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Working Paper

Would you Marry me? The Effects of Marriage on German Couples? Allocation of Time

ZEW Discussion Papers, No. 07-024

Provided in Cooperation with:

ZEW - Leibniz Centre for European Economic Research

Suggested Citation: El Lahga, AbdelRahmen; Moreau, Nicolas (2007) : Would you Marry me? The Effects of Marriage on German Couples? Allocation of Time, ZEW Discussion Papers, No. 07-024, Zentrum für Europäische Wirtschaftsforschung (ZEW), Mannheim

This Version is available at:

<https://hdl.handle.net/10419/24584>

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Non-Technical Summary

Living arrangements have undergone considerable change in recent decades. In most Western countries marriage is no longer the exclusive context of family formation. In the former Federal Republic of Germany, about 38.3 percent of the women born between 1954 and 1958 started their first union outside of a formal marriage. The figure increases to 67.9 percent for the cohort 1964 - 1968 (Le Goff, 2002). These demographic trends challenge the microeconomic literature in which couples living in consensual unions are implicitly assumed to act exactly as married couples.

A closer look at the literature reveals, however, growing evidence of the link between marital status and household behavior with respect to many outcomes. For instance, DeLeire and Kalil (2005) find that cohabiting-parent families spend a greater amount on alcohol and tobacco, and a smaller amount on education and health care than married-parent families. Cohen (2002), Mamun (2004) and Stratton (2002) examine the wage differential for married and cohabiting men. Higher wages are observed for married men, *ceteris paribus*. There is also evidence, based on US cross section data, that married couples exhibit a more 'traditional' division of domestic and market work than cohabitants (South and Spitze 1994, Stratton, 2005).

The purpose of this paper is to examine whether the shift from cohabitation to marriage is associated with a significant change in household market and non-market labor supply. More specifically, we use a long German panel (GSOEP) to test whether the transition from cohabitation to marriage reinforces the degree of specialization among couples. We estimate a model that relates married life to the female-to-male domestic and market work hours log ratios. Other regressors of the log ratios are the female relative earnings, the number of children and the duration of the relationship. We account for selection bias in the presence of endogenous regressors following the procedure advocated in Semykina and Wooldridge (2005), that we adapt to system GMM estimation.

Our results suggest that marriage increases female specialization in home-based activities. Importantly, marriage leads to a fall in women's leisure, particularly for couples with pre-school children. The results also exhibit a fall in married men's leisure coming from either a rise in market hours or an increase in domestic hours depending on the specification.

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1 Introduction

Living arrangements have undergone considerable change in recent decades. In most Western countries marriage is no longer the exclusive context of family formation. In the United States (US) for instance, the number of unmarried couples nearly doubled in the 1990s, from 3.2 million couples in 1990 to 5.5 million couples in 2000 (source: U.S. Census Bureau). Le Goff (2002) reports that in the case of French women born between 1944 and 1948, 22 percent started their first union as a cohabiting union. For the cohort 1964 – 1968, this applies to 81 percent. In the former Federal Republic of Germany, about 38.3 percent of the women born between 1954 and 1958 started their first union outside of a formal marriage. The figure increases to 67.9 percent for the cohort 1964 – 1968 (Le Goff, 2002). These demographic trends challenge the microeconomic literature in which couples living in consensual unions are implicitly assumed to act exactly as married couples.

A closer look at the literature reveals, however, growing evidence of the link between marital status and household behavior with respect to many outcomes. Using data from the US consumer expenditure survey, DeLeire and Kalil (2005) find that cohabiting-parent families spend a greater amount on alcohol and tobacco, and a smaller amount on education and health care than married-parent families. Cohen (2002), Mamun (2004) and Stratton (2002) examine the wage differential for married and cohabiting men. Higher wages are observed for married men, *ceteris paribus*. McConocha et al. (1993) compare financial decisions between cohabiting, remarried and married couples. More recently, Rangel (2006) uses a sample of married Brazilian couples as a control group to identify the effect of extending alimony rights to cohabitants on adults' time allocation and investments in the education of children. There is also evidence, based on US cross section data, that married couples exhibit a more 'traditional' division of domestic and market work than cohabitants (South and Spitze 1994, Stratton, 2005).¹

The purpose of this paper is to examine whether the shift from cohabitation to marriage is associated with a significant change in household market and non-market labor supply. More specifically, we use a long German panel (GSOEP) to test whether the transition from cohabitation to marriage reinforces the degree of specialization among couples. We estimate a model that relates married life to the female-to-male domestic and market work hours log ratios. Other regressors of the log ratios are the female relative earnings, the number of children and the duration of the relationship. We account for selection bias in the presence of endogenous regressors following the procedure advocated in Semykina and Wooldridge (2005), that we adapt to system GMM estimation.

Our results suggest that marriage increases female specialization in home-based activities. Importantly, marriage leads to a fall in women's leisure, particularly for couples with pre-school children. The results also exhibit a fall in married men's leisure coming from either a rise in market hours or an increase in domestic hours depending on the specification.

The paper is structured as follows. Section 2 exposes the theoretical underpinning for the effects of marriage and cohabitation on market and non-market labor supply. Section 3 discusses the empirical specifications. Section 4 exposes the econometric issues. The data are described in Section 5. The results are presented in Section 6, and Section 7 concludes.

¹Waite (1995) offers a survey of the effects of marriage and cohabitation on health, mortality, children's well-being, assets and labor.

2 Theoretical background

The economic motivations which lie behind the existence of the household have been extensively discussed in the literature. At least since Becker (1973), it has been commonly argued that one of the reasons for household formation is that it allows household members to specialize efficiently on activities in which each has a comparative advantage. One partner can specialize in non-market household activities while the other specializes in market work. The distinction between legal marriage and consensual union is not formally stated and the word “spouses” usually refers to two individuals living together. However, several arguments exist in the economic literature to predict that cohabitants specialize less than married couples.

First, cohabitants are often seen as playing non-cooperatively (Nordblom, 2004 and references therein). Cohabitations are usually shorter lived than marriages (Brien et al., 2006; Bumpass and Sweet, 1995) and there is consensus to admit that cooperation is more likely to occur in stable couples, committed in a long term relationship. Stratton (2005) also puts forward the hypothesis that specialization is closely related to perceived household stability. Using US data, she presents some empirical evidence that the degree of specialization is greater within married couples compared to cohabitants. In contrast to cooperative settings, efficient specialization is less likely for couples playing strategically. For instance, Lundberg (2002) considers a bargaining model of intrahousehold allocation in a multi-period setting with limited commitment. Members are unable to make credible promises regarding future behavior.² Inefficient levels of specialization and underprovision of household public good are likely outcomes. Basu (2006) specifies a bargaining model with an endogenous balance of power between partners and no intertemporal commitment. He shows that strategic considerations can lead to inefficient outcomes.

Second, consensual unions offer less legal protection than marriage. Married spouses are often obliged to care for each other and spousal maintenance is anticipated after divorce. Cohabitation provides individuals with less risk-sharing opportunities than marriage. That may prevent individuals from specializing in home-based activities and household production skills. Cohabitants are thus less likely to specialize in household specific human capital. In this vein, Nordblom (2004) considers a model where married couples have legal restrictions on their relationship that force them to act cooperatively, while cohabitants with limited commitment act non-cooperatively. This makes precautionary savings greater for cohabitants than for married couples.³ In Germany, the article 6 of the Constitution obliges the state to promote the institutions of marriage and the family through its legislation and to prevent any situation which could disadvantage these institutions (Stintzing, 1999). One spouse is obliged to support the other before the latter is entitled to subsidies from the state and this support is tax-deductible. This is not the case for cohabitants. In addition, the economic consequences of partnership dissolution are different for married couples and cohabiting couples. The German Constitution does not impose maintenance payments after non-marital separation. However, child support payment is anticipated. In 1994, suggestions to extend the protection of marriage to any form of long-term cohabitation were not approved by Parliament.

Finally, income tax distorts the allocation of time between married and cohabiting couples whenever they are subjected to different income tax schedules. In Germany, married couples can opt for the splitting system. Spouses’ income is aggregated and halved, and the tax schedule is applied to this tax base. Married couples thus profit from a more

²See also Wells and Maher (1998).

³Kotlikoff and Spivak (1981) and Anderberg (2003) also study risk sharing between spouses.

favorable taxation in the case of an asymmetric earnings situation between the spouses (Gustafsson, 1992). As a result, tax saving is maximized for one earner households, or if partners combine full-time/part-time employment. From a theoretical perspective, Wrede (2003) analyzes among others things, the effect of joint taxation on specialization. Under the assumption that partners allocate their time efficiently between market and non-market activities through a Nash bargaining process, he shows that family members specialize more in reaction to joint taxation. Most importantly, only married couples can opt for the splitting system in Germany, while cohabitants face individual taxation. To the extent that it implies higher marginal tax rates on the first earner and lower marginal tax rates on the secondary earner, cohabitants have fewer incentives for specialization, *ceteris paribus*.

3 Empirical Specifications

We consider a household i consisting of a female (f) and a male (m), that makes decisions about market work, non-market work, marital status and consumption. Let *Married* be a dummy variable denoting the marital status of the couple, with *Married* = 1 if married and 0 if cohabiting. Each partner p ($p = m, f$) offers $l^p > 0$ hours on the labor market at wage rate w^p and spends $h^p > 0$ hours in domestic work. Following Pollak and Wachter (1975) and Kooreman and Kapteyn (1987), we model the demand for consumption and the allocation of time by assuming that households or individuals maximize a utility function with goods and time spent on market and non-market activities as arguments. The allocation of time among market and non-market activities is thus expressed as a function of prices, wage rates and non-labor incomes.

There are several ways of taking marital status into account in the utility function. In Couprie (2007), Gray (1997) and Lundberg and Rose (2002), to name a few, marital status is a predetermined variable. Current shocks on labor supply do not influence current marital status. The utility function at time t is a function of current marital status but is maximized with respect to consumption goods and leisure only. Also, marital status can be seen as a preference parameter that may vary over time. Couples may move from cohabitation to marriage and this change is likely to modify the allocation of time among market and non-market activities but marital status is not a choice variable per se. However, it could be that marital status and the allocation of time are interrelated choices. Van der Klaauw (1996) explicitly studies their interdependence. In a life-cycle setting, the interaction between female labor force participation (not hours) and marital decision (married or single) is examined. Participation and marital status are the choice variables over which the utility function is maximized. Brien et al. (2006) estimate a model of non-marital cohabitation, marriage and divorce. They examine union formation and dissolution in the presence of uncertain match quality. They do not consider labor supply issues. Brien et al. (1999) focus on the interrelationship of cohabitation, marriage and childbearing while Lillard et al. (1995) analyze premarital cohabitation and the stability of later marriages.

In this paper, we estimate models with marriage assumed predetermined and models with endogenous marriage. We do not specify a structural model for the interaction between marital status and time use but take this relation in reduced form via instrumental variables use in the market and non-market hours equations.

Also, we ignore the issue of union formation and dissolution. Our estimates may then suffer from selection bias but this problem is general to the labor supply literature. The analysis is usually done conditionally on household formation. We follow the bulk

of the literature on this matter.⁴ Our results must be interpreted conditionally on couple formation.

Finally, conditioning the analysis on couples with strictly positive labor supplies also raises the problem of selectivity bias. We test and correct for it.

We now present our base model. To analyze the extent of specialization within households, we specify a two-equation system that relates marriage, relative earnings and a set of preference factors to the female-to-male domestic and market work hours log ratios.

3.1 The domestic and market work hours (log) ratios

The base specification is one that is often encountered in the literature. It was already used in a unitary framework (Kiker and Ng, 1990) as well as in a collective setting (Browning and Gørtz, 2006)⁵ and explains the shares of domestic and market work:

$$\begin{aligned}\ln(r_{it}^h) &= \alpha_h \text{Married}_{it} + \beta_h \ln\left(\frac{w^f}{w^m}\right) + \gamma_h Y_{it} + \delta'_h Z_{it} + \varepsilon_{it}^h \\ \ln(r_{it}^l) &= \alpha_l \text{Married}_{it} + \beta_l \ln\left(\frac{w^f}{w^m}\right) + \gamma_l Y_{it} + \delta'_l Z_{it} + \varepsilon_{it}^l,\end{aligned}\quad (1)$$

where $r^h = \frac{h^f}{h^m}$ is the female-to-male domestic (home) work hours ratio, $r^l = \frac{l^f}{l^m}$ the female-to-male labor market work hours ratio, Y is the household non-labor income, Z is a vector of household characteristics including the constant term, ε is a conformable error term, and α , β , γ and δ are the parameters to be estimated. The subscript i indicates household and t indicates time.

Now we turn to the expected signs of the variables included in our empirical specification. As pointed out above, we expect a negative impact of marriage on women's relative hours on paid work (r^l) and a positive effect on women's relative hours on unpaid work (r^h). The overall effect on female relative leisure is undetermined a priori. Relative earnings ($\frac{w^f}{w^m}$) are expected to be positively related to r^l and negatively to r^h . Again, the overall effect on relative leisure is undetermined. Non-labor income has a positive effect on partners' leisure (if leisure is a normal good) but its impact on r^h and r^l is ambiguous.

We expect the specialization in home-based activities to increase with the duration of the relationship. To capture this effect we include a series of dummy variables $Dur2$, $Dur3$, $Dur4$ and $Dur5$ indicating the relationship duration in number of years.⁶ Also, the number and age of children are likely to influence the extent of specialization within the family. In line with the effect of children on female labor supply documented in the literature, we expect children to have a positive incidence on r^h and a negative impact on r^l . We include the number of children under five and the number of children older than four in our specification.

One might argue that time allocation within the household is sensitive to generational effects. More precisely, younger cohorts might exhibit a more equal division of domestic work and paid labor. To test this hypothesis we include three cohort dummies in our model. Other factors such as nationality, regional disparities and religion may influence

⁴See Blundell et al. (1998), Chiappori et al. (2002), and Pencavel (2006) among others.

⁵In the unitary framework, it is assumed that households, irrespective of the number of household members, behave as single decision makers. The collective setting introduced by Chiappori (1988) and Apps and Rees (1988) takes into account several decision makers and the bargaining process. See also the seminal paper of McElroy and Horney (1981) for a Nash-bargaining model applied to household behavior.

⁶ $Dur2 = \mathbf{1}[5 \leq \text{duration} < 10]$, $Dur3 = \mathbf{1}[10 \leq \text{duration} < 15]$, $Dur4 = \mathbf{1}[15 \leq \text{duration} < 20]$ and $Dur5 = \mathbf{1}[\text{duration} \geq 20]$.

the intrafamily allocation of time. The dummy variable *German00* is defined as 1 for non-German couples and 0 otherwise, *German10* takes on the value 1 for couples with a German male partner and a non-German female partner whereas *German01* = 1 for couples with a non-German male partner and a German female partner. Our control group is German couples. Included regions of residence are the southern and middle states of West Germany, Baden-Württemberg, Rheinland Pfalz, Saarland and Bavaria. In contrast to the Protestant northern states, these regions include a majority of Catholics.⁷

3.2 Alternative formulations

We also focus on the effect of marriage on the allocation of time between domestic and market work. We first consider a two-equation system that relates marital status, the female wage and the female unearned income (household non-labor income together with male labor earnings) to the female non-market and market labor supply. This setting is consistent with the theoretical models that assume the male allocation of time to be rigid.⁸ Actually, as exposed in Section 5, German men exhibit little variation in their number of domestic and market hours. This alternative specification is:

$$\begin{aligned} h_{it}^f &= \alpha_l \text{Married}_{it} + \beta_l \log(w_{it}^f) + \gamma_l Y_{it}^f + \delta'_l Z_{it} + \varepsilon_{it}^l \\ l_{it}^f &= \alpha_h \text{Married}_{it} + \beta_h \log(w_{it}^f) + \gamma_h Y_{it}^f + \delta'_h Z_{it} + \varepsilon_{it}^h, \end{aligned} \quad (2)$$

where Y^f is the female unearned income. Other variables are the same as before. The female wage is expected to have a positive effect on market hours and a negative effect on domestic hours. The female unearned income is expected to reduce women's hours in paid work.

Although men's labor supply is more concentrated than women's, there is some variation in the male allocation of time. We then specify a four-equation model where the male market and non-market labor supply are supposed to be flexible. Formally, we estimate the following system:

$$\begin{aligned} h_{it}^f &= \alpha_{lf} \text{Married}_{it} + \beta_{lf} \log(w_{it}^f) + \eta_{lf} \log(w_{it}^m) + \gamma_{lf} Y_{it} + \delta'_{lf} Z_{it} + \varepsilon_{it}^{lf} \\ l_{it}^f &= \alpha_{hf} \text{Married}_{it} + \beta_{hf} \log(w_{it}^f) + \eta_{hf} \log(w_{it}^m) + \gamma_{hf} Y_{it} + \delta'_{hf} Z_{it} + \varepsilon_{it}^{hf} \\ h_{it}^m &= \alpha_{lm} \text{Married}_{it} + \beta_{lm} \log(w_{it}^f) + \eta_{lm} \log(w_{it}^m) + \gamma_{lm} Y_{it} + \delta'_{lm} Z_{it} + \varepsilon_{it}^{lm} \\ l_{it}^m &= \alpha_{hm} \text{Married}_{it} + \beta_{hm} \log(w_{it}^f) + \eta_{hm} \log(w_{it}^m) + \gamma_{hm} Y_{it} + \delta'_{hm} Z_{it} + \varepsilon_{it}^{hm}, \end{aligned} \quad (3)$$

that corresponds to a general unrestricted model of household behavior. The covariates are the same as those of the base specification.

4 Econometric issues

In this section we discuss econometric issues that arise with our specification and we present our estimation method. We draw heavily upon Semykina and Wooldridge (2005). For the sake of simplicity, we consider one equation of interest to be estimated. All results can be easily generalized to a system of equations. A formal derivation is in Appendix A. We have:

$$y_{it} = x_{it}\beta + \varepsilon_{it}, \quad (4)$$

⁷Religious preferences are asked in the GSOEP but contain a lot of missing values.

⁸Blundell et al. (2005), Donni (2007) and Donni and Moreau (2007) deal with the rigidity of the man's behavior in a collective framework.

where x_{it} is a $1 \times K$ vector of explanatory variables, β is a $K \times 1$ vector of parameters to be estimated, and ε_{it} is the error term.

First, we control for household-specific fixed effects c_i , which captures all unobserved household-specific heterogeneity in the labor supply that remains constant over time. The error term is then expressed as $\varepsilon_{it} = c_i + u_{it}$, where u_{it} are the idiosyncratic errors.

We allow for arbitrary correlation between the unobserved effect and the explanatory variables. In addition, we allow some elements of x_{it} to be endogenous (that is, correlated with the idiosyncratic error, u_{it}). Let z_{it} a $1 \times L$ ($L \geq K$) vector of instruments which are strictly exogenous conditional on c_i .

As previously mentioned, we focus our analysis on a sample of couples with strictly positive labor supply. Let s_{it} a binary selection rule that takes on the value 1 if the couple exhibits strictly positive market and non-market labor supply at period t , and 0 otherwise. Whether s_{it} equals 1 or 0, x_{it} and z_{it} are always observed.

Semykina and Wooldridge (2005) –SW hereafter– show that applying the usual fixed effect two-stage least squares (FE-2SLS) estimator to the selected sample yields consistent estimates if s_{it} is completely random—technically s_{it} is independent of (u_{it}, z_{it}, c_i) in all periods—or if s_{it} is a deterministic function of (z_i, c_i) , where $z_i = (z_{i1}, \dots, z_{iT})'$. Thus, in order to obtain consistent estimates one should carry out a formal test for sample selection and apply a correction method if necessary.

In what follows, we briefly sketch the procedure proposed by SW for linear fixed effects models.⁹ The selection indicator s_{it} is generated by means of a latent variable s_{it}^* such that:

$$s_{it} = \mathbf{1}[s_{it}^* > 0] = \mathbf{1}[z_{it}\delta + a_i + \tau_{it} > 0],$$

where $\mathbf{1}[\cdot]$ is the indicator function, a_i is an unobserved effect and τ_{it} is an error term defined such that $\tau_{it} \mid z_i, e_i \sim N(0, 1)$, so that s_{it} follows an unobserved effect probit model. To allow a_i to be correlated with z_i , SW specify, following Mundlak (1978),

$$a_i = \eta + \xi \bar{z}_i + f_i, \quad (5)$$

where \bar{z}_i is a vector of individual exogenous variables averaged across periods of time, and $f_i \mid z_i$ has a zero mean normal distribution. Hence, the selection indicator s_{it} is rewritten as:

$$s_{it} = \mathbf{1}[\eta + z_{it}\delta + \xi \bar{z}_i + v_{it} > 0], \quad (6)$$

where $v_{it} = (f_i + \tau_{it})$ has a zero mean normal distribution.

Now, suppose that (ε_{it}, v_i) is independent of (z_i, c_i) , where $v_i = (v_{i1}, \dots, v_{iT})'$, and $(\varepsilon_{it}, v_{it})$ is independent of $(v_{i1}, \dots, v_{i,t-1}, v_{i,t+1}, \dots, v_{iT})$. If $E(\varepsilon_{it} \mid v_{it})$ is linear, then:

$$E(\varepsilon_{it} \mid z_i, c_i, s_i) = \rho E(v_{it} \mid z_i, c_i, s_i) = \rho E(v_{it} \mid z_i, s_{it}). \quad (7)$$

Under the previous assumptions, we can write the primary equation (4) as:

$$y_{it} = x_{it}\beta + c_i + \rho E(v_{it} \mid z_i, s_{it}) + e_{it}, \quad (8)$$

where e_{it} is an idiosyncratic error term verifying $E(e_{it} \mid z_i, c_i, s_i) = 0$ by construction.

As noted above, the FE-2SLS yields a consistent estimation of the parameters of interest if the expectation given by (7) is 0. Then, an immediate test for sample selection bias is obtained by testing $\mathbf{H}_0 : \rho = 0$ in (8), which can be estimated by FE-2SLS. For the

⁹Related papers are Gonzalez-Chapela (2004), Dustmann and Rochina-Barrachina (2000) and Kyriazidou (1997).

selected sample (that is, $s_{it} = 1$) we only need $E(v_{it} | z_i, s_{it} = 1)$, which can be obtained from the usual probit estimation:

$$E(v_{it} | z_i, s_{it} = 1) = \lambda(\eta + z_{it}\delta + \xi\bar{z}_i), \quad (9)$$

where $\lambda(\cdot)$ denotes the inverse Mill's ratio. Let $\hat{\lambda}_{it}$ denote a consistent estimate. To test for selection one simply has to estimate $P(s_{it} = 1 | z_i) = \Phi(\eta + z_{it}\delta + \xi\bar{z}_i)$ with a reduced-form probit at each period t , to plug $\hat{\lambda}_{it}$ into the primary equation, to estimate the augmented primary equation by FE-2SLS and to test for $\mathbf{H}_0 : \rho = 0$ with a t -statistic. To add more flexibility to the model, it is possible to interact $\hat{\lambda}_{it}$ with time dummies to allow the coefficient ρ to be different across t . In this case, a Wald statistic is used to test the joint significance of the T coefficients ρ_t . In our empirical specification, we use a FE-GMM estimator instead of the FE-2SLS and allow for different ρ_t .

SW offer a correction procedure for the sample selection problem when the null is rejected. It amounts to estimating equation (8) by Pooled 2SLS using a decomposition of the household-specific effect c_i that follows Mundlak (1978). Under the previous assumptions about the selection rule and the unobserved effects, the primary equation of interest (4) can be rewritten as:

$$y_{it} = \alpha + x_{it}\beta + \bar{z}_i\xi + \gamma_t E(v_{it} | z_i, s_{it}) + e_{it}. \quad (10)$$

SW show that applying the Pooled 2SLS estimator to (10) after replacing $E(v_{it} | z_i, s_{it})$ by the estimated inverse Mill's ratio $\hat{\lambda}_{it}$ yields a consistent estimator of the parameters.

We adapt the procedure presented in SW to system estimation and propose a more efficient GMM estimator (Pooled GMM hereafter). Its derivation and the computation of the asymptotic variance estimator that accounts for the effect of using estimated rather than actual values for λ_{it} are collected in Appendix A.

Now, we address the question of the endogeneity of the regressors and the choice of the instruments. It is likely that hourly earnings and household non-labor income are not independent of hours worked. Therefore, we have chosen to instrument the woman's wage rate, the man's wage rate and the household non-labor income. One might also argue that the effect of marriage on labor supply can not be distinguished from the effect of pre-school children on parental time use. Indeed, the presence of children is more frequent among married couples in comparison with cohabitants who may enter marriage to begin childbearing or to legitimate the birth of a child. Child dependency on their mothers (breastfeeding for example) coupled with the virtual absence of child-care facilities in West Germany for small children¹⁰ create a strong incentive for specialization in conjunction with motherhood. To limit the extent of this problem, we focus on observations with no children under two. We account for this potential endogenous selection rule in our estimation procedure but assume that older children can be regarded as strictly exogenous after conditioning on the unobserved effect. This approach is commonly used in the literature (for instance Lundberg and Rose, 2002). Moreover, given that we ignore the issue of union formation and dissolution, we also consider the duration of the relationship to be strictly exogenous once we condition on the unobserved effect. Marital status will be either exogenous or endogenous depending on the estimations.

The set of excluded instruments that do not appear in the labor supply equations consists of the following variables: male and female years of schooling and their squares,

¹⁰In West Germany full-day child-care is rare (Deutsches Jugendinstitut, 2002). Approximately 5 percent of children under age three are enrolled in formal child-care. Among three-five years old, 74.6 percent are enrolled in kindergarten (Gornick and Meyers, 2003).

male and female age and their squares, product of partners' age and education, and time dummies. Our intuition is that these variables have an impact on the various sources of the household income.¹¹ Therefore, there are 26 excluded instruments from the labor supply equations.

Finally, it is important to have at least one instrument that affects only the selection equation, otherwise the parameters of the primary equation are identified through the nonlinearity in the inverse Mill's ratios. We use the female and male unemployment rates as exclusive regressors for the probit model.

5 The data

Our data is drawn from the first 21 waves of the German Socio-Economic Panel (GSOEP) for the years 1984-2004. We extract a sample of observations that correspond to couples living in the former Federal Republic of Germany who are between 25 and 55 years old, who have finished their education and are available for the labor market. Households where adults are retired or students are excluded. We also exclude households where adults are self-employed. Excluded from the sample are also couples who gave incomplete or incoherent information. Finally, when an individual appears in more than one couple, we select the one with the longest duration in the data. Therefore, to each individual corresponds only one household. In all, these selection criteria lead to a sample of 4,762 couples resulting in 28,167 observations. On average, a couple appears six times over the 21 years period (the minimum and maximum are 2 and 21, respectively). Of the 4,762 couples in our sample, 3,796 (79.71 percent) are always married, 364 (7.64 percent) always cohabiting and 602 (12.65 percent) make the transition from a cohabiting to a married couple. The couples who always live in a consensual union are observed for a shorter number of periods. Their average number of waves is four within our observation window.

Also, to estimate the labor supply systems (1), (2) and (4) conditional on fixed effects, we need at least two observations for the same couple. Therefore, the data we use for estimating these equations uses all couples without children under two who report, for both partners, a strictly positive amount of domestic work and market work in at least two periods. This leaves us with 12,925 observations from 2,762 households.

Measuring time use

Time spent on non-market work is computed as the sum of hours spent on housework (washing, cooking and cleaning), childcare, gardening and repairs in a typical weekday. Not all domestic work time is covered by this variable as weekend non-market work is not included.¹² However, we probably account for a larger set of activities than that used when the question is only about time spent on housework (and not about childcare) in a normal week.

We measure time spent in market work as the annual work hours on all jobs divided by 365 (and by 366 for leap years). It corresponds to an average number of hours worked per day. It means that market work and non-market work have the same unit of time.

Measuring earnings

Total labor earnings include wages and salaries from all jobs including overtime and secondary jobs. The wage rate is the average hourly earnings defined by dividing total labor income over annual hours of work on all jobs. Non-labor income includes income

¹¹These are standard instruments in the literature. See for instance Chiappori et al. (2002) among others.

¹²Complete information on weekend domestic work is asked only with years 93, 95, 97, 99, 01 and 03.

from assets, rent, private transfers, public transfers, etc. All these income variables are in euros adjusted for inflation with the price index provided by the GSOEP. Non-labor income is in thousands of euros.

Measuring marital status

The couple’s marital status is represented by the dummy variable *Married* that takes on a value of 1 if the partners are legally married at the time of the interview, and 0 if they are cohabiting.¹³

Measuring duration of conjugal life

Duration of conjugal life can be computed from a retrospective data file on marital history that contains yearly information on marital status. The data include the beginning and the ending of each marital status spell.

Table 1: Descriptive Statistics for the Whole Sample

	Total	Marital status		
		AM	CM	AC
Age (<i>f</i>)	38.70 (7.58)	39.69 (7.39)	33.88 (6.43)	35.91 (7.85)
Age (<i>m</i>)	41.21 (7.77)	42.33 (7.53)	35.98 (6.53)	37.35 (7.74)
Market work participation (<i>f</i>)	0.66 (0.47)	0.62 (0.48)	0.80 (0.40)	0.88 (0.33)
Market work participation (<i>m</i>)	0.95 (0.22)	0.95 (0.22)	0.96 (0.20)	0.92 (0.27)
Non-market work participation (<i>f</i>)	0.99 (0.08)	0.99 (0.07)	0.99 (0.11)	0.98 (0.14)
Non-market work participation (<i>m</i>)	0.87 (0.34)	0.87 (0.34)	0.90 (0.30)	0.85 (0.35)
Number of children (<i>f</i>)	1.37 (1.16)	1.54 (1.15)	0.72 (0.91)	0.44 (0.77)
Education level (<i>f</i>)	10.99 (2.41)	10.74 (2.33)	12.04 (2.35)	12.27 (2.67)
Education level (<i>m</i>)	11.44 (2.56)	11.27 (2.51)	12.26 (2.68)	12.00 (2.61)
Duration of conjugal life	16.98 (9.03)	18.96 (8.61)	10.26 (5.63)	7.44 (5.79)
Number of observations	28,167	22,882	3,891	1,394

Notes: Sample standard deviations are in parentheses below sample means. “AM” refers to couples that are always married, “CM” to couples that transit from consensual union to marriage and “AC” to couples that are always cohabiting. The duration of conjugal life is the maximum duration observed per couple. Its average is computed over the 4,762 couples, not over the 28,167 observations.

Summary statistics for the whole sample that includes nonparticipation in the labor market and in the house are given in Table 1. Sample characteristics are classified by marital status. The average age for men is 41 years and 39 years for women. From the 28,167 observations, 34 percent are women that do not participate in the labor market and 13 percent are men that do not work in the house. Of the 4,762 females (males) in our sample, 47.84 (88.70) percent always participate in the labor market and 18.5 (1.83)

¹³The distinction between legal marriage and cohabitation is made in the original data file, not in the GSOEP cross-national equivalent file that is often used by researchers.

percent never. Only 1.81 percent of the men never work in the house.¹⁴

Women who live in a consensual union participate more in the labor market than married women. On average, cohabiting couples are younger, have fewer children and are shorter lived than married couples. These findings could illustrate the transitional status of cohabitation but could also result from the composition of our sample. Cohabitation is indeed increasing over time.¹⁵ The variables related to consensual unions are thus more likely to suffer from right censoring.

Table 2: Descriptive Statistics for two-earner households without young children where both partners do some housework

	Total	Number of children				Marital status		
		0	1	2	3	AM	CM	AC
Age (f)	39.15 (7.39)	37.02 (8.84)	40.19 (6.97)	40.13 (5.75)	40.53 (5.38)	40.52 (6.96)	34.27 (6.62)	35.53 (7.66)
Age (m)	41.70 (7.56)	39.27 (8.76)	42.69 (7.16)	42.87 (6.01)	43.87 (5.75)	43.24 (7.04)	36.45 (6.79)	37.16 (7.40)
Domestic hours (h^f)	5.21 (3.41)	2.67 (1.63)	5.91 (3.16)	6.84 (3.50)	7.24 (3.28)	5.57 (3.31)	4.44 (3.73)	2.91 (2.46)
Domestic hours (h^m)	2.69 (1.85)	2.02 (1.24)	2.94 (1.99)	3.07 (1.96)	3.21 (2.17)	2.77 (1.87)	2.58 (1.81)	2.08 (1.52)
Domestic hours ratio ($r^h = \frac{h^f}{h^m}$)	2.45 (2.06)	1.61 (1.19)	2.69 (2.11)	2.99 (2.34)	3.13 (2.41)	2.61 (2.08)	2.07 (2.06)	1.63 (1.39)
Market hours (l^f)	4.31 (1.65)	5.29 (1.28)	4.03 (1.56)	3.66 (1.59)	3.62 (1.61)	4.15 (1.63)	4.73 (1.63)	5.22 (1.44)
Market hours (l^m)	5.93 (1.11)	5.95 (1.13)	5.92 (1.08)	5.93 (1.13)	5.93 (1.13)	5.93 (1.08)	5.94 (1.19)	5.98 (1.28)
Market hours ratio ($r^l = \frac{l^f}{l^m}$)	0.77 (0.42)	0.94 (0.41)	0.72 (0.39)	0.66 (0.41)	0.65 (0.37)	0.74 (0.40)	0.85 (0.45)	0.93 (0.44)
Wage rate (w_f)	11.27 (4.55)	12.13 (4.35)	11.22 (4.57)	10.54 (4.59)	10.55 (4.61)	11.06 (4.57)	11.91 (4.38)	12.16 (4.52)
Wage rate (w_m)	15.01 (4.78)	14.84 (4.79)	15.02 (4.81)	15.31 (4.74)	14.71 (4.72)	15.11 (4.76)	14.73 (4.74)	14.61 (5.11)
Relative wage rate ($\frac{w_f}{w_m}$)	0.81 (0.41)	0.88 (0.42)	0.80 (0.41)	0.74 (0.38)	0.77 (0.39)	0.78 (0.39)	0.88 (0.46)	0.90 (0.44)
Share of total work (f)	0.52 (0.07)	0.50 (0.07)	0.52 (0.07)	0.53 (0.08)	0.54 (0.07)	0.52 (0.07)	0.51 (0.08)	0.50 (0.07)
Share of domestic work (f)	0.64 (0.15)	0.56 (0.15)	0.66 (0.14)	0.68 (0.14)	0.69 (0.14)	0.65 (0.15)	0.59 (0.16)	0.55 (0.15)
Share of market work (f)	0.41 (0.12)	0.47 (0.09)	0.39 (0.11)	0.37 (0.12)	0.37 (0.12)	0.40 (0.12)	0.43 (0.11)	0.46 (0.09)
Number of observations	12,925	4,325	3,821	3,642	1,137	9,914	2,184	827

Notes: Sample standard deviations are in parentheses below sample means. “AM” refers to couples that are always married, “CM” to couples that transit from consensual union to marriage and “AC” to couples that are always cohabiting.

¹⁴These percentages are computed over the corresponding 28,167 observations.

¹⁵About 4.32 (9.19) percent of the households in our sample live in a consensual union in 1984 (2004). The highest proportion is observed for 1995 with 12.32 percent of cohabiting couples.

Table 2 reports statistics on the sample of households with no children under two and with both partners working in the labor market and in the house. The reported sample characteristics are classified by the number of children in the household and by marital status. On average, women work more hours in the house than men but less in the labor market. The variability of hours is lower for men: from the 12,925 observations, 77 (38) percent are men (women) that work at most three hours in the house per day and 75 (46) percent are men (women) that work between 35 and 45 hours per week in paid jobs.

The domestic workload increases with the number of children. Also, married women work more in the house than cohabiting women and less in the labor market. In all, the female share of “total” work (total work is market work together with domestic work) is slightly greater than 50 percent for married women and for couples with children. It seems that men do not fully compensate for the fewer hours of domestic work by working more in the market when they have children. Finally, women have a lower average wage rate than men.

6 Results

6.1 The effects of marriage on the domestic work and market work hours (log) ratios

Conditioning the sample on households with working partners (that is, $h^f > 0$, $l^f > 0$, $h^m > 0$ and $l^m > 0$) and no children under two years of age may induce a selectivity bias. To account for all these selection rules we estimate 21 reduced-form participation equations and include the 21 inverse Mill’s ratios into the market and non-market work equations. The results show a significant effect of the unemployment rates.¹⁶ Hence the parameters of the domestic work and the market work equations, which exclude the latter variables, are non-parametrically identified.

The estimates of the base model (1) with exogenous marriage and using the fixed effect estimator FE-GMM are shown in Table 3. We report a Wald statistic to test the overall significance of the 21 inverse Mill’s ratios added to test for selection bias.

At the five percent level, there is statistically significant evidence of selection bias for the log ratio of domestic hours but not for the log ratio of market hours. These results seem contradictory as our selection rule affects mainly couples with women that do not work in the labor market. Consequently, we decide to correct for contemporaneous selection in the equation related to domestic work but also in the equation related to market work.¹⁷ For this purpose, we use the Pooled GMM estimator that models the unobserved effect as a linear combination of the time averages of the exogenous variables. In comparison with the preceding estimation, we add eight time-constant regressors to explain the log ratios. Their effect could not be identified with the FE-GMM estimator.

The results are shown in Table 4. We present two Wald statistics to test the joint significance of the 13 coefficients related to the time averages of the exogenous variables and to test the joint significance of the 21 inverse Mill’s ratios.

Columns 1 and 2 in Table 4 report the estimates when marriage is assumed to be exogenous. At the five percent level, the time averages used to model the unobserved effect are jointly significant for both equations. Like the estimates obtained with the FE-GMM

¹⁶To save space the 21 probits are not reported. Results are available upon request.

¹⁷We also estimate the model that accounts for selection bias only for the female relative domestic workload. The estimates are very similar.

Table 3: The effects of married life on the female-to-male domestic and market work hours (log) ratios - FE-GMM estimates (Base specification, Model 1)

	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$
Married	0.112	-0.084
	(0.055)	(0.032)
Log wage ratio	-1.74	1.008
	(0.432)	(0.291)
Non-labor income	-0.002	-0.000
	(0.009)	(0.006)
Dur2 ^a	0.019	-0.072
	(0.039)	(0.025)
Dur3	0.03	-0.087
	(0.06)	(0.039)
Dur4	0.048	-0.116
	(0.07)	(0.045)
Dur5	0.079	-0.106
	(0.078)	(0.05)
Children under 5	0.102	-0.223
	(0.064)	(0.044)
Children 5+	0.069	-0.079
	(0.029)	(0.019)
Wald test $\hat{\lambda}_{it}$	35.337	9.46
P-value	0.026	0.985
Sargan test	31.896	
P-Value	0.964	

^aSee footnote 6

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

Table 4: The effects of married life on the female-to-male domestic and market work hours (log) ratios - Pooled GMM estimates (Base specification, Model 1)

	Marriage is exogenous		Marriage is endogenous	
	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$
Married	0.153 (0.049)	-0.099 (0.030)	0.375 (0.184)	-0.436 (0.115)
Log Wage ratio	-1.034 (0.342)	0.726 (0.220)	-1.067 (0.38)	0.632 (0.239)
Non-labor income	-0.020 (0.010)	0.006 (0.006)	-0.034 (0.013)	0.012 (0.007)
Dur2	-0.002 (0.033)	-0.066 (0.020)		
Dur3	0.012 (0.045)	-0.078 (0.029)		
Dur4	0.006 (0.059)	-0.098 (0.038)		
Dur5	-0.035 (0.068)	-0.064 (0.043)		
Children under 5	0.276 (0.049)	-0.280 (0.033)	0.287 (0.053)	-0.299 (0.034)
Children 5+	0.129 (0.024)	-0.096 (0.015)	0.135 (0.027)	-0.108 (0.016)
Middle regions	0.072 (0.032)	0.009 (0.022)	0.080 (0.035)	-0.004 (0.023)
Southern regions	0.087 (0.025)	0.002 (0.017)	0.097 (0.027)	-0.003 (0.017)
German00	0.113 (0.089)	0.026 (0.069)	0.111 (0.097)	0.067 (0.072)
German10	-0.010 (0.050)	-0.037 (0.040)	0.006 (0.055)	-0.042 (0.044)
German01	0.123 (0.058)	-0.116 (0.047)	0.139 (0.061)	-0.111 (0.047)
Cohort 1931-45	0.159 (0.064)	-0.018 (0.044)	0.108 (0.062)	0.009 (0.042)
Cohort 1946-55	0.089 (0.036)	0.003 (0.024)	0.073 (0.037)	0.013 (0.024)
Cohort 1966-79	-0.108 (0.039)	0.012 (0.024)	-0.076 (0.043)	-0.014 (0.026)
Intercept	0.473 (0.370)	0.502 (0.253)	0.323 (0.411)	0.584 (0.264)
Wald test $\hat{\lambda}_{it}$	34.624	24.857	36.598	29.598
P-value	0.031	0.253	0.019	0.100
Wald test \bar{z}_{it}	49.336	49.103	33.976	29.618
P-value	0.000	0.000	0.001	0.003
Sargan test		44.098		52.413
P-Value		0.633		0.381

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

estimator, there is only evidence of sample selection for the log ratio of domestic hours. Also, the Sargan test does not reject the validity of the instruments and the overidentifying restrictions. We now turn to the parameters of main interest.

The results indicate that marriage has a significant effect on both domestic and market work. When women are married, their domestic workload increases relative to men whereas the ratio of female-to-male market hours falls. All else being equal, married women are more likely to specialize in domestic work than cohabiting women.

Columns 3 and 4 in Table 4 report the estimates when married life is instrumented. The duration of the relationship is used as a supplementary excluded instrument.¹⁸ The effects of marriage on the ratios of domestic work hours and market work hours are still significant when it is instrumented. However, the estimated coefficients of marriage exhibit a severe change in magnitude. It is possible that the effects of married life are seriously underestimated when marriage is supposed exogenous. On the other hand, such large variations in the estimated coefficients of marriage can denote a problem of weak instruments.¹⁹ In such a case, and with no other relevant instruments for marriage, it may be better to just assume its exogeneity. Nevertheless, being married still raises the female relative domestic workload and decreases the female relative market workload, just as before.

The effects of the other explanatory variables are in line with those reported when marriage is supposed to be exogenous. Relative earnings have a negative effect on the ratio of non-market hours. A one percent rise in relative earnings leads roughly to a one percent decrease in the ratio of non-market hours. Women with a high relative wage are less likely to specialize in domestic activities. On the contrary, relative earnings have a positive impact on the log ratio of market hours. The division of market work between partners is more equal for high female relative wages. Also, non-labor income has a significant and negative impact on the log ratio of unpaid hours. It could be that wealthy couples buy more market substitutes for home-based activities.

The presence of children in the household, especially of young children, raises women's domestic workload relative to men and decreases their share of market work. This is in accordance with the negative correlation between children and female market labor supply usually observed in empirical studies. Moreover, some regional disparities explain the division of domestic work between men and women. The female relative domestic workload is higher for households living in the southern states of West Germany. The ratio of female-to-male domestic hours is also higher for German born women living with a non-German partner. Also, the results exhibit a cohort effect. Younger women are less likely to specialize in unpaid work than their elders.

6.2 The effects of marriage on domestic work and market work hours

In this section the two alternative specifications of section 3.2 are considered. Two supplementary explanatory variables are female age in the equations related to women and male age in the equations related to men.

¹⁸If we allow the duration of the relationship to appear in the labor supply equations, the estimates for *Married* become very imprecise. We therefore maintain this exclusion restriction. Also, we test for the endogeneity of the duration of conjugal life whether marital status is assumed to be exogenous or endogenous. In both cases we do not reject the null (that is, the exogeneity of the duration of conjugal life).

¹⁹We estimate the model with other instruments such as the female-to-male age and education ratios. It does not change the estimates. Also, as mentioned in Browning (1992), the usual practice of treating dummies as unbounded and continuous in the auxiliary equation may cause problems. We may be predicting values outside the (zero, one) interval that may in turn affect the estimates of the parameter of interest.

6.2.1 The effects of marriage on women's domestic and market work hours

To save space, only the Pooled GMM estimates of the alternative model (2) are shown. Columns 1 and 2 in Table 5 report the results when marriage is assumed to be exogenous. On the whole, the results are in accordance with those obtained for the base model. Also, there is evidence of selection in both domestic hours and market hours. Marriage has a positive effect on the female number of domestic hours and a negative effect on the female number of market hours. Interestingly, the former effect offsets the latter so that "pure" leisure falls for married women, especially for couples with pre-school children.

The results show a strong effect of pre-school children and a significant effect of the duration of conjugal life on women's time use. Women that have been living with a partner for a certain period of time spend more hours on home-based activities and fewer hours on market activities than the others. Clearly, role specialization tends to increase with time. This is consistent with Becker's notion of increasing relationship-specific capital.

Own wage elasticities of market work and domestic work are shown in Table 6. These are not intertemporal elasticities as we do not account for intertemporal budgeting in our model. However, the amplitude of the market work elasticities is in line with that found in other studies using German data. For example, Laisney et al. (1993) report median intertemporal wage elasticities that range from 0.46 to 0.60 depending on the estimation. The wage elasticity is 0.40 for West German couples in Steiner and Jacobebbinghaus (2003). A similar pattern is found in Davies et al. (2000).

Finally, columns 3 and 4 in Table 5 report the estimates when married life is instrumented. The estimates of the effect of marriage are very different to those in columns 1 and 2. Marriage now increases by 7.23 the number of hours per day that women spend in home-based activities, which is greatly overestimated. It is very likely that this variable suffers from weak instruments. Moreover, none of the wage and non-female income parameters are significant. However, the qualitative impact of marriage on female labor supplies remains the same. Pure leisure decreases for married women. We now turn to the estimates of the second alternative specification.

6.2.2 The effects of marriage on female and male domestic and market work hours

Table 7 reports the estimates of the alternative model (4) when marriage is assumed to be exogenous. We note that the parameters of the male equations have large standard errors. This lack of precision is explained by the little variation in the male allocation of time.

Most of the parameter estimates of the female equations are in line with those in Table 5. The overall qualitative results remain the same. In particular, married life increases women's number of domestic hours and decreases their number of market hours such that the total effect is a fall in women's pure leisure.

Own and cross wage elasticities are shown in Table 8. Women exhibit a wage elasticity of domestic work lower than that of Table 6. Their cross wage elasticity of domestic work is significant and positive. Again, domestic work is more elastic than market work for women. On the contrary, men's domestic labor supply is inelastic.

Once again, the instrumentation of marriage leads to a huge variation in the estimated effects of this variable but the qualitative impact of marriage on women's allocation of time is the same (see Table 10, in Appendix B). Married women spend more time on domestic work and fewer hours on market work. Also, it seems that marriage now has a significant and positive effect on male domestic work whereas its impact on male market work is not significant anymore.

Table 5: The effects of married life on female domestic and market work hours - Pooled GMM estimates (Alternative specification, Model 2)

	Marriage is exogenous		Marriage is endogenous	
	h^f	l^f	h^f	l^f
Married	0.723	-0.271	7.229	-2.242
	(0.207)	(0.099)	(0.73)	(0.343)
Log female wage	-4.477	1.629	-0.796	0.559
	(1.917)	(0.893)	(1.891)	(0.853)
Female non-labor income $\times 10^{-3}$	0.098	-0.021	0.024	0.000
	(0.024)	(0.011)	(0.026)	(0.000)
Dur2	0.763	-0.277		
	(0.144)	(0.066)		
Dur3	1.632	-0.554		
	(0.215)	(0.096)		
Dur4	1.316	-0.589		
	(0.268)	(0.123)		
Dur5	0.442	-0.462		
	(0.359)	(0.168)		
Children under 5	2.842	-0.898	2.699	-0.865
	(0.218)	(0.095)	(0.226)	(0.099)
Children 5+	0.933	-0.344	1.103	-0.399
	(0.105)	(0.046)	(0.103)	(0.045)
Age $\times 10^{-2}$	-0.109	-0.012	-0.224	0.009
	(0.032)	(0.015)	(0.035)	(0.015)
Middle regions	0.230	0.004	0.506	-0.082
	(0.128)	(0.068)	(0.144)	(0.075)
Southern regions	0.269	0.001	0.363	-0.021
	(0.11)	(0.052)	(0.118)	(0.059)
German00	-0.393	0.442	-1.262	0.663
	(0.413)	(0.194)	(0.458)	(0.215)
German10	-0.399	0.005	-0.227	-0.032
	(0.266)	(0.125)	(0.318)	(0.152)
German01	0.169	-0.166	-0.067	-0.109
	(0.278)	(0.122)	(0.316)	(0.132)
Cohort 1931-45	1.263	-0.199	1.072	-0.219
	(0.294)	(0.156)	(0.295)	(0.169)
Cohort 1946-55	0.435	0.030	0.346	0.025
	(0.169)	(0.082)	(0.172)	(0.090)
Cohort 1966-79	-0.396	0.101	-0.078	0.038
	(0.171)	(0.076)	(0.208)	(0.090)
Intercept	12.172	3.827	3.845	6.333
	(4.145)	(2.000)	(4.258)	(1.974)
Wald test $\hat{\lambda}_{it}$	90.281	45.355	122.552	64.247
P-value	0.000	0.002	0.000	0.000
Wald test \bar{z}_{it}	120.53	60.326	108.813	16.868
P-value	0.000	0.000	0.000	0.155
Sargan test		53.198		66.316
P-Value		0.217		0.041

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

Table 6: Female wage elasticities of labor supply

Estimates	Asymptotic Std. Err.	P-Values
Estimated wage elasticity of the female domestic labor supply		
-1.119	0.479	0.019
Estimated wage elasticity of the female market labor supply		
0.343	0.188	0.068

Notes: Elasticities are computed at the median. Marriage is assumed to be exogenous.

6.3 Robustness Checks

6.3.1 Estimates on alternative samples

As previously mentioned, it could be difficult to disentangle the observed effect of marriage on labor supply from the effect of children. To give more robustness to our results, we re-estimate the base model on the sample of couples with no children under five and on the sample of couples with no children under eleven. This leaves us with 11,727 observations from 2,579 households if we include all observations with no children under five or with 8,657 observations from 2,041 couples if we include all observations with no children under eleven.

Columns (1) and (2) of Table 9 are the estimates of the base model (1) with no children under five. Columns (3) and (4) are the estimates with no children under eleven. Marriage is assumed to be exogenous. On the whole, the coefficients of the sample with no children under eleven are less precisely estimated.

The marital status coefficients exhibit a substantial fall in magnitude when we move from couples with pre-school children to couples with no children under eleven. It suggests that the effect of marriage on the parental allocation of time is higher for couples with young children. It is also possible that the marital status coefficients capture part of the effect of children on time use as married couples tend to have more children. However this may be, the effect of marriage is significant and has the expected sign. Married life increases women's specialization in home-based activities. This effect remains when marriage is instrumented; though with a large variation in the point estimates (see Table 11, in Appendix B).

For couples with no children under five, relative earnings still have a significant and negative effect - though smaller in magnitude - on the ratio of non-market hours and a significant and positive impact on the ratio of market hours. The effect of relative earnings is insignificant for couples with no children under eleven. Also, whatever the children's age, non-labor income continues to have a significant and negative impact on the log ratio of domestic hours.

6.3.2 Does marriage cause this outcome?

The observed effects of marriage may be due to selectivity if married couples exhibit systematic different characteristics than cohabitants. To check this assumption we estimate the base model (1) with a different marital status variable (*Change* hereafter) that takes on the value one for observations that correspond to married couples previously cohabiting. It is zero for couples that are either cohabiting or always married. This variable is for the permanent effect of the transition from cohabitation to marriage on time use. It does not capture intrinsic differences between married couples and cohabitants. Interestingly, its effect on domestic and market work hours is significant and very similar to those in

Table 7: The effects of marriage on female and male domestic and market work hours - Pooled GMM estimates, marriage is exogenous (Alternative specification, Model 4)

	No children under two			
	h^f	l^f	h^m	l^m
Married	0.631	-0.262	0.116	0.156
	(0.204)	(0.098)	(0.125)	(0.087)
Log female wage	-3.535	1.368	-1.645	-0.662
	(1.949)	(0.942)	(1.233)	(0.801)
Log male wage	4.431	-0.960	-0.098	1.882
	(1.296)	(0.623)	(0.823)	(0.543)
Non-labor income	0.018	0.006	0.071	-0.021
	(0.042)	(0.019)	(0.034)	(0.018)
Dur2	0.695	-0.261	0.443	-0.034
	(0.131)	(0.066)	(0.087)	(0.063)
Dur3	1.400	-0.503	0.719	-0.249
	(0.214)	(0.105)	(0.131)	(0.093)
Dur4	1.137	-0.535	0.528	-0.136
	(0.266)	(0.137)	(0.164)	(0.117)
Dur5	0.266	-0.411	0.221	-0.125
	(0.336)	(0.175)	(0.214)	(0.149)
Children under 5	2.966	-0.938	0.981	0.068
	(0.214)	(0.098)	(0.139)	(0.084)
Children 5+	1.009	-0.370	0.175	-0.000
	(0.106)	(0.05)	(0.068)	(0.042)
Middle regions	0.242	-0.012	-0.101	-0.026
	(0.129)	(0.069)	(0.080)	(0.054)
Southern regions	0.239	0.007	-0.081	-0.028
	(0.105)	(0.052)	(0.065)	(0.043)
German00	-0.263	0.428	-0.089	0.385
	(0.433)	(0.208)	(0.281)	(0.168)
German10	-0.276	-0.026	-0.174	0.199
	(0.255)	(0.126)	(0.167)	(0.091)
German01	0.285	-0.196	-0.042	0.276
	(0.286)	(0.131)	(0.185)	(0.116)
Cohort 1931-45	1.111	-0.148	0.180	-0.098
	(0.281)	(0.154)	(0.174)	(0.118)
Cohort 1946-55	0.391	0.035	0.017	0.031
	(0.159)	(0.081)	(0.101)	(0.062)
Cohort 1966-79	-0.351	0.085	-0.102	0.066
	(0.163)	(0.074)	(0.102)	(0.068)
Age $\times 10^{-2}$	-0.139	-0.007	-0.031	-0.023
	(0.030)	(0.015)	(0.018)	(0.012)
Intercept	0.891	6.261	6.568	2.049
	(3.916)	(1.958)	(2.499)	(1.703)
Wald test $\hat{\lambda}_{it}$	101.364	41.591	128.051	42.965
P-value	0.000	0.005	0.000	0.003
Wald test \bar{z}_{it}	136.553	62.038	56.806	32.776
P-value	0.000	0.000	0.000	0.001
Sargan test		102.681		
P-Value		0.136		

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

Table 8: Female and male wage elasticities of labor supply

Estimates	Asymptotic Std. Err.	P-Values
Estimated wage elasticity of the female domestic labor supply		
-.884	0.487	0.069
Estimated cross wage elasticity of the female domestic labor supply		
1.108	0.324	0.001
Estimated wage elasticity of the female market labor supply		
0.288	0.198	0.147
Estimated cross wage elasticity of the female market labor supply		
-.202	0.131	0.124
Estimated wage elasticity of the male domestic labor supply		
-.049	0.411	0.905
Estimated cross wage elasticity of the male domestic labor supply		
-.823	0.617	0.182
Estimated wage elasticity of the male market labor supply		
0.331	0.096	0.001
Estimated cross wage elasticity of the male market labor supply		
-0.116	0.141	0.408

Notes: Elasticities are computed at the median. Marriage is assumed to be exogenous.

Tables 4 and 9. For couples with no children under two for instance, the transition from cohabitation to married life, when assumed to be exogenous, increases the log ratio of domestic work hours by 0.150 and decreases the log ratio of market work hours by -0.099 . We interpret this result as evidence of no significant selection into marriage. Table 12 in Appendix C reports the results.

6.3.3 Testing exclusion restrictions

We now consider whether education, age and unemployment rates (which only appear in the selection equation) provide valid exclusion restrictions. Including these variables in the log ratios equations does not have significant effects on the original parameters estimates and their coefficients are insignificant. The effect of the duration of conjugal life on the log ratio of market work remains when age is a regressor. The t -values for the coefficients of the female education (age) are below 1.40 (1.33) whereas the t -values for the male education (age) are below 1.17 (1.33). We hence maintain these exclusion restrictions. We also test for interaction terms between marital status and wages and between marital status and the duration of conjugal life. Whatever the sample used, none of them is significant.²⁰ Similar conclusions are drawn for the two alternative specifications (2) and (4) of section 3.2.²¹

²⁰These figures and conclusions concern the estimation with marriage being exogenous. They still hold with endogenous marriage.

²¹Results are available upon request.

Table 9: Alternative samples: the effects of married life on the female-to-male domestic and market work hours (log) ratios - Pooled GMM estimates, marriage is exogenous (Base specification, Model 1)

	No children under five		No children under eleven	
	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$
Married	0.126	-0.099	0.094	-0.078
	(0.049)	(0.032)	(0.047)	(0.024)
Log wage ratio	-0.893	0.757	-0.090	-0.125
	(0.341)	(0.227)	(0.318)	(0.198)
Non-labor income	-0.019	0.010	-0.025	0.000
	(0.011)	(0.007)	(0.013)	(0.008)
Dur2	-0.016	-0.057	-0.071	-0.005
	(0.034)	(0.022)	(0.036)	(0.02)
Dur3	0.039	-0.079	-0.041	-0.013
	(0.046)	(0.032)	(0.046)	(0.029)
Dur4	0.012	-0.080	0.102	-0.158
	(0.059)	(0.041)	(0.058)	(0.04)
Dur5	-0.021	-0.051	0.084	-0.160
	(0.070)	(0.047)	(0.072)	(0.049)
Children 5+	0.140	-0.099	0.099	-0.065
	(0.023)	(0.017)	(0.022)	(0.013)
Middle regions	0.087	-0.002	0.103	-0.019
	(0.031)	(0.022)	(0.031)	(0.026)
Southern regions	0.107	-0.004	0.105	0.013
	(0.025)	(0.018)	(0.026)	(0.02)
German00	0.122	-0.010	-0.061	0.157
	(0.093)	(0.074)	(0.092)	(0.087)
German10	0.024	-0.053	-0.000	0.004
	(0.048)	(0.042)	(0.053)	(0.062)
German01	0.110	-0.128	0.017	-0.081
	(0.059)	(0.050)	(0.059)	(0.051)
Cohort 1931-45	0.126	-0.023	0.007	0.01
	(0.064)	(0.045)	(0.066)	(0.052)
Cohort 1946-55	0.076	-0.002	0.003	0.023
	(0.035)	(0.026)	(0.038)	(0.03)
Cohort 1966-79	-0.098	0.004	-0.143	0.022
	(0.041)	(0.027)	(0.046)	(0.027)
Intercept	0.629	0.343	0.995	-0.063
	(0.366)	(0.264)	(0.382)	(0.274)
Wald test $\hat{\lambda}_{it}$	32.629	28.658	19.158	34.855
P-value	0.050	0.122	0.575	0.029
Wald test \bar{z}_{it}	29.214	29.168	29.643	27.188
P-value	0.002	0.002	0.002	0.004
Sargan test		41.692		52.643
P-Value		0.728		0.299

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

7 Conclusion

In this paper, we have estimated the effects of the transition from cohabitation to marriage on household market and non-market labor supply using a German sample of working couples. Our results indicate that marriage raises women's specialization in domestic work with a greater impact on couples with preschool children. We find that specialization in unpaid work is less likely for women with a high market wage.

These findings are robust across three different specifications. Actually, marriage increases women's specialization in home-based activities whether we consider the female-to-male domestic and market work hours (log) ratios or women's market and non-market labor supply or men's and women's market and non-market labor supply.

Interestingly, we find that married women enjoy less leisure than cohabiting women. Marriage decreases women's market work and increases their domestic work so that the overall effect is a fall in their leisure. We also found evidence that married men enjoy less leisure than cohabiting men but the effect of marriage on men's labor supply is less clear cut due to little variation in the male allocation of time.

Finally, the magnitude of the estimated coefficients of marriage changes drastically when marital status is assumed to be endogenous but the qualitative results remain the same. Marriage still increases women's specialization in unpaid work. It could also be that marriage influences the decision to participate in the labor market. In this case, the idea would be to estimate structural participation equations with marital status as an explanatory variable. This is a topic of future work.

A The GMM Estimator

Here we derive the GMM estimator and its variance for a general model that controls for unobserved heterogeneity and sample selection bias. We begin with some notation. For couple i at period t we have a joint observation on (Y_{it}, X_{it}) , where $Y_{it} = \begin{pmatrix} Y_{1it} \\ \vdots \\ Y_{Git} \end{pmatrix}$ and $X_{it} = \begin{pmatrix} X_{1it} \\ \vdots \\ X_{Git} \end{pmatrix}$ are bloc-diagonal matrices that stand for the G regressands $Y_{it}^1, \dots, Y_{it}^G$, and the G vectors of regressors $X_{it}^1, \dots, X_{it}^G$, respectively.

The second-step system has the form:

$$Y_{it} = X_{it}\theta + e_{it} = \widehat{X}_{it}\theta + (X_{it} - \widehat{X}_{it})\theta + e_{it}, \quad (11)$$

where \widehat{X}_{it} is the bloc-diagonal matrix of regressors including the generated inverse Mill's ratios from the first-step probit estimation, $\theta = (\theta'_1, \dots, \theta'_G)'$ is the vector of interest, and e_{it} is a G -vector of disturbance terms.

Let \widehat{Z}_{it} denoted the vector of instruments that include the generated inverse Mill's ratios, and π_t the vector of unknown coefficients of the first-step probit equation. For the Fixed Effect model, we have $Y_{it} = y_{it} - \bar{y}_i$, $\widehat{X}_{it} = (x_{it} - \bar{x}_i, \widehat{\lambda}_{it} - \bar{\lambda}_i)$ and $\widehat{Z}_{it} = \widehat{z}_{it} - \bar{z}_i$ whereas $Y_{it} = y_{it}$, $\widehat{X}_{git} = (\alpha_g, x_{git}, \bar{z}_i, \widehat{\lambda}_{it})$ and $\widehat{Z}_{it} = \widehat{z}_{it}$ for the Pooled model.

The GMM estimator of (11) on the selected sample is then given by:

$$\widehat{\theta} = [C'W^{-1}C]^{-1} C'W^{-1}E, \quad (12)$$

where C is a bloc-diagonal matrix with $C_g = N^{-1} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{X}'_{git} \widehat{Z}_{it}$ on the diagonal, E is a stacked vector of elements $N^{-1} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{Z}'_{it} y_{git}$, and W is a bloc-diagonal matrix with $N^{-1} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{Z}'_{it} \widehat{Z}_{it}$ on the diagonal, for $g = 1, \dots, G$.

We now turn to the asymptotic variance estimator of $\widehat{\theta}$. Plugging equation (11) in (12) we can write:

$$\sqrt{N}(\widehat{\theta} - \theta) = [C'W^{-1}C]^{-1} C'W^{-1}A,$$

where $A = \left(N^{-1/2} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{Z}'_{Git} \left[(X_{it} - \widehat{X}_{it})\theta + e_{it} \right] \right)$, with \widehat{Z}'_{Git} an appropriate stack of \widehat{Z}'_{it} to have compatible dimensions.²²

The term $(X_{it} - \widehat{X}_{it})\theta$ in A can be approximated to first order around $\pi = \widehat{\pi}$, the estimates of the probit equation, by the following expression:

$$\widehat{X}_{it}\theta = X_{it}\theta + \theta' \nabla_{\pi} X'_{it} (\widehat{\pi}_G - \pi_G) + o_p(1),$$

where $\nabla_{\pi} X'_{it}$ is the Jacobian of X'_{it} with respect to π . Again, the vector $\widehat{\pi}_G - \pi_G$ is just a stack of $\widehat{\pi} - \pi$ to have compatible dimensions. Following Semykina and Wooldridge (2005), we write:

$$\sqrt{N}(\widehat{\pi} - \pi) = N^{-1/2} \sum_{i=1}^N r_i(\pi) + o_p(1),$$

²²For instance, $\widehat{Z}'_{2it} = (\widehat{Z}_{it}, \widehat{Z}_{it})'$ for a two-equation system.

where $r_i(\pi)$ is a term that depends on the expected hessian and scores of the probit log-likelihood.

Using some algebra it follows that:

$$A \simeq N^{-1/2} \sum_{i=1}^N \left[\sum_{t=1}^T s_{it} \widehat{Z}'_{Git} e_{it} - F \sum_{i=1}^N r_i(\pi) \right], \quad (13)$$

where $F = \begin{pmatrix} F_1 \\ \vdots \\ F_G \end{pmatrix}$, with $F_g = \frac{1}{N} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{Z}'_{it} \theta'_g \nabla_{\pi} X'_{git}$, $g = 1, \dots, G$. The asymptotic variance estimator of $\widehat{\theta}$ is thus given by:

$$[C'(A'A)^{-1}C]^{-1}.$$

In practice, the unknown term e_{it} , F and r_i are replaced with their respective consistent estimators $\widehat{e}_{it} = (y_{it} - \widehat{X}_{it}\widehat{\theta})$, $\widehat{F}_g = \left(N^{-1} \sum_{i=1}^N \sum_{t=1}^T s_{it} \widehat{Z}'_{it} \widehat{\theta}'_g \nabla_{\pi} \widehat{X}'_{git} \right)$ and \widehat{r}_i . Let $q_{it} = (\eta, z_{it}, \bar{z}_i)$ the vector of regressors that appear in the first-step probit. Then,

$$\widehat{r}_{it} = \widehat{H}_t^{-1} \{ \Phi(q_{it}\widehat{\pi}_t) [1 - \Phi(q_{it}\widehat{\pi}_t)] \}^{-1} \phi(q_{it}\widehat{\pi}_t) q'_{it} [s_{it} - \Phi(q_{it}\widehat{\pi}_t)],$$

where:

$$\widehat{H}_t^{-1} \equiv \frac{1}{N} \{ \Phi(q_{it}\widehat{\pi}_t) [1 - \Phi(q_{it}\widehat{\pi}_t)] \}^{-1} [\phi(q_{it}\widehat{\pi}_t)]^2 q'_{it} q_{it}$$

is the consistent estimator of minus the expected Hessian (see Semykina and Wooldridge, 2005). For each couple i , the \widehat{r}_{it} are stacked to obtain the \widehat{r}_i used in equation (13). The estimated variance estimator is therefore:

$$[C'(\widehat{A}'\widehat{A})^{-1}C]^{-1}.$$

B Estimation with endogenous marriage

Table 10: The effects of marriage on female and male domestic and market work hours - Pooled GMM estimates, marriage is endogenous (Alternative specification, Model 4)

	No children under two			
	h^f	l^f	h^m	l^m
Married	6.831	-1.982	2.293	0.109
	(0.801)	(0.377)	(0.414)	(0.269)
Log female wage	-1.618	1.144	-0.448	-0.974
	(2.023)	(0.923)	(1.125)	(0.717)
Log male wage	2.893	-1.101	-0.553	1.212
	(1.433)	(0.663)	(0.832)	(0.553)
Non-labor income	-0.042	0.022	0.051	-0.018
	(0.043)	(0.020)	(0.028)	(0.017)
Children under 5	2.813	-0.911	0.914	0.158
	(0.229)	(0.101)	(0.135)	(0.081)
Children 5+	1.114	-0.403	0.247	-0.013
	(0.11)	(0.049)	(0.06)	(0.037)
Middle regions	0.449	-0.061	-0.022	-0.018
	(0.149)	(0.075)	(0.082)	(0.053)
Southern regions	0.334	-0.013	-0.057	-0.006
	(0.118)	(0.057)	(0.063)	(0.041)
German00	-0.929	0.502	-0.418	0.344
	(0.493)	(0.222)	(0.276)	(0.162)
German10	-0.243	-0.011	-0.184	0.261
	(0.316)	(0.142)	(0.175)	(0.091)
German01	0.166	-0.216	-0.161	0.249
	(0.329)	(0.137)	(0.186)	(0.109)
Cohort 1931-45	0.934	-0.169	0.075	-0.111
	(0.298)	(0.163)	(0.164)	(0.111)
Cohort 1946-55	0.329	0.022	-0.041	0.021
	(0.17)	(0.085)	(0.098)	(0.059)
Cohort 1966-79	-0.081	0.046	-0.005	0.079
	(0.204)	(0.085)	(0.105)	(0.067)
Age $\times 10^{-2}$	-0.241	0.014	-0.069	-0.007
	(0.031)	(0.015)	(0.016)	(0.011)
Intercept	-1.697	7.671	4.454	4.771
	(4.057)	(1.937)	(2.325)	(1.459)
Wald test $\hat{\lambda}_{it}$	120.481	51.919	128.912	47.715
P-value	0.000	0.000	0.000	0.001
Wald test \hat{z}_{it}	116.848	21.15	78.424	26.223
P-value	0.000	0.048	0.000	0.010
Sargan test		119.711		
P-Value		0.028		

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

Table 11: Alternative samples: the effects of married life on the female-to-male domestic and market work hours (log) ratios - Pooled GMM estimates, marriage is endogenous (Base specification, Model 1)

	No children under five		No children under eleven	
	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$
Married	0.516	-0.506	0.382	-0.431
	(0.172)	(0.113)	(0.155)	(0.096)
Log wage ratio	-0.767	0.625	-0.254	-0.105
	(0.357)	(0.23)	(0.327)	(0.214)
Non-labor income	-0.041	0.020	-0.039	0.006
	(0.014)	(0.009)	(0.016)	(0.009)
Children 5+	0.159	-0.111	0.108	-0.076
	(0.026)	(0.018)	(0.024)	(0.013)
Middle regions	0.106	-0.019	0.112	-0.027
	(0.033)	(0.023)	(0.033)	(0.027)
Southern regions	0.117	-0.010	0.119	0.004
	(0.027)	(0.019)	(0.029)	(0.021)
German00	0.089	0.051	-0.032	0.179
	(0.098)	(0.076)	(0.097)	(0.089)
German10	0.053	-0.052	0.023	0.011
	(0.052)	(0.046)	(0.056)	(0.063)
German01	0.115	-0.120	0.035	-0.080
	(0.061)	(0.050)	(0.062)	(0.053)
Cohort 1931-45	0.067	0.019	0.002	-0.001
	(0.059)	(0.042)	(0.066)	(0.050)
Cohort 1946-55	0.053	0.016	0.014	0.014
	(0.036)	(0.026)	(0.040)	(0.031)
Cohort 1966-79	-0.052	-0.034	-0.114	-0.010
	(0.044)	(0.029)	(0.051)	(0.032)
Intercept	0.488	0.408	0.800	0.039
	(0.412)	(0.283)	(0.425)	(0.296)
Wald test $\hat{\lambda}_{it}$	36.461	31.775	20.602	39.982
P-value	0.019	0.062	0.483	0.007
Wald test \bar{z}_{it}	21.05	26.105	23.813	28.003
P-value	0.033	0.006	0.014	0.003
Sargan test	45.517		43.039	
P-Value	0.654		0.747	

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

C The permanent effects of the transition to marriage

Table 12: The permanent effects of the transition to marriage on the female-to-male domestic and market work hours (log) ratios - Pooled GMM estimates, marriage is exogenous

	No children under two		No children under five		No children under eleven	
	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$	$r^h = \frac{h^f}{h^m}$	$r^l = \frac{l^f}{l^m}$
Change	0.150	-0.099	0.113	-0.091	0.086	-0.082
	(0.049)	(0.029)	(0.049)	(0.030)	(0.048)	(0.025)
Log wage ratio	-1.026	0.690	-0.946	0.633	-0.293	-0.099
	(0.345)	(0.219)	(0.331)	(0.211)	(0.326)	(0.195)
Non-labor income	-0.021	0.005	-0.017	0.005	-0.028	-0.005
	(0.011)	(0.006)	(0.011)	(0.007)	(0.015)	(0.009)
Dur2	0.006	-0.075	-0.007	-0.064	-0.062	-0.008
	(0.033)	(0.020)	(0.034)	(0.021)	(0.037)	(0.02)
Dur3	0.025	-0.090	0.063	-0.105	-0.036	-0.014
	(0.046)	(0.030)	(0.047)	(0.030)	(0.046)	(0.029)
Dur4	0.024	-0.114	0.042	-0.114	0.101	-0.164
	(0.059)	(0.038)	(0.059)	(0.038)	(0.059)	(0.040)
Dur5	-0.014	-0.078	0.014	-0.081	0.089	-0.161
	(0.068)	(0.042)	(0.069)	(0.043)	(0.073)	(0.050)
Children under 5	0.263	-0.281				
	(0.050)	(0.033)				
Children 5+	0.128	-0.096	0.131	-0.093	0.096	-0.061
	(0.024)	(0.015)	(0.023)	(0.016)	(0.022)	(0.013)
Middle regions	0.072	0.010	0.088	-0.004	0.101	-0.016
	(0.032)	(0.022)	(0.032)	(0.022)	(0.031)	(0.026)
Southern regions	0.089	0.005	0.107	0.002	0.107	0.019
	(0.025)	(0.016)	(0.025)	(0.017)	(0.027)	(0.020)
German00	0.116	0.031	0.136	0.021	-0.023	0.166
	(0.090)	(0.069)	(0.093)	(0.071)	(0.093)	(0.086)
German10	-0.009	-0.039	0.016	-0.046	-0.001	0.024
	(0.051)	(0.041)	(0.049)	(0.042)	(0.052)	(0.062)
German01	0.127	-0.111	0.117	-0.109	0.036	-0.076
	(0.059)	(0.046)	(0.060)	(0.047)	(0.059)	(0.051)
Cohort 1931-45	0.146	-0.017	0.107	-0.018	-0.011	0.0184
	(0.065)	(0.044)	(0.064)	(0.043)	(0.065)	(0.051)
Cohort 1946-55	0.081	0.006	0.064	0.008	-0.001	0.026
	(0.036)	(0.024)	(0.036)	(0.024)	(0.037)	(0.029)
Cohort 1966-79	-0.104	0.012	-0.091	-0.004	-0.145	0.026
	(0.039)	(0.024)	(0.041)	(0.025)	(0.046)	(0.026)
Intercept	0.513	0.438	0.685	0.208	0.965	-0.135
	(0.378)	(0.256)	(0.372)	(0.259)	(0.382)	(0.273)
Sargan test	47.324		50.049		52.838	
P-Value	0.500		0.392		0.293	

Notes: Asymptotic standard errors are in parentheses.

Coefficients in bold are significant at the five percent level.

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