%	13%	56%	71%	82%	0%	0%	1%	25%
Belgium	1%	4%	12%	47%	59%	70%	0%	0%	1%	37%
Austria	4%	9%	24%	40%	58%	74%	0%	1%	3%	32%
Finland	3%	11%	27%	27%	38%	51%	27%	42%	55%	9%
Greece	4%	11%	26%	20%	34%	49%	0%	2%	11%	53%
Ireland	1%	4%	11%	37%	50%	63%	0%	2%	7%	44%
Portugal	3%	9%	20%	29%	39%	49%	3%	10%	30%	43%
UK	7%	20%	42%	2%	6%	14%	45%	68%	86%	6%
Sweden	2%	6%	20%	0%	2%	8%	61%	82%	93%	10%
Denmark	24%	44%	67%	12%	24%	42%	8%	19%	37%	13%
Switzerland	19%	36%	60%	8%	18%	32%	21%	39%	59%	7%
Norway	8%	24%	51%	2%	7%	20%	0%	1%	21%	68%
US	2%	6%	17%	26%	55%	80%	0%	2%	7%	37%
Canada	1%	2%	9%	57%	86%	96%	0%	2%	6%	10%
Japan	0%	1%	5%	0%	2%	7%	2%	6%	15%	91%
Australia	13%	29%	52%	12%	34%	59%	3%	9%	22%	29%
New Zealand	8%	20%	40%	14%	34%	57%	2%	7%	17%	40%
Korea	1%	5%	15%	6%	17%	39%	10%	21%	38%	57%
Av. Euro Area	3%	7%	19%	37%	49%	61%	5%	9%	15%	34%
Av. other Europe	12%	26%	48%	5%	11%	23%	27%	42%	59%	21%
Av. North-America	1%	4%	13%	42%	71%	88%	0%	2%	7%	24%
Av. Asia	1%	3%	10%	3%	10%	23%	6%	14%	26%	74%
Av. Oceania	11%	24%	46%	13%	34%	58%	3%	8%	19%	34%
Average ALL	5%	12%	27%	25%	38%	51%	10%	16%	26%	34%

Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of nominal effective exchange rate growth over the whole sample period 1971Q1-2016Q4. \* Idiosyncratic contribution derived as residual.

**Chart P – Variance decompositions: contributions of uncertainty components to the volatility of nominal effective exchange rate growth over time**



Notes: Contributions of the global, region-specific, country-specific and idiosyncratic components to the variance of nominal effective exchange rate growth. 1970s: average 1971Q1-1979Q4, 1980s: average 1980Q1-1989Q4, 1990s: average 1990Q1-1999Q4, 2000s: average 2000Q1-2009Q4, 2010s: average 2010Q1-2016Q4.

# School of Economics and Finance



This working paper has been produced by  
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