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## Article

# Measuring Inflation: Testing Carl Menger's Concept of the Inner Value of Money

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## **Measuring Inflation: Testing Carl Menger's Concept of the Inner Value of Money**

By M. M. G. Fase and C. K. Folkertsma, Amsterdam\*

### **I. Introduction**

The main objective of this article is to identify the component of observable price changes, which is due to monetary as opposed to productivity and other real shocks. In other words, its main concern is the measurement of genuine inflation. In order to achieve this objective, it applies an old concept of Austrian economics and its monetary theory in particular. It attempts to operationalize Carl Menger's old concept of the inner value of money as the true measure for inflation. This operationalization is applied to the Netherlands and the European Union, yielding a measure of price changes, which reflects more closely the theoretical notion of inflation as a monetary phenomenon.

According to the definition adopted here, inflation is any increase in prices induced by monetary factors. Contrary to Friedman's well-known definition of inflation as "a steady and sustained rise in prices" (Friedman (1963, p. 1)), a non-recurring price change is considered as (short-term) inflation as long as it is due to monetary influences. Clearly, without stating that the price changes are induced by monetary factors, inflation would not be "always and everywhere a monetary phenomenon" (Friedman (1963, p. 17)), since short periods of rising prices may, after all, be due to real factors such as a change in productivity alone. The adoption of the broader inflation concept seems justified since in economic theory the important distinction is not between the effects of a temporary and a sustained price increase but between anticipated and unanticipated changes. Furthermore, from the perspective of monetary policy, it is interesting to measure any movement in prices brought about by monetary shocks, irrespective of whether the movements are temporary or sustained.

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This article is organised as follows. Section II. argues why the change in the consumer price index (CPI) and comparable indices are inappropriate for measuring inflation. Section III. goes into Carl Menger's distinction between "innerer" and "äußerer Tauschwert des Geldes" which is our main inspiration for this research (Fase (1986, pp. 9–10)). Section IV. proposes a decomposition of price changes. Four possible inflation gauges are examined, the aim being to establish which components of price changes they identify. Section V. discusses a method to identify the change in the inner value of money insofar as price rises are caused by monetary shocks. Section VI. applies this method to Dutch and European data. Conclusions are drawn in the final section.

## II. Changes in Price Indices as Inflation Gauge

The change in the CPI published monthly is seen by the public and by politicians as the measure of inflation. The change in this index gauges the increase in expenditure on a basket of goods and services consumed by the representative household. Roughly, the rise in the CPI reflects the loss of purchasing power of money as the representative household experiences it. Application of this index is justified if the aim is to determine changes in the spending potential of households, which are caused by price movements ensuing from changes in monetary policy, or from changes in fiscal policy, productivity increases and other real causes. However, this gauge is inappropriate for measuring inflation or changes of the *general* price level brought about by monetary shocks as the CPI reflects every change in consumer prices, whether they are caused by real or monetary shocks.

Several attempts have been made to restructure the CPI into a better measure of inflation and several of the resulting inflation measures are used for monetary policy purposes. For instance, in the Netherlands as well as other countries, a derived consumer price index is published, which has been adjusted for changes in indirect taxation and subsidies. Another index that is quite commonly used for monetary policy purposes is the CPI excluding the prices for food and energy. Moreover, various statistical methods have been developed to identify the trend component in the change in prices. Nevertheless, even if the index is corrected for e.g. the direct influence of changed indirect taxes and subsidies, the index still reacts to price changes that have been generated by second-order effects of tax and subsidy adjustments. Moreover, the derived CPI or the CPI excluding food and energy prices will react to all other real

influences. Furthermore, the weighting scheme of the price index means that some prices will determine the general price level thus measured to a greater extent than others. For an assessment of changes in purchasing power, the weighting scheme has a theoretical foundation in consumer theory but there is no clear reason to weight prices by budget shares when gauging monetary inflation. Finally, the basket of goods consumed by households is but a subset of the goods marketed within the economy as a whole. Notably the prices of the various factors of production are left out of account.

Its partial nature, the weighting aspect, and its inability to distinguish between real and monetary causes of price changes make the CPI an unsuitable instrument for gauging monetary inflation. For similar reasons, other frequently used price indices such as the producer price index or the implicit deflator of gross domestic product do not constitute better tools for measuring inflation either.

### **III. Menger's Concept of "Innerer Tauschwert des Geldes"**

At the end of the nineteenth century, Carl Menger (Menger (1923)) introduced the concepts of the "innerer" (i. e. inner) and the "äußerer Tauschwert" (i. e. outer value) of a commodity, and of money in particular. By the outer value of a commodity, he meant the price of that commodity or the amount of money, which is to be exchanged for the commodity in equilibrium. Analogously, the outer value of money is its purchasing power, viz. the commodity bundle that can be exchanged for one unit of account. In Menger's terminology, the CPI is thus a measure of the outer value of money. While Menger stressed that the ratio at which two goods are exchanged in equilibrium is ultimately determined by the (marginal) subjective valuation of the goods involved, he avers that a change in the relationship may be caused by changes affecting only one of the goods. He calls these change movements in the inner value of a good. Analogously, changes in the inner value of money are those price changes, which are due to purely monetary factors. These monetary factors are changes of the money stock and shocks to the velocity if interpreted in terms of the equation of exchange.

According to Menger, a decrease in the inner value of money must lead to a proportional increase in all goods prices. After all, if the changes relate solely to money, the relative goods prices will, in his view, remain unchanged. However, he does acknowledge that a proportional rise in all

prices need not necessarily constitute a fall in the inner value of money, because this may also be caused by real factors affecting the production of all commodities simultaneously. That is why Menger is sceptical about the possibility of measuring changes in the inner value of money. He mentions measurement based on the distribution of price changes as a possible way of operationalization. If all prices rise by the same percentage, the hypothesis that the inner value of money has fallen is more likely than the hypothesis that the inner value of all goods has gone up to the same extent. The likelihood of this conclusion rests on the fact that the first explanation relates to the changes in the value of fewer objects of exchange. If not all goods prices go up by the same percentage, then the change in the inner value of money could, based on the same argument, be estimated with the aid of the mode of the frequency distribution of the price changes. However, Menger indicates himself that the method becomes less convincing as the spread of the price changes increases.

Although Menger introduced the concept of the inner value of money, it should be noted that the underlying notion is closely related to the distinction made by earlier scholars and later writers in the Austrian economic tradition between price changes brought about by causes “on the commodity side” and on “the money side” (see von Mises (1934) and Jordan (1998)). Already at Menger’s time, the distinction between these two factors of price changes had a long history, which goes at least back to the work of Samuel Pufendorf (1672, p. 698), a German law professor of the seventeenth century. Selgin’s (1995) recent reconsideration of the productivity norm, cherished by the proponents of the neutral money tradition (see Fase (1994)) belongs to the same line of thoughts.

Clearly, Menger’s concept of the inner value of money is closely related to the definition of inflation used in this paper and in macroeconomic theory in general. Inflation, defined here as any increase in prices induced by monetary factors, is the change in the inner value of money. In macroeconomic theory, inflation is a *sustained* change of the general price level, which being a monetary phenomenon is equivalent to a sustained fall of the inner value of money. Thus, Menger’s classical concept of the inner value of money turns out to have a very modern interpretation. This was already observed by Hayek (1934, p. XXXI) who noted that “the actual terms employed are somewhat misleading” but “the underlying concept of the problem is extra-ordinarily modern”.

However, the measurement of Menger’s changes of the inner value of money would provide one not only with a closer empirical counterpart of

inflation in macroeconomic theory, but it would also be useful for monetary policy purposes. Indeed, using Menger's concept of the inner value of money could prevent that the central bank is held accountable for price rises which are not caused by monetary policy. Even more important, however, is that Menger's concept may be useful for the conduct of monetary policy. Guided by the CPI, monetary policy may take unintended measures since changes of the CPI are difficult to interpret. For example a sudden increase of the CPI may be due to an expansion of money supply or may be caused by a negative productivity shock affecting the producers of a specific consumer good. The optimal reaction of the central bank to the observed hike of the CPI will depend on its cause. If the hike is caused by a monetary expansion, the central bank may consider reducing the money supply, whereas it would not necessarily want to adjust money supply to a price shock on a single market. Contrary to the changes of the CPI, changes of the inner value of money have an unambiguous interpretation and may therefore serve as a useful input for monetary policy.

In the light of the relevance of Menger's concept, it is interesting to search for a more convincing operationalization. The main characteristic of Menger's concept should, however, remain intact. This characteristic is that a change in the inner value of money should ultimately lead to a proportional rise in all commodity prices. As suggested by Menger, a suitable starting point for operationalization is the frequency distribution of price changes. The following section deals with this approach and discusses possible gauges for the change in the inner value of money.

#### IV. Inner Value of Money: A Framework for the Decomposition of Price Changes

The observed change in the price of a good may be caused by various factors. The change in relative and absolute prices may be due to monetary or real causes or an error of measurement may have occurred. If  $P_{kt}$  is the price of good  $k$  at time  $t$  and

$$\pi_{kt} = \ln P_{kt} - \ln P_{k(t-1)}$$

is the increase in the price of good  $k$ , then the observed price change may be broken down into

$$(1) \quad \pi_{kt} = \alpha_t^M + \alpha_{kt} + \beta_{kt} + \varepsilon_{kt} \quad k = 1, \dots, K.$$

$\alpha_t^M + \alpha_{kt}$  is the price rise at time  $t$  of good  $k$ , which is underlain by monetary factors, such as an expansion of the money supply. Although an expansion of the money supply should, at least in the long run, lead to a proportional rise in all prices, the transmission of monetary shocks will, at least temporarily, disturb relative prices.  $\alpha_t^M$  is the change in the inner value of money, i.e. the proportional rise in all goods prices as a result of a monetary shock following the completion of all adjustment processes.  $\alpha_{kt}$  reflects the temporary deviation of the relative prices from the new long-run equilibrium during the transmission of a monetary shock<sup>1</sup>.  $\varepsilon_{kt}$  is the error of measurement, which may arise in the observation of prices.  $\beta_{kt}$  is the price change in period  $t$ , which is caused by real factors. Real shocks may effect a change in supply and demand in all markets. This disturbance of the general equilibrium of the economy will, if the equilibrium is stable, lead to new relative prices.

The component of the price changes that must be identified is the change in the inner value of money  $\alpha_t^M$ . The decomposition of price changes according to (1) may help to examine to what extent the measuring results obtained with the aid of various inflation gauges will approximate the change in the inner value of money. The first gauge to be considered here is the change in the CPI as the gauge most commonly used in practice. For the change in the CPI, say  $\pi_t^C$ , which is defined as the weighted sum of individual price changes by

$$(2) \quad \pi_t^C = \sum_k w_{kt} \pi_{kt},$$

with  $w_{kt} = \frac{x_{k0} P_{k(t-1)}}{\sum_i x_{i0} P_{i(t-1)}} > 0$  and  $\sum_k w_{kt} = 1$ , one has, in view of the decomposition (1) that

$$(3) \quad \pi_t^C = \alpha_t^M + \sum_k w_{kt} \alpha_{kt} + \sum_k w_{kt} \beta_{kt} + \sum_k w_{kt} \varepsilon_{kt}.$$

We see that the change in the CPI does not simply measure the change in the inner value of money  $\alpha_t^M$ . We note that, generally speaking, neither the weighted sum of the relative price effects of transmission  $\sum_k w_{kt} \alpha_{kt}$ , nor the sum of the budget-share weighted price changes caused by real

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<sup>1</sup> In his discussion of the inner value of money, Menger abstracted from the problem that monetary shocks might lead to a temporary disturbance in relative prices.

factors  $\sum_k w_{kt}\beta_{kt}$  equal 0<sup>2</sup>. Finally, it cannot be ruled out that measurement errors – the term  $\sum_k w_{kt}\varepsilon_{kt}$  – affect  $\pi_t^C$ .

However, other inflation gauges based on, for instance, the frequency distribution of price changes, may be considered. As the change in the inner value of money is a component of the general price rise, the average, and the median or, as Menger proposed, the modal price changes form alternative ways of measuring inflation.

The average price change  $\pi_t^A$  would only identify the change in the inner value of money if it may be assumed that the average price rise caused by real factors and transmission equals nil. After all  $\pi_t^A = \frac{1}{K} \sum_k \pi_{kt}$ , and, on the basis of decomposition (1)

$$\pi_t^A = \alpha_t^M + \frac{1}{K} \sum_k \alpha_{kt} + \frac{1}{K} \sum_k \beta_{kt} + \frac{1}{K} \sum_k \varepsilon_{kt}$$

or, if the calculation of the average price changes is based on a large number of goods

$$(4) \quad \pi_t^A \approx \alpha_t^M + \frac{1}{K} \sum_k \alpha_{kt} + \frac{1}{K} \sum_k \beta_{kt}.$$

The latter equation follows under mild conditions from the law of large numbers<sup>3</sup>. There are, however, no arguments in economic theory to justify the hypothesis that the relative price changes caused by real or monetary factors average nil. In fact, it is extremely unlikely that  $\frac{1}{K} \sum_k \beta_{kt}$  equals zero after an increase in VAT by 1%. Therefore, the average price change as such is not a suitable statistic for the change in the inner value of money.

The *median* and the *modal* price change, too, lead to a breakdown of the changes in the inner value of money and price changes caused by real factors only based on certain ad hoc assumptions. For the median price change  $\pi_t^M$ , and the modal price change  $\pi_t^X$ ,

$$(5) \quad \pi_t^M = \alpha_t^M + z_t$$

<sup>2</sup>  $\sum_k w_{kt}\beta_{kt} = 0$  holds only if the demand and supply functions of the economy fulfill highly exceptional conditions.

<sup>3</sup> Where the change in the CPI is concerned, the law of large numbers applies only under highly implausible assumptions with regard to the budget shares  $w_{kt}$ .



and

$$(6) \quad \pi_t^X = \alpha_t^M + s_t$$

respectively, with  $z$  the median and  $s$  the mode of the joint distribution of  $\alpha_{kt}$ ,  $\beta_{kt}$  and the measurement errors  $\varepsilon_{kt}$ . Like the change in the CPI and the average price change, the median and the modal price change are also unable to identify the change in the inner value of money. It is clear from this discussion that the changes in the inner value of money cannot be gleaned with the aid of purely descriptive statistics. None of the gauges is capable of distinguishing between general price level increases caused by monetary factors and those resulting from real shocks. In addition, all gauges, except the average price change, are sensitive to errors of measurement. Therefore, we follow another route to identify the changes in the inner value of money.

## V. A Model to Identify the Inner Value of Money

### 1. The model

To attack the problem of measuring monetary inflation, we employ a specific structural VAR model. The basic assumption underlying this model is that in the long run inflation, being a monetary phenomenon, is output-neutral, with the proviso that unexpected inflationary shocks in the short and medium term may influence real income. Measuring inflation by means of the CPI or other price indices can, however, be misleading as has been shown in the previous section, since price changes brought about by real factors are not eliminated. Therefore, we decompose measured inflation into core inflation and a residual.<sup>4</sup> Core inflation is defined as the component of measured inflation, which is output-neutral in the long run.

We assume that the first differences of (the logarithm of) output and measured inflation are stationary stochastic processes. Furthermore, we assume that the change in measured inflation,  $\Delta\pi$ , and the growth rate of output,  $\Delta y$ , can be explained by contemporaneous and lagged effects of two types of shocks  $\varepsilon_1$  and  $\varepsilon_2$ . Therefore, the model can be written as

$$(7) \quad \begin{pmatrix} \Delta\pi_t \\ \Delta y_t \end{pmatrix} = A_0 \begin{pmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{pmatrix} + A_1 \begin{pmatrix} \varepsilon_{1t-1} \\ \varepsilon_{2t-1} \end{pmatrix} + A_2 \begin{pmatrix} \varepsilon_{1t-2} \\ \varepsilon_{2t-2} \end{pmatrix} + A_3 \begin{pmatrix} \varepsilon_{1t-3} \\ \varepsilon_{2t-3} \end{pmatrix} + \dots$$

<sup>4</sup> A similar approach is followed by *Quah* and *Vahey* (1995).

The shocks  $\varepsilon_{it}$  of this structural VAR model are serially and contemporarily uncorrelated with zero expectation and unit variance<sup>5</sup>. Finally, it is assumed that only one of the shocks, the “core inflation shock”,  $\varepsilon_{1t}$ , does not affect the level of output in the long run. The change in the output-neutral component of measured inflation, i.e. the change in core inflation, is then defined as  $\Delta\pi_t^c = \sum_{j=0}^{\infty} A_{j,11}\varepsilon_{1,t-j}$ , with  $A_{j,11}$  the element (1,1) of matrix  $A_j$ .

The parameters of the stochastic process generating inflation and output are, however, unknown and must be determined empirically. Here an identification problem arises: only the reduced form of the vector autoregressive representation of (7)

$$(8) \quad \begin{pmatrix} \Delta\pi_t \\ \Delta y_t \end{pmatrix} = D_1 \begin{pmatrix} \Delta\pi_{t-1} \\ \Delta y_{t-1} \end{pmatrix} + D_2 \begin{pmatrix} \Delta\pi_{t-2} \\ \Delta y_{t-2} \end{pmatrix} + \dots + D_p \begin{pmatrix} \Delta\pi_{t-p} \\ \Delta y_{t-p} \end{pmatrix} + \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix}$$

can be estimated. The moving average representation of (7), however,

$$(9) \quad \begin{pmatrix} \Delta\pi_t \\ \Delta y_t \end{pmatrix} = \begin{pmatrix} v_{1t} \\ v_{2t} \end{pmatrix} + C_1 \begin{pmatrix} v_{1t-1} \\ v_{2t-1} \end{pmatrix} + C_2 \begin{pmatrix} v_{1t-2} \\ v_{2t-2} \end{pmatrix} + C_3 \begin{pmatrix} v_{1t-3} \\ v_{2t-3} \end{pmatrix} + \dots$$

reveals that the shocks  $\varepsilon$  in the structural form (7) are not identified. Indeed, comparing the coefficients in (7) and (9) shows that  $v_t = A_0\varepsilon_t$  and  $C_i A_0 = A_i$ ,  $i = 1, 2, \dots$  with the matrix  $A_0$  unknown.

However, with the aid of the estimated covariance matrix of the reduced-form disturbances  $E v_t v_t^T = \Omega$  and the hypothesis that in the long run core inflation is output-neutral, all elements of  $A_0$  can be identified. After all, one has  $E v_t v_t^T = A_0 A_0^T$  so that the covariance matrix yields three restrictions for the four elements of  $A_0$ . The neutrality of core inflation implies that the model parameters must meet a fourth restriction: after  $k$  periods, a core inflation shock leads to a change in the output level of size  $\varepsilon_{1t+k} \sum_{j=0}^k A_{j,21}$ . On the basis of neutrality,  $\sum_{j=0}^{\infty} A_{j,21} = 0$  should therefore hold. In other words, the element (2,1) of matrix  $\sum_{i=0}^{\infty} C_i A_0$  must equal zero.

Once the matrix  $A_0$  has been determined with the aid of these restrictions, the structural form (7) can be constructed using the residuals and the estimated parameters from the reduced form (8). Subsequently, core inflation or the output-neutral component of measured inflation, which

<sup>5</sup> The normalization of the variance of the structural shocks does not have any consequences for the estimations or other outcomes of the model.

is not directly observable, may be derived from the parameters and the shocks of the structural form.

## 2. A closer look at the model

Our basic assumption is that observed inflation and output are explained by no more than two types of exogenous shocks. The reasons why core inflation depends on just one type of exogenous shock are that it is due entirely to monetary influences, that there is a single monetary base and that monetary policy is conducted by a single institution, viz. the central bank.

The assumption that all other changes in measured inflation and output may be explained by a single second type of shock which invariably influences the two endogenous variables in the same way may be seen as no more than an approximation. The latter assumption can, however, be relaxed if the number of endogenous variables in the model is increased by the number and nature of possible structural shocks. A desirable extension of the model, would e.g. consist of the explicit treatment of indirect tax rate changes. It seems unlikely that the effect of a changed VAT rate is identical to that of an oil price change or of a variation in government spending.

In the VAR model, the identification of the structural shocks is underlain by the economic hypothesis that in the long run inflation does not affect output. This hypothesis corresponds to the view that the long-run Phillips curve is vertical, a view commonly held by economists, especially for moderate levels and small variations of the inflation rate<sup>6</sup>. The bone of contention lies not with the long-run but with the short-run effects of inflation or the speed with which the short-run turns into the long-run Phillips curve. However, the impact of inflation on the output level in the short and the medium run remains unrestricted by the identification method. The VAR model also permits the validity of the identification method to be tested informally. As inflation is a monetary phenomenon, the second type of shocks, viz. the real or output shocks, should, in the long run, not affect measured inflation. However, should measured inflation be found to be influenced by output shocks even in

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<sup>6</sup> Note, however, that even if monetary inflation is not exactly output-neutral in the long-run, the identification scheme is approximately correct if the long-run effect of inflation on output is small compared to the effect of non-core shocks on output, see *Blanchard and Quah* (1989, p. 668–669).

the long run, doubts would arise about the validity of our identification procedure.

As the model is estimated in first differences of the endogenous variables, it is not core inflation itself that is recovered, but the change in core inflation. The level of core inflation itself remains unknown and undetermined.

### *3. The relationship between the inner value of money and core inflation*

The main question is what conceptual relationship exists between the change in the theoretical notion or the inner value of money and the statistical concept of core inflation employed in our VAR approach.

Core inflation is that part of measured inflation, which is output-neutral in the long run. The decompositions of two possible inflation gauges, viz. the change in the CPI of equation (3) and the average price change of equation (4), indicate that in the long run three components do not affect the level of output, and may therefore be identified as part of core inflation.

These components are

- the change in the inner value of money,
- the (weighted) average of temporary relative price changes brought about by monetary shocks, and
- measurement errors.

Of course, the (weighted) average of the relative price changes generated by monetary shocks is output-neutral in the long run because these price effects will disappear if the equilibrium is stable.

Consequently, it may be concluded that our method is, in theory, capable of decomposing the influence of real and monetary shocks on inflation, measured by one of these two gauges. However, for both gauges, our definition of core inflation does not correspond wholly to the change in the inner value of money. Core inflation derived from the CPI or the average price change at time  $t$  is the (weighted) average of the price changes at that time, insofar as caused by monetary factors, but not the change in the inner value of money, i.e. the proportional change in all prices following a monetary shock after the new long-run equilibrium is reached. Thus, in the absence of measurement errors, the difference between core inflation and the decrease in the inner value of money depends on transitory relative price changes due to monetary shocks.

The use of the unweighted average price change as the inflation series which is to be decomposed by our model probably yields the least distorted estimation of the change in the inner value of money because, on the one hand, errors of measurement have a negligible effect on this inflation gauge and, more importantly, on the other, for the measurement of monetary inflation there is no firm theoretical foundation for weighting price changes by budget shares. Moreover, when calculating the average price, one is in principle not limited to consumer commodities only. For the other two gauges, the median and the modal price change, it is not possible to determine, without the aid of further and highly detailed assumptions, which components would be identified as core inflation.

Although core inflation, as it has been defined here, does not exactly correspond to the change in the inner value of money, core inflation derived from the average price change is thus far the best available operationalization to measure Menger's concept. In the next section, we use this operationalization to calculate the change of the inner value of money for the Netherlands and for the European Union.

## VI. Measuring the Inner Value of Money

### 1. The Netherlands

The VAR model (8) is estimated for the Netherlands with monthly data from the period 1991–1995. For real output the deseasonalized average daily output of the production industries excluding construction was chosen. For observed inflation, we used the average price change, calculated from the 200 price series which also underlie the CPI.

Before estimating the VAR model, we tested whether the non-stationarity assumptions are indeed satisfied by the Dutch data. The results of the tests for the non-stationarity of the average price change and real output, summarised in Table 1 of Appendix I, indicate that the series are integrated of order one. Completing the specification of the model, we determined the order of the VAR model. Based on preliminary estimations we included 3 lagged variables in our final model<sup>7</sup>. The results of the estimation are presented as impulse-response functions shown in Figure 1 and 2, which indicate how real output and measured inflation respond to the shocks. Note that these impulse-responses show the movements in the *level of measured* inflation and output.

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<sup>7</sup> Details are provided in Appendix I.

Figure 1 shows that a monetary shock leads to a permanent increase in inflation, while after less than a year output has returned to its initial level. The speed with which the effect of an unanticipated inflation impulse on real output wears off is not determined by the identification method. Indeed, the identification implies solely that core inflation has become output-neutral after an infinite number of periods. It is noteworthy that an inflation shock decreases output in the first month, while the opposite was to be expected on the basis of the short-run Phillips curve. The confidence intervals are, however, so large that this effect is not significant.

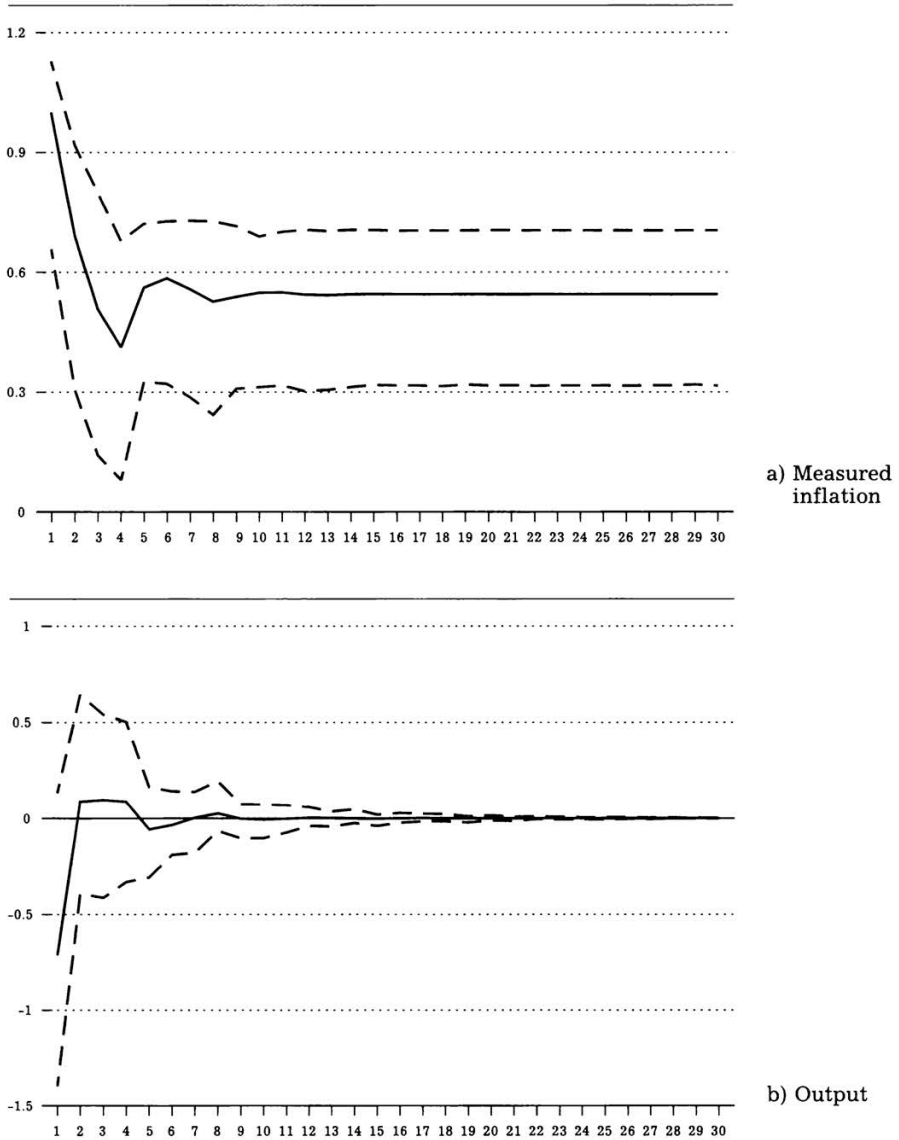
Figure 2 shows that an output shock leads to a permanent rise in measured inflation, too. This effect is, however, not significant. This confirms the hypothesis that core inflation shocks do indeed reflect monetary influences.

The ultimate objective of the model is the identification of price changes, which have been caused by monetary factors. Figure 3 shows measured inflation, i.e. the average price change, the core inflation derived from the average price change and the conventional measure of inflation based on the CPI. Phrased differently, using core inflation based on the average price change to operationalize the decrease in the inner value of money, Figure 3 depicts the movement in the outer and inner value of money.

The remarkable thing about the pattern in Figure 3 is that the discrepancy between measured inflation and core inflation does not show any trend over time. This means that the average price change is either relatively insensitive to price changes generated by real factors or – and this is more likely – that over the sample period real shocks had but a relatively small influence on the price level. This second interpretation is also supported by the breakdown of the shocks to *measured* inflation  $v_{1t}$ . The shocks  $v_{1t}$  relating to measured inflation, i.e. the average price change, are, after all, related to the structural shocks  $\varepsilon_{1t}$  and  $\varepsilon_{2t}$  through  $v_{1t} = A_{0,11}\varepsilon_{1t} + A_{0,12}\varepsilon_{2t}$ . Figure 4 shows, for every month of the sample period, the monetary component  $A_{0,11}\varepsilon_{1t}$ , and the real component  $A_{0,12}\varepsilon_{2t}$  of the inflation shock  $v_{1t}$ . The chart shows that the effect of real shocks on measured inflation has indeed been fairly small over the past three years by comparison with the effect of monetary shocks.

The movements in the average price change and derived core inflation in Figure 3 shows that a number of periods stand out where the average price change over- or underestimates the monetary influences on infla-

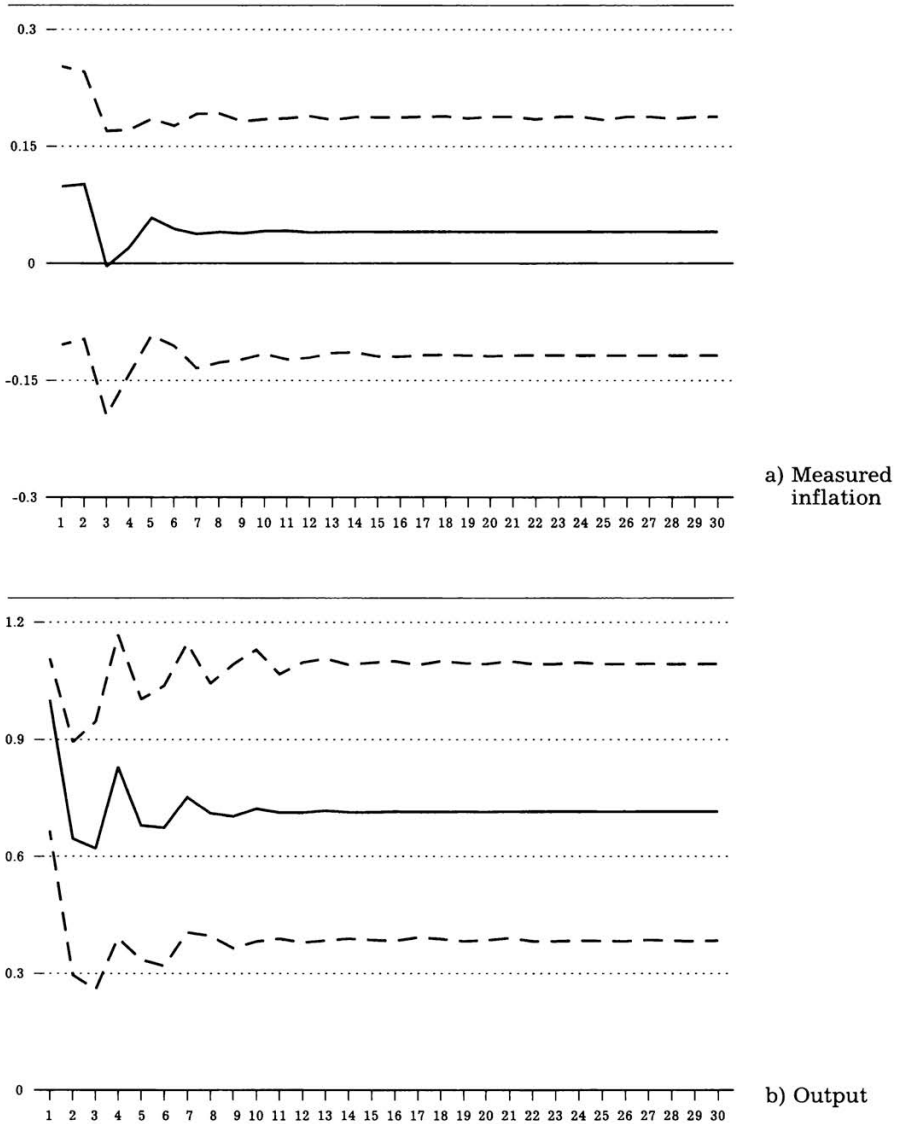
## Response to core inflation shock



Explanatory note: The horizontal axis shows time in months. The vertical axis shows the deviation in percent of inflation and (log-)output, respectively, from the initial level. The core inflation shock  $\varepsilon_1$  and the output shock  $\varepsilon_2$  have been chosen such that in the first period measured inflation would be up by one percentage point, and the level of (log-)output by one percent. The shock lasts but one period. The 95% confidence intervals – based on 1000 replications – for the impulse-response functions are also shown (see Runkle (1987)) for details.

*Figure 1*

## Response to output shock



Explanatory note: The horizontal axis shows time in months. The vertical axis shows the deviation in percent of inflation and (log-)output, respectively, from the initial level. The core inflation shock  $\varepsilon_1$  and the output shock  $\varepsilon_2$  have been chosen such that in the first period measured inflation would be up by one percentage point, and the level of (log-)output by one percent. The shock lasts but one period. The 95% confidence intervals – based on 1000 replications – for the impulse-response functions are also shown (see Runkle (1987)) for details.

*Figure 2*



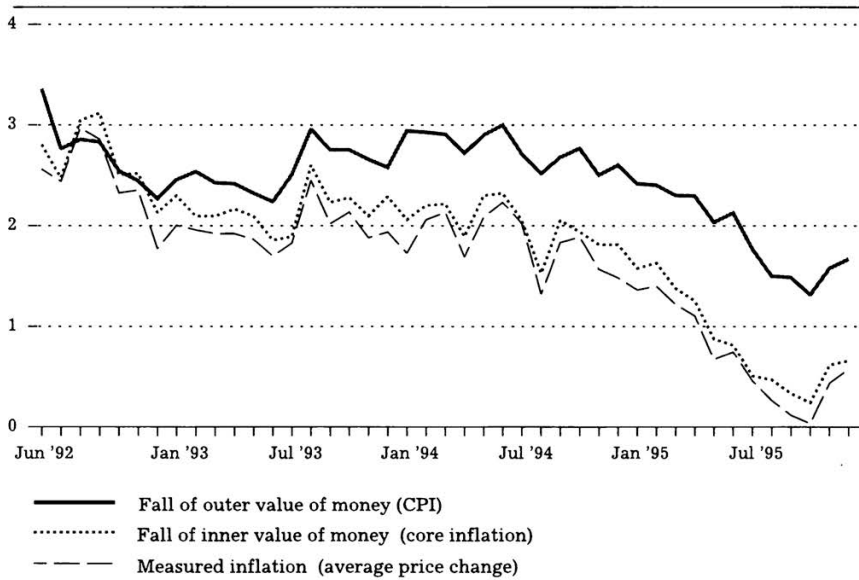


Figure 3: Change of inner and outer value of money (percent)

tion. Figure 4 shows, for example, that the drop in the average price change in December 1992 is only partially the result of monetary policy. Simultaneous real impulses lead to a drop in the average price change as well. On the other hand, the increase in the average price change in February and March 1994 is not caused by a monetary shock only, but real factors drove prices up as well. Finally, in the second half of 1995, the average price change first dropped more substantially and then rose more considerably than core inflation. Again real and monetary impulses worked in the same direction leading to an overestimation of the monetary effects on measured inflation.

From a comparison of the movements in the inner and the outer value of money, i.e. core inflation and the change in the CPI, it becomes evident that a notable difference between the two is that from July 1993 onwards the fall in the outer value of money is much more pronounced than that in the inner value of money. It goes without saying that the weighting of the CPI explains this phenomenon, because certain goods and services whose prices continued to rise after July 1993 figure prominently in the CPI, such as actual and imputed rents. In the case of the average price change and derived core inflation, the marked rise in the

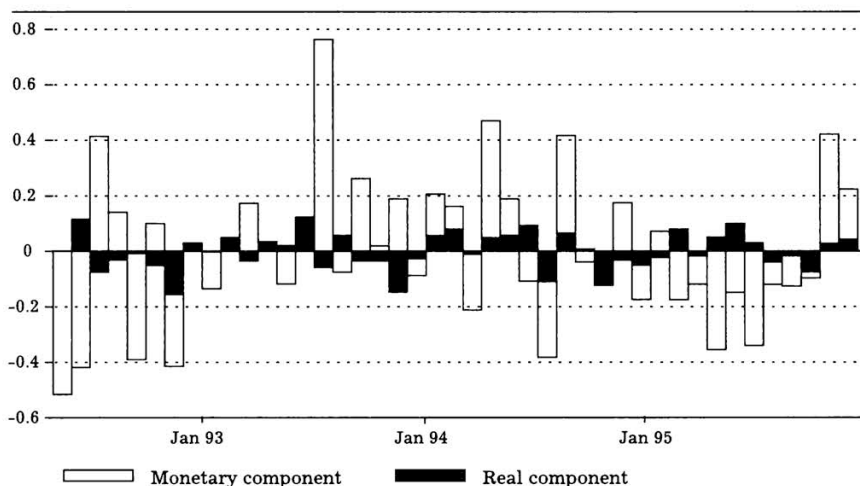


Figure 4: Historical decomposition of inflation shocks  $v_{1t}$

prices of these items is partially offset by the smaller increase or even fall in prices of the bulk of goods and services.

## 2. European Union

Attempting to measure the inner and outer value of money on a European level by the methods described so far presents problems of its own, most notably the problem that a common European currency does not yet exist. Therefore, in order to measure the value of money, one first has to define a European concept of money. Here we define European money as the aggregated money stocks of the various nations using purchasing power parities to convert all nominal values into ECU. Thus, the European money stock at time  $t$  is defined as

$$(10) \quad M_{tEU} = \sum_i e_{ti} M_{ti},$$

with  $e_{ti}$  denoting the purchasing power parity of country  $i$  at time  $t$  and  $M_{ti}$  the money stock of country  $i$  at time  $t$ . A matching definition of the outer value of European money takes the form

$$(11) \quad P_t^L = \sum_i w_i \frac{e_{ti}}{e_{0i}} P_{ti}^L, \quad \text{with} \quad w_i = \frac{\sum_j e_{0i} p_{0ij} x_{0ij}}{\sum_{lj} e_{0l} p_{0lj} x_{0lj}},$$

which is a weighted average of the national CPI's  $P_{ii}^L$  of the various countries, with weights  $w_i$  equal to the countries' shares in aggregate European final consumption in 1985. Finally, as European average price change we use

$$(12) \quad \pi(\text{avg.}) = \frac{1}{IJ} \sum_{ij} \ln e_{ii} p_{tij} - \ln e_{(t-1)i} p_{(t-1)ij}$$

where  $I$  denotes the number of countries considered and  $J$  the number of commodities per country.

In order to estimate the VAR model (8) for Europe we used monthly data for the period January 1985 to December 1995 for Austria, Belgium, France, Germany, Italy, the Netherlands, Spain, Sweden, and the United Kingdom. As the series reflecting real output, the deseasonalized average daily output of the national production industries, excluding construction, was chosen. European real output was constructed as the weighted average of national real output, with weights equal to each country's share in European gross added value based on factor prices. Measured inflation in this application is the European average price change, calculated based on the 11 price series per country which also underlie the national CPI's.

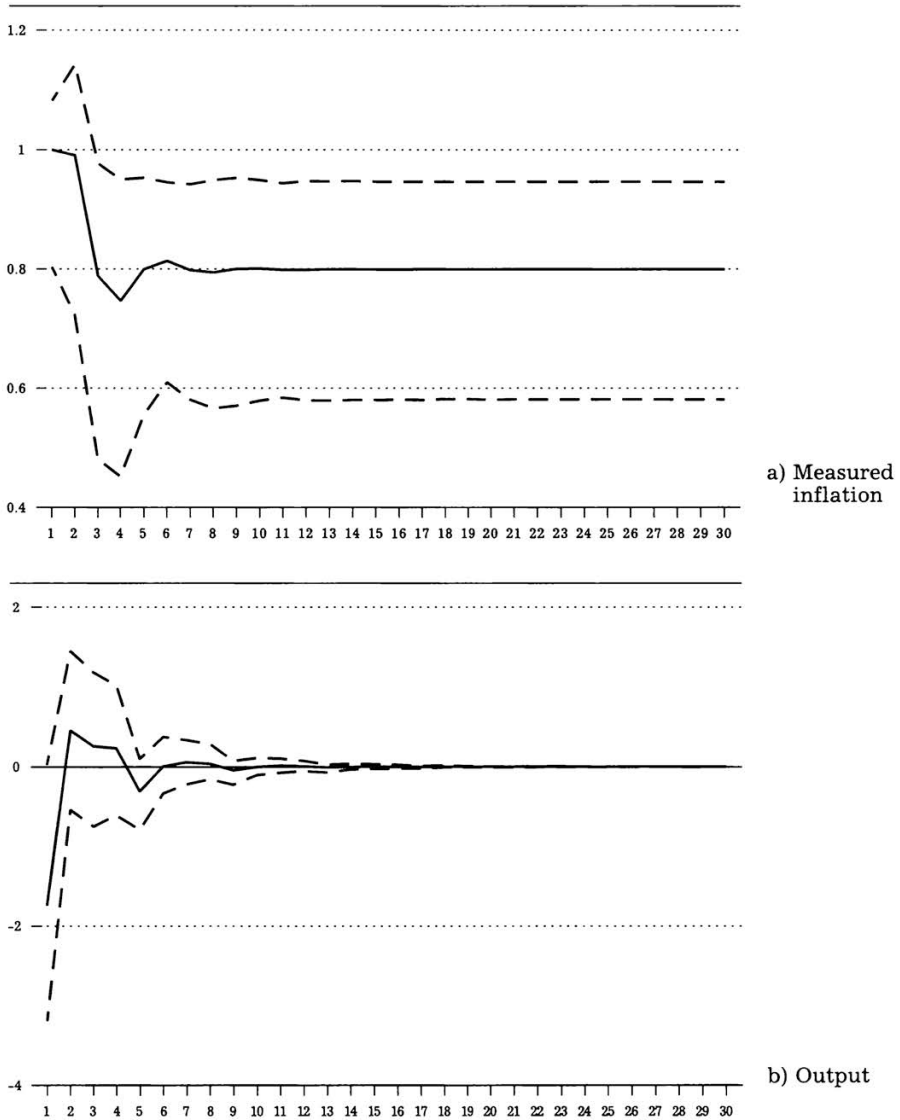
Before we estimated the model for the European Union we tested whether the series for measured inflation and real output show non-stationary behaviour. The test results indicate that for the European Union these conditions for the identification method are satisfied. For the model specification to be complete, the lag order of the VAR model must be chosen. Based on preliminary estimates and several standard criteria we choose a lag length of three<sup>8</sup>. The estimation results for the structural VAR model are again presented as impulse-response functions for measured inflation, i.e. European average price change, and European (log-)output, shown in Figures 5 and 6.

A core inflation shock leads to a permanent increase in inflation. In the same way as for the Netherlands, it is observed that the effect of an inflation impulse on output wears off quickly and that output returns to its initial level within 12 months. Similarly, the European results suggest that an inflation shock may decrease output in the first month. The confidence intervals reveal, however, that this effect is not significant.

Contrary to the model for the Netherlands, however, the model for the European Union implies that a real shock has a significant and permanent effect on inflation. This casts some doubt on the hypothesis that

<sup>8</sup> Details are presented in Appendix I.

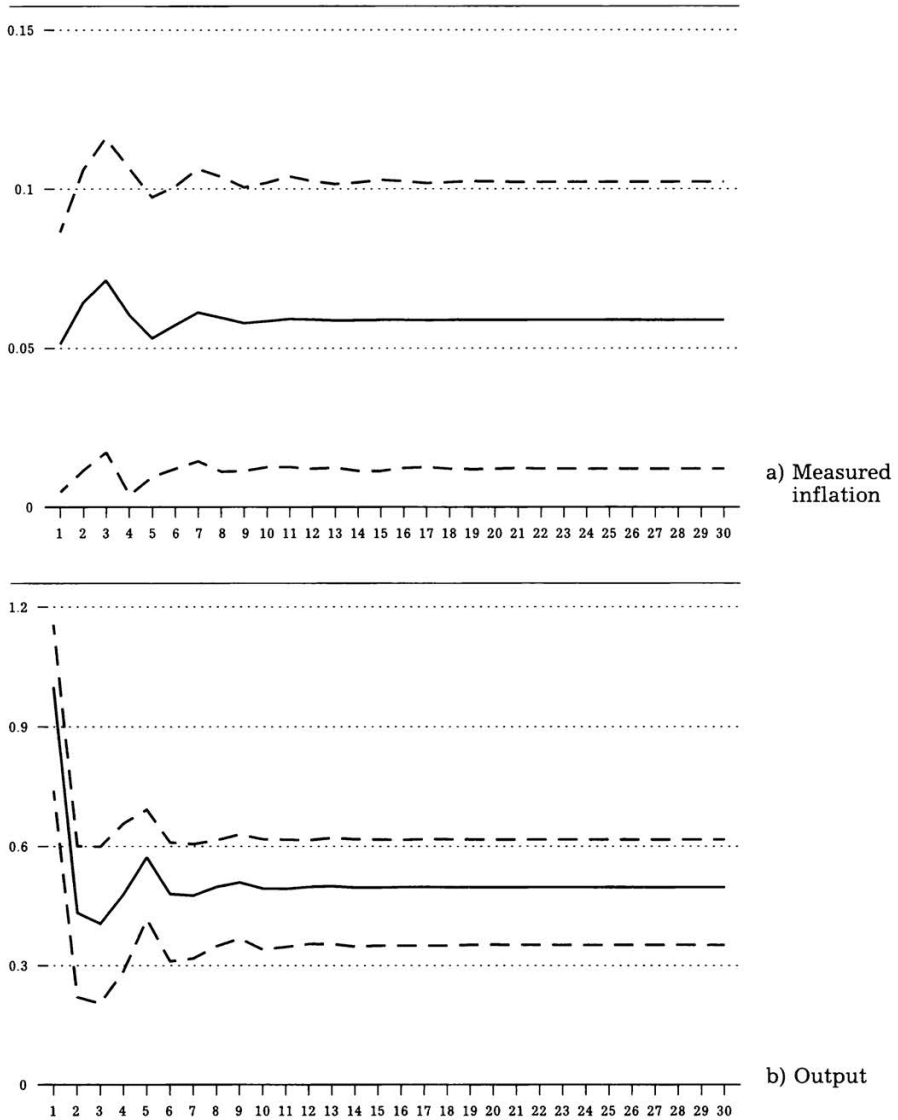
Response to core inflation shock



Explanatory note: The horizontal axis shows time in months. The vertical axis shows the deviation in percent of inflation and (log-)output, respectively, from the initial level. The core inflation shock  $\varepsilon_1$  and the output shock  $\varepsilon_2$  have been chosen such that in the first period measured inflation would be up by one percentage point, and the level of (log-)output by one percent. The shock lasts but one period. The 95 % confidence intervals for the impulse-response functions are also shown.

Figure 5

Response to output shock



Explanatory note: The horizontal axis shows time in months. The vertical axis shows the deviation in percent of inflation and (log-)output, respectively, from the initial level. The core inflation shock  $\varepsilon_1$  and the output shock  $\varepsilon_2$  have been chosen such that in the first period measured inflation would be up by one percentage point, and the level of (log-)output by one percent. The shock lasts but one period. The 95 % confidence intervals for the impulse-response functions are also shown.

Figure 6

monetary effects are correctly identified by this approach applied to European data. Indeed, the prediction that a permanent rise in the output level by 0.5 percentage implies a permanent rise in inflation in the absence of any monetary effects, contradicts economic theory.

This difficulty of the model to decompose the European average price change into a purely monetary and a purely real component is explained by the assumptions underlying the structural VAR model. Implicitly the model assumes that only two types of shocks drive inflation and output. Furthermore, it is assumed that each shock has a similar effect on the endogenous variables. If applied to one country, the assumption of a typical inflation shock is justifiable because of the existence of a single monetary base. For Europe, however, a single monetary base does not yet exist. Moreover, the transmission of monetary shocks may differ between countries due to diverging institutional arrangements. Both facts may imply that a monetary shock originating in e.g. Italy leads to a different effect on European output than an unanticipated inflation shock in, for instance, Germany.

Although the identification of the monetary component of the average price change by the structural VAR model applied to the European Union is less convincing compared to the application to the Netherlands, we present in Figure 7 the change in the inner and outer value of European

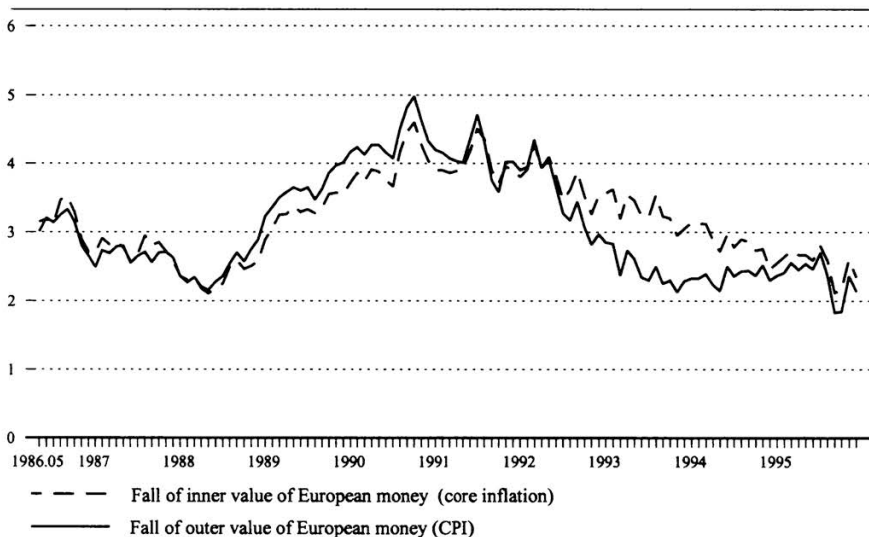


Figure 7: Change of inner and outer value of European money

money. The chart shows that the inner value of money decreased less than the outer value of money in 1988–89, catching up in 1990–91. In the third and fourth quarters of 1992, however, the fall in the inner value of money was more marked than that in the outer value. In the last quarter of 1993 and 1994 the fall of the inner value decelerated once more, catching up with the decrease in the inner value of European money. Incidentally, in 1992 the year when the changes in the inner and the outer value diverged quite sharply, the UK and Italy moved out of the EMS.

## VII. Conclusions

This article was motivated by the fact that the most commonly used measure of inflation, the change in the CPI, does not match the concept of inflation used in economic theory: it cannot distinguish between monetary and real causes leading to price changes nor between a one-off and a permanent price rise. From the point of view of a monetary authority which aims to stabilise the value of money, the former shortcoming is especially disturbing since the central bank may be held accountable for price rises which are not caused by monetary policy or it may take inappropriate policy actions on the basis of a biased inflation measure.

This article investigated possible operationalizations of Carl Menger's concept of the inner value of money. A change in the inner value of money is defined as the change in prices, which is solely brought about by monetary causes. On examining different descriptive statistics of price changes more closely, we found that neither the change in the CPI, nor the average price change or the mode or median of the price change frequency distribution is capable of identifying changes in the inner value of money. Furthermore, we also tried to decompose the price changes measured by the average price change into a real and monetary component using the economic hypothesis that inflation is output-neutral in the long run. It was argued that this approach does indeed identify the monetary component of price changes but not exactly the inner value of money. The difference between the two is that the former responds to price changes which are caused by the transmission of monetary shocks, whereas the latter is defined in terms of price changes following a monetary shock after all adjustment processes have been completed. Despite this difference, core inflation or the monetary component of the average price change is the best available operationalization of the change in the inner value of money, which is wholly in accordance with economic theory.

Applying the approach to the Netherlands and the European Union, we found that the change in the inner and that in the outer value of money have diverged considerably and persistently in the sample periods. This finding indicates that using the CPI as a gauge for monetary inflation is not only theoretically inappropriate but that even in practical applications it yields distorted information on the actual inflationary tendencies. The change in the inner value of money may be seen as an alternative measure that matches the concept of inflation used in economic theory more closely than the change in the CPI. Moreover, from the point of view of monetary policy, it seems to be the more adequate measure of inflation in terms of accountability.

## Appendix I

### *I.1. The model for the Netherlands: non-stationarity tests and lag order*

With the aid of the two augmented Dickey-Fuller tests, the stationarity of the base series and their first differences were examined. For the (log-) output series  $y$ , the test statistic  $T(\rho - 1)$ , with  $T$  the sample size and  $\rho$  the autocorrelation between successive observations, and the Dickey-Fuller  $t$ -test  $\frac{\rho - 1}{\sigma_\rho}$  indicate the existence of non-stationarity; the first differences  $\Delta y$ , however, do form a stationary process. The hypothesis

*Table 1*  
**Results of the Augmented Dickey-Fuller Test**

Variable	Lag length	Excluding trend		Including trend	
		$T(\rho - 1)$	$(\rho - 1)/\sigma_\rho$	$T(\rho - 1)$	$(\rho - 1)/\sigma_\rho$
$y$	5	0.46***	0.05***	-9.91***	-0.94***
$\Delta y$	4	-231.39	-6.41	-238.87	-6.78
$\pi(\text{aver.})$	1	-3.04***	-1.16***	-15.36***	-2.68***
$\Delta\pi(\text{aver.})$	1	-66.66	-6.64	-66.32	-6.51

Notes: The number of lagged variables in the test regression has been chosen such that the disturbance term is not serially correlated. \*\*\*/\*\*/\* means that the hypothesis that a unit root is present cannot be rejected at a significance level of 10%/5%/1%, respectively.



that the output series is integrated of order one is thus confirmed. The hypothesis that the average price change is also integrated of order one may also be accepted.

The number of lagged variables to be included in the VAR model is determined with the aid of various criteria and test statistics. The criteria of Akaike, Hannan-Quinn and Schwartz indicate a lag length of 1 to 3. Although the Box-Pierce Portmanteau test and Godfrey's Lagrange multiplier test do not indicate serial correlation of the residuals if the model includes but one lagged variable, and neither does the log-likelihood ratio test reject a lag length of one in favour of a lag length of three, this specification seems overly restrictive. It seems unlikely that the change in inflation and the growth rate of output can be explained by current inflation and output as well as inflation and output of the previous month only. Therefore, three lagged variables were included in the final model version. It also turned out that a deterministic trend is not significant, so that, apart from the lagged variables, only a constant term was added to the model.

### *I.2. The model for the European Union: non-stationarity tests and lag order*

The results of the test for the non-stationarity of the European average price change and European real output are summarised in Table 2. With the aid of the two augmented Dickey-Fuller tests, the stationarity of the

*Table 2*  
**Results of the Augmented Dickey-Fuller Test**

Variable	Lag length	Excluding trend		Including trend	
		$T(\rho - 1)$	$(\rho - 1)/\sigma_\rho$	$T(\rho - 1)$	$(\rho - 1)/\sigma_\rho$
$y$	3	-5.21***	-1.58***	-7.89***	-1.43***
$\Delta y$	3	-307.70	-8.19	-311.77	-8.22
$\pi(\text{aver.})$	2	-3.58***	-1.16***	-4.03***	-1.28***
$\Delta\pi(\text{aver.})$	2	-140.69	-7.69	-141.03	-7.69

Notes: The number of lagged variables in the test regression has been chosen such that the disturbance term is not serially correlated. \*\*\*/\*\*/\* means that the hypothesis that a unit root is present cannot be rejected at a significance level of 10%/5%/1%, respectively.

base series and their first differences were examined. For the (log-) output series  $y$ , the test statistic  $T(\rho - 1)$ , and the Dickey-Fuller  $t$ -test indicate non-stationarity of the data; the first differences  $\Delta y$ , however, do form a stationary process. The hypothesis that the output series is integrated of order one is thus confirmed. The hypothesis that the average price change is also integrated of order one may also be accepted.

As in the model for the Netherlands, the number of lagged variables to be included in the VAR model is determined with the aid of the criteria of Akaike, Hannan-Quinn and Schwartz. The criterion of Schwartz points towards a lag length of 1, whereas the other two criteria indicate a lag length of 3. The model with 1 lagged variable is, however, not correctly specified since the Box-Pierce Portmanteau test and Godfrey's Lagrange multiplier test indicates serial correlation of the residuals. The log-likelihood ratio test, too, rejects a lag length of 1 maintaining the model with three lagged variables.

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## Summary

### **Measuring Inflation: Testing Carl Menger's Concept of the Inner Value of Money**

This paper attempts to operationalize Carl Menger's concept of the "innerer Tauschwert des Geldes", i.e. the inner value of money. Since the change in the inner value of money is the component of price movements which is due to monetary influences, the operationalization provides an alternative measure of inflation. We consider several approaches for gauging the change in the inner value of money. Of these, we use a structural VAR model to identify the price movements in the Netherlands and the EU due to monetary shocks. (JEL E 31)

## Zusammenfassung

### **Inflationsmessung: Das Konzept des Inneren Tauschwertes des Geldes von Carl Menger**

In diesem Beitrag wird versucht, Carl Mengers Konzept des „Inneren Tauschwertes des Geldes“ zu operationalisieren. Da Veränderungen des inneren Tauschwertes des Geldes Bestandteil von Preisschwankungen aufgrund monetärer Einflüsse sind, beinhaltet die Operationalisierung eine alternative Möglichkeit der Inflationsmessung. Wir betrachten mehrere Vorgehensweisen zur Messung der Veränderungen des inneren Tauschwertes des Geldes. Von diesen bedienen wir uns eines strukturellen VAR-Modells zur Identifizierung von Preisschwankungen in den Niederlanden und in der EU, die auf monetäre Schocks zurückzuführen sind.

## Résumé

### **Mesure de l'inflation: un test de la thèse de Carl Menger sur la valeur interne de la monnaie**

Cette analyse essaie d'opérationnaliser la thèse de Carl Menger sur la «valeur d'échange interne de la monnaie». Puisque le changement dans la valeur interne de la monnaie est l'élément des mouvements de prix dus aux influences monétaires, l'opérationnalisation fournit une mesure alternative de l'inflation. Les auteurs considèrent différentes approches pour mesurer le changement dans la valeur interne de la monnaie. Ils utilisent, entre autres, un modèle structurel VAR pour identifier les mouvements de prix aux Pays-Bas et aux Etats-Unis provoqués par des chocs monétaires.