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**Article**

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Schmollers Jahrbuch – Journal of Applied Social Science Studies. Zeitschrift für Wirtschafts- und Sozialwissenschaften

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*Suggested Citation:* Bartels, Charlotte (2012) : Redistribution and Insurance in the German Welfare State, Schmollers Jahrbuch – Journal of Applied Social Science Studies. Zeitschrift für Wirtschafts- und Sozialwissenschaften, ISSN 1865-5742, Duncker & Humblot, Berlin, Vol. 132, Iss. 2, pp. 265-295, <https://doi.org/10.3790/schm.132.2.265>

This Version is available at:

<https://hdl.handle.net/10419/292369>

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## Redistribution and Insurance in the German Welfare State

By Charlotte Bartels\*

### Abstract

Welfare states redistribute both between individuals reducing annual inequality and over the life-cycle insuring against income risks. But studies measuring redistribution often focus only on a one-year period. Using German SOEP data from 1984 to 2009, long-term inequality over a 20-year period is computed and then decomposed into an inter- and intra-individual component. Results show that annual inequality is higher than long-term inequality, but redistribution is also larger from an annual perspective. In the long-term, the German welfare state clearly prioritizes insurance over redistribution. This gets even more pronounced at later stages of the life-cycle through the payment of social security pensions.

*JEL Classification: D31, D63, H53, H55*

### 1. Introduction

Welfare states redistribute both between individuals (inter-individual redistribution) and over the life-cycle of an individual (intra-individual redistribution). Transfers such as social assistance and housing benefits are clearly aimed at reducing cross-sectional inequality between individuals. In contrast, transfers such as sickness benefits, unemployment benefits and social security pensions follow motives of inter-temporal redistribution within an individual's life-cycle (Pettersson/Pettersson, 2007) insuring individuals against risks. As people are generally assumed to be risk averse preferring a stable over an unstable income stream, state provided insurance aimed at smoothing income profiles generates a welfare gain.

There is a large literature on inequality and welfare states' redistribution. The effect of redistribution is mostly measured as the reduction of annual economic

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\* I thank Giacomo Corneo and two anonymous referees for most valuable advice. Furthermore, I thank Timm Bönke, Anders Björklund, Carsten Schröder, Viktor Steiner, and participants of the IIPF and ECINEQ conferences 2011 and the FDZ-RV-Workshop 2011 for helpful comments and suggestions.

inequality before and after state intervention using gross and net annual incomes. Annual inequality provides a snapshot of the inequality in one year revealing how the welfare state reduces inter-individual income differences in that particular year. Benefits are viewed as pure transfers and social security contributions as pure taxes (Burkhauser/Warlick, 1981). But redistribution between individuals measured annually will probably exceed redistribution taking place in the long-term, since most of the current beneficiaries do not remain beneficiaries throughout their lives and will support other people at a different stage of their life-cycle or of their fortunes (Layard, 1977; Liebman, 2001). On top of this, the influence of welfare state insurance can only be taken into account in the long-term. This is particularly important, since most of the annual aggregate redistribution volume is probably due to intra-individual income smoothing (Sandmo, 1999). In the long-term, social security contributions can be viewed as insurance premiums as well. And pension benefits from German social security can also be seen as annuities equivalent to the contributions paid during working life rather than as inter-generational transfers (Börsch-Supan/Reil-Held, 2001).

This study aims at broadening the view on redistribution from the annual to the long-term perspective. The most suitable approach to measure long-term inequality is to look at lifetime income. However, the number of studies enquiring lifetime inequality is limited due to the lack of data encompassing the entire life-cycle of individuals. An example for Germany is Bönke et al. (2011). They use German social security data to estimate lifetime earnings inequality. But since these data do not include all information relevant to quantify taxes and transfers, it is not possible to draw conclusions about redistribution in the long-term. This study uses equivalent incomes of West German households documented by SOEP data from 1984 to 2009, which allows to consider different taxes and transfers contingent upon the household context. Long-term income is computed as the Net Present Value (NPV) of income streams over a 20-year period discounted by the yield curve.

Long-term inequality is then decomposed into an inter- and intra-individual component. Redistribution is measured as the distance between the components of pre- and post-government income, where the distance between the between-group components provides a measure for redistribution between individuals and the distance between the within-group components for income smoothing. In the long-term, the German welfare state clearly prioritizes insurance over redistribution. Income taxes, social security contributions, public transfers and social security pensions enable to tackle their potential importance for either redistribution between individuals or income smoothing. Children and elderly are particularly interesting in an analysis of welfare schemes as they depend on intra-household transfers or public benefits not being part of the work force yet or anymore. Therefore, individuals are divided into six age groups reaching from 1–10 years to 51–60 years in the first year of observation. Hereby, the

impact of different welfare state schemes depending on the stage of the life-cycle can be analyzed. When individuals are young, state intervention reduces income differences between individuals mainly through the progressive tax system. Getting older and reaching retirement age income-smoothing redistribution via social security pensions becomes central.

The paper is organized as follows. Section 2 describes SOEP data and income concepts used for the analysis and lays out the methodology. Long-term inequality as measured by the Theil coefficient is decomposed into a between-group (inter-individual) and a within-group (intra-individual) component. Section 3 offers a discussion of the results. Section 4 concludes.

## 2. Data and Methodology

### 2.1 Database

The analysis is based on a subsample from the SOEP survey years 1984 to 2009. The SOEP is a representative panel study containing individual and household data in Germany from 1984 onwards. After the German reunification in 1990 the study was expanded to the New German Laender. All household members are interviewed individually once they reach the age of 16. The sample design ensures representativeness by oversampling special subpopulations. These include subsamples of guest workers from 1984 onwards, immigrants starting in 1994 and high income households from 2002 on. A critical variable in the calculation of taxable income is the year in which the reported income is received. In the SOEP as well as in most other surveys, yearly income is inquired retrospectively, e.g., the income reported in 1984 belongs to 1983.<sup>1</sup>

### 2.2 Sample Selection

To analyze the long-term income distribution, a balanced panel is required providing a complete sequence of annual incomes. Seven balanced panels are constructed, each encompassing 20 years: 1983–2002, 1984–2003, 1985–2004, 1986–2005, 1987–2006, 1988–2007, 1989–2008. Hence, the analysis only includes individuals who participated 20 years in a row. Consequently, the sample only includes individuals who live in West German households and East German households entering the SOEP after 1990 are excluded. Members of these households are split into six age groups such that in the first year persons in the sample are 1–60 years old. In the 20th year, individuals are 20–79 years old. Additionally to observed income, imputed values provided by the

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<sup>1</sup> See Haisken-DeNew/Frick (2005), Frick (2006) and Wagner et al. (2007) for further details.

SOEP are used. Item non-response on income questions in the SOEP is concentrated in the tails of the income distribution (Frick/Grabka, 2005), but only weakly associated with observable variables such as human capital variables, marital status, firm size, being foreign, and being employed in public service (Biewen, 2001). According to Frick and Grabka (2005) income inequality and income instability would be underestimated when restricting the sample to observed income components only. Households are dropped if they exhibit a missing income source not replaced by an imputed value. Because certain inequality measures are not defined for zero and negative values, zero market incomes are replaced by 0.01 in order to include them in the analysis.

Each age group is treated as a subsample. Thus, individuals are considered at different stages of their life cycle reaching from childhood to retirement. Table 1 presents the numbers of individuals within each age group observed in each single year of the 20-year-period.

*Table 1*  
**Number of Observed Individuals**

age group	age	1983–2002	1984–2003	1985–2004	1986–2005	1987–2006	1988–2007	1989–2008
1	1–10	579	542	535	517	499	488	452
2	11–20	713	648	590	514	464	429	373
3	21–30	791	783	797	765	740	722	692
4	31–40	798	775	785	761	738	698	630
5	41–50	830	833	783	736	690	661	632
6	51–60	515	486	487	483	489	485	477

Source: SOEP.

### 2.3 Income Concepts

The computation of long-term income follows the NPV method. Each individual  $i$  could sell the promise to a market participant today ( $t = 0$ ) of paying him his future income  $x_{i,t}$  at time  $t$ . If  $P$  is the price that the market is willing to pay for this promise, then  $\frac{x_{i,t}}{1+i_{0,t}}$  gives  $P$ . The term  $i_{0,t}$  is the interest rate for a safe investment today with time to maturity  $t$ . The NPV indicates what future income streams accumulated over the 20-year period are worth today ( $t = 0$ ) and is defined as

$$NPV_i = \sum_{t=0}^{T=19} \frac{x_{i,t}}{d_{0,t}}$$

with

$$d_{0,t} = 1 + i_{0,t}.$$

The NPV depends crucially on the discount rate  $d_{0,t}$  chosen. One approach is to take into account market participants' expectations today on future interest rates and inflation. The relation between interest rates  $i_{0,t}$  and different times to maturity  $t$  of zero-coupon bonds without default risk is provided by the so-called yield curve. Since the yield curve allows interest rates to vary with maturity, it is a better approximation for expected market conditions than a constant discount factor would be. But the yield curve function is only known with certainty for a few specific maturity dates, because only very few zero-coupon bonds exist. Hence, the other maturities have to be estimated. Serving as a tool for monetary policy decisions the yield curve is provided by the Bundesbank in Germany.<sup>2</sup> The Bundesbank applies the method of Svensson (1994) which is an extended version of Nelson and Siegel's (1987) approach. Following Svensson (1994) the interest rate is the sum of a constant and three exponential terms to allow for sufficient flexibility:

$$\begin{aligned}
 i(t, \beta) &= \beta_0 + \beta_1 \left( \frac{1 - \exp(-t/\tau_1)}{(t/\tau_1)} \right) \\
 &+ \beta_2 \left( \frac{1 - \exp(-t/\tau_1)}{(t/\tau_1)} - \exp\left(-\frac{t}{\tau_1}\right) \right) \\
 &+ \beta_3 \left( \frac{1 - \exp(-t/\tau_2)}{(t/\tau_2)} - \exp\left(-\frac{t}{\tau_2}\right) \right)
 \end{aligned}$$

where  $\beta_0, \beta_1, \beta_2, \beta_3, \tau_1$  and  $\tau_2$  are estimated parameters. Yield curves for the years 1983, 1984, 1985, 1986, 1987, 1988 and 1989 from one up to 19 years to maturity are given in the Appendix.<sup>3</sup> Since the NPV is sensitive to the discount rate, deflated long-term incomes using the Consumer Price Index (CPI) with base year 2005 are computed alternatively.

As household members usually share their income, individuals have to be analyzed in the context of their household and equivalent household income is used in the analysis.<sup>4</sup> Neglecting their household context and looking solely at

<sup>2</sup> However, the German yield curve is only available as of 1972. Other studies use average interest rates of government bonds (e.g., Bönke et al., 2011) or ad-hoc constant discount rates such as 3 percent (e.g., Börsch-Supan/Reil-Held, 2001).

<sup>3</sup> Yield curves from 1983 to 1988 present a positive, normal slope. But the yield curve of the year 1989 reveals an inverse shape because interest rates are taken from month December, hence, succeeding the event of German reunification in November 1989. The German yield curve after reunification is an often cited example for an inverse shape not being the result of an expected recession but of extraordinary circumstances. It is usually explained by the great demand for liquidity to finance urgent investments. As those were mainly short-term financed, short-term interest rates rose ("einigungsbedingter Zinsanstieg"). Additionally, uncertainties regarding the scope of transfers from the Old to the New German Laender and inflationary pressures because of the ongoing economic boom contributed to increased interest rates in the end of 1989 (Bundesbank, 1991).

their individual income would overestimate inequality in a first stage and insurance and redistribution in a second stage. Income pooling within a household contributes to stabilizing individual incomes provided that there is less than perfect correlation between income positions of the household members (Björklund/Palme, 2002). Furthermore, a low-wage spouse would appear poorer and a high-income earner richer than their standard of living in a joint household. The elevated difference between individual income inequality and net household income inequality would be mistakenly attributed to welfare state redistribution alone. Indeed, the extent of redistribution through the U.S. social security system is halved, when benefits and taxes are analyzed at the family level because of the large intra-household transfers from men to women (Gustman/Steinmeier, 2001). Additionally, public transfers and income taxes can be considered using household income, because they are contingent upon the household context in Germany.

Table 2 presents the income concepts used in the analysis. Social security pensions are considered separately because pensions state the greatest item in Germany's social budget amounting to about one third of overall expenses. As German social security pensions can also be seen as annuities equivalent to the contributions, they are comparable with private pensions and, hence, an element of gross household income before redistribution.<sup>5</sup> Gross household income including social security pensions is called modified gross household income. Subtracting social security contributions and income taxes and adding public transfers articulates different stages of government intervention. Table 3 gives details on income, tax and transfer measures in the SOEP.

Table 2

**Income Concepts**

1	gross household income (ghi)
2	modified ghi = ghi + social security pensions
3	modified ghi – social security contributions
4	modified ghi – social security contributions – income taxes
5	net household income = modified ghi – social security contributions – income taxes + public transfers

Source: SOEP.

<sup>4</sup> Equivalent household income is derived using the OECD modified equivalence scale that assigns a value of 1 to the household head, 0.5 to each additional adult member and 0.3 to each child.

<sup>5</sup> Households could decide to invest in housing as alternative to a private pension plan and to enjoy non-monetary returns from this investment (Canberra Group, 2001). Adding imputed rental value to gross household income gives slightly lower inequality estimates than for gross household income, especially for intra-individual inequality. The difference tends to grow with age reflecting the rising significance of housing investment, but overall the effect is quite small. Results are available from the author upon request.

Table 3

**From Gross to Net Income in the SOEP**

<b>gross household income</b>	labor earnings, asset income, private transfers, private pensions
<b>+ social security pensions</b>	payments from old age, disability and widowhood pension schemes
<b>– social security contributions</b>	pension insurance, health-care insurance, and unemployment insurance estimated by routines described by Schwarze (1995)
<b>– income taxes</b>	estimated by routines described by Schwarze (1995)
<b>+ public transfers</b>	housing allowances, child benefits, subsistence assistance, special circumstances benefits, government student assistance, maternity benefits, unemployment benefits, unemployment assistance, unemployment assistance allowance, nursing care
<b>= net household income</b>	

Source: SOEP.

## 2.4 Measurement of Redistribution and Insurance

Inequality measures are computed to answer two questions. First, how does redistribution change when extending the measurement period from one to 20 years? Second, to what extent does the German welfare state prioritize insurance over redistribution in the long-term?

The first question is addressed by computing age-specific inequality measures based on long-term and annual incomes. First, all inequality measures are computed using long-term incomes constructed as explained in 2.2. These measures provide inequality levels over a 20-year period. Second, inequality levels are computed for every single year and then averaged over the 20-year period. Comparing results for the averaged cross-section and for the long-term gives the extent to which inequality changes when extending the measurement period. A simple, implicit measure for redistribution is provided by the difference between inequality measures of pre- and post-government income concepts. Redistribution is higher annually, if the difference between pre- and post-government income inequality is higher for the averaged cross-section than in the long-term. However, deploying this simple method may violate the Pigou-Dalton principle as it does not control for reranking due to transfers.

To measure redistribution and insurance in the long-run, an approach by decomposing inequality is chosen. An advantage of the Theil coefficient is its simple decomposability. The Theil coefficient can be rewritten as an additive function of between-group and within-group inequality. Each individual can be interpreted as a subgroup  $i$  ( $i = 1, 2 \dots n$ ) consisting of 20 observations per in-



dividual during each 20-year period. Thus, the Theil's between-group component provides a measure for inter-individual inequality and the within-group component for intra-individual inequality. Björklund and Palme (2002) use a similar method deploying Swedish data. Again, inequality measures are calculated separately for age-specific subgroups. The Theil coefficient ranges from 0 to  $\sqrt{\ln(K)}$ , where larger values indicate higher inequality, and it can be decomposed as follows:

$$\begin{aligned} T &= \frac{1}{K} \sum_{k=1}^K \frac{y_k}{\bar{y}} \ln \left( \frac{y_k}{\bar{y}} \right) \\ &= T_{between} + T_{within} \\ &= \sum_{i=1}^n \nu_i \ln \frac{\nu_i}{w_i} + \sum_{i=1}^n \nu_i T_i \end{aligned}$$

where  $K$  is the number of observations,  $y_k$  is individual equivalent household income and  $\bar{y}$  is the mean of individual equivalent household income.  $n$  is the number of individuals equal to the number of subgroups,  $w_i$  the individual's weight in the total population and  $\nu_i$  the individual's share of age group specific total income. Individuals are weighted using selection and staying probabilities as given in the SOEP.<sup>6</sup>

The between-group component is equivalent to before-measured long-term inequality. Again, redistribution is measured as the distance between the components of pre- and post-government income, where the distance between the between-group components provides a measure for redistribution between individuals and the distance between the within-group components for income smoothing.

To verify the statistical significance of the inequality measures, the bootstrap method is used (Mills/Zandvakili, 1997). 100 random samples with replacement are drawn from all observations within a 20-year period. Each bootstrap sample contains as many sampling units as the original sample. Moreover, stratified bootstrap sampling is implemented to take the different selection probabilities of the SOEP into account. The variable "strat" documented in the SOEP indicates the strata associated with sampling units. Per stratified bootstrap sample, inequality measures are computed. This gives 100 different values of the inequality measures for each income concept and each 20-year-period. Hall's (1994) percentile confidence intervals for the point estimates are then calculated.

<sup>6</sup> To properly account for selection and staying probabilities, individuals are weighted by cross-sectional weights when computing cross-sectional inequality and by longitudinal weights when computing long-term inequality.

### 3. Results

First, results for annual and long-term redistribution are presented and discussed. Second, results for Theil decomposition of long-term inequality are shown uncovering the role of insurance and redistribution.

#### 3.1 Annual and Long-term Redistribution

Long-term and average cross-sectional inequality as measured with the Theil coefficient is computed for six age groups and for seven 20-year periods. Results for age group 3 over the 20-year period 1984–2003 are given Table 5. In 1984, group members are aged 21–30 years and in 2003 40–49 years. Long-term inequality is based on the NPV of income streams in the 20-year-period. Average cross-sectional inequality is the 20-year average of annual inequality within age group 3.

Comparing Theil coefficients of long-term income to cross-sectional income shows that inequality within the age group is lower in the long-term than annually regardless of the underlying income concept. The equalizing effect of time is in line with many studies confirming that inequality declines when extending the measurement period.<sup>7</sup> When there is mobility in income over time, long-term inequality will be lower than annual inequality as moving up and down the income distribution will make the distribution of long-term income more equal (Shorrocks, 1978). Additionally, post-government income distribution is more equal than pre-government income distribution confirming the results of Blomquist (1981).

To understand how redistribution changes when extending the measurement period from one to 20 years, the distance between gross and net household income inequality is compared for both annual and long-term income. In the long-term, inequality is reduced by about 38 percent through state intervention, whereas annually the reduction amounts to 44 percent. Hence, the long-term redistributive impact is smaller than the annual incidence insinuates, which coincides with Nelissen (1998).

Figures 1 and 2 show long-term and average cross-sectional inequality. Each graph depicts the age-group-specific results for annual and long-term inequality over seven 20-year periods. Three graphs per age group contrast results for gross household income, modified gross household income and net household income. All graphs in Figures 1 and 2 confirm that extending the measurement period decreases the measured income inequality, since long-term inequality is

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<sup>7</sup> See, for example, Björklund (1993), Bönke et al. (2011), Burkhauser/Poupore (1997), Creedy (1991), Finkel et al. (2006), Gibson et al. (2001), Kopczuk et al. (2010), Nelissen (1998) and Shorrocks (1978). See Wodon and Yitzhaki (2003) for a formal proof of the occurrence of this fact.

lower than cross-sectional inequality in all seven periods for all age groups and all income concepts.

Table 4

**Annual and Long-term Redistribution 1984–2003, Age Group 3**

	1984–2003	annual inequality	long-term inequality
gross household income (ghi)		0.16	0.08
modified ghi = ghi + social security pensions		0.14	0.08
net household income*		0.09	0.05

*Source:* SOEP, own calculations.

*Note:* All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

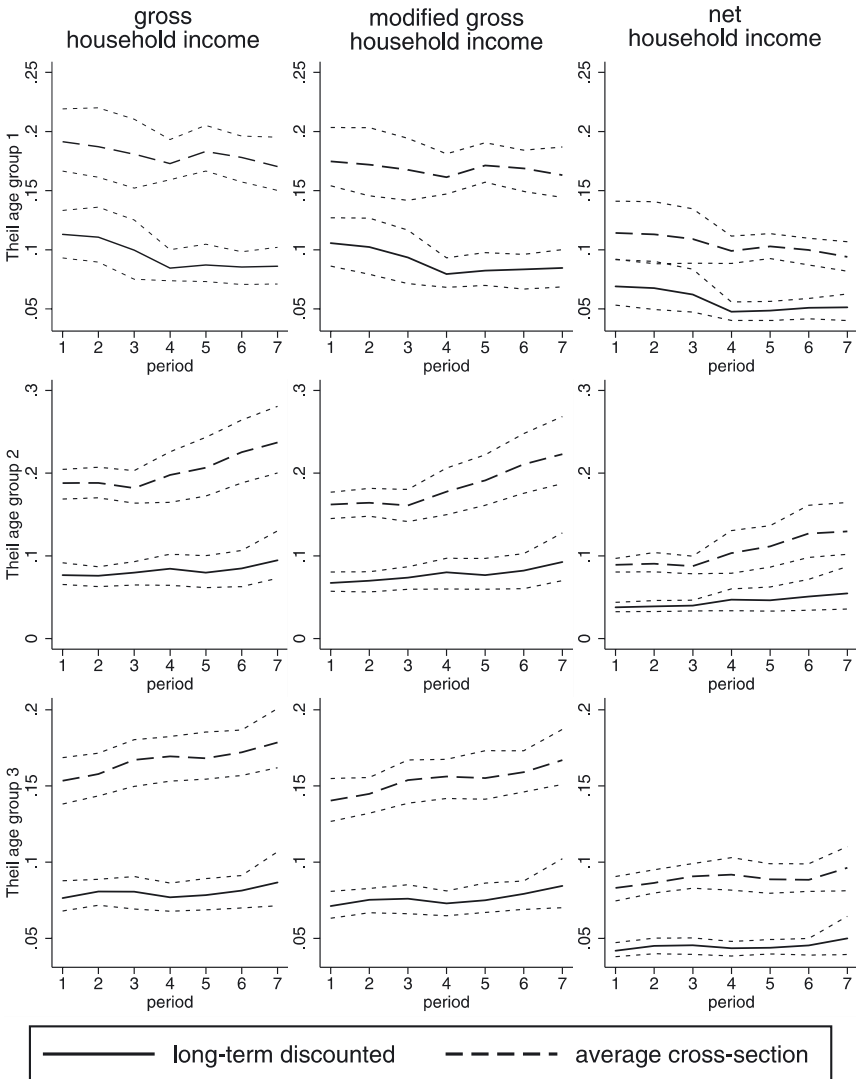
\* Net household income = modified ghi – social security contributions – income taxes + public transfers.

Furthermore, gross household income is both annually and in the long-term increasingly unequally distributed with age. The oldest age group – aged between 51 and 60 years in each period’s beginning and between 70 and 79 years at the end – experiences the highest levels of gross household income inequality, which can partly be explained by some individuals still working and others already being retired receiving no gross household income.

Children and younger people mostly receive income and transfers through their parents. Gross household income inequality for children and younger persons is reduced from about 0.18 to about 0.1 in the cross-section and from about 0.1 to about 0.06 in the long-term. Income inequality between children and younger persons is reduced through income taxes paid by the parents and through public transfers such as child benefits. Public transfers are most important for the youngest age group compared to the others as they contribute more than ten percent to long-term net household income (see Table 5).

The oldest age group not only experiences the highest level of gross household income inequality, but also the highest reduction of both annual and long-term inequality through state intervention. Income inequality between elderly is strongly reduced through state intervention from about 0.65 to about 0.1 in the cross-section and from about 0.25 to about 0.08 in the long-term. The main effect can be attributed to the payment of social security pensions which adds to modified gross household income (see middle graph of age group 6 in Figure 2). Social security pensions provide almost one half to net household income of elder persons. Since social security pensions have an effective minimum and maximum level, even though they reflect earnings from former earnings, they are much more compressed. But still some further inequality reduction takes place when adding public transfers and subtracting income taxes and social security contributions. In the view of quite similar levels of net house-

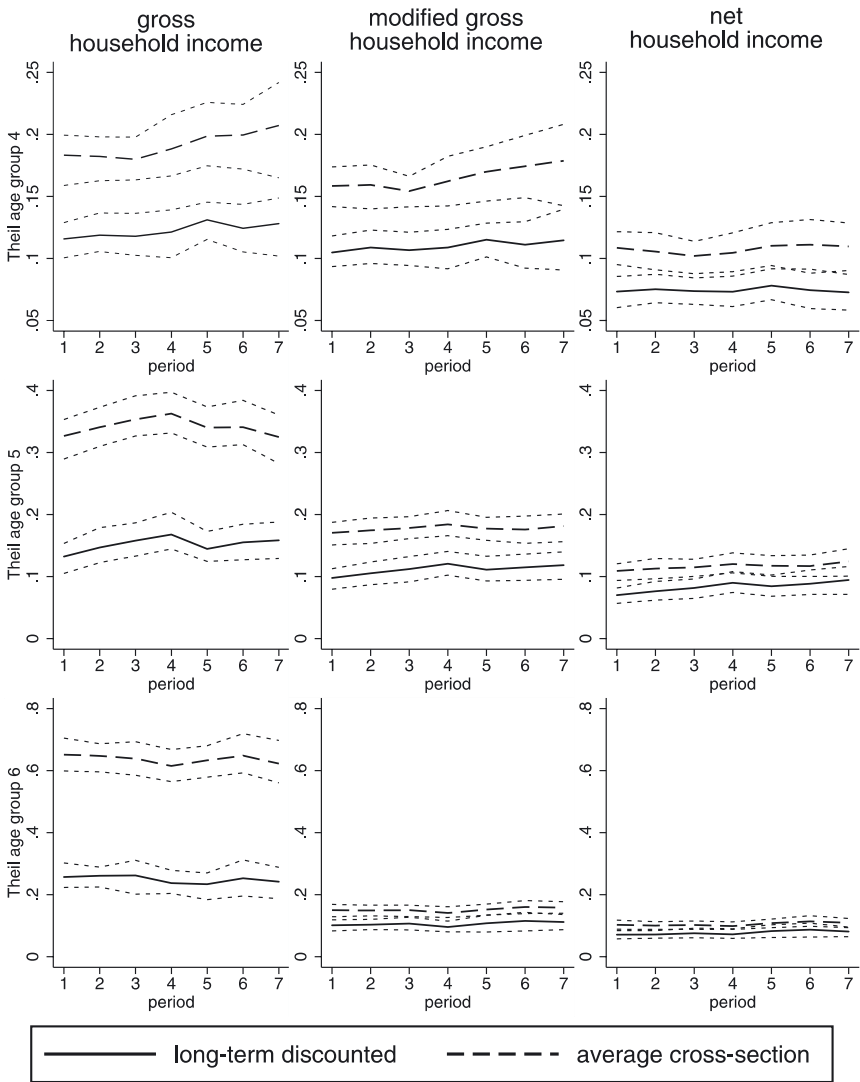
hold income inequality throughout the age groups, the welfare state seems to successfully reduce inequality among age groups, especially for older people.



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale. Dotted lines denote Hall's bootstrap confidence intervals at 95%-level.

Figure 1: Theil Coefficients for Gross and Net Household Income, Age Groups 1–3



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified scale. Dotted lines denote Hall's bootstrap confidence intervals at 95%-level.

Figure 2: Theil Coefficients for Gross and Net Household Income, Age Groups 4–6

We refrain from the discussing the development over time, since time effects could not be separated properly from cohort effects. Theil coefficients based on

CPI-deflated long-term incomes show that the use of deflated long-term incomes states slightly higher long-term inequality for both gross and net household income than when using the yield curve (see Appendix). On the whole, Gini coefficients given in the Appendix reproduce results of Theil coefficients.

The importance of the income components in long-term net equivalent household income for different age groups in the first 20-year period 1984–2003 is given in Table 5.<sup>8</sup> Because of the use of equivalent income also very young persons display a high share of labor earnings which is generated by their parents. Labor earnings are the most important income source, especially for the middle age groups. Asset income, private pensions and social security pensions become increasingly important with age for the most part. The opposite is true for private and public transfers that loose importance with age.

Table 5

**The Share of Long-term Income Components, 1984–2003**

age group	1	2	3	4	5	6
labor earnings	87.19	87.63	89.80	89.86	80.99	50.61
asset income	2.62	1.87	2.06	2.52	3.14	4.29
private transfers	0.76	0.52	0.43	0.29	0.38	0.31
private pensions	0.17	0.34	0.16	0.19	0.99	3.33
public transfers	7.44	6.21	5.55	4.34	3.89	3.20
social security pensions	1.82	3.43	2.01	2.80	10.62	38.26
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-12.39	-13.54	-14.26	-14.71	-13.06	-8.46
social security contributions	-13.18	-14.28	-14.16	-13.81	-13.00	-9.96

Source: SOEP, own calculations.

Note: All numbers are in percent and rounded such that they do not add up to 100 in all cases.

However, equivalent income is sensitive to household composition. Any changes in the household composition lead to changes in equivalent income. If, for example, a child is born into a household, gross equivalent household income of all household members declines. At the same time, equivalent social security contributions also decline and do not rise because medical care for a child is granted for free.

<sup>8</sup> Tables for the other 20-year periods are given in the Appendix since the shares do not vary much across periods.

### 3.2 Redistribution and Insurance

Table 6 provides a snapshot of the results and shows total, between-group and within-group inequality for five income concepts over the 20-year period 1984–2003 when all ages are considered jointly.<sup>9</sup> Gross household income inequality between persons lies at 0.15 and is thus higher than individuals' income variation at 0.13. The effect of specific welfare state measures is revealed by subtracting taxes or adding transfers stepwise. The change of the between group component gives a measure for the reduction of inter-individual income redistribution. The progressive tax system contributes most to even out income differences. Between-group inequality decreases from 0.14 to 0.11 when subtracting income taxes. This pattern is also found by Björklund and Palme (2002) in the Swedish case. Public transfers play the second biggest role in reducing inter-individual inequality from 0.11 to 0.10. Taxes and public transfers are similarly important to smoothen individual income. Subtracting income taxes reduces within-group inequality from 0.08 to 0.07 and adding transfers induces a reduction from 0.07 to 0.06. In contrast, the results of Björklund and Palme (2002) give public transfers a much higher weight than income taxes on smoothing income. Overall, income inequality between persons is reduced by 37 percent, whereas 56 percent of individual income variation is evened out through state intervention. The larger reduction of intra-individual inequality suggests that the German welfare state possibly puts more emphasis on insurance than on redistribution. Indeed, Bartels and Bönke (2012) find that insurance against transitory labor market risks outweighs the reduction of permanent earnings differences in Germany.

Table 6

#### Theil Decomposition, All Ages

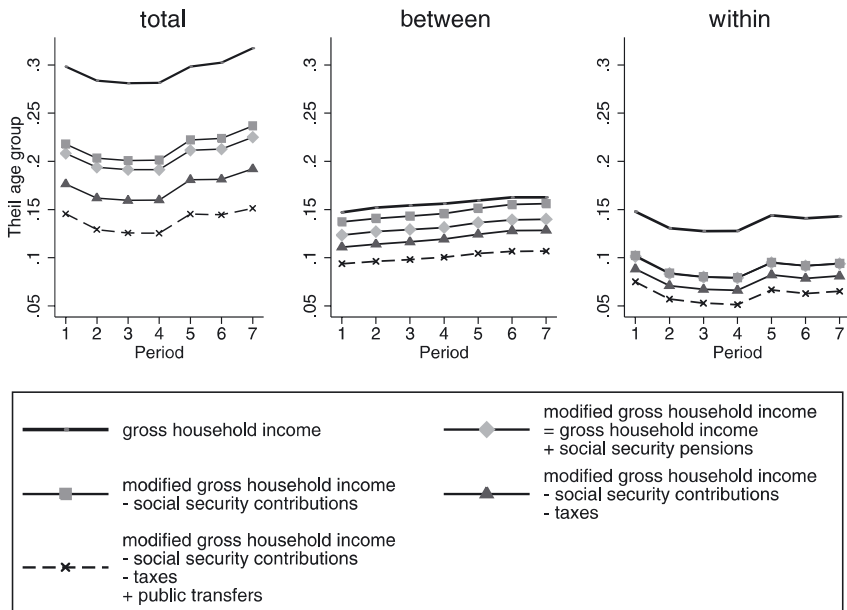
1984–2003	total	change	between	change	within	change
gross household income	0.28		0.15		0.13	
modified ghi = ghi + social security pensions	0.19	-0.32	0.13	-0.16	0.08	-0.36
modified ghi – social security contributions	0.20	-0.28	0.14	-0.07	0.08	-0.36
modified ghi – social security contributions – income taxes	0.16	-0.43	0.11	-0.25	0.07	-0.46
net household income*	0.13	-0.54	0.10	-0.37	0.06	-0.56

Source: SOEP, own calculations.

Note: All income concepts are defined as equivalent income using the OECD-modified equivalence scale. Changes refer to inequality of gross household income. \* Net household income = modified ghi – social security contributions – income taxes + public transfers.

<sup>9</sup> A Table including results for the remaining six 20-year periods is given in the Appendix.

Figures 3 to 7 present the results of Theil decomposition for all seven 20-year periods. Results for the decomposition over the whole population and an adult group 25 years and older are given in Figures 3 and 4, respectively. When considering only the adult population, as presented in Figure 4, income variation within a life-cycle becomes more pronounced than is the case for the whole population in Figure 3. Hence, the welfare states' role as an insurer gains importance as the weight of older age groups increases. For the adult population the within-group component is reduced even more by the welfare state intervention.



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

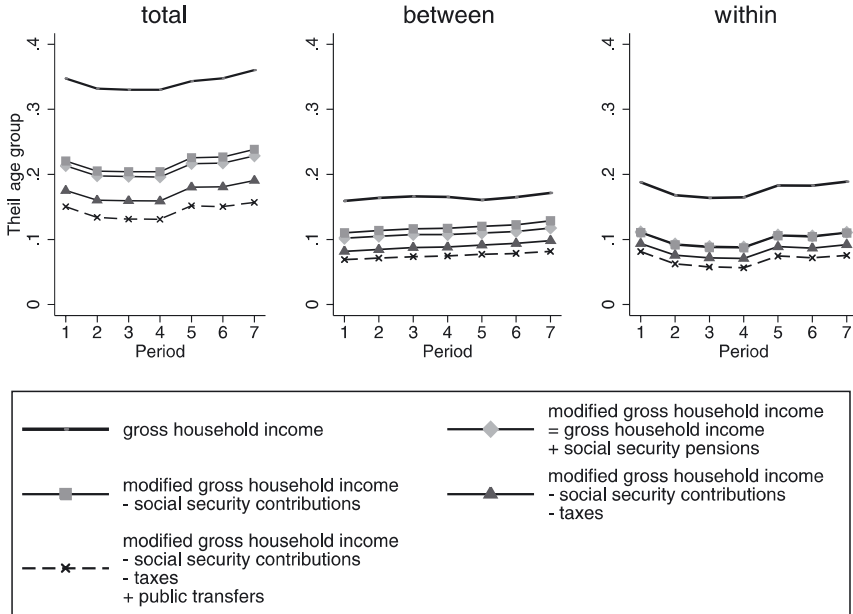
Figure 3: Theil Decomposition, All Ages

For almost all age groups total inequality of gross household income can be attributed to a greater extent to intra-individual inequality as Figure 5 and 6 present. However, this pattern gets more pronounced with age. About two thirds of total inequality within the oldest age group is explained by intra-individuals' differences meaning that differences within an individual's income stream over 20 years outweighs the differences between persons.

Comparing gross and net household income inequality reveals that between- and within-group inequality is reduced quite differently depending on age.<sup>10</sup>



Inter-individual inequality is reduced more (age group 1 and 4) or quite similarly to intra-individual inequality (age group 2 and 3). The role of progressive taxes and public transfers does not differ much from the pattern of the whole population seen in Figure 3.



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

Figure 4: Theil Decomposition, Above 25 Years

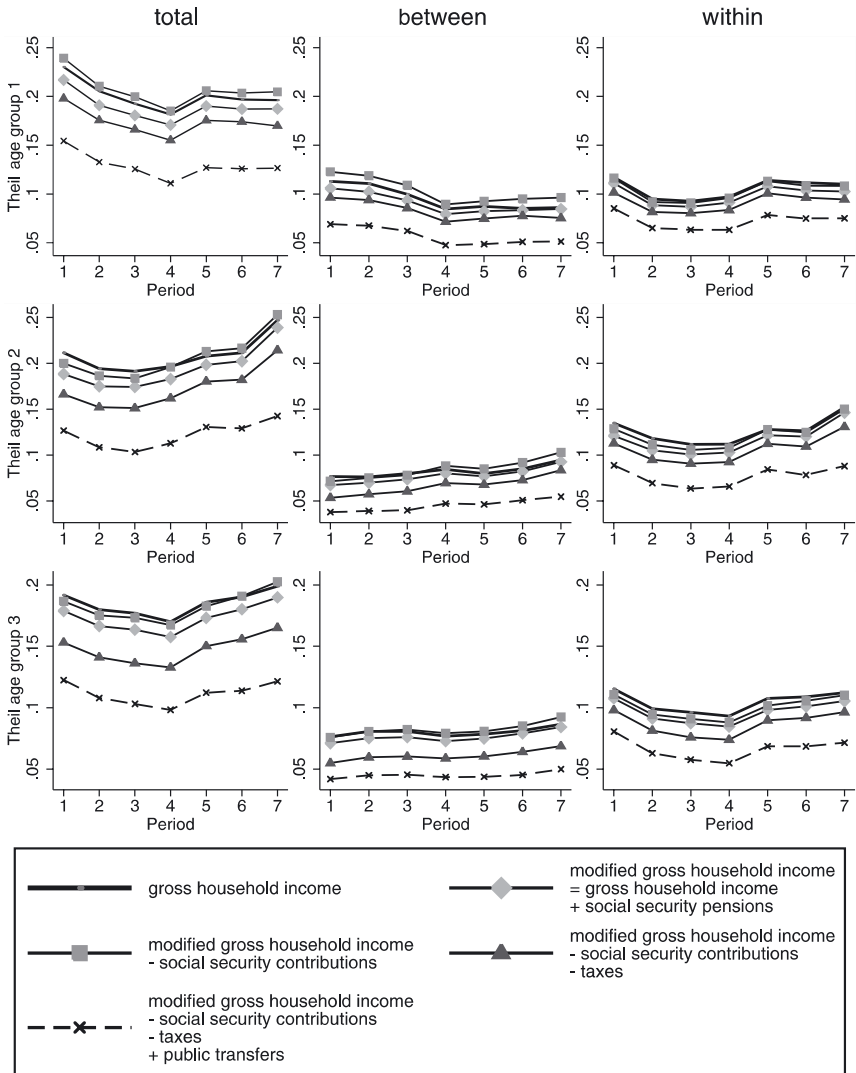
But the picture changes for the two oldest age groups. For age group 5 and 6 state intervention focusses on insurance reducing the within-group component far more than the between-group component. Not surprisingly, social security pensions become central for the two oldest age groups reaching retirement age within the period. Adding social security pensions to gross household income brings down intra-individual inequality by as much as 70 percent for the oldest age group. This seems plausible since the German pension system is insurance-oriented. German social security pensions include insurance benefits that depend on accumulated relative income points. But the system also includes a redistributive component. Non-insurance benefits are granted for time of unem-

<sup>10</sup> Between-group inequality in Figures 5 and 6 corresponds to long-term inequality presented in Figures 1 and 2.

ployment, military service, schooling, child-raising or because of early retirement or a pension below minimum income. Börsch-Supan and Reil-Held (2001) state that 86 percent of male social security pensions are attributable to insurance benefits whereas only 13 percent stem from non-insurance benefits. For women, insurance benefits are 70 percent and non-insurance benefits amount to 30 percent where non-insurance benefits are mainly due to minimum benefits. Although the authors find that some non-insurance benefits, such as those for education, are given to almost every worker in the sample, they confirm redistribution between permanent earnings quintiles of the pensioners through non-insurance benefits. Indeed, the results for the oldest age group in Figure 6 clearly show the redistributive component: Social security pensions bring down inter-individual inequality by 57 percent. Nelissen (1998) does not distinguish between age groups, but also finds that Dutch old-age pensions have the highest impact on reducing lifetime inequality even though the Dutch system largely provides flat pensions. Nelissen (1998) finds the redistributive impact of Dutch pensions to be smaller over the lifetime than the annual incidence suggests, a result that is confirmed for Japan by Oshio (2005). As opposed to younger age groups, public transfers and income taxes play only a minor role for age groups 5 and 6.

Interestingly, adding social security contributions to modified gross household income increases both inter- and intra-individual income differences, except for the two oldest age groups. Social security contributions are paid as a fixed percentage of earnings such that contributions increase with earnings, but only up to a maximum amount. In 2008, the monthly earnings cap was at 3600 Euro for health care and 5300 for unemployment insurance and social security pensions. Wagstaff et al. (1999) confirm that social security contributions for health care have a regressive effect in Germany. But one should note that cash values of in-kind transfers such as medical care and utility gains through, e.g., unemployment insurance are not accounted for in the analysis. Research on this subject documented that including the cash value of in-kind transfers reduces inequality measures since almost half of the welfare state transfers in rich nations consists of in-kind transfers (see, e.g., Garfinkel et al., 2006, and Paulus et al., 2009).

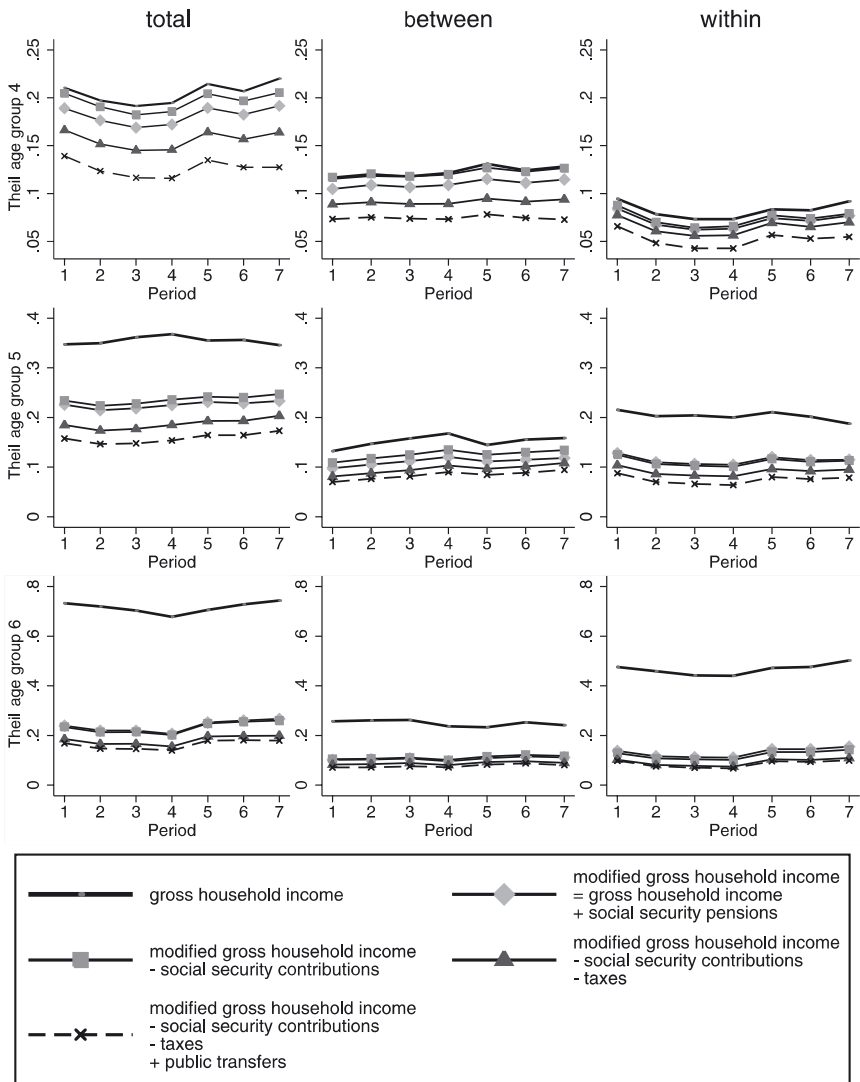
If only one-person-households of all ages are considered, inequality of gross household income is far higher as can be taken from Figure 7. First, income differences between one-person-households are far higher than between larger households. Second, intra-household income pooling cannot occur to reduce individual income variations. But the result that the welfare states aims more at smoothing income than at redistributing income also remains for one-person-households.



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

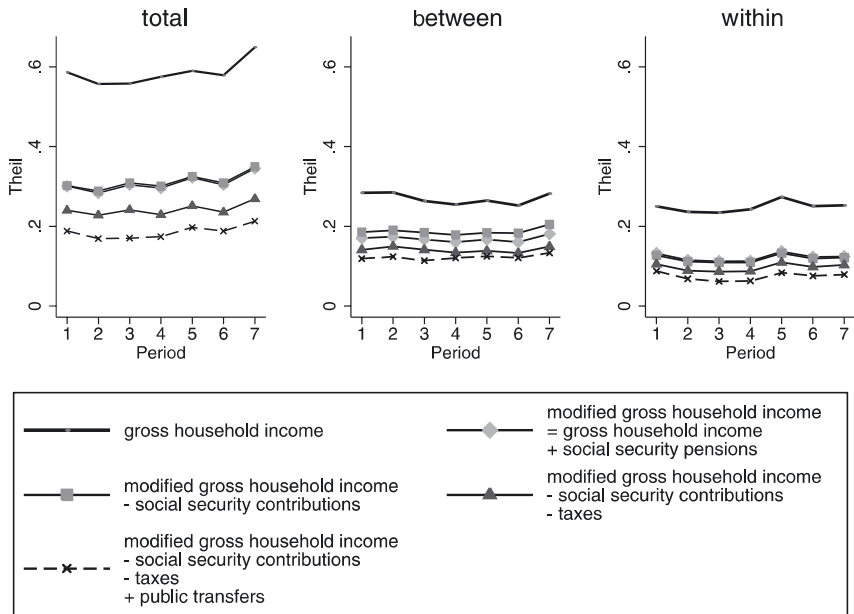
Figure 5: Theil Decomposition, Age Groups 1–3



Source: SOEP, own calculations,

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

Figure 6: Theil Decomposition, Age Groups 4–6



Source: SOEP, own calculations.

Notes: All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

Figure 7: Theil Decomposition, One-person-households, All Ages

#### 4. Conclusion

This paper analyzes the impact of a redistributing welfare state on inter- and intra-individual redistribution by including taxes and transfers such as income taxes, social security contributions and public transfers. Long-term income is measured as the NPV of equivalent household income streams over a 20-year period discounted by the yield curve. The results confirm the literature that inequality in the long-term is lower than annually, but the effect of redistribution is also lower when measured in the long-term. Switching from the annual to the long-term perspective allows to gain additional insights on impacts such as income smoothing via insurance payments of the social security system. A decomposition approach is used to identify inequality between people and income variation within an individual's life (inter- and intra-individual inequality) in a first stage and redistribution between people and income smoothing within a life-cycle in a second stage. Results show that the German welfare state clearly gives priority to insurance over redistribution. The scope of this priority depends on the stage of the life-cycle. When persons are young, state intervention also notably redistributes between people through the progressive tax system

and public transfers. Getting older and reaching retirement age intra-individual redistribution via social security pensions becomes central. Social security pensions reduce intra-individual inequality by 70 percent for the oldest age group. In an individual's life-cycle perspective one could thus conclude that the welfare state evolves from being a poverty reliever in earlier years to an insurer in later years. However, the presumably inequality reducing effect of in-kind transfers such as health care and education is not included in the analysis. Overall, in his role as an insurer, social security pensions is the most important instrument of the welfare state in smoothing income over time.

## References

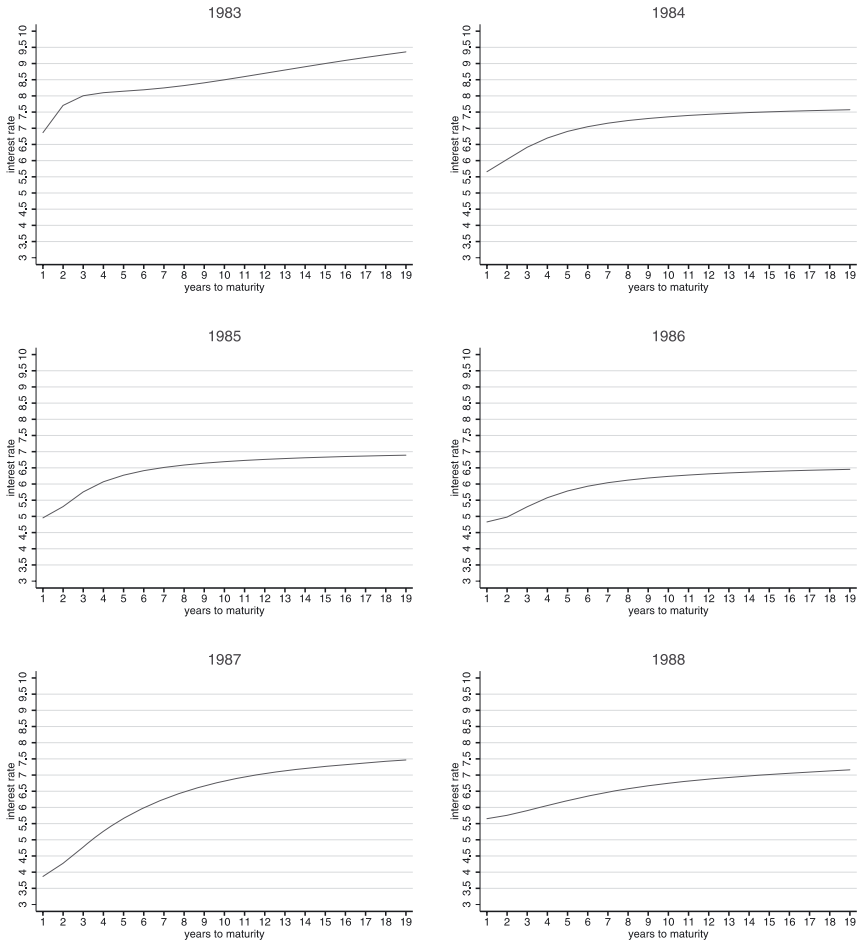
- Bartels, C./Bönke, T.* (2012): Can households and welfare states mitigate rising earnings instability?, *Review of Income and Wealth*, forthcoming.
- Biewen, M.* (2001): Item non-response and inequality measurement: Evidence from the German earnings distribution, *Allgemeines Statistisches Archiv* 85, 409–425.
- Björklund, A.* (1993): A comparison between actual distributions of annual and lifetime income: Sweden 1951–89, *Review of Income and Wealth* 39, 377–386.
- Björklund, A./Palme, M.* (2002): Income redistribution within the life cycle versus between individuals: Empirical evidence using Swedish panel data, in: D. Cohen/T. Piketty/G. Saint-Paul (eds.): *The economics of rising inequality*, Oxford.
- Blomquist, S.* (1981): A comparison of distributions of annual and lifetime income: Sweden around 1970, *Review of Income and Wealth* 2, 243–264.
- Bönke, T./Corneo, G./Lüthen, H.* (2011): Lifetime earnings inequality in Germany, *DIW-Discussion Paper No. 1160/2011*.
- Börsch-Supan, A./Reil-Held, A.* (2001): How much is transfer and how much is insurance in a pay-as-you-go system? The German case, *Scandinavian Journal of Economics* 103, 505–524.
- Bundesbank* (1991): *Monatsbericht der Deutschen Bundesbank Juli 1991*, Frankfurt am Main.
- Burkhauser, R. V./Poupore, J. G.* (1997): A cross-national comparison of permanent inequality in the United States and Germany, *Review of Economics and Statistics* 79, 10–17.
- Burkhauser, R. V./Warlick, J.* (1981): Disentangling the annuity from the redistributive aspects of social security in the United States, *Review of Income and Wealth* 27, 401–421.
- Canberra Group* (2001): *Expert Group on Household Income Statistics: Final Report and Recommendations*, Ottawa.
- Creedy, J.* (1991): Lifetime earnings and inequality, *Economic Record* 67, 46–58.
- Finkel, Y./Artsev, Y./Yitzhaki, S.* (2006): Inequality measurement and the time structure of household income in Israel, *Journal of Economic Inequality* 4, 153–179.

- Frick, J.* (2006): A general introduction to the German Socio-Economic Panel Study (SOEP) - Design, contents and data structure (waves A–V, 1984–2005).
- Frick, J./Grabka, M.* (2005): Item nonresponse on income questions in panel surveys: incidence, imputation and the impact on inequality and mobility, *Allgemeines Statistisches Archiv* 89, 49–61.
- Garfinkel, I./Rainwater, L./Smeeding, T. M.* (2006): A re-examination of welfare states and inequality in rich nations: How in-kind transfers and indirect taxes change the story, *Journal of Policy Analysis and Management* 25, 897–919.
- Gibson, J./Huang, J./Rozelle, S.* (2001): Why is income inequality so low in China compared to other countries? The effect of household survey method, *Economics Letters* 71, 329–333.
- Gustman, A. L./Steinmeier, T. L.* (2001): How effective is redistribution under the social security benefit formula?, *Journal of Public Economics* 82, 1–28.
- Haisken-DeNew, J./Frick, J.* (2005): Desktop companion to the German Socio-Economic Panel Study (GSOEP), Version 8.0 – Update to Wave 21. German Institute for Economic Research, Berlin.
- Hall, P.* (1994): Methodology and theory for the bootstrap, in: R. Engle/D. McFadden (eds.): *Handbook of Applied Econometrics* 4, Amsterdam.
- Kopczuk, W./Saez, E./Song, J.* (2010): Earnings inequality and mobility in the United States: Evidence from social security data since 1937, *The Quarterly Journal of Economics* 125, 91–128.
- Layard, R.* (1977): On measuring the redistribution of lifetime income, in: M. Feldstein/R. Inman (eds.): *The Economics of Public Services*, London.
- Liebman, J. B.* (2001): Redistribution in the current U.S. social security system, NBER Working Paper No. 8625.
- Mills, J./Zandvakili, S.* (1997): Statistical inference via bootstrapping for measures of inequality, *Journal of Applied Econometrics* 12, 133–150.
- Nelissen, J. H. M.* (1998): Annual versus lifetime income redistribution by social security, *Journal of Public Economics* 68, 223–249.
- Nelson, C. R./Siegel, A. F.* (1987): Parsimonious modeling of yield curves, *Journal of Business* 60, 473–489.
- Oshio, T.* (2005): Intra-age, inter-age and lifetime income redistribution, *Japanese Economic Review* 56, 85–106.
- Paulus, A./Sutherland, H./Tsakloglou, P.* (2009): The distributional impact of in kind public benefits in European countries, IZA-Discussion Paper No. 4581.
- Pettersson, T./Pettersson, T.* (2007): Lifetime redistribution through taxes, transfers and non-cash benefits, in: A. Harding/A. Gupta (eds.): *Modelling our future: population ageing social security and taxation*, International symposia in economic theory and econometrics, Amsterdam.
- Sandmo, A.* (1999): The public economics of redistribution and the welfare state, *Review of Population and Social Policy* 8, 139–154.

- Schwarze, J.* (1995): Simulating German income and social security tax payments using the SOEP. Cross-national studies in aging, Programm Project Paper No. 19, All-University Gerontology Center Maxwell School of Citizenship and Public Affairs, Syracuse University, New York.
- Shorrocks, A.* (1978): Income inequality and income mobility, *Journal of Economic Theory* 19, 376–393.
- Svensson, L.E.O.* (1994): Estimating and interpreting forward interest rates: Sweden 1992–1994, NBER Working Paper No. 4871.
- Wagner, G./Frick, J./Schupp, J.* (2007): The German Socio-Economic Panel Study (SOEP) – Scope, evolution and enhancements, *Schmollers Jahrbuch* 127, 139–170.
- Wagstaff, A./van Doorslaer, E./van der Burg, H./Calonge, S./Christiansen, T./Citoni, G./Gerdtham, U./Gerfin, M./Gross, L./Häkinnen, U./Johnson, P./John, J./Klavus, J./Lachaud, C./Lauritsen, J./Leu, R./Nolan, B./Perán, E./Pereira, J./Propper, C./Puffer, F./Rochaix, L./Rodríguez, M./Schellhorn, M./Sundberg, G./Winkelhake, O.* (1999): Equity in the finance of health care: some further international comparisons, *Journal of Health Economics* 18, 263–290.
- Wodon, Q./Yitzhaki, S.* (2003): Inequality and the accounting period, *Economics Bulletin* 4, 1–8.



Appendix



Note: Own calculations based on listed German Federal Treasury bonds.

Source: [http://www.bundesbank.de/statistik/statistik\\_zeitreihen.php?func=list&tr=www.s300\\_it03c](http://www.bundesbank.de/statistik/statistik_zeitreihen.php?func=list&tr=www.s300_it03c)  
(16. 12. 2011)

Figure A1: Yield Curves for the Years 1983 to 1989

Table A1

**The Share of Long-term Income Components**

age group	1	2	3	4	5	6
1983–2002						
labor earnings	87.60	88.12	90.33	90.36	82.57	52.37
asset income	2.38	1.71	1.97	2.37	2.79	4.25
private transfers	0.67	0.54	0.44	0.30	0.35	0.28
private pensions	0.16	0.27	0.15	0.16	0.89	2.96
public transfers	7.30	5.94	5.22	4.13	3.87	3.23
social security pensions	1.89	3.43	1.90	2.67	9.53	36.91
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-12.34	-13.50	-14.08	-14.94	-13.06	-8.71
social security contributions	-13.10	-14.12	-14.14	-13.52	-13.17	-9.91
1985–2004						
labor earnings	87.34	87.64	89.26	89.58	79.86	50.73
asset income	2.59	1.85	2.07	2.72	3.38	4.39
private transfers	0.72	0.56	0.42	0.32	0.40	0.25
private pensions	0.19	0.26	0.21	0.21	1.08	3.40
public transfers	7.51	6.44	5.84	4.25	3.95	3.53
social security pensions	1.65	3.24	2.19	2.93	11.33	37.70
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-12.34	-13.42	-14.31	-14.77	-13.08	-8.36
social security contributions	-13.29	-14.52	-14.08	-14.05	-12.88	-10.12
1986–2005						
labor earnings	86.86	87.60	89.01	89.15	79.48	51.51
asset income	2.60	1.93	2.14	2.75	3.51	4.69
private transfers	0.78	0.55	0.43	0.32	0.39	0.23
private pensions	0.17	0.26	0.18	0.22	1.09	3.28
public transfers	8.00	6.49	6.11	4.57	3.92	3.08
social security pensions	1.58	3.17	2.13	2.99	11.61	37.21
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-12.08	-13.39	-14.35	-14.64	-13.16	-8.57
social security contributions	-13.40	-14.49	-14.15	-14.13	-12.81	-10.28

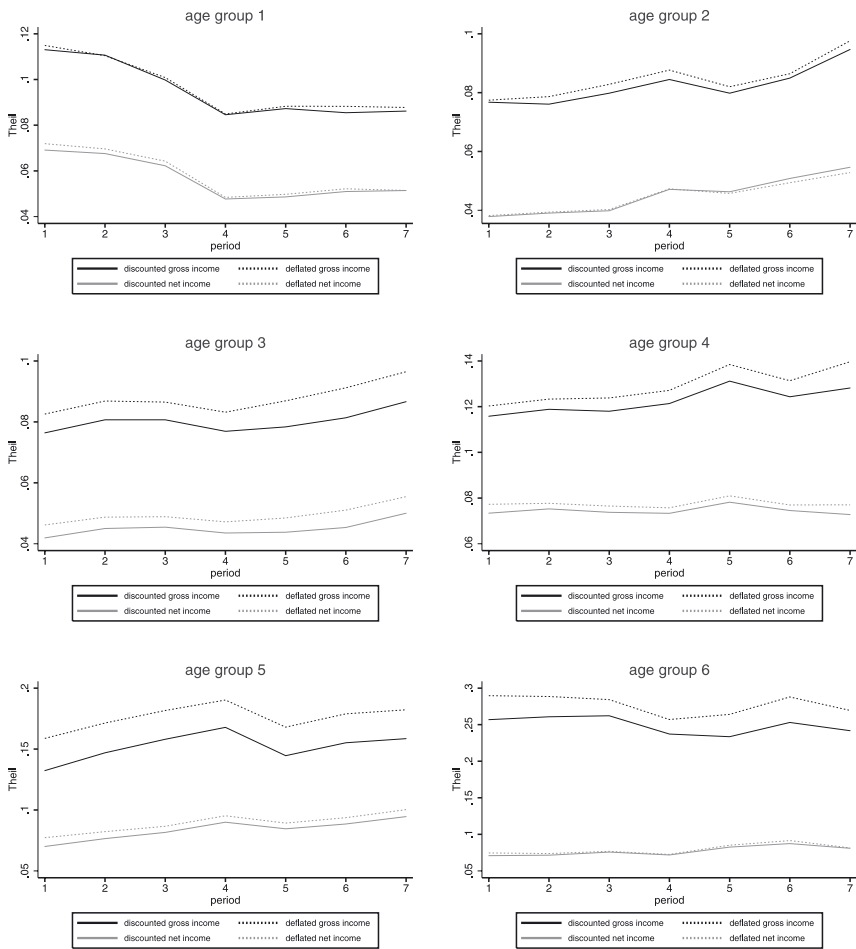
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age group	1	2	3	4	5	6
1987–2006						
labor earnings	86.10	87.99	88.60	88.60	80.61	53.64
asset income	2.75	2.05	2.15	2.99	3.66	4.71
private transfers	0.76	0.59	0.45	0.31	0.46	0.24
private pensions	0.17	0.24	0.18	0.22	0.98	3.15
public transfers	8.58	6.38	6.49	4.77	3.81	3.19
social security pensions	1.64	2.75	2.12	3.11	10.47	35.06
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-11.96	-13.41	-14.23	-14.61	-13.54	-8.82
social security contributions	-13.24	-14.66	-14.27	-14.00	-12.87	-10.49
1988–2007						
labor earnings	85.46	87.65	88.29	88.48	80.22	53.12
asset income	2.77	2.22	2.14	3.10	3.58	4.93
private transfers	0.83	0.54	0.47	0.36	0.32	0.51
private pensions	0.21	0.25	0.19	0.26	0.96	3.38
public transfers	9.05	6.65	6.78	5.11	3.97	2.93
social security pensions	1.68	2.69	2.13	2.70	10.96	35.13
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-11.80	-13.29	-14.21	-14.78	-13.50	-8.80
social security contributions	-13.52	-14.63	-14.48	-13.93	-13.05	-10.54
1989–2008						
labor earnings	85.75	87.32	88.36	87.90	81.06	53.97
asset income	2.56	2.20	2.10	3.23	3.82	4.76
private transfers	0.90	0.59	0.50	0.34	0.33	0.47
private pensions	0.17	0.22	0.21	0.27	0.93	3.25
public transfers	9.12	7.04	6.93	5.38	3.84	3.18
social security pensions	1.49	2.63	1.90	2.88	10.01	34.37
total	100.00	100.00	100.00	100.00	100.00	100.00
income taxes	-11.87	-13.24	-14.39	-14.69	-13.95	-8.86
social security contributions	-13.77	-14.44	-14.66	-13.71	-13.07	-10.71

Source: SOEP, own calculations.

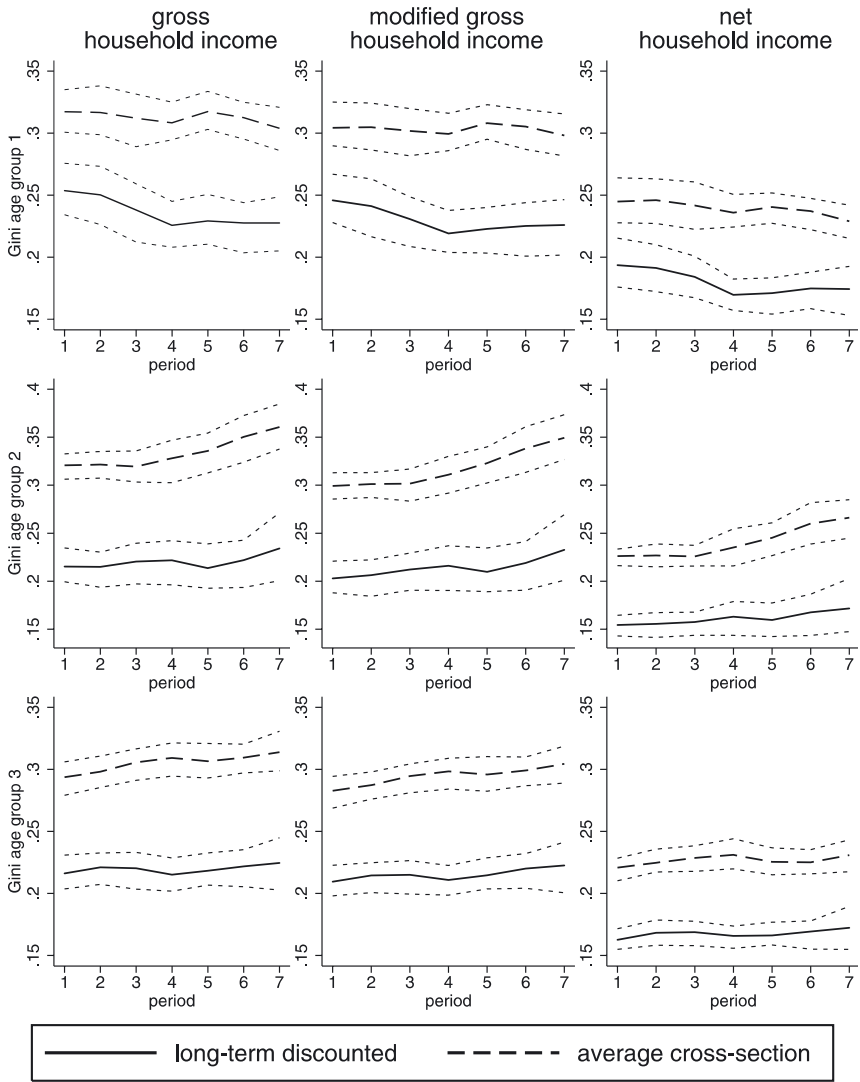
Note: All numbers are in percent and rounded such that they do not add up to 100 in all cases. All income concepts are defined as equivalent income using the OECD-modified equivalence scale.



Source: SOEP, own calculations.

Note: All numbers are in percent and rounded such that they do not add up to 100 in all cases. All income concepts are defined as equivalent income using the OECD-modified equivalence scale.

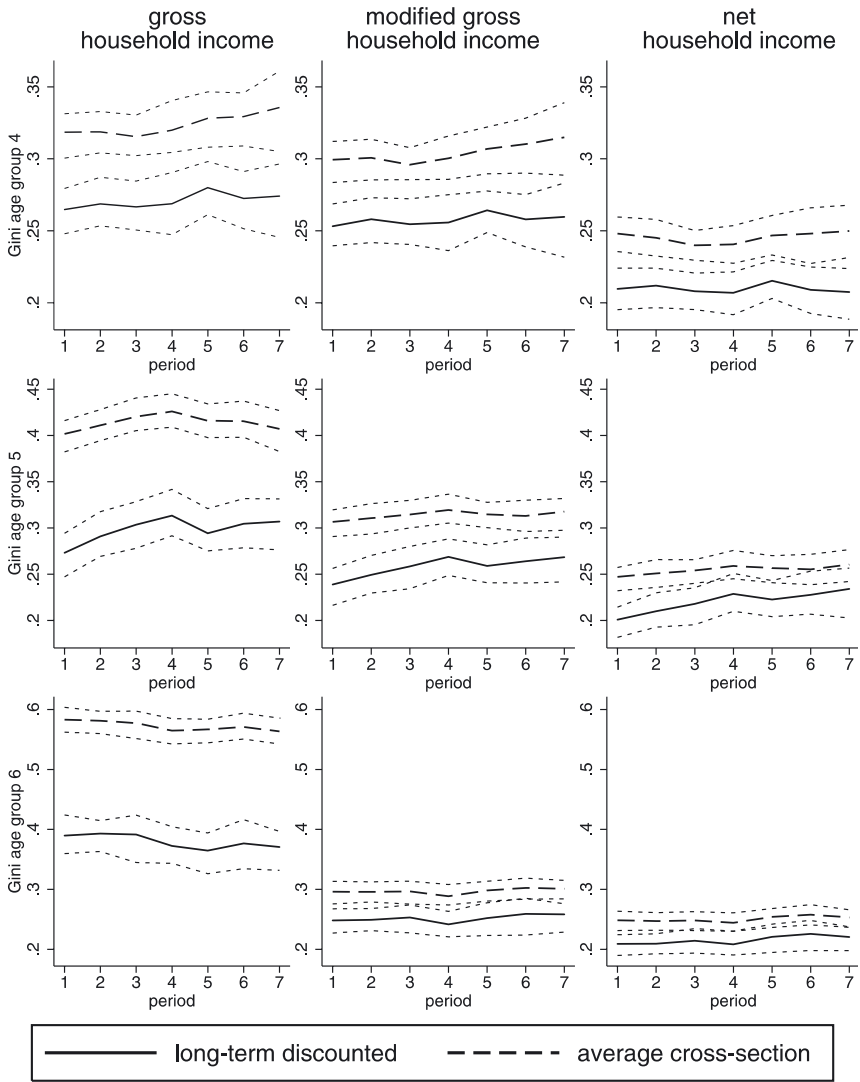
Figure A2: Theil Coefficients for Deflated and Discounted Long-term Household Income



Source: SOEP, own calculations.

Note: All numbers are in percent and rounded such that they do not add up to 100 in all cases. All income concepts are defined as equivalent income using the OECD-modified equivalence scale. Dotted lines denote Hall's bootstrap confidence intervals at 95%-level.

Figure A3: Gini Coefficients for Gross and Net Household Income, Age Groups 1–3



Source: SOEP, own calculations

Note: All numbers are in percent and rounded such that they do not add up to 100 in all cases. All income concepts are defined as equivalent income using the OECD-modified equivalence scale. Dotted lines denote Hall's bootstrap confidence intervals at 95%-level.

Figure A4: Gini Coefficients for Gross and Net Household Income, Age Groups 4–6

Table A2

**Theil Decomposition, All Ages**

income concept	total	change	between	change	within	change
1983–2002						
gross household income	0.30		0.15		0.15	
modified ghi = ghi + social security pensions	0.21	-0.30	0.12	-0.16	0.10	-0.31
modified ghi – social security contributions	0.22	-0.27	0.14	-0.07	0.10	-0.31
modified ghi – social security contributions – income taxes	0.18	-0.41	0.11	-0.25	0.09	-0.40
net household income*	0.15	-0.51	0.09	-0.36	0.08	-0.49
1984–2003						
gross household income	0.28		0.15		0.13	
modified ghi = ghi + social security pensions	0.19	-0.32	0.13	-0.16	0.08	-0.36
modified ghi – social security contributions	0.20	-0.28	0.14	-0.07	0.08	-0.36
modified ghi – social security contributions – income taxes	0.16	-0.43	0.11	-0.25	0.07	-0.46
net household income*	0.13	-0.54	0.10	-0.37	0.06	-0.56
1985–2004						
gross household income	0.28		0.15		0.13	
modified ghi = ghi + social security pensions	0.19	-0.32	0.13	-0.16	0.08	-0.37
modified ghi – social security contributions	0.20	-0.29	0.14	-0.07	0.08	-0.37
modified ghi – social security contributions – income taxes	0.16	-0.43	0.12	-0.24	0.07	-0.47
net household income*	0.13	-0.55	0.10	-0.36	0.05	-0.58
1986–2005						
gross household income	0.28		0.16		0.13	
modified ghi = ghi + social security pensions	0.19	-0.32	0.13	-0.16	0.08	-0.38
modified ghi – social security contributions	0.20	-0.28	0.15	-0.07	0.08	-0.38
modified ghi – social security contributions – income taxes	0.16	-0.43	0.12	-0.23	0.07	-0.48
net household income*	0.13	-0.55	0.10	-0.36	0.05	-0.60

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(Continue Table A2)

income concept	total	change	between	change	within	change
1987–2006						
gross household income	0.30		0.16		0.14	
modified ghi = ghi + social security pensions	0.21	-0.29	0.14	-0.14	0.09	-0.34
modified ghi – social security contributions	0.22	-0.25	0.15	-0.05	0.10	-0.34
modified ghi – social security contributions – income taxes	0.18	-0.39	0.12	-0.22	0.08	-0.43
net household income*	0.15	-0.51	0.10	-0.34	0.07	-0.54
1988–2007						
gross household income	0.30		0.16		0.14	
modified ghi = ghi + social security pensions	0.21	-0.30	0.14	-0.14	0.09	-0.35
modified ghi – social security contributions	0.22	-0.26	0.16	-0.05	0.09	-0.35
modified ghi – social security contributions – income taxes	0.18	-0.40	0.13	-0.21	0.08	-0.44
net household income*	0.14	-0.52	0.11	-0.34	0.06	-0.55
1989–2008						
gross household income	0.32		0.16		0.14	
modified ghi = ghi + social security pensions	0.23	-0.29	0.14	-0.14	0.09	-0.34
modified ghi – social security contributions	0.24	-0.25	0.16	-0.04	0.09	-0.34
modified ghi – social security contributions – income taxes	0.19	-0.39	0.13	-0.21	0.08	-0.43
net household income*	0.15	-0.52	0.11	-0.34	0.07	-0.54

Source: SOEP, own calculations.

Note: All income concepts are defined as equivalent income using the OECD-modified equivalence scale. Changes refer to inequality of gross household income. \* Net household income = modified ghi – social security contributions – income taxes + public transfers.