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Foster, Dean P.

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Discussion Paper No. 1182

A PROOF OF CALIBRATION VIA BLACKWELL'S APPROACHABILITY THEOREM

by
Dean P. Foster*
Department of Statistics
Wharton School, The University of Pennsylvania

February 27, 1997

^{*} Foster is also a visiting professor in the Center for Mathematical Studies in Economics and Management Science at Northwestern University, Winter 1997.

Dean P. Foster*

February 27, 1997

Abstract

Over the past few years many proofs of calibration have been presented (Foster and Vohra (1991, 1997), Hart (1995), Fudenberg and Levine (1995), Hart and Mas-Colell (1996)). Does the literature really need one more? Probably not, but this algorithm for being calibrated is particularly simple and doesn't require a matrix inversion. Further the proof follows directly from Blackwell's approachability theorem. For these reasons it might be useful in the class room.

This work was done while I was visiting the Center for Mathematical Studies in Economics and Management Science, Northwestern University. Permanent affiliation: Dept. of Statistics. The Wharton School, University of Pennsylvania, Philadelphia, PA 19104. Email:foster@hellspark.wharton.upenn.edu.

Thanks to Sergiu Hart who provided the proof of the only result in the paper. The algorithm is a modification of the original algorithm in Foster and Vohra (1991).

Suppose at time t a forecast, f_t , is made which takes on the value of the midpoint of each of the intervals [0, 1/m], [1/m, 2/m], ..., $[\frac{m-1}{m}, 1]$, namely, $\frac{2i-1}{2m}$ for i equals 1 to m. Let A_t^i the vector of indicators as to which forecast is actually made:

$$A_t^i = \begin{cases} 1 & \text{if } f_t = \frac{2i-1}{2m} \\ 0 & \text{otherwise} \end{cases}$$

Let X_t be the outcome at time t. We can now define the empirical frequency ρ_t^i as:

$$\rho_T^i = \frac{\sum_{t=1}^T X_t A_t^i}{\sum_{t=1}^T A_t^i}$$

Hopefully, ρ_T^i lies in the interval $\left[\frac{i-1}{m}, \frac{i}{m}\right]$. If so, the forecast is approximately calibrated. If not, I will measure how far outside the interval it is by two distances: \overline{d}_t^i and $\overline{\epsilon}_t^i$ (for deficit and excess) which are defined as:

$$\overline{d}_{T}^{i} = \frac{1}{T} \sum_{t=1}^{T} (\frac{i-1}{m} - X_{t}) A_{t}^{i} = [\frac{i-1}{m} - \rho_{T}^{i}] \overline{A}_{T}^{i}$$

$$\overline{\epsilon}_{T}^{i} = \frac{1}{T} \sum_{t=1}^{T} (X_{t} - \frac{i}{m}) A_{t}^{i} = [\rho_{T}^{i} - \frac{i}{m}] \overline{A}_{T}^{i}$$

where $\overline{A}_T^i = \sum A_t^i/T$. I will show that the following forecasting rule will drive both of these distances to zero:

- 1. If there exist an i^* such that $\overline{e}^{i^*} \leq 0$ and $\overline{d}^{i^*} \leq 0$, then forecast $\frac{2i^*-1}{2m}$.
- 2. Otherwise, find an i^* such that $\overline{d}_T^{i^*} > 0$ and $\overline{e}_T^{i^*-1} < 0$ then randomly forecast either $\frac{2i^*-1}{2m}$ or $\frac{2i^*+1}{2m}$ with probabilities:

$$P\left(f_{T+1} = \frac{2i^* - 1}{2m}\right) = 1 - P\left(f_{T+1} = \frac{2i^* + 1}{2m}\right) = \frac{\overline{d}_T^{i^*}}{\overline{d}_T^{i^*} + \overline{\epsilon}_T^{i^* - 1}}$$

It is clear that an i^* can be found in step 2, since i=1 always under forecasts and i=m always over-forecasts.

The L-1 calibration score:

$$C_{1,T} \equiv \sum_{i=0}^{m} |\rho_t^i - \frac{2i+1}{m}| \overline{A}_T^i = \frac{1}{2m} + \sum_{i=1}^{m} \max(\overline{d}_T^i, \overline{c}_T^i)$$

so showing that all the $\overline{\epsilon}_T^i$ and \overline{d}_T^i converge to zero, implies that $C_{1,T}$ converges to $\frac{1}{2m}$.

Theorem 1 (Foster and Vohra) For all $\epsilon > 0$, there exists a forecasting method which is calibrated in the sense that $C_{1,T} < \epsilon$ if T is sufficiently large. In particular the above algorithm will achieve this goal if $m \geq \frac{1}{\epsilon}$.

Consider this as a game between a statistician and nature. The statistician picks the forecast f_t and nature picks the data sequence X_t . The statisticians goal is to force all of the \bar{c}^i and \bar{d}^i to be negative (or at least approach this in the limit). Nature's goal is to keep the statistician from doing this. This set up is a game of "approachability" which was studied by Blackwell. He found a necessary and sufficient condition for a set to be approachable.

Theorem 2 (Blackwell 1956) Let L_{ij} be a vector valued payoff taking values in \mathbb{R}^n , where the statistician picks an i from \mathcal{I} at round i and nature picks a strategy j from \mathcal{J} at time t. Let G be a convex subset of \mathbb{R}^n . Let $a \in \mathbb{R}^n$ and let $c \in G$ be the closest point in G to the point a. Then G is approachable by the statistician if for all such a, there exist a weight vector w_i such that for all $j \in \mathcal{J}$.

$$\left(\sum_{i\in\mathcal{I}} w_i L_{ij} - c\right)'(a - c) \le 0. \tag{1}$$

To prove Theorem 1, we need to translate the calibration game into a Blackwell approachability game. The set of strategies for the statistician, \mathcal{I} ,

is the set of the m different forecasts. The set of strategies for nature, \mathcal{J} , is the set $\{0,1\}$. Define

$$\begin{array}{rcl} \epsilon_X^i & = & (X - \frac{i}{m})A^i \\ d_X^i & = & (\frac{i-1}{m} - X)A^i \end{array}$$

The vector loss is the vector of all the (d^i, ϵ^i) 's, in otherwords, it is a point in R^{2m} . The goal set $G \subset R^{2m}$ is $G = \{x \in R^{2m} | (\forall k) x_k \leq 0\}$. Let $\overline{\epsilon}_T^i = \frac{1}{T} \sum_{t=1}^T \epsilon_{X_t}^i$ and $\overline{d}_T^i = \frac{1}{T} \sum_{t=1}^T d_{X_t}^i$. The $(d_X^i, \epsilon_X^i)_i$ will be our L_{ij} in the Blackwell game, and $(\overline{d}^i, \overline{\epsilon}^i)_i$ will be the point c. The closest point in G to the current average $a = (\overline{d}^i, \overline{\epsilon}^i)_{i \in \mathcal{I}}$ is

$$c = \left((\overline{d}^i)^-, (\overline{\epsilon}^i)^- \right)_{i \in \mathcal{I}}.$$

where we have defined the positive and negative parts as $x^+ = \max(0, x)$ and $x^- = \min(0, x)$. The weight vector w is the vector of probability of forecasting i/k.

Proof: (Hart 1996) Now to check equation (1). Writing it in terms of d^{i*} s and ϵ^{i*} s equation (1) is:

$$\sum_{i=1}^{m} \left((w^i d^i - (\overline{d}^i)^-) (\overline{d}^i - (\overline{d}^i)^-) + (w^i \epsilon^i - (\overline{\epsilon}^i)^-) (\overline{\epsilon}^i - (\overline{\epsilon}^i)^-) \right) \le 0$$

from $x - x^- = x^+$ equation (1) is equivalent to

$$\sum_{i=1}^{m} \left((w^{i} d^{i} - (\overline{d}^{i})^{-}) (\overline{d}^{i})^{+} + (w^{i} \epsilon^{i} - (\overline{\epsilon}^{i})^{-}) (\overline{\epsilon}^{i})^{+} \right) \leq 0$$

Since, $(x^{-})(x^{+}) = 0$, it is sufficient to show:

$$\sum_{i=1}^{m} w^{i} (\epsilon^{i} (\overline{\epsilon}^{i})^{+} + d^{i} (\overline{d}^{i})^{+}) \leq 0.$$

If step 1, of the algorithm is used the weight vector is just $w^{i^*} = 1$ if i^* is the forecast chosen and zero otherwise. So $w^i \neq 0$ only when both $(\overline{d}^i)^+$ and $(\overline{c}^i)^+$ are zero, so the entire sum is zero.

If step 2, is used, the non-zero terms are w^{i^*} and w^{i^*-1} . But, $(\overline{e}^{i^*})^+$ is zero and $(\overline{d}^{i^*-1})^+$ is zero. So, it is sufficient to show:

$$w^{i^*}d^{i^*}(\overline{d}^{i^*})^+ + w^{i^*-1}e^{i^*-1}(\overline{e}^{i^*-1})^+ \le 0$$

But, $d^{i^*} = -e^{i^*-1}$, so it is sufficient to show:

$$w^{i^*}(\overline{d}^{i^*})^+ - w^{i^*-1}(\overline{c}^{i^*-1})^+ \le 0$$

But, this follows (with equality) from the definition of our probabilities. \Box

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