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Predicative Ability of Similarity-based Futures Trading Strategies

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Predicative Ability of Similarity-based Futures Trading Strategies

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September 21, 2018

Abstract

A trading rule that draws on the empirical similarity concept is proposed to simulate the technical trading mentality—one that selectively perceives structural resemblances between market scenarios of the present and the past. In more than half of the nineteen futures markets that we test against for profitability of this similarity-based trading rule, we find evidence of predictive ability that is robust to data-snooping and transaction-cost adjust-ments. When aided by an exit strategy that liquidates the trader's positions across some evenly-spaced time points, this rule generates the most robust returns.

Keywords: empirical similarity; technical trading; futures markets; analogical reasoning JEL classification: G11; G12

1 Introduction

Indeed, as Sharpe (1975) pointed out, it is said that the military is usually well prepared to fight the previous war—an intriguing old saying that reminds us of how susceptible we are to past experiences of our own when projecting the future. In this paper, such experience driven projection of the future takes place in technical trading. We assume that technical traders are "similarity-based", meaning that their judgments of present market conditions and projections of probable future returns are derived from recognizing vivid, concrete patterns of similarities portrayed by a multitude of technical indicators between market scenarios of the present and the past. To the best of our knowledge, the assumption itself is the first attempt in the literature to render the behavioral motive behind technical trading with similarity-based analogical reasoning (Hume, 1748).

Specifically, we resort to the Case-based Decision Theory of Gilboa and Schmeidler (1995), Billot, Gilboa, Samet and Schmeidler (2005), and Gilboa, Lieberman and Schmeidler (2006, 2011) in conceptualizing the technical trading mentality where belief formation leading to the execution of eventual buy-sell decisions comes as a direct consequence of traders selectively perceiving structural resemblances between scenario-analogies. Unlike Barber and Odean (2000) and Grinblatt and Keloharju (2009) who challenge about the exact nature of trading (it can be hazardous to your health), this paper postulates about how people trade and what lies behind their decision making process. Most important, using the futures markets as one particular application of economic interest, we add to the technical trading literature a novel prospective depicting how reason by analogy/similarity in one's trading decision making relates to his/her trading outcomes.

The similarity-based trading rule (SBTR) is thus devised to manifest the case-based logic. To initiate the SBTR, the trader must be capable of constructing a "mental reference" of information entities. This amounts to devising a vector of explanatory variables that represents the qualitative characteristics of financial information. Because his/her preferred way of processing information can be very unique, the SBTR is naturally adoptable as a univariate or a multivariate strategy depending on the trader's preferred parameterization for the vector of qualitative characteristics. At the same time, we assume that the trader is in possession of a knowledge-/data- base comprising historical cases of the vector of qualitative characteristics. The historical cases of the vector serve to depict the market scenarios of the past, while the vector's present case portrays the market scenario today. To decide if any past market scenarios resemble the present, and to ask to what extent they resemble the present, a similarity-based trader is equipped with a similarity function that embodies a user-definable measure of distance so as to quantify the extent of two objects being similar. Using this similarity function, the trader then conducts similarity searches, under moving windows of specifically chosen time lengths prior to the present date, by drawing analogies between market scenarios of the past and the present. Probable future returns under the current market scenario are then predicted using the similarity-weighted average of past returns would indicate a buy (sell) trading signal. This similarity-weighted average of past returns is hereafter referred to as the "stochastic averaging predictor" of the SBTR.

In terms of time frame, the SBTR is applicable daily or weekly. In the former, the SBTR adopts the day-trade mechanism as depicted by Kuo and Lin (2013) and Barber, Lee, Liu, and Odean (2014), where a similarity-based trader enters/closes his/her trading positions at the opening/closing price on the same date. As a weekly strategy, we assume that the trader enters (closes) his/her trading positions at the beginning (end) of each week.

We address market timing by allowing the trader to unload his/her trading positions under different exit strategies. We follow Lu, Chen and Hsu (2015) to consider the MYR (Marshall, Young and Rose, 2006) and the CL (Caginalp and Laurent, 1998) exit strategies. While the MYR exit strategy is characterized by a predetermined date and condition to liquidate the trader's trading positions once and for all, the CL exit strategy relies on an average exit price for the holding period so that the trader can gradually unload his/her trading positions across some evenly-spaced time points.

We test these daily and weekly SBTRs for profitability against historical prices from nineteen futures markets. To ensure that the (daily or weekly) best-performing rules detected are indeed genuine, we apply the superior predictive ability test of Hansen (2005) throughout to control for data-snooping bias. For robustness checks, we conduct out-of-sample experiments to see if the (daily or weekly) best-performing SBTRs selected from the first sub-period sample—based on their individual nominal *p*-values—would continue to perform in the second sub-period sample. To address transactions costs, we follow Qi and Wu (2006) and Park and Irwin (2010) to consider the maximum one-way break-even transaction costs and the commissions per round-trip trade.

Our main findings can be summarized as follows. First, after data snooping adjustment, we find that there exists a best daily SBTR that generates significantly positive profit in ten out of the nineteen futures markets. On the other hand, the best weekly SBTRs with the MYR exit strategy are profitable in six of the nineteen futures markets. The best weekly SBTRs with the CL exit strategy are profitable in all futures markets. The sub-sample analyses confirm that the weekly SBTRs with the CL exit strategy generate the most robust excess returns. Second, we find that common to these best SBTRs is the application of a multivariate strategy that utilizes feasible combinations of the moving average, relative strength index, trading range breakthrough, and trading-volume moving average indicators in the construction of the vector of qualitative characteristics. Consistent with Park and Irwin's (2010) finding, we document also favorable results for the technical trading rules being applied to the commodity futures relative to the financial futures. Third, we find that, for an SBTR to outperform, the required moving time lengths for conducting similarity searches indeed matter, and they tend to vary across markets. Structural resemblances of the recent past—rather than those of the distant past—seem to induce greater impacts on a similarity-based trader's decision making. Lastly, to examine the after-cost performance of the SBTRs, we report the maximum one-way costs to break even for the best-performing SBTRs traded in each futures market. We show that the SBTRs still generate positive returns when a moderate transaction cost is assumed. Most of the best performing SBTRs robust to the data-snooping bias are profitable based on a one-way transaction cost of 0.04% or a one-way commission of \$6.25. When the one-way commission is relatively high at \$50, most of the SBTRs are no longer profitable except for the weekly SBTRs with the MYR exit strategy.

The remainder of this paper is organized as follows. Section 2 introduces the similarity-based technical trading rule, and outlines the classes of technical indicators employed by the SBTR.

Section 3 describes the empirical test with data-snooping adjustments. Section 4 presents our empirical findings. In Section 5, we conclude this study.

2 Methodology

This section introduces the SBTRs. We first define the stochastic averaging predictor for an outcome variable of interest, and present the decision rule based on the predictor. Then, we prescribe the vector of qualitative characteristics with the classes of technical indicators that are most prevalent in real-life trading. Finally, we construct the set of SBTRs.

2.1 The similarity-based technical trading rule

We begin by assuming that a similarity-based trader is in possession of a set of information signals conveyed by a vector of qualitative characteristics of dimension k, i.e., $x_t = (x_{1,t}, \ldots, x_{k,t})$, where $x_{j,t}$ can be a specific trading indicator, such as the difference between the shorter-term and the longer-term moving average, the value of the relative strength index over the past 10 days, or the difference between the opening and closing prices of one Candlestick. Conditional on the available information signals at time t, he/she will then attempt to forecast future returns, which we refer to as a (real-valued) outcome variable y_t . In this study, y_t indicates either the next 1-day (daily) or the next 5-day (weekly) future returns. Define \hat{y}_t^s as the stochastic averaging predictor of y_t such that

$$\hat{y}_{t}^{s} = \frac{\sum_{i=1}^{m} s(x_{t}, x_{t-i}) \cdot y_{t-i}}{\sum_{i=1}^{m} s(x_{t}, x_{t-i})} \tag{1}$$

where $s: R^k \times R^k \to (0, \infty)$ is a similarity function. The value of $s(x_t, x_{t-i})$ measures how similar x_t and x_{t-i} are, and the more similar that x_t and x_{t-i} are, the larger the value of $s(x_t, x_{t-i})$. The similarity function we use is $s(x_t, x_{t-i}) = \exp\left(-d(x_t, x_{t-i})\right)$, where $d(x_t, x_{t-i}) = \sqrt{\sum_{j=1}^k \left(\frac{x_{j,t}-x_{j,t-i}}{\sigma_j}\right)^2}$ defines a standardized Euclidean norm and σ_j is the sample standard deviation of the element j in the vector of qualitative characteristics. The standardized Euclidean norm scales each coordinate difference between x_t , and x_{t-i} by dividing it by the standard deviation of the values

in the corresponding coordinate over the moving window.

The \hat{y}_t^s so defined in (1) is a stochastic weighted averaging of the past returns y_{t-i} over the last m periods, and hence known as a "stochastic averaging predictor" (Gilboa and Schmeidler, 1995). In this paper, the SBTR based on \hat{y}_t^s has the following trading rule:

If
$$\hat{y}_t^s > 0$$
, enter a long position;
If $\hat{y}_t^s < 0$, enter a short position. (2)

In other words, the SBTR would recommend that the investor buys when $\hat{y}_t^s > 0$, because based on the past experience of the investor, the future return is more likely to be positive when the signal vector is x_t . A similar argument applies to the case where $\hat{y}_t^s < 0$.

One can certainly opt for other measures of similarity—for example the Mahalanobis metric to quantify the distance between any two vectors of qualitative characteristics. In our case, we use the standard Euclidean norm because it is simple to implement and more appropriate as a distance measure when different scales and sources of qualitative characteristics are involved.

2.2 Prescribing the Vector of Qualitative Characteristics

Because attention allocation is costly and requires effort (Kahneman, 1973; Hirshleifer and Teoh, 2003; Peng and Xiong, 2006; Barber and Odean, 2008), a similarity-based trader must be selective in his/her intended focus. We thus illustrate the SBTR using five prevalent classes of trading indicators to prescribe the vector of qualitative characteristics: the moving average (MA), relative strength index (RSI), trading range breakthrough (TRB), trading-volume moving average (VMA), and past 5-day Candlestick patterns. The information contents of these technical indicators are described as follows:

Moving Average (MA): MAs enable traders to identify current price trends, trend reversals, and the associated support and resistance levels. Traditionally, a technical trading rule based on the MA triggers buy/sell signals through the upward/downward crossovers of longer-term MAs by shorter-term MAs, where the MAs are calculated as the average closing prices over a specified period of time prior to the trading date, see, for example Brock, Lakonishok and LeBaron (1992), Sullivan, Timmermann and White (1999), and Hsu and Kuan (2005).

Relative Strength Index (RSI): The RSI measures the changes in and speed of price movements and helps traders to identify overbought or oversold conditions for an asset. The RSI is calculated as the ratio of average upward price changes to average downward price changes over a predetermined time period as in Wong, Manzur and Chew (2003) and Park and Irwin (2010). Conventionally, a technical trading rule based on the RSI generates sell signals when the RSI values are relatively high and buy signals when they are relatively low.

Trading Range Breakthrough (TRB): For TRB which is associated with the notion of support and resistance, we define the resistance level and the support level as the local maximum and the local minimum closing price over a predetermined time period, respectively, following Brock, Lakonishok and LeBaron (1992), Sullivan, Timmermann and White (1999), Hsu and Kuan (2005), and Park and Irwin (2010). A TRB trading rule generates buy signals when the price penetrates a resistance level and sell signals are generated by the crossing of the price over a support level.

Trading-volume Moving Average (VMA): VMAs are calculated as the average trading volume over a specified period of time that allow traders to identify trends in trading volume. A technical trading rule exploiting the VMA is typically constructed by entering a long position when the shorter-term VMA crosses above the longer-term VMA, and by entering a short position when the opposite occurs. Alternatively, a high reading in the difference between the shorter- and longer-term VMAs can also be considered as an overbought signal (Pring, 1993). Karpoff (1987), Campbell, Grossman, and Wang (1993), Blume, Easley and O'Hara (1994), and Lee and Rui (2002) examine the informational position of trading volume on the contemporaneous and expected price changes.

5-day Candlestick patterns: Candlestick patterns, commonly illustrated as a series of open-high-low-close prices, reflect the changing balance between supply and demand, in-

vestor sentiment and psychology. See e.g., Caginalp and Laurent (1998), Marshall, Young, and Rose (2006), Lu, Shiu, and Liu (2012), Lu (2014), and Lu, Chen, and Hsu (2015). To make feasible the adapation of past 5-day Candlestick patterns to the SBTR. We characterize each Candlestick pattern by their three vital elements: the difference between the closing price and opening price, the difference between the highest price and the maximum of the closing price and opening price (the upper shadow), and the difference between the lowest price and the minimum of the closing price and opening price (the lower shadow). In addition, we include the differences between the average price of the closing price and opening price of two consecutive Candlesticks. Therefore, using the past 5-day Candlestick patterns to depict the current market condition will give rise to a vector of characteristics containing a total of 19 elements.

2.3 Universe of SBTRs

Noting the common observation that technical traders rarely rely the sole use of one particular trading indicator while making investment decisions (Hsu and Kuan, 2005), a similarity-based trader is allowed to employ the SBTR as a univariate or multivariate strategy in which the vector of qualitative characteristics is prescribed with the values of a single type of trading indicator, or, that of a panel of mixture types.

In the following, we illustrate how to construct the vector of qualitative characteristics using the five types of technical indicators under a univariate SBTR and a multivariate SBTR.

2.3.1 Univariate SBTRs

Univariate SBTRs refer to the case where an SBTR prescribes the vector of qualitative characteristics with one and only one type of technical indicator at a time. For example, a similarity-based trader may decide that he/she is more comfortable with the MA rules, and his/her judgments of how a past market scenario resembles the present depend solely on how similar—as indicated by the similarity function—their MA qualitative characteristics match. As a consequence, the stochastic averaging predictor will be solely specified by the MA rules, with the vector of qualitative characteristics x_t comprising only the differences between the s-day and the l-day MAs. We follow Brock, Lakonishok and LeBaron (1992) for the choice of time periods to calculate the moving averages. The combinations of shorter- and longer-term MAs are denoted by MAs - l, where s=1, 5, 10, 20, 50, 100, l=5, 10, 20, 50, 100, 150, and s < l. Thus, the univariate SBTRs with the MA will include a total of 21 (6 + 5 + 4 + 3 + 2 + 1) shorter- and longer-term combinations.

For univariate SBTRs with the RSI, the vector of qualitative characteristics is then prescribed by the past p-day RSI values, where p corresponds to 5, 10, 20, 50, 100, and 150, respectively. This amounts to a total of 6 univariate SBTRs with the RSI. For univariate SBTRs with the TRB, we use the difference between the current closing price and the local maximum price over a specific period of time prior to the current date and the difference between the closing price and the local minimum price as two elements in the vector. In this study, we calculate the local maximum and minimum prices based on the past 5, 10, 20, 50, 100, and 150 days prior to the current date, and the univariate SBTR with the TRB is thus denoted as TRBp, where p = 5, 10, 20, 50, 100, 150. The total is then 6 for the univariate SBTRs with the TRB. As for the combinations of shorter- and longer-term VMAs, these are denoted by VMAs -l, where s=1, 5, 10, 20, 50, 100, l=5, 10, 20, 50, 100, 150, and s < l. The value for VMAs -l is the log difference between the shorter- and longer-term VMAs. Similar to the univariate SBTRs with the MA, the total number of univariate SBTRs with the VMA is 21. Lastly, when applying the 5-day Candlestick patterns, the vector of qualitative characteristics contains 19 elements as mentioned in Section 2.2.

Overall, the total number of univariate SBTRs is 55 (21 for the MA, 6 for the RSI, 6 for the TRB, 21 for the VMA, and 1 for the 5-day Candlestick pattern).

2.3.2 Multivariate SBTRs

Multivariate SBTRs refer to the case where a similarity-based trading mentality is capable of processing a heterogeneous pool of information entities. In making conditional forecasts based on stochastic averaging—of future returns, the trader chooses to prescribe the vector of qualitative characteristics with a multitude of technical indicators varying in terms of their types. The set of the SBTRs is expandable upon any feasible combination of technical indicators; it is constrained only by the limit of one's computational power.

In this paper, we demonstrate multivariate SBTRs using two sets of experiments. First, from any two of the MA, RSI, TRB, and VMA types, the trader chooses one indicator from each of them (981 rules); he/she chooses one indicator from each of the MA, RSI and TRB types (756 rules); and he/she chooses one indicator from each of the MA, RSI, TRB, and VMA types (15,876 rules). This yields a total of 17,613 multivariate SBTRs. In addition, the abovementioned rules can be mixed with the 5-day Candlestick pattern by expanding the vector of qualitative characteristics with the 19 elements characterizing the 5-day Candlestick pattern. This yields another 17,613 SBTRs.

Second, we also use the following mixed strategies where the vector contains one of the following: the 5-day Candlestick pattern coupled with an indicator chosen from the MA, RSI, TRB, and VMA; all MA indicators; all RSI; all TRB; all VMA; all indicators of the MA, RSI, TRB, and VMA; the past 5-day Candlestick pattern with all MA indicators; the past 5-day Candlestick pattern with all RSI; the past 5-day Candlestick pattern with all TRB; the past 5-day Candlestick pattern with all TRB; the past 5-day Candlestick pattern with all TRB; the past 5-day Candlestick pattern with all VMA indicators; and the past 5-day Candlestick pattern with all indicators. This yields a total of 64 SBTRs. Overall, the total number of the SBTRs with multiple types of technical rules is 35,290 (17,613+17,613+64).

3 Test with Data-snooping Adjustments

Data-snooping bias arises when the same data set is repetitively used to test for the statistical significance of the technical trading rules individually. In order to control for such bias, we employ Hansen's (2005) superior predictive ability (SPA) test. In the following, we first discuss how to calculate the strategic returns of the SBTRs and then how to adjust for data-snooping bias when examining the predictive ability of the SBTRs.

3.1 Strategic Returns of SBTRs

Let sign(.) be an indicator function such that $sign(\hat{y}_t^s)$ indicates whether a similarity-based trader should enter a long or short position according to the stochastic averaging predictor \hat{y}_t^s . Specifically, $sign(\hat{y}_t^s) = 1$ or -1 indicates a buy or sell signal when the stochastic averaging predictor \hat{y}_t^s predicts a positive or negative next 1-day (daily) or 5-day (weekly) future returns y_t . The strategic returns of the SBTRs, denoted by μ_t , can thus be defined by $\mu_t = sign(\hat{y}_t^s)y_t$.

For the three types of strategies considered in this paper—daily SBTRs, weekly SBTRs with the MYR exit strategy, and weekly SBTRs with the CL exit strategy—the strategic returns, realized at observation time t are computed as follows:

$$y_t^d = \ln(\frac{C_t}{O_t}) \tag{3}$$

$$y_{n,t}^{MYR} = \ln(\frac{C_{t+n-1}}{O_t}) \tag{4}$$

$$y_{n,t}^{CL} = \frac{\sum_{i=0}^{n-1} \ln(\frac{C_{t+i}}{O_t})}{n}$$
(5)

where y_t^d denotes the realized return of a daily SBTR and is computed as the log difference between the closing price and the opening price on the same trading day t. For a weekly SBTR with the MYR exit strategy, the realized weekly return, $y_{n,t}^{MYR}$, is the log difference between the closing price C_{t+n-1} on the last trading day of the week and the opening price O_t on the first trading day of the week, where n denotes the number of trading days within a week. For a weekly SBTR with the CL exit strategy, the similarity-based trader is assumed to enter a long or short position with n futures contracts and close them out one by one across some evenly spaced, predetermined time points t, ..., t + n - 1, and the realized weekly return, $y_{n,t}^{CL}$, is calculated as the average daily realized returns within a week.

Finally, once the strategy returns are constructed, the performance measure for an SBTR will be its mean return $\overline{\mu}$ calculated as follows,

$$\overline{\mu} = \sum_{t=1}^{N} \mu_t / N \tag{6}$$

where $\mu_t = sign(\hat{y}_t^s)y_t$ with $y_t \in \{y_t^d, y_{n,t}^{MYR}, y_{n,t}^{CL}\}$ as defined by (4), (5), and (6); N is the actual number of trading days or weeks involved.

3.2 Data-snooping adjustments

To test the null hypothesis that, among the set of SBTRs, none of them can generate significantly positive returns, one can employ White's (2000) reality check (RC) test to control for datasnooping bias. However, the RC test may be conservative because its null distribution is obtained under a least favorable configuration of parameter values. Moreover, the RC test may lose power when too many poor and irrelevant technical trading rules are included (see, e.g., Hansen (2005), Hsu, Hsu and Kuan (2010) and Hsu, Kuan and Yen (2014)). To avoid the least favorable configuration and to improve the power property of the RC test, Hansen (2005) proposes the SPA test based on a re-centering method.

To conduct the SPA test in our case, we formally let the performance measure of the k-th SBTR be the mean return of it, $\overline{\mu}_k = \sum_{t=1}^N \mu_{k,t}/N$ for $k = 1, \ldots, K$ following the literature on the profitability of technical trading rules (Sullivan, Timmermann and White, 1999; Hsu and Kuan, 2005; Park and Irwin, 2010). The test statistic of Hansen's (2005) SPA test is defined as

$$t^{SPA} = \max\left(\max_{k=1,\dots,K} \frac{\sqrt{N}\overline{\mu}_k}{\widehat{\sigma}_k}, 0\right) \tag{7}$$

where $\hat{\sigma}_k^2$ is a consistent estimator of $\sigma_k^2 = var\left(\sqrt{N}\overline{\mu}_k\right)$.¹ Following Hansen (2005), we apply the stationary bootstrap of Politis and Romano (1994) to approximate the null distributions. To be specific, for $b = 1, \ldots, B$, let $\overline{\mu}_k^b$ denote the sample mean of the *b*-th bootstrap sample. Define the recentering mean as $\hat{\mu}_k = \overline{\mu}_k 1(\sqrt{N}\overline{\mu}_k \leq -\hat{\sigma}_k\sqrt{2\log\log(N)})$, where 1(E) denotes the indicator function of the event *E*. The bootstrapped null distribution is given by:

$$t_b^{SPA} = \max\left(\max_{k=1,\dots,K} \frac{\sqrt{N}(\overline{\mu}_k^b - \overline{\mu}_k + \widehat{\mu}_k)}{\widehat{\sigma}_k}, 0\right).$$
(8)

The *p*-value of the SPA test is then approximated by $p^{SPA} = \sum_{b=1}^{B} \frac{1(t_b^{SPA} > t^{SPA})}{B}$, the proportion of times when the resampling statistics are larger than the test statistic. For a given significance level α , we will reject the null when $p^{SPA} < \alpha$ and conclude that there exists at least one SBTR

¹We follow Hsu and Kuan (2005) to compute the consistent estimator of σ_k^2 based on the bootstrap resample: $\hat{\sigma}_k^2 = \frac{1}{B} \sum_{b=1}^{B} \left(\overline{\mu}_k^b - \overline{\mu}_k^b\right)^2$, where $\overline{\mu}_k^b$ is the average of $\overline{\mu}_k^b$ over b.

that can generate a positive mean return.

Regarding the choice of the total number of bootstrap resamples B, and the probability parameter q for the stationary bootstrap, we follow Sullivan, Timmermann and White (1999), and Park and Irwin (2010) to set B = 500 and q = 0.1.² Note that changing these parameters yields similar empirical results.

4 Empirical Results

4.1 Data

Because the sample periods of the available data do not match, we test the profitability of the SBTRs against the historical prices of nineteen futures markets over two different sample periods. The first sample period ranges from 1995/01/01 to 2015/12/31, and forms our study sample "Group 1". Study sample Group 1 comprises futures markets of the corn, soybean, wheat, live cattle, lumber, cocoa, sugar, silver, oil, S&P 500, T-Bills, Eurodollar, YEN and GBP. On the other hand, Study sample "Group 2" represents the second sample period ranging from 2005/01/01 to 2015/12/31, and consists of the E-mini S&P, E-mini NASDAQ, EUR and AUD futures.³

Group 1 has historical price data available for at least ten years before 1995/01/01, and Group 2 has historical prices available for at least five years before 2005/01/01. This amounts to Group 1 allowing for a 10-year maximum moving time window and Group 2 allowing for a 5-year maximum moving time window during similarity searches. Thus, for Group1, the choice of time length, m, as in equation (1), will take on a value of 250, 750, 1250 or 2,500, so as to represent the 1-, 3-, 5- and 10-year moving windows, respectively. Similarly for Group 2, the choice of time length for a moving window is set to 250, 750, or 1250, which correspond to 1, 3, and 5 years. With all feasible combinations of moving-window time lengths accounted for, the

²The number of boostrap samples B may influence the accuracy of the estimated p-values. However, Brock, Lakonishok, and LeBaron (1992) and Kho (1996) show that their estimated bootstrap p-values are insensitive to the bootstrap size B when the number of bootstrap samples goes beyond 500.

³The list of futures markets that we choose is close to that of Park and Irwin (2010). Although we exclude the pork bellies and Mark futures due to price unavailability, we include in addition two financial futures (E-mini S&P and E-mini NASDAQ) and two currency futures (EUR and AUD).

numbers of SBTRs total 141,380 ($(55 + 35, 290) \times 4$) and 106,035 ($(55 + 35, 290) \times 3$) for Group1 and Group 2, respectively.

Table 1 reports summary statistics for the annualized means and standard deviations of unconditional daily returns across the nineteen futures markets and across the two sub-periods. Daily returns are calculated as the log differences between the opening and closing prices of the same trading day, so as to comply with the daily strategies of SBTRs. Differential characteristics in the return distributions therein, across the futures markets, can clearly be identified: The annualized mean returns range from -30.29% for lumber to 14.08% for soybean; and the annualized standard deviations range from 0.68% for the Eurodollar to 31.06% for sugar. Within each market, the statistical properties of the return distributions across different sub-periods also vary. Sugar futures, for example, have their annualized mean returns ranging from -29.34% to 25.43% while their annualized standard deviations range from 30.03% to 32.05% across the sub-periods.

4.2 Performance Evaluation

Table 2 presents, for each futures market, the annualized mean returns of the best-performing daily SBTRs identified using the whole sample period. The second column of Table 2 depicts the qualitative characteristics of these winning strategies, which include the types of technical indicators in use, and the adopted time lengths for the moving windows. For example, the winning SBTR for corn futures is found to exhibit the following qualitative characteristics: it uses a 10-year moving window, and adopts a multivariate strategy that involves an MA indicator defined by the difference between the past 10-day (short-term) MA and the past 150-day (long-term) MA; the past 150-day RSI; the difference between the closing price and the past 50-day maximum price; and the difference between the closing price and the past 50-day minimum price.

In addition, according to Table 2, the best SBTR for corn futures is found to generate a positive annualized mean return of 24.29% with a 25.24% annualized standard deviation. The unconditional annualized mean return following the buy (sell) signals generated by the stochastic averaging predictor is 26.70% (-20.89%) with an annualized standard deviation of 26.14% (24.94%). The total number of buy (sell) signals comes to 3,089 (2,177), and the fraction of buy (sell) signals followed by positive unconditional returns is 53.32% (50.30%). The winner across the markets is the SBTR applied to lumber futures, which generates the best annualized mean return (39.61%) with an annualized standard deviation of 25.05%. This win-out relies on a multivariate strategy that facilitates the joint-work of one MA, one VMA, one RSI, and one TRB indicator to identify scenario resemblances of the present and to carry out stochastic averaging predictions based on a 3-year moving window.

The rightmost column presents p-values of Hansen's (2005) SPA test for those best-performing daily SBTRs. The p-values of Hansen's (2005) SPA test are defined in Section 2.2. The results of Hansen's (2005) SPA test in Table 2 show that in ten of the nineteen futures markets, the best-performing daily SBTRs are robust to the data-snooping adjustments at the 10% significance level.⁴

Tables 3 and 4 examine the return predictive ability of the best-performing weekly SBTRs under the MYR and CL exit strategies, respectively. Table 3 shows that among six of the nineteen futures markets, the best-performing weekly SBTRs using the MYR exit strategy generate positive returns at the 10% significance level based on Hansen's SPA test, which are weaker than the results for the daily SBTRs in Table 2. Across the markets, the best-performing SBTR is found for sugar futures where a multivariate strategy of one MA, one VMA, one RSI and one TRB indicator is adopted under a moving window of 5 years. Table 4 reports the annualized mean returns of the best-performing weekly SBTRs using the CL exit strategy. All strategies are found to generate profits that are significant at the 10% level after we adjust for the data-snooping biases. Compared to the results in Tables 2 and 3, weekly SBTRs with the CL exit strategy of Table 4 seem to generate the most profound returns. This finding, consistent with Lu, Chen and Hsu (2015), is likely to be a direct consequence of strategy returns exhibiting lower standard deviations—pointing to the role played by the CL exit strategy as a mechanism for risk sharing.

Several interesting findings can be drawn from Tables 2, 3, and 4. First, the ways in which

⁴For a one-sided test, a t-statistic has to be greater than 1.29 (1.64) to be claimed significantly greater than zero at 10% (5%) significance level.

different types of trading indicators synergize in order to search (resemblances), identify (similarities) and predict (stochastic averaging returns) seem to vary across markets. However, we find that common to most of the futures markets, winning SBTRs seem to rely on multivariate strategies that exploit different combinations of the MA, VMA, RSI, and TRB indicators in order to construct the vector of qualitative characteristics. For example, there are twelve futures markets in Table 2 where the best-performing daily SBTRs use the combinations of the MA, VMA, RSI, and TRB indicators. In addition, in four futures markets, the best-performing daily SBTRs exploit the 5-day Candlestick patterns combined with the MA, VMA, RSI, and TRB indicators. For the rest of these markets, SBTRs that win out are also found to be multivariate strategies of heterogeneous indicator types. This finding suggests that similarity-based traders may consider it unwise, or even brutal, to trade and rely solely on a single type of trading indicator. The similarity-based technical trading mentality is likely to be at its best upon for those decision makers whose mental capacity allows them to process information based on a multitude of technical indicators.

Second, the fraction of buy signals followed by positive unconditional returns exceeds the fraction of sell signals followed by positive unconditional returns for all futures markets. For example in Table 2, the fraction for buy signals ranges from 51.54% to 71.63%, while the fraction for sell signals ranges from 45.58% to 65.67%. According to Brock, Lakonishok and LeBaron (1992), the fraction of positive unconditional returns should be the same for both buy and sell signals if the signals generated by technical rules are indeed useless. Using binomial tests, it can be confirmed that the differences in the fraction of positive returns between buy and sell signals are significant, thereby rejecting the null hypothesis for their equality. This result indicates that the buy/sell signals generated by SBTRs are of predictive ability over future probable returns.

Third, Tables 2, 3 and 4 provide supporting evidence for the predictive ability of stochastic averaging predictors over future probable returns. Consistent with Park and Irwin (2010), our finding documents also favorable results of technical trading rules applied to commodity futures relative to financial futures. For example, in Table 2, the best-performing daily SBTR traded among financial futures markets yields an annualized mean return of 15.76% when trading on E-mini Nasdaq futures. On the other hand, the best-performing daily SBTR traded among commodity futures generates an annualized mean return of 39.61% for lumber futures. In addition, the average mean return across financial futures markets is 8.78%, while the average mean return across commodity futures is 25.99%.

Finally, in searching for past market scenario-resemblances of the present, we find that the adopted time lengths in fact vary across different markets. Similar experiences of the recent past—rather than those of the distant past—seem to induce greater impacts on the similarity-based traders' decision making. For example, in Table 2, the best-performing SBTRs for wheat, live cattle, and Yen futures utilize a moving window based on recent 1-year period (relative to a 5-year maximum), whilst for lumber, sugar, copper, silver, the S&P 500, and the Eurodollar, the best-performing SBTRs entail conducting a similarity search based on a moving window for the recent 3-year period (relative to a 10-year maximum).

4.3 Sub-sample analysis

For robustness checks, we perform two kinds of sub-sample analyses over two sub-periods of approximately equal lengths. (In the case of Study sample Group 1, the two sub-periods are 1995-2005 and 2006-2015. For Group 2, the two sub-periods are 2005-2010 and 2011-2015.)

We first follow Qi and Wu (2006) and Park and Irwin (2010) to conduct an in- and out-ofsample test. We use the first sub-period to select the best strategies for each futures market and then evaluate the performance of these best strategies using the second sub-period. Table 5 presents the in- and out-of-sample performance for the best-performing daily SBTRs selected from the first sub-period. The second column reports the qualitative characteristics behind these winning strategies including the types of technical indicators in use and the length of time adopted for the moving windows. We also report annualized mean returns, standard deviations, and p-values based on Hansen's (2005) test for each best-performing SBTR in the first subperiod. For the out-of-sample results, we report annualized mean returns, standard deviations, and nominal p-values of Hansen's (2005) test. Table 5 shows that the best-performing daily SBTRs generate significant profits among eight futures markets in the first subperiod at the 10% significance level. For the out-of-sample results, we the annualized mean returns, standard deviations, and t-statistics of those significant best-performing daily SBTRs selected in the first sub-period. We find that there are four out of eight futures markets where the best daily SBTRs selected from the first sub-period can generate significant profits in the out-of-sample period at 10% significance level.

Tables 6 and 7 report the in- and out-of-sample results for the weekly SBTRs with the MYR and CL exit strategies, respectively. In Table 6, there are five best-performing weekly SBTRs with the MYR exit strategy selected from the first sub-period that pass Hansen's (2005) SPA test at the 10% significance level. Most of the strategies cannot generate significant profits in the second sub-period. On the other hand, the weekly SBTRs with the CL exit strategy in Table 7 produce robust results in eighteen futures markets in the first sub-period, and in fourteen futures markets, the best-performing SBTRs selected from the first sub-period still generate profits in the second sub-period at the 10% significance level. The results of the in- and outof-sample experiments suggest that the weekly SBTRs with the CL exit strategy dominate the daily SBTRs and the weekly SBTRs with the MYR exit strategy.

We also perform a sub-sample analysis to examine the profitability of the best-performing SBTRs selected from the second sub-period. Table 8 presents the performance of the bestperforming daily SBTRs selected from the second sub-period. There are seven futures markets where the best-performing SBTRs are found to generate profits at the 10% significance level. Tables 9 and 10 report the sub-sample analyses for the weekly SBTRs with the MYR and CL exit strategies, respectively. Based on Hansen's (2005) nominal *p*-value, there are only four futures markets in which the best-performing weekly SBTRs with the MYR exit strategy generate profits at the 10% significance level during the second sub-period, and there are sixteen when the CL exit strategy is applied. The performances of the weekly SBTRs with the CL exit strategy are still the most robust based on the sub-sample analysis, which is consistent with the results of the data-snooping test in Table 4.

4.4 Transaction costs

Although testing the profitability of technical trading rules against index prices would seem to benefit from the data's longer available time series, testing against the futures markets generates more convincing results because the associated transactions costs are transparent and hence easy to control. Furthermore, the complications imposed by short-sale constraints (Brock, Lakonishok and LeBaron, 1990; Sullivan, Timmermann and White, 1999) can be avoided.

To analyze the impact of transaction costs on the best-performing SBTRs, we consider two types of transaction cost measures, namely, the break-even transaction costs and the applicable commissions per round-trip trade. For the break-even transaction costs, we follow Qi and Wu (2006) to calculate the maximum one-way transaction costs for the best-performing SBTRs to break even, i.e., to eliminate all possible returns or performance (Bessembinder and Chan, 1998). The first four columns of Table 11 report the annualized mean returns, the number of round-trip trades and the maximum one-way transaction cost of these strategies for the best daily SBTRs selected from the overall sample period. The length of the overall sample period is 11 years for copper, E-mini S&P, E-mini NASDAQ, EUR and AUD futures (Group 2) and 21 years for the rest of the futures markets (Group 1). Taking the corn futures as an example, the annualized mean return and the number of round-trip trades for the best daily SBTR are 24.29% and 5,266, respectively. The number of round-trip trades is equal to the number of trading days over the sample period. The maximum one-way cost to break even is 21*24.29%/(5,266*2)=0.0484%, where 5,266 multiplied by 2 is the number of trades. That is, if the one-way trading cost for an investor is greater than 0.0484%, the SBTR may not generate a positive return. To test the profitability of technical strategies in currency markets, Qi and Wu (2006) adopt one-way costs in the range of 0.025%-0.04% as suggested by Bessembinder (1994) who measures the bid-ask spread in the inter-bank market. We find that the range of the one-way cost is also applicable to the futures market, as Wang, Yau, and Baptiste (1997) and Wang and Yau (2000) also report a similar range of the one-way cost for seven futures markets. The fifth and sixth columns report the annualized mean returns after we consider the transaction costs of 0.025%and 0.04%, respectively. The daily SBTRs are still profitable if the transaction cost is 0.025%in thirteen of the nineteen futures markets, and the after-cost returns are still positive in eight futures markets when the transaction cost is assumed to be 0.04%.

Alternatively, we can consider the commission per round-trip trades to account for the transaction costs. Park and Irwin (2010) consider a range of commission costs of \$12.5-\$100 per futures contract per round-trip trade. The transaction cost of \$12.5 per round-trip is documented by Lukac and Brorsen (1990) who suggest that such a low transaction cost is possible because commissions through discount brokers are around \$12.5 and even lower for high volume trades or electronic trades. To convert the dollar commission costs to the percentage deducted, we first estimate the average prices of these futures markets during the overall sample period and calculate the average contract sizes as the average prices multiplied by the point value of the contract. Then we divide the one-way commission cost (which is the commission per round-trip trade divided by 2) by the average contract size to obtain the ratio of the one-way commission to average contract size. For example, since the average price of the soybean futures during the overall sample period is 846.92 and the point value is \$50, the average contract size is therefore \$42,346. If the commission per round-trip trade is \$12.5, the percentage of the one-way commission is 6.25/42,346=0.0148%, which is lower than the maximum one-way cost for the best SBTR to break even over the sample period. When the one-way commission is \$6.25, all the SBTRs still have positive returns because the ratios of one-way commission to average contract size are all lower than the maximum one-way cost for each futures market with the exception of the Eurodollar futures. However, when we assume a one-way commission of \$50, then none of the best SBTRs generate positive returns except for the S&P 500 futures.

Table 12 reports the performance of the best weekly SBTRs using the MYR exit strategy after the transaction cost is considered. The maximum one-way costs that are tolerable are much higher than those of the daily SBTRs in Table 11, since the trading frequency of the weekly SBTRs with the MYR exit strategy is much lower. Almost all strategies are profitable after we consider the one-way transaction costs in the range of 0.025%-0.04% or the one-way commission of \$6.25. When the one-way commission is \$50, there are thirteen futures markets where the ratio of the one-way commission to the average contract size is lower than the maximum one-way cost. Although the after-cost performance of the weekly SBTRs using the MYR exit strategy seems better than that of the daily SBTRs, the performance is not robust to the sub-sample and out-of-sample analyses.

Table 13 shows that the after-cost performance of the best weekly SBTRs with the CL exit strategy is similar to the result of the daily SBTRs as their trading frequencies are similar. In ten of the nineteen futures markets, the weekly SBTRs with the CL exit strategy generate positive profits under the assumption of a 0.04% one-way cost, and when the one-way cost is 0.025%, the SBTRs are profitable in thirteen futures markets. When the one-way commission is \$6.25, all strategies survive. However, when the one-way commission is \$50, none of the best SBTRs can generate positive returns except for the S&P 500 futures.

5 Conclusion

By devising the SBTR, this paper adds to the technical trading literature a novel prospective depicting how technical-trading decision making relates to similarity-based analogical reasoning. The SBTR, while allowing for a univariate setting as its degenerate case, is a multivariate technical trading strategy. In this paper, indicator values alone do not initiate a similaritybased trader's trading decisions. Yet the loss-gain experiences of past returns, which bring about pain and pleasure, have impacts on his/her decision making. His/her buy (sell) decisions are determined by the stochastic averaging predictor—an indicator-assisted forecasting process of probable future returns—whose positivity (negativity) triggers a buy (sell) signal.

Our key findings are as follows. First, both the daily SBTRs and the weekly SBTRs with the CL exit strategy are capable of generating profits in most of the futures markets at the 10% significance level even after we control for data-snooping bias. The performance of the weekly SBTRs with the CL exit strategy, in particular, is robust to the sub-period and out-of-sample tests. The results indicate the predictive ability of the SBTR. Second, winners are those who use SBTR as a multivariate strategy comprising trading indicators of different types. That is, to trade and win under reasoning by similarity/analogy seems to entail exploiting feasible combinations of the MA, RSI, TRB, and VMA indicators in the construction of the vector of qualitative characteristics. Third, we find that the trader's adopted time lengths of moving time windows in conducting similarity searches indeed matter. The optimal choice of time lengths for the moving window is seldom the distant past. Instead, the performance returns of the recent past play a major role in decision making process of the similarity-based trader. Fourth, transactions costs are as important as anticipated in prior studies. After considering a wide range of one-way transaction costs, most of the best-performing SBTRs remain profitable when the maximum one-way transaction cost to break even is 0.04%, or when the commission per round-trip trade is assumed to be \$6.25.

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