

A Service of



Leibniz-Informationszentrum Wirtschaft Leibniz Information Centre for Economics

Mumford, Kevin J.; Patterson, Richard; Yim, Anthony

Working Paper College Course Shutouts

IZA Discussion Papers, No. 16859

Provided in Cooperation with: IZA – Institute of Labor Economics

Suggested Citation: Mumford, Kevin J.; Patterson, Richard; Yim, Anthony (2024) : College Course Shutouts, IZA Discussion Papers, No. 16859, Institute of Labor Economics (IZA), Bonn

This Version is available at: https://hdl.handle.net/10419/295882

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.



WWW.ECONSTOR.EU



Initiated by Deutsche Post Foundation

DISCUSSION PAPER SERIES

IZA DP No. 16859

College Course Shutouts

Kevin J. Mumford Richard W. Patterson Anthony Yim

MARCH 2024



Initiated by Deutsche Post Foundation

DISCUSSION PAPER SERIES IZA DP No. 16859 **College Course Shutouts** Kevin J. Mumford Purdue University **Richard W. Patterson** Brigham Young University, IZA and CESifo Anthony Yim Brigham Young University **MARCH 2024**

Any opinions expressed in this paper are those of the author(s) and not those of IZA. Research published in this series may include views on policy, but IZA takes no institutional policy positions. The IZA research network is committed to the IZA Guiding Principles of Research Integrity.

The IZA Institute of Labor Economics is an independent economic research institute that conducts research in labor economics and offers evidence-based policy advice on labor market issues. Supported by the Deutsche Post Foundation, IZA runs the world's largest network of economists, whose research aims to provide answers to the global labor market challenges of our time. Our key objective is to build bridges between academic research, policymakers and society.

IZA Discussion Papers often represent preliminary work and are circulated to encourage discussion. Citation of such a paper should account for its provisional character. A revised version may be available directly from the author.

ISSN: 2365-9793

IZA – Institute of Labor Economics

| Schaumburg-Lippe-Straße 5–9 | Phone: +49-228-3894-0 | |
|-----------------------------|-----------------------------|-------------|
| 53113 Bonn, Germany | Email: publications@iza.org | www.iza.org |

ABSTRACT

College Course Shutouts^{*}

What happens when college students are not able to enroll in the courses they want? We use a natural experiment at Purdue University in which first-year students are conditionally randomly assigned to oversubscribed courses. Compared to students who are assigned a requested course, those who are shut out are 40% less likely to ever take the oversubscribed course and 30% less likely to ever take a course in the same subject. While a course shutout is equally likely to occur to female and male students who requested the course, shutouts are much more disruptive for female students. In the short run, shutouts decrease the credits female students earn as well as their GPA. In the long-run, shutouts increase the probability female students drop out of school in the first year, decrease the probability they choose majors in STEM fields (Science, Technology, Engineering, and Math), decrease cumulative GPA, and decrease the probability of graduating within four years. In contrast, shutouts have no effects on short-run credits earned, dropout, majoring in STEM, cumulative GPA, or four-year graduation for male students. Shutouts do have one large measurable long-run impact on male students—shutouts significantly increase the probability that men choose a major from the business school.

JEL Classification:I23, J16, J24Keywords:course shutouts, major choice

Corresponding author: Richard W. Patterson Brigham Young University Provo, Utah 84602 USA E-mail: rich patterson@byu.edu

^{*} All opinions expressed in this manuscript are those of the authors and do not represent the opinions of Purdue University. We thank Tim Bond, Jillian Carr, Jeff Denning, and Victoria Prowse along with seminar participants at APPAM, New York Federal Reserve Bank, NBER Education Program meeting, Purdue University, Utah Applied Economics conference, and Utah State University for their helpful comments. We also thank Tomas Muller for helping us understand the course schedule assignment algorithm and how to simulate it.

1 Introduction

Across the country, states are reducing spending on higher education. In real terms, state appropriations per student are 40% lower now than they were in 1990.¹ One way many colleges respond to this increasing budgetary pressure is by reducing course offerings which causes more courses to be oversubscribed. This increases the number of students who are not able to enroll in (shut out of) courses they want to take.² On average, public institutions have a significantly higher student-faculty ratio than private non-profit institutions (see Figure 1) and one important difference between public and private non-profit institutions is the extent to which a student is able to enroll in desired courses. While researchers have speculated that these course shutouts contribute to negative student outcomes, including increased dropout rates and longer time-to-degree, there is limited and conflicting evidence of the effects of course shutouts.³

We contribute to the nascent literature on course shutouts by exploiting random variation in course shutouts at a large public university. In 2018, Purdue University introduced a "batch registration" algorithm to assign first-year student course schedules.⁴ The number of first-year students in fall 2018 was about 2,000 more than typical and there was little corresponding increase in course capacity. University administrators were concerned that using the existing course registration system would concentrate course shutouts disproportionately on first-generation and disadvantaged students. An advantage of the batch

¹See Chakrabarti *et al.* (2020).

²See Deming (2017); Mitchell *et al.* (2016); Bahr *et al.* (2015) for discussions of the potential impacts of shutouts. We define a shutout in our setting as a primary course request (i.e. non-contingent request) that is not granted, either because no class is assigned or a secondary/contingent course is assigned in its place. This is comparable to a student in a typical enrollment system failing to enroll in a desired course because the course was full or the course conflicted with their schedule. In a typical enrollment system, we would consider a student to be shut out from a course if they formally requested a desired course and were not assigned (e.g. they were waitlisted and never assigned the course) or they were discouraged from making a formal request (e.g. the class was full at the time a student engaged with the enrollment system or the course was not offered at a time that was compatible with their schedule).

³Robles *et al.* (2021) find that course shutouts in community colleges significantly increase the probability that students take zero courses in a semester or transfer to a lower-quality two-year college. In contrast, Kurlaender *et al.* (2014) find that course shutouts only affect time-to-degree when shutouts occur in semesters when students could have otherwise graduated.

⁴We limit our analysis to entering first-year students in Fall 2018 for four reasons: (1) in 2018, the batch registration process only applied to entering first-year students and not continuing students (2) batch registration was suspended for spring and summer terms, and reintroduced for fall 2019, which precludes us from studying second-semester freshman, (3) there were changes to the Batch algorithm process after 2018, and (4) Limiting to the 2018 entering cohort allows us to examine long-run outcomes.

registration system is that students are conditionally randomly assigned to oversubscribed courses, which spreads out course shutouts across students.⁵ Among first-year students in their first semester, 49% were assigned their preferred course schedule, while the other 51% were shut out from at least one of their top six requested courses. We find that course rationing significantly changes both short- and long-term course-taking behavior. First-year students who are initially shut out from a course are 35 percentage points less likely to ever complete the course and 25 percentage points less likely to take a course in the same subject.

In addition to examining the effects of shutouts on course-taking behavior, we also investigate the impact of shutouts on a number of short- and long-run outcomes including credits earned, GPA, major choice, dropouts, and on-time graduation. We find that each first-term shutout reduces first-term credits earned by 0.2 credits, with no evidence that students compensate with more credits in later terms. Additionally, we find that shutouts do not affect short-run GPAs, but may reduce GPAs in the second semester of students' senior year. We also find that each shutout decreases the probability that students major in STEM by 1.6 percentage points and increases the probability that students choose a major from the business school by 1.1 percentage points. Next, we find that shutouts reduce the probability that students drop out in their first term by 0.5 percentage points but have no long-term effects on dropouts. Finally, we estimate that shutouts have a negative, but statistically insignificant effect on whether students graduate within 4 years of enrollment.

Overall, these results suggest that shutouts affect student course taking, but have mixed effects on broader student outcomes. However, these overall results mask important differences by gender. When we look at the effects separately for female and male students, we find that shutouts lead to large negative effects for female students while having little or even modestly positive effects for male students. For female students, we find that each first-semester freshman shutout reduces first-semester credits earned by 0.4 credits and first-semester GPA by 0.06 grade points. Each shutout reduces the probability that women major in a STEM field by 2.9 percentage points (or 5.0%) and reduces the probability that

⁵The algorithm explicitly randomizes the order in which students are matched to courses after an initial sorting of students by scheduling constraints. See Appendix III for a full description of the algorithm.

women graduate within 4 years by 5.1 percentage points (or 7.6%). In contrast, for male students, shutouts do not have a significant effect on first-term credits earned, first-term GPA, choosing a STEM major, or on-time graduation. However, each shutout does increase the probability that male students choose a major from the business school by 1.9 percentage points (or 24%).

Our results contribute to the existing literature in several ways. First, our research contributes directly to the emerging literature on college course shutouts. Robles *et al.* (2021) find that shutouts decrease course-taking at community colleges and increase transfers to lower-ranked community colleges. Despite the difference in institutional setting, our estimates of the effects of shutouts on course-taking and leaving the institution are similar. Our also paper pushes the literature forward by identifying the effects of shutouts for a broader range of students and by examining the effects of shutouts on previously unexplored outcomes including course-taking patterns, major choice, GPA, and on-time graduation.⁶

Two important new relationships we explore in this paper are the the effects of course shutouts on course-taking patterns and major choice. Because college majors have large effects on long-term earnings, career trajectories, and lifestyles, (Altonji *et al.*, 2012, 2014; Bleemer & Mehta, 2022; Chevalier, 2011; Hastings *et al.*, 2013; Kirkeboen *et al.*, 2016; Patnaik *et al.*, 2020; Webber, 2014) significant policy attention has focused on steering students towards high-returns majors (Bleemer & Mehta, 2021; Denning & Turley, 2017; Sjoquist & Winters, 2015). Our paper suggests that shutouts can significantly influence students' major choices. In particular, we find that shutouts in STEM courses reduce the likelihood that students major in Engineering and reduce the likelihood that students choose a major that

⁶Additionally, Kurlaender *et al.* (2014) examine the effects of shutouts at the University of California-Davis and find shutouts do not impact student outcomes. We do not highlight their study because we believe their null results are due to limitations in their identification strategy and an unusual definition of course shutouts. Kurlaender *et al.* (2014) use randomization in how often students get early access to the course registration site as an instrument for course shutouts, but this instrument leads to an insignificant first stage. When they create a dichotomous variable for "extremely unlucky" or being at the 10th percentile of early registration, they get a statistically significant, but underpowered first stage (F-stat of 7.32). Furthermore, they do not show whether their instrument is balanced across treatment and control groups. Finally, their definition of a course shutout differs from how we define a course shutout. Specifically, Kurlaender *et al.* (2014) define a course shutout as any instance where a student attempts to register for a specific course section and finds that it is full. This means that many of the course "shutouts" in their data could be instances of students who enrolled in a desired course (i.e. not shut out) but had queried the registration site to find which sections of a course have open seats in their enrollment process. This is corroborated by the fact that students average four shutouts per term and that some students average over 40 shutouts per term.

corresponds to the subject of the shutout course. We also find that shutouts push male students toward the business school and push female students out of STEM majors.

Our finding that shutouts push female students away from STEM majors contributes to research on the "leaky STEM pipeline" (e.g. Buckles, 2019; Griffith, 2010; Price, 2010). Approximately 50% more students initially declare a STEM major than graduate with a STEM degree, with disproportionally larger shifts away from STEM majors for female and under-represented minority students (Speer, 2023). Our finding that first-semester freshman shutouts explain 8.4% of the female-male gap in STEM degrees suggests shutouts are an important factor to consider when addressing the leaky STEM pipeline.

More broadly, our findings contribute to a body of research into the factors that influence college major choice. Student preferences, expected earnings, peer effects, subject ability, and costs have all been shown to affect college major choice (Elsner *et al.*, 2021; Patnaik *et al.*, 2020; Wiswall & Zafar, 2014; Zölitz & Feld, 2021). However, recent research has found that seemingly small changes in student experience, such as the time of day (Haggag *et al.*, 2021; Yim, 2023) or semester (Patterson *et al.*, 2021) a student takes a course, can meaningfully influence major choice as well. Our finding that a course shutout significantly alters students' major demonstrates another way in which seemingly small schedule changes can have large effects.

Our results also relate to research on factors that influence on-time graduation. Fewer than half of graduating low-income students in the United States graduate within four years of college entry (Denning *et al.*, 2022). Delays to graduation increase both the direct costs of college attendance (e.g. tuition, room, and board) and indirect costs (foregone wages). Researchers have investigated how financial incentives affect on-time graduation and have found mixed results.⁷ Our finding, that each course shutout during a female student's first semester decreases her on-time graduation likelihood by 5.1%, suggests that addressing course shutouts may be an important and potentially low-cost way to increase on-time graduation rates.

⁷For example, Scott-Clayton (2011) finds that a merit aid program in Georgia increased on-time graduation whereas Angrist *et al.* (2022) find that a merit aid program in Massachusetts decreased on-time graduation. Garibaldi *et al.* (2012) find that a 1000-Euro bump in continuation tuition increased on-time graduation by 5.2%.

Finally, our findings that women are disproportionately harmed by course shutouts contribute to a rich literature investigating differences in how female and male students respond to changing educational circumstances. In particular, there is evidence female students may be disproportionately responsive to changes in financial aid (Bartik *et al.*, 2021), learning incentives (Angrist *et al.*, 2009; Kremer *et al.*, 2009), student resources (Angrist *et al.*, 2009; Evans *et al.*, 2020), mentoring programs (Carrell & Sacerdote, 2017), instructor characteristics (Carrell *et al.*, 2010; Fairlie *et al.*, 2014), and course grades (Bleemer & Mehta, 2021; Kugler *et al.*, 2021). While most of these studies show that improving circumstances tend to disproportionately benefit women, our study suggests that the converse is also true: worsening educational circumstances are likely to disproportionately harm women.

The remainder of the paper is structured as follows. In Section 2 we describe our study environment and data. In Section 3 we describe our empirical approach. In Section 4 we report our primary results and explore the potential mechanisms for our findings. In Section 5 we conclude.

2 Study Environment and Data

Data for this study come from administrative records at Purdue University, which include 15,112 student-course observations in 241 oversubscribed courses in the 2018 fall semester. Of the 8,566 first-year students in fall 2018, our study follows 7,646 traditional non-athlete students who requested to enroll in one or more of these oversubscribed courses in their first semester. Our analysis sample excludes Division I scholarship athletes (who receive special scheduling treatment) and students over the age of 23 at entry. Table 1 shows that 44% of students in our analysis sample are female, 11% are Asian, 3% are Black, 6% are Hispanic, and 66% are White.⁸ Additionally, 17% of our sample are first-generation college students. The average Math SAT score in our sample is 664 and the average Verbal SAT score is 651 (both out of 800).

As part of the enrollment process, students submit their course preferences to the univer-

⁸The remaining 15%, categorized as "Other race/ethnicity", are primarily composed of international students. While the data do not designate the race/ethnicity of international students, many of these students come from Asia.

sity by completing a course request form (see Figure C.1).⁹ After all students submit their preferences, they are assigned a course schedule by Purdue's batch registration algorithm (Müller *et al.*, 2010). The algorithm uses each student's preference ranking of courses and preferred class times as inputs in generating schedules for all students.¹⁰ The algorithm uses these inputs, along with course availability and student schedule constraints, to assign each student a schedule. While 45% of students in the analysis sample are assigned each of the courses they request, 55% are shut out from at least one requested course with 9% being shut out from two or more courses. Of the 241 oversubscribed courses, the two most common shutout courses are required English writing and communications courses.¹¹ The remaining 239 courses with shutouts come from nearly every subject area offered at the university (See Appendix Table B.1 for a complete list of oversubscribed courses).

In Columns 3 and 4 of Table 1 we explore the characteristics of students that have and have not been shut out of at least one course as a first-semester freshman, respectively. In these raw data, we find that female students and students from a race/ethnicity other than Asian, Black, Hispanic, and White (predominantly international students for whom we do not observe race/ethnicity) are less likely to be shut out of courses. Otherwise, we find similar characteristics for students who are and are not shut out of courses.

3 Empirical Approach

Our empirical approach takes advantage of Purdue's course assignment algorithm that assigns student course schedules based on student preferences and schedule constraints. Whether students with similar course preferences are assigned a slot in an oversubscribed course depends on randomization within the algorithm process. By running 1,000 simulations of the exact algorithm used to assign course schedules in Fall 2018, we are able to estimate the probability that each course request will result in a shutout (See Appendix III for a detailed

 $^{^{9}}$ Students can submit up to nine preferences, but in practice only 12% of students submit more than six preferences. Our analysis sample, therefore, includes the first 6 preferences requested by students.

¹⁰An academic advisor can also indicate whether a course is required for a student, which is typically the result of students being required to take a course with a specific group (e.g. band, sports team, or honors section). We exclude course requests with a "required" designation from our analysis.

¹¹Both English and communication requirements can be met with several alternative courses.

description of the course assignment algorithm). Figure 2, which shows the distribution of shutout probabilities for course requests that resulted in shutout and assignment separately, illustrates overlap in shutout probabilities and the role that randomness plays in whether a student is assigned a course. Once we control for shutout probability, whether a student is actually shut out of a course should be uncorrelated with student characteristics, including their potential outcomes (Rosenbaum & Rubin, 1983). Our identification assumption is that, after accounting for simulated shutout probability, whether a student experiences a shutout is random, potential outcomes are uncorrelated with shutout status, and differences between shut-out and non-shut-out students can be causally attributed to course shutouts.¹² While the course assignment algorithm is not strategy-proof (e.g. a student is more likely to be assigned their entire desired schedule by listing their hard-to-match courses in top priority), our identification strategy is robust to strategic behavior. Strategic behavior will be reflected in preference rankings, captured in our algorithm simulations, and, therefore, accounted for in our identification strategy. Furthermore, because the batch algorithm process was introduced in the same semester we observe and there was little information provided to students about how the batch algorithm process worked, there may have been little scope for strategic behavior in our sample.

Evidence of conditional random assignment can be examined by estimating balance in observable characteristics with the following equation:

$$shutout_{ic} = \boldsymbol{\theta} \boldsymbol{X}_{i} + \delta Shutout Probability_{ic} + \gamma_{c} + \epsilon_{ic} \tag{1}$$

where $shutout_{ic}$ is an indicator for whether student *i* is shut out of (not assigned to) oversubscribed course *c*. X_i is a vector of individual characteristics including sex, race/ethnicity, first-generation student status, and SAT math/verbal test scores. *ShutoutProbability_{ic}* is the estimated probability that student *i* will be shut out from course *c*.¹³ γ_c is a course fixed

¹²Note that we are not using an instrumental variables approach for identification. This is because actual shutouts and simulated shutout probability are both correlated with student preferences and potential outcomes, thus simulated shutout probability is not a valid instrument for shutouts. Instead, our estimation approach takes advantage of the fact that estimated shutout probability is a sufficient statistic for potential outcomes (i.e. individuals with identical preferences will have indistinguishable simulated shutout probabilities) and accounting for shutout probability is as good as observing conditional random assignment.

 $^{^{13}}$ Shutout probability is calculated as the fraction of our 1,000 simulations where student *i* is not assigned requested

effect, which does not affect balance. Finally, ϵ_{ic} is an individual-by-course idiosyncratic error term. In our estimates of Equation 1, we cluster our standard errors at the student level.¹⁴

In column (1) of Table 2 we show why we need to account for shutout probability when estimating Equation 1. In Column (1) we do not control for the simulated shutout probability and find that three student characteristics are significantly correlated with course shutout (indicators for Asian, Other Race, and whether the course is in the student's Preenrollment Major). Additionally, when all variables are considered jointly, we strongly reject the hypothesis that observable characteristics are the same among shutout and non-shutout students (p < 0.01). If we do not use simulated shutout probability, but instead flexibly control for the course preference inputs into the assignment algorithm,¹⁵ we still observe some imbalance in observable characteristics. Specifically in column (2) of Table 2 we observe an imbalance in the fraction of Asian and first-generation students shut out after accounting for algorithm inputs and we still reject the hypothesis that observable characteristics are equal among shutout and non-shutout students at the 10% level. Finally, in column (3) we show how controlling for shutout probability is likely to fully account for differences in potential outcomes between shut-out and non-shut-out students. When we control for the simulated course shutout probability and find strong evidence of balance: no characteristics vary significantly by whether a student is shut out and the coefficients are jointly insignificant.¹⁶ Note that our simulated shutout probability strongly predicts shutouts—a 1 percentage point increase in simulated shutout probability corresponds to a 0.99 percentage point increase in actual shutout probability.

Our balance in column (3) of Table 2 motivates our individual-by-course level analysis of course shutouts. In this analysis, we examine how shutouts affect course-related outcomes,

course. The simulations are obtained from the exact algorithm used to assign students to courses in the Fall of 2018. ¹⁴We cluster at the student level, because a student i treatment status in course c is not independent of the other courses student i requests.

¹⁵these controls include course-by-preference fixed effects, an indicator for whether the student's pre-enrollment major had reserved slots for majors in the course, and a measure of how difficult the student's other courses are to match.

¹⁶Because we also focus on gender-specific estimates, we also show balance separately for male and female students in Table A.1. The results in Table A.1 show that observable characteristics balance for both female and male students after conditioning on shutout probability.

including completing the requested course, taking courses in the subject of the shutout course, and choosing a major that corresponds to the shutout course. We do so by estimating the following equation:

$$Y_{ic} = \beta shutout_{ic} + \boldsymbol{\theta} \boldsymbol{X}_{i} + \delta ShutoutProbability_{ic} + \epsilon_{ic}$$
(2)

where Y_{ic} is a course-specific outcome for individual *i* that has expressed preference for taking course *c* and all other variables are as previously specified in Equation 1. The parameter of interest is β , which measures the effect of being shutout of a course in a student's first semester on their subsequent academic outcomes. We estimate this equation with ordinary least squares, clustering standard errors at the course-by-preference level.

In addition to examining the effects of shutouts at the individual-course level, we also investigate the effects of shutouts at the individual level. We do so by estimating the following equation:

$$Y_i = \beta shutouts_i + \boldsymbol{\theta} \boldsymbol{X}_i + \delta SummedShutoutProbabilities_i + \epsilon_i \tag{3}$$

where Y_i is an individual-level outcome, such as credits earned in each semester and on-time graduation. The variable of interest *shutouts_i* is a measure of the number of courses an individual has been shut out from in their first semester of college. X_i is a vector of individual characteristics. Additionally, *SummedShutoutProbabilities_i* is the sum of the shutout probabilities across all courses requested by individual *i* (*SummedShutoutProbabilities_i* = $\sum_{c=1}^{6} ShutoutProbability_{ic}$).

To evaluate whether the number of shutouts is likely to be conditionally independent of potential outcomes after accounting for the sum of shutout probabilities, we estimate the following balance equation:

$$shutouts_i = \boldsymbol{\theta} \boldsymbol{X}_i + \delta SummedShutoutProbabilities_i + \epsilon_i \tag{4}$$

where all variables are defined at the student level as in Equation 3. Table 3 reports estimates

of Equation 4, which allows us to examine whether the number of shutouts balances across individual characteristics after controlling for *SummedShutoutProbabilities_i*. While student characteristics do not unconditionally balance in column (1), they do balance conditional on algorithm inputs (Course-by-preference order fixed effects and number of reservation courses) in column (2), and on the sum of conditional shutout probabilities in column (3).¹⁷ The results of this balance exercise support our assumption that the number of shutouts a student experiences is independent of potential outcomes, after conditioning on the sum of shutout probabilities.

4 Results

4.1 Effects of a Course Shutout on Course-taking and Major Choice

We begin by estimating Equation 2 to examine the effects of a first-semester freshman course shutout on course-taking outcomes including course enrollment, course-taking patterns, and major choice. In column (1) of Table 4, we find that a course shutout reduces the probability of attending the course in the semester it is requested by 72 percentage points relative to the non-shutout mean attendance of 85%.¹⁸ Decomposing the imperfect compliance with course assignment, we find that 8% of students who are shut out of a course end up attending the course with the remaining non-compliance coming from students who are assigned their requested course but drop the course prior to the add/drop deadline. In column (2) we examine the effects of course shutout on course completion and find nearly identical results: students who are shut out of a course in their first semester are 71 percentage points less likely to complete the course that semester. For some students, a course shutout only delays their course completion to the next semester. However, column (3) of Table 4 reports that being shut out of a course in the first semester of freshman year reduces the probability of taking the course as a freshman by 44 percentage points and column (4) reports that a course shutout reduces the probability of ever completing the course by 35 percentage points.

¹⁷We also show balance separately for male and female students in Table A.2.

¹⁸We define course attendance as being enrolled in the course immediately after the semester add/drop deadline.

Given that a course shutout reduces the likelihood of ever taking the course, a natural question is: Are students substituting the shutout course with a similar course in the same subject area? Table 5 suggests that this is not usually the case. Panel A of Table 5 reports that a first-semester course shutout decreases the probability that the student completes a course in the shutout course's subject by 61 percentage points in the first semester and 35 percentage points in the first year.¹⁹ If a student does not take a course in the shutout course's subject within a year, it is unlikely that they ever will; a course shutout decreases the probability that a student is ever exposed to the subject by 25 percentage points. In Panel B of Table 5, we find that a course shutout changes the subjects the student is exposed to increases by 55 percentage points in the first semester, 32 percentage points by the end of the first year, and 25 percentage points overall. Panel C of Table 5 examines if a course shutout changes the total number of subjects students are exposed to and finds that it reduces the number of subjects in the first semester by 0.061, but has no effects on the long-run total number of subjects.

That a course shutout affects course-taking behavior suggests that it may also affect the major the student chooses. One way a course shutout may impact major choice is by reducing the likelihood a student chooses a major that corresponds to the shut-out course. We investigate this question in Table 6. In column (1), we estimate the effect of a course shutout on choosing a major that corresponds to the course subject and find a negative, but statistically insignificant effect.²⁰ In columns (2) and (3) we examine the effects separately by STEM and Non-STEM courses. The estimate in column (2) shows that a STEM course shutout has an economically meaningful and marginally significant effect: a STEM course shutout reduces the probability that students choose a major in a corresponding subject by 20% (2.5 percentage points). In contrast, in column (3) we find that non-STEM shutouts have no impact on whether students choose a major in a corresponding subject. In columns

¹⁹The course subject is defined by the subject code that proceeds the course number (e.g. ECON or EDUC) which often corresponds to the department offering the course. However, some departments use several different subject codes and some subject codes correspond to no department.

²⁰We define choosing a major that corresponds to the course subject as selecting a major offered by the same department as the course was offered.

(4) and (5) we find similar, and statistically insignificant, effects of shutouts on choosing a corresponding major for both top three and bottom three priority courses, respectively.²¹

4.2 Heterogeneous Effects of a Course Shutout

In results reported in the Appendix, we explore the heterogeneous effects of shutouts on course-taking behavior in Table A.3 and find that the effects of shutouts on initial course attendance are 4% stronger for female students than male students, 6% stronger for underrepresented minority students than non-underrepresented students, 5% larger for first-generation students than non-first generation students and 6% larger for students with Low SAT scores than those with high SAT scores. This is consistent with evidence that those in privileged positions are more willing to ask for exceptional treatment. For example, male college students are significantly more likely than female students to ask for grade changes (Li & Zafar, 2023). However, these short-term effects disappear and sometimes reverse in the long run. In particular, the effects of shutouts on exposure to courses and subjects are approximately 20% smaller in absolute magnitude for students with low SAT scores relative to students with high SAT scores. One reason for this difference is that students with low SAT scores may be less aware of substitute courses that could replace the shutout required course.

In Table A.4, we explore whether the response to shutouts differs by course characteristics and find that a course shutout reduces attendance and completion much more for courses that fulfill a general education requirement compared to courses that do not. One reason for this difference could be that many students who requested a particular general education course are doing so only to fulfill a general education requirement and are much more willing to substitute into another course (possibly in a different subject area) that fulfils the same general education requirement. Requested courses that do not fulfil a general education

²¹In Figure A.1 we find that a shutout in a student's top-priority course significantly decreases the probability that a student majors in a corresponding subject. However, given the effects of 2nd through 6th priority shutouts are all indistinguishable from zero, we are reluctant to draw any strong conclusions from this result. Figure A.2 shows that the effects of a course shutout on course-taking behavior do not systematically differ by priority ranking of courses. For example, the effects of shutouts on whether students ever take a course in a subject area are larger for first- and fourth-priority courses than second- and third-priority courses.

requirement are more likely to relate to the student's interests or desired major and may be why a non-general education course shutout causes less course substitution.²² We also find that a course shutout in a high-difficulty class is less likely to reduce exposure to the corresponding subject than a low-difficulty class.²³ High-difficulty courses are more likely to be a prerequisite for related upper-division courses and therefore students may be less able to substitute to another course if they wish to remain in their chosen major.

4.3 Effect of Course Shutouts on Student-Level Outcomes

To estimate the effect of the number of shutouts in the first semester on student-level outcomes including credits earned, GPA, on-time graduation, and major choice, we estimate Equation 3 using data at the student level rather than the student-course level. The estimated effects are generally small and insignificant when using the full sample, but there are important gender differences that we describe after presenting the overall effects. Panel A of Figure 3, shows that each shutout during the first semester decreases the number of credits earned in that semester by 0.2 credits, but shutouts have no effect on credits earned in subsequent semesters. When we look at the effects of shutouts on total credits earned in column (1) of Table 7, we find that shutouts have no overall effect. Then, in Panel B of Figure 3, we find that first-semester shutouts slightly decrease the probability that students drop out in their first semester, but have no effect on the likelihood of having dropped out in subsequent semesters. Next, in Panel C of Figure 3, we find that first-semester shutouts have no short- or long-term effects on GPA, with the exception of a GPA drop in the second semester of students' fourth year. We similarly find no overall effect of first-semester shutouts on cumulative GPA in column (2) of Table 7. We examine the effect of shutouts on graduating within 4 years in column 3 of Table 7 and find that each shutout leads to an economically meaningful, but statistically insignificant 3.0% (1.8 percentage point) decrease

 $^{^{22}}$ In columns (5) and (6) of Table A.4, we find that a course shutout is much more likely to reduce exposure to subjects in bottom-3 priority requests relative to top-3 preference requests.

 $^{^{23}}$ We construct the course difficulty measure by (1) regressing course grade onto student observable characteristics (e.g. SAT scores, sex, race), (2) generating average residual GPA by course, and (3) designating a course as high difficulty if it has below median residual average GPA and as low difficulty if it has above median residual average GPA.

in the probability that students graduate within 4 years.

4.4 Gender Differences for Student-Level Outcomes

The results above mask considerable heterogeneity by gender. Shutouts have large negative effects on female students' academic outcomes and somewhat positive effects on male students' outcomes. In Figure 4 we compare the effects of first-semester shutouts over time separately for female and male students. In column (1) of Panel A we find that, among female students, each shutout reduces the number of credits earned in their first semester by 0.4 and, while not always statistically significant, leads to similar reductions in credits earned in each subsequent semester. In contrast, Panel B shows that shutouts have a generally positive but statistically insignificant effect on credits earned for male students. In Panel C of Figure 4 we find that each shutout increases the probability that female students drop out of the university by the end of their first year by 1.8 percentage points, and while the estimates become less precise over time, the results for later semesters suggest that shutouts cause an increase in female dropouts on the extensive margin and not just in timing. In contrast, Panel D shows that shutouts for male students make them, if anything, less likely to drop out of the university. In Panel E of Figure 4 we find that shutouts have a negative effect on female students' GPA that persists through their first four semesters. In Panel F, we find no short-term effects of shutouts on GPA for male student with evidence suggesting a negative effect on GPA only in the second semester of their fourth year.

In Table 8 we report the effect of shutouts on the same student-level outcomes examined in Table 7, but do so separately by gender. In column (1) we find that female students who have no course shutouts in their first semester tend to complete more cumulative credits than male students with no course shutouts (113 vs 105), but course shutouts may reduce this advantage for women. While statistically imprecise, each shutout decreases cumulative attainment for women by 1.8 credits and increases cumulative attainment by 1.2 credits for male students. Similarly, in column (2) non-shutout female students earn higher GPAs than non-shutout male students (3.33 vs. 3.18), but shutouts reduce this advantage. Specifically, we find that each shutout decreases a female student's cumulative GPA by 0.05 grade points but has no effect on a male student's GPA, meaning each shutout reduces women's GPA advantage over men by 33%.

In column (3) of Table 8 we estimate that each shutout decreases the probability that women graduate within four years by 7% (5 percentage points). In contrast, we find no effects on graduation for men. This difference is economically substantial. Even if we make the conservative assumption that all women affected by shutouts graduate the following semester with a wage equal to what they would have had with seven months of job experience, each shutout costs women approximately \$1,500 in foregone wages²⁴ and \$800 in additional tuition and housing costs.²⁵

In column (4) of Table 8, we show that shutouts have differential effects on majoring in STEM, by gender. Each shutout decreases the probability that female students major in STEM by 5.1% (or 2.9 percentage points). We do not find any effect of shutouts on majoring in STEM for male students. To put these results in context, first-semester shutouts can explain 8.4% of the female-male STEM major gap at this university.²⁶

While we estimate the effects of shutouts linearly in Table 8, it is possible that the effects of shutouts are non-linear. In Table 9 we estimate the effect of one shutout vs. two or more shutouts on student outcomes (the average number of shutouts among students with at least two shutouts is 2.11). While we cannot reject linear effects of shutouts among female students for any outcomes in Panel A, our estimates are consistent with compounding negative effects of shutouts on cumulative credits and GPA and more linear effects on four-year graduation and choosing a STEM major. Relative to one shutout, women who have 2+ shutouts experience effects that are 252% more negative for cumulative credits, 675% more negative for cumulative GPA, 118% more negative for four-year graduation, and 64%

 $^{^{24}}$ We arrive at this value by estimating 7 months of forgone wages due to a December graduation instead of a May graduation at a wage of \$50,256. This wage amount is calculated by taking the weighted average first wages by major from Purdue via Purdue's survey of graduates, where the weights are number of (non-shutout) female students from the 2018 entering cohort in each major.

²⁵The average 2-semester net price for in-state students, accounting for financial aid and scholarships, is \$14,619 (source: https://nces.ed.gov/ipeds/datacenter/ accessed 11/15/2023) and Purdue's estimated net price for one semester for out-of-state and international students is \$21,947 and \$24,002 respectively (source: https://www.purdue.edu/treasurer/finance/bursar-office/tuition/fee-rates-2023-2024/undergraduate-2023-2024 accessed 11/15/2023. We use IPEDS estimates of 45% in-state, 45% out-of-state, and 10% international students to calculate an average semester cost of attendance of \$15,566.

 $^{^{26}}$ 57.9% Female students major in STEM compared to 79.6% male students. Female students have 0.63 shutouts on average, meaning that shutouts explain 1.83 percentage points of the 21.7 percentage point gap.

more negative for majoring in STEM. In Panel B of Table 9 we explore potential non-linear effects of shutouts on credits, GPA, four-year graduation, and majoring in STEM for male students and, while we find more negative effects of 2+ shutouts than for one shutout, each estimated effect is small and statistically insignificant.

Our finding that each shutout makes female students 5.1% less likely to major in STEM motivates a more detailed examination of the differential effects of shutouts on major choice by gender. In Figure 5 we examine how course shutouts affect major choice (organized by university college) and, while imprecise, these estimates suggest that shutouts are moving female students away from majors in technology and science and into agriculture.²⁷ Shutouts move male students, but not female students, into business majors. Each shutout increases the probability that a male student majors in business by 1.9 percentage points. This has important implications for gender parity in business majors. At this university, men are 19% more likely than women to major in business and this entire gender gap can be explained by course shutouts.²⁸

4.4.1 Can Gender Differences be Explained by Course-Taking Patterns?

Given the large differences in the effects of shutouts for male and female students, a natural question to ask is whether female students request courses where shutouts are disproportionately likely to result in adverse outcomes. To investigate this potential mechanism for our results, we first investigate differences in the types of courses female and male students are shut out from in Table 10. While female students have a smaller share of shutouts coming from STEM and more difficult courses and a greater share of shutouts coming from upperdivision and required courses, none of these differences are greater than three percentage points. When we examine gender differences in shutout subject areas, we find small differences in shutouts in the colleges where shutouts are most common (Liberal Arts and Science colleges), but do find a disproportionate share of shutouts come from courses in Agriculture

 $^{^{27}}$ See Panel A of Figure A.4 for the overall effects of shutouts on major sorting.

 $^{^{28}8.6\%}$ of men major in business. Each shutout decreases the probability that men major in business by 1.9 percentage points and men have an average of 0.7 shutouts in their first semester. This means that 7.2% of men would have majored in business in the absence of shutouts, which is the same fraction of women (7.2%) that major in business.

(5% vs. 2%), Education (5% vs. 1%), and Health Science colleges (3% vs. 1%) for female students and a disproportionate share of shutouts come from courses in Engineering (4% vs. 1%) and Polytechnic colleges (11% vs. 4%) for male students.

In Table A.6 we explore to what extent differences in the courses male and female students are shut out from could explain the differences in the effects of shutouts we observe in Table 8. Table 8 shows that shutouts in courses that women are disproportionately shut out from do not systematically lead to worse outcomes for students. For example, STEM shutouts, which are slightly more common among male students, appear to lead to worse outcomes than non-STEM shutouts, but shutouts in required courses, which are slightly more common among female students appear to lead to somewhat worse outcomes than non-required shutouts. Back-of-the-envelope calculations suggest that there is no difference in the composition of course shutouts that can explain more than 4% of the gender gap in the effects of shutouts on any individual outcome examined in Table 8 and a majority of compositional differences in shutouts predict shutouts should be slightly less impacting for women.²⁹ Given that differences in the composition of shutouts are unlikely explain much of why female students are more negatively affected by shutouts than male students, our findings are most consistent with a growing literature that finds that female students are simply more responsive than male students to changes in higher education environments (e.g. Angrist *et al.*, 2009; Bartik et al., 2021; Bleemer & Mehta, 2021; Evans et al., 2020; Kugler et al., 2021; Kremer et al., 2009).

²⁹Our back-of-the-envelope calculations are constructed as follows: (1) we collect the effects of shutouts in a specific course type (e.g. STEM shutouts) on a specific outcome (e.g. cumulative credits), (2) we collect the effects of shutouts in the reciprocal course type (e.g. Non-STEM shutouts), (3) we take the difference in the effect sizes from step (1) and (2), (4) we multiply the difference from (3) by the difference in gender composition for the course type, and (5) divide our value from (4) by the overall female-male difference in the effects of shutouts for that outcome. For this specific example, (1) the effect of a STEM shutout is a -0.204 reduction in cumulative credits, (2) the effect of a non-STEM shutout is a 0.216 increase in cumulative credits, (3) the difference in STEM vs. Non-STEM is -0.420 credits. (4) 52% of female shutouts are in STEM courses, 54% of male shutouts are in STEM courses, so Female-Male difference in STEM composition can account for -0.420(0.52-0.54)=0.0084 cumulative credits, (5) the Female-Male difference in the effects of shutouts is -1.778-1.250=-3.028 credits, so differences in STEM shutouts explains -0.2% of the Female-male differences in STEM shutouts.

4.5 Additional Effect Heterogeneity for Student-Level Outcomes

While the starkest differences in the effects of shutouts are between female and male students, we explore other potential differences in the Appendix (see Figures A.3 - A.10 and Table A.5). In general, the effects of shutouts do not appear to systematically differ by request priority or demographic differences other than gender. However, there is one notable exception. In Table A.5, we find that students with high SAT scores seem unaffected by course shutouts, while students with low SAT scores experience negative effects on credits earned, GPA, on-time graduation, and the likelihood of STEM major, though the only statistically significant difference is for the STEM major result. This suggests that course shutouts are more disruptive for students with low SAT scores.

5 Conclusion

Private non-profit universities generally offer fewer majors than public universities but students are typically able to enroll in whichever courses they desire. In contrast, students in public universities have a large number of majors to choose from, but frequently find that a course they would like to take is full and are unable to register. In this paper, we examine what happens when college students are not able to enroll in the courses they request. Using data from a large public university where students were conditionally randomly assigned to oversubscribed courses, we find that being shut out from a course in a student's first semester changes the types of courses taken and can even cause a change in the student's major.

Consistent with recent evidence that women are more responsive to changes in educational environments than men (e.g. Angrist *et al.*, 2009; Bartik *et al.*, 2021; Evans *et al.*, 2020), we find that shutouts are particularly disruptive for women. Women who experience course shutouts earn worse grades, take longer to graduate, and are less likely to choose a major in a STEM field. Our findings show that course shutouts can have large effects on student academic outcomes. In an environment where institutions are interested in widening the path to high-return majors, decreasing gender gaps in STEM fields, improving student GPAs, and reducing time to graduation, our estimates suggest that reducing course shutouts can be an effective way to improve these student outcomes.

References

- Altonji, Joseph G, Blom, Erica, & Meghir, Costas. 2012. Heterogeneity in human capital investments: High school curriculum, college major, and careers. Annu. Rev. Econ., 4(1), 185–223.
- Altonji, Joseph G, Kahn, Lisa B, & Speer, Jamin D. 2014. Trends in earnings differentials across college majors and the changing task composition of jobs. *American Economic Review*, **104**(5), 387–393.
- Angrist, Joshua, Lang, Daniel, & Oreopoulos, Philip. 2009. Incentives and services for college achievement: Evidence from a randomized trial. American Economic Journal: Applied Economics, 1(1), 136–163.
- Angrist, Joshua, Autor, David, & Pallais, Amanda. 2022. Marginal effects of merit aid for low-income students. The Quarterly Journal of Economics, 137(2), 1039–1090.
- Bahr, Peter Riley, Gross, Jillian Leigh, Slay, Kelly E, & Christensen, Rebecca D. 2015. First in line: Student registration priority in community colleges. *Educational Policy*, 29(2), 342–374.
- Bartik, Timothy J, Hershbein, Brad, & Lachowska, Marta. 2021. The effects of the Kalamazoo Promise Scholarship on college enrollment and completion. *Journal of Human Resources*, 56(1), 269–310.
- Bleemer, Zachary, & Mehta, Aashish. 2021. College major restrictions and student stratification. Available at SSRN 3981921.
- Bleemer, Zachary, & Mehta, Aashish. 2022. Will studying economics make you rich? A regression discontinuity analysis of the returns to college major. American Economic Journal: Applied Economics, 14(2), 1–22.
- Buckles, Kasey. 2019. Fixing the leaky pipeline: Strategies for making economics work for women at every stage. *Journal of economic perspectives*, **33**(1), 43–60.

- Carrell, Scott, & Sacerdote, Bruce. 2017. Why do college-going interventions work? American Economic Journal: Applied Economics, 9(3), 124–151.
- Carrell, Scott E, Page, Marianne E, & West, James E. 2010. Sex and science: How professor gender perpetuates the gender gap. The Quarterly journal of economics, 125(3), 1101– 1144.
- Chakrabarti, Rajashri, Gorton, Nicole, & Lovenheim, Michael F. 2020. State investment in higher education: Effects on human capital formation, student debt, and long-term financial outcomes of students. Tech. rept. National Bureau of Economic Research.
- Chevalier, Arnaud. 2011. Subject choice and earnings of UK graduates. Economics of Education Review, 30(6), 1187–1201.
- Deming, David J. 2017. Increasing college completion with a federal higher education matching grant.
- Denning, Jeffrey T, & Turley, Patrick. 2017. Was that SMART? Institutional financial incentives and field of study. *Journal of Human Resources*, 52(1), 152–186.
- Denning, Jeffrey T, Eide, Eric R, Mumford, Kevin J, & Sabey, Daniel J. 2022. Decreasing time to baccalaureate degree in the United States. *Economics of Education Review*, 90, 102287.
- Elsner, Benjamin, Isphording, Ingo E, & Zölitz, Ulf. 2021. Achievement rank affects performance and major choices in college. *The Economic Journal*, **131**(640), 3182–3206.
- Evans, William N, Kearney, Melissa S, Perry, Brendan, & Sullivan, James X. 2020. Increasing community college completion rates among low-income students: Evidence from a randomized controlled trial evaluation of a case-management intervention. *Journal of Policy Analysis and Management*, **39**(4), 930–965.
- Fairlie, Robert W, Hoffmann, Florian, & Oreopoulos, Philip. 2014. A community college instructor like me: Race and ethnicity interactions in the classroom. *American Economic Review*, **104**(8), 2567–2591.

- Garibaldi, Pietro, Giavazzi, Francesco, Ichino, Andrea, & Rettore, Enrico. 2012. College cost and time to complete a degree: Evidence from tuition discontinuities. *Review of Economics* and Statistics, 94(3), 699–711.
- Griffith, Amanda L. 2010. Persistence of women and minorities in STEM field majors: Is it the school that matters? *Economics of Education Review*, **29**(6), 911–922.
- Haggag, Kareem, Patterson, Richard W, Pope, Nolan G, & Feudo, Aaron. 2021. Attribution bias in major decisions: Evidence from the United States Military Academy. *Journal of Public Economics*, 200, 104445.
- Hastings, Justine S, Neilson, Christopher A, & Zimmerman, Seth D. 2013. Are some degrees worth more than others? Evidence from college admission cutoffs in Chile. Tech. rept. National Bureau of Economic Research.
- Kirkeboen, Lars J, Leuven, Edwin, & Mogstad, Magne. 2016. Field of study, earnings, and self-selection. The Quarterly Journal of Economics, 131(3), 1057–1111.
- Kremer, Michael, Miguel, Edward, & Thornton, Rebecca. 2009. Incentives to learn. The Review of Economics and Statistics, 91(3), 437–456.
- Kugler, Adriana D, Tinsley, Catherine H, & Ukhaneva, Olga. 2021. Choice of majors: Are women really different from men? *Economics of Education Review*, 81, 102079.
- Kurlaender, Michal, Jackson, Jacob, Howell, Jessica S, & Grodsky, Eric. 2014. College course scarcity and time to degree. *Economics of Education Review*, 41, 24–39.
- Li, Cher Hsuehhsiang, & Zafar, Basit. 2023. Ask and you shall receive? gender differences in regrades in college. *American Economic Journal: Economic Policy*, **15**(2), 359–394.
- Mitchell, Michael, Leachman, Michael, & Masterson, Kathleen. 2016. Funding down, tuition up: State cuts to higher education threaten quality and affordability at public colleges.
- Müller, Tomáš, Barták, Roman, & Rudová, Hana. 2004. Iterative forward search: Combining local search with maintaining arc consistency and a conflict-based statistics. Page 1 of: 1st International Workshop on Local Search Techniques in Constraint Satisfaction. Citeseer.

- Müller, Tomáš, Murray, Keith, et al. 2010. Comprehensive approach to student sectioning. Annals of Operations Research, 181(1), 249–269.
- Patnaik, Arpita, Venator, Joanna, Wiswall, Matthew, & Zafar, Basit. 2020. The Role of Heterogeneous Risk Preferences, Discount Rates, and Earnings Expectations in College Major Choice. Tech. rept. National Bureau of Economic Research.
- Patterson, Richard W, Pope, Nolan G, & Feudo, Aaron. 2021. Timing matters: Evidence from college major decisions. *Journal of Human Resources*, 0820–11127R1.
- Price, Joshua. 2010. The effect of instructor race and gender on student persistence in STEM fields. *Economics of Education Review*, **29**(6), 901–910.
- Robles, Silvia, Gross, Max, & Fairlie, Robert W. 2021. The effect of course shutouts on community college students: Evidence from Waitlist Cutoffs. *Journal of Public Economics*, 199, 104409.
- Rosenbaum, Paul R, & Rubin, Donald B. 1983. Assessing sensitivity to an unobserved binary covariate in an observational study with binary outcome. *Journal of the Royal Statistical Society: Series B (Methodological)*, 45(2), 212–218.
- Scott-Clayton, Judith. 2011. On money and motivation a quasi-experimental analysis of financial incentives for college achievement. *Journal of Human resources*, **46**(3), 614–646.
- Sjoquist, David L, & Winters, John V. 2015. State merit aid programs and college major: A focus on STEM. Journal of Labor Economics, 33(4), 973–1006.
- Speer, Jamin D. 2023. Bye Bye Ms. American Sci: women and the leaky stem pipeline. Economics of Education Review, 93, 102371.
- Webber, Douglas A. 2014. The lifetime earnings premia of different majors: Correcting for selection based on cognitive, noncognitive, and unobserved factors. *Labour economics*, 28, 14–23.
- Wiswall, Matthew, & Zafar, Basit. 2014. Determinants of college major choice: Identification using an information experiment. The Review of Economic Studies, 82(2), 791–824.

- Yim, Anthony LokTing. 2023. How Early Morning Classes Change Academic Trajectories: Evidence from a Natural Experiment. Available at SSRN 4618616.
- Zölitz, Ulf, & Feld, Jan. 2021. The effect of peer gender on major choice in business school. Management Science, 67(11), 6963–6979.

Tables and Figures

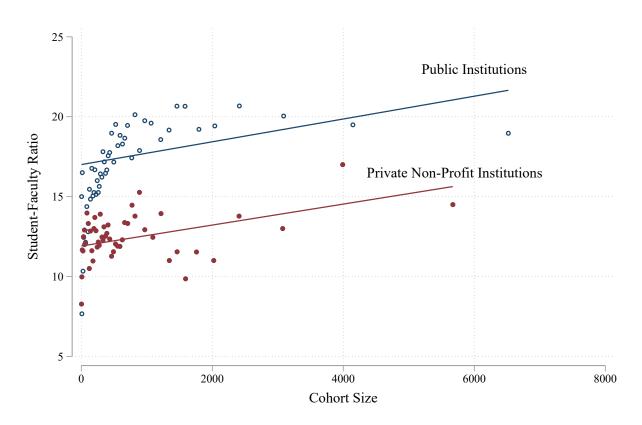
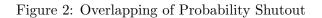
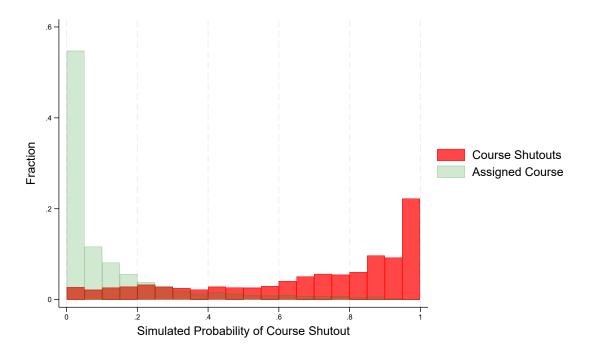


Figure 1: Student-Faculty Ratio at Public and Private Institutions

This figure shows the average student-faculty ratio by the size of the incoming class bins as reported in IPEDS for the 2019 academic year separately for public institutions and private non-profit institutions.





This figure plots the density of course shutout probabilities separately for requests that result in shutouts and course assignments. The probability of course shutouts is defined as the fraction of course requests from 1000 algorithm simulations that result in a shutout.

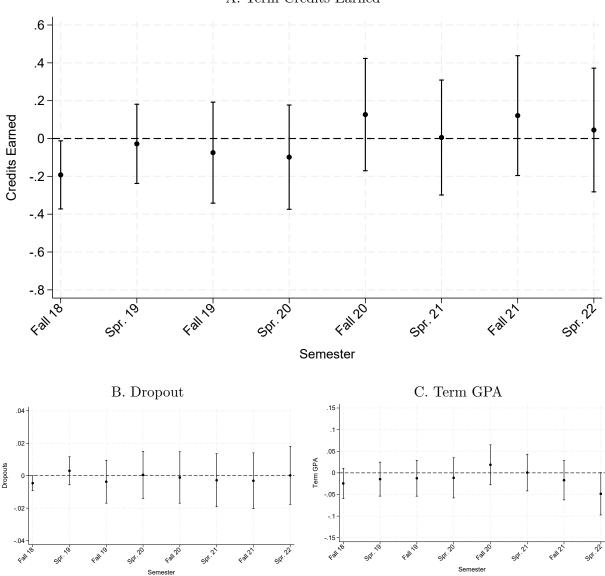


Figure 3: Effects of First-Semester Total Shutouts by Semester

A. Term Credits Earned

All estimates are at the student level as outlined in Equation 3. Panel A reports the effect of first-semester (Fall 2018) shutouts on credits earned in Fall 2018 and each of the subsequent 7 semesters. Panel B reports the effect of first-term shutouts on whether individuals have dropped out of Purdue by the referenced semester. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent 7 semesters. Panel C reports the effect of first-term shutouts on GPAs earned in Fall 2018 and each of the subsequent 7 semesters. GPA is omitted if the student is not enrolled in the term. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

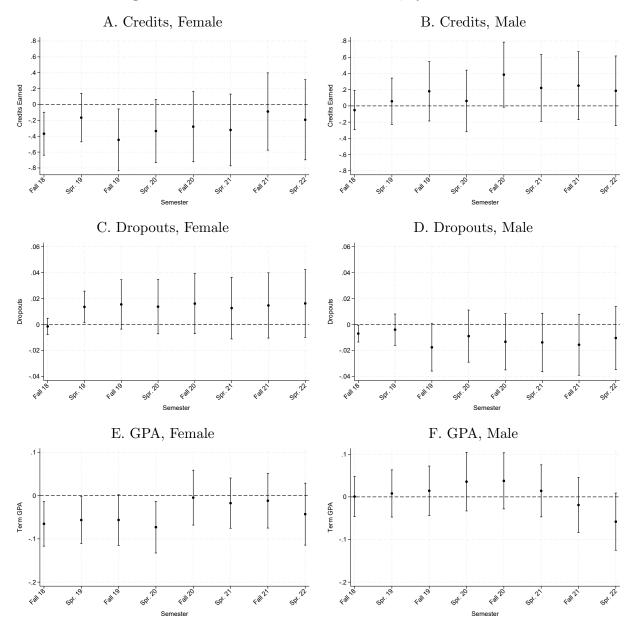


Figure 4: Effects of First-Semester Shutouts, by Term and Sex

All estimates are at the student level as outlined in Equation 3. Panels A, C, and E estimate the effects of shutouts on outcomes for female students while Panels B, D, and F estimate the effects of shutouts on outcomes for male students. Panels A and B report the effects of first-term (Fall 2018) shutouts on credits earned in Fall 2018 and each of the subsequent 7 semesters. Panels C and D report the effects of first-term shutouts on whether individuals have dropped out of Purdue by the referenced semester. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent semesters. Finally, Panels E and F report the effects of first-term shutouts on GPAs earned in Fall 2018 and each of the subsequent 7 semesters. GPA is omitted if the student is not enrolled in the term. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

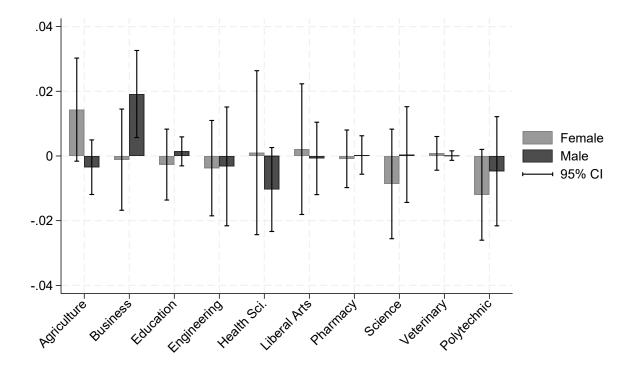


Figure 5: Effects of Shutouts on Major Choice

This Figure estimates the effect of shutouts on choosing a major from each of the 10 colleges at Purdue University. Engineering, Health Science, Pharmacy, Science, and Polytechnic Colleges are primarily comprised of STEM majors while Agriculture, Business, Education, Liberal Arts, and Veterinary colleges are primarily comprised of Non-STEM majors. Estimates are at the student level, as outlined in Equation 3 and are reported by gender. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

| | All | Analysis Sample | 1+ Shutouts | No Shutouts |
|----------------------|-------|-----------------|-------------|-------------|
| | (1) | (2) | (3) | (4) |
| Any shutouts | 0.51 | 0.55 | 1.00 | 0.00 |
| Total shutouts | 0.62 | 0.67 | 1.20 | 0.00 |
| Female | 0.43 | 0.44 | 0.41 | 0.47 |
| First generation | 0.17 | 0.17 | 0.17 | 0.17 |
| Asian | 0.11 | 0.11 | 0.12 | 0.10 |
| Black | 0.03 | 0.03 | 0.02 | 0.03 |
| Hispanic | 0.06 | 0.06 | 0.06 | 0.05 |
| White | 0.65 | 0.66 | 0.67 | 0.66 |
| Other race/ethnicity | 0.17 | 0.15 | 0.13 | 0.17 |
| Math SAT | 662 | 664 | 664 | 663 |
| Verbal SAT | 649 | 651 | 652 | 651 |
| Observations | 8,566 | 7,646 | 4,241 | $3,\!405$ |

Table 1: Summary Statistics

Summary Statistics are for the 8,566 freshmen from the Fall 2018 entering cohort. Our analysis sample of 7,646 excludes 213 students who are over age 23, 108 Division I athletes (who receive special treatment during scheduling due to unique practice and game restrictions), 588 students who did not request any potentially over-subscribed courses, 5 individuals who have degenerate probabilities of shutout for each of their requested courses, and 6 individuals who do not fulfill the requirement to declare a major prior to enrollment. ($P[Shutout_{ic}] = 1$ or $P[Shutout_{ic}] = 0$). We use 1,000 schedule algorithm simulations to determine the probability of shutout for each course request.

| | (1) | (2) | (3) |
|------------------------------------|---|-------------------------|--------------------------|
| Female | -0.008 (0.007) | -0.002 (0.006) | 0.003 (0.006) |
| Black | -0.009 (0.018) | -0.014 (0.019) | -0.003 (0.016) |
| Hispanic | -0.012 (0.012) | -0.011 (0.012) | -0.003 (0.011) |
| Asian | $\begin{array}{c} 0.024^{***} \\ (0.009) \end{array}$ | 0.018^{**} (0.008) | $0.013 \\ (0.008)$ |
| Other Race/Ethnicity | 0.022^{**} (0.009) | 0.011 (0.008) | $0.011 \\ (0.008)$ |
| First Generation | $0.013 \\ (0.009)$ | 0.017^{*} (0.009) | $0.011 \\ (0.007)$ |
| SAT Math | $0.009 \\ (0.006)$ | $0.005 \\ (0.005)$ | $0.000 \\ (0.004)$ |
| SAT Verbal | -0.001 (0.004) | -0.006 (0.004) | -0.001 (0.004) |
| Course in Pre-enrolled Major | -0.083^{***} (0.023) | -0.016 (0.031) | -0.003 (0.018) |
| Simulated Shutout Probability | | | 0.991^{***} (0.012) |
| F-stat P-Value | 0 | .077 | .529 |
| Observations | $15,\!121$ | $15,\!121$ | $15,\!121$ |
| R^2 | 0.308 | 0.384 | 0.559 |
| Course FE Course-by-Priority FE | X _ | \mathbf{X}^{-} | X _ |

Table 2: Student-by-Course Level Balance Test

* p < 0.10, ** p < 0.05, *** p < 0.01. The outcome in this regression is an indicator for being shutout of a potentially oversubscribed course. This sample includes all course-by-individual observations in potentially oversubscribed courses. F-stat p-value comes from a joint test of significance for sex, race/ethnicity, first-generation status, SAT score variables, and whether the course is in the student's pre-enrollment major. Column (3) corresponds to our primary individual-by-course level specification. Columns (2) and (3) additionally control for an indicator for a potential "reservation course". Certain majors reserve slots in classes for at least some students pre-enrolled in a corresponding major. Individuals are in a "reservation course" if they (1) are pre-enrolled in the major offering the course and (2) the major is one of the majors that reserve slots for some students. if they Standard errors clustered at the individual level are reported in parentheses.

| | (1) | (2) | (3) |
|-------------------------------|---------------------------|------------------------|---|
| Female | -0.081^{***} (0.017) | -0.007 (0.018) | $0.012 \\ (0.011)$ |
| Black | -0.015 (0.050) | -0.018 (0.049) | -0.008 (0.032) |
| Hispanic | -0.004 (0.034) | -0.005 (0.033) | -0.005 (0.022) |
| Asian | 0.078^{***} (0.026) | $0.031 \\ (0.024)$ | $0.019 \\ (0.016)$ |
| Other Race/Ethnicity | -0.062^{***} (0.023) | 0.045^{*} (0.025) | $0.017 \\ (0.015)$ |
| First Generation | -0.009 (0.021) | $0.002 \\ (0.020)$ | $0.021 \\ (0.014)$ |
| SAT Math | -0.029^{**} (0.012) | -0.003 (0.012) | -0.000 (0.008) |
| SAT Verbal | $0.006 \\ (0.012)$ | $0.002 \\ (0.012)$ | $0.000 \\ (0.008)$ |
| Summed Shutout Probabilities | | | $\begin{array}{c} 0.994^{***} \\ (0.010) \end{array}$ |
| F-stat P-Value | 0 | .75 | .656 |
| Observations | 7646.000 | 7646.000 | 7646.000 |
| R^2 | 0.006 | 0.530 | 0.610 |
| Pre-enrolled Major | — | Х | Х |
| Number of Reservation Courses | — | Х | X |
| Simulated Shutout Probability | - | — | Х |

Table 3: Student-Level Balance Test

* p < 0.10, ** p < 0.05, *** p < 0.01. The outcome in this regression is the number of courses students being shut out. This sample includes all individual observations in potentially oversubscribed courses. F-stat p-value comes from a joint test of significance for sex, race/ethnicity, first-generation status, and SAT score variables. Column (3) corresponds to our primary individual-level specification. Certain majors reserve slots in classes for at least some students pre-enrolled in a corresponding major. Individuals are in a "reservation course" if they (1) are pre-enrolled in the major offering the course and (2) the major is one of the majors that reserve slots for some students. Robust standard errors are reported in parentheses

| | Attend First Semester (1) | Complete First Semester (2) | Complete First Year (3) | Ever Complete (4) |
|-------------------------------|------------------------------|--------------------------------|-------------------------|----------------------|
| Shutout of Course | -0.723*** | -0.712*** | -0.440*** | -0.352*** |
| | (0.009) | (0.009) | (0.011) | (0.010) |
| Observations | 15,121 | 15,121 | 15,121 | 15,121 |
| R^2 | 0.699 | 0.667 | 0.512 | 0.479 |
| Non-Shutout Mean | 0.852 | 0.836 | 0.850 | 0.857 |
| Simulated Shutout Probability | Х | Х | Х | Х |
| Demographic Characteristics | Х | Х | Х | Х |
| Course FE | Х | Х | Х | Х |

Table 4: Effect of Course Shutout on Attendance and Completion in Oversubscribed Courses

* p < 0.10, ** p < 0.05, *** p < 0.01. Each outcome in this table relates to the exact course a student was potentially shut out from. Column (1) reports the effects of a shutout on attending a requested course in the semester the request is made. Attendance is defined as being enrolled in the course after the add/drop deadline. Column (2) reports the effects of a shutout on completing a requested course in the semester the request is made. Column (3) reports the effects of a shutout on completing a requested course in Fall, Spring, or Summer semester of the 2018/2019 school year. Column (4) reports the effect of a shutout on completing a requested course by the Fall 2022 semester. Observations are at the student-course level. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects. Standard errors that are clustered at the individual level are reported in parentheses.

| Panel A: Requested | Subject | | | | | | | |
|--|----------------|------------|------------|--|--|--|--|--|
| | First Semester | First Year | Ever | | | | | |
| | (1) | (2) | (3) | | | | | |
| Shutout of Course | -0.608*** | -0.343*** | -0.254*** | | | | | |
| | (0.010) | (0.010) | (0.010) | | | | | |
| Observations | $15,\!121$ | $15,\!121$ | $15,\!121$ | | | | | |
| R^2 | 0.581 | 0.431 | 0.398 | | | | | |
| Non-Shutout Mean | 0.879 | 0.896 | 0.905 | | | | | |
| Panel B: Subjects Outside of Requested Subject | | | | | | | | |
| | First Semester | First Year | Ever | | | | | |
| | (1) | (2) | (3) | | | | | |
| Shutout of Course | 0.554^{***} | 0.322*** | 0.255*** | | | | | |
| | (0.023) | (0.040) | (0.089) | | | | | |
| Observations | $15,\!121$ | $15,\!121$ | $15,\!121$ | | | | | |
| R^2 | 0.256 | 0.154 | 0.062 | | | | | |
| Non-Shutout Mean | 3.647 | 6.043 | 11.790 | | | | | |
| Panel C: All Subject | ts | | | | | | | |
| | First Semester | First Year | Ever | | | | | |
| | (1) | (2) | (3) | | | | | |
| Shutout of Course | -0.061*** | -0.016 | 0.040 | | | | | |
| | (0.022) | (0.040) | (0.090) | | | | | |
| Observations | $15,\!121$ | $15,\!121$ | $15,\!121$ | | | | | |
| R^2 | 0.197 | 0.138 | 0.062 | | | | | |
| Non-Shutout Mean | 4.592 | 7.001 | 12.758 | | | | | |

Table 5: Effect of Course Shutout on Exposure to Subjects

* p < 0.10, ** p < 0.05, *** p < 0.01. Observations are at the student-course level. In Panel A the outcome is whether students take at least one course from the requested subject. In Panel B the outcome is the number of subjects outside of the requested course in which the student takes at least one course . In Panel C the outcome is the total number of subjects in which a student takes at least one course. Column (1) reports the effects of a shutout on exposure to the relevant subject(s) in the semester the request is made. Column (2) reports the effects of a shutout on on exposure to the relevant subject(s) in the Fall, Spring, or Summer semester of the 2018/2019 school year. Column (3) reports the effect of a shutout on exposure to the relevant subject(s) by the Fall 2022 semester. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects. Standard errors that are clustered at the individual level are reported in parentheses.

| | All Courses (1) | $\begin{array}{c} \text{STEM} \\ (2) \end{array}$ | Non-STEM (3) | $\begin{array}{c} \text{Top } 3\\ (4) \end{array}$ | Bottom 3 (5) |
|--|---|---|---------------------------|--|---------------------------|
| Shutout of Course | -0.008 (0.006) | -0.025^{*} (0.013) | $0.002 \\ (0.005)$ | -0.008 (0.011) | -0.006 (0.007) |
| Observations R^2 Non-Shutout Mean | $\begin{array}{c} 12,309 \\ 0.605 \\ 0.097 \end{array}$ | | $5,796 \\ 0.629 \\ 0.138$ | $5,146 \\ 0.648 \\ 0.120$ | $7,163 \\ 0.591 \\ 0.142$ |

Table 6: Effect of Course Shutout on Choosing a Corresponding Major

* p < 0.10, ** p < 0.05, *** p < 0.01. The outcome in each column is whether a student chooses a major that corresponds to the course they request. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. Majoring in a corresponding subject is defined by choosing a major that shares the same subject code as the course. Observations from courses that do not correspond to a major (e.g. subjects only offered as a minor) or do not fulfill any potential credits in the major (e.g. remedial math courses) are omitted. Column (1) reports results for all potential shutouts. Columns (2) and (3) report results for requested STEM and non-STEM courses, respectively. Columns (4) and (5) report results for top 3 and bottom 3 priority requests, respectively. Observations are at the student-course level. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects. Standard errors that are clustered at the individual level are reported in parentheses.

| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
|--|-----------------------------|-------------------------|---------------------------|---------------------------|
| Total Shutouts | $0.048 \\ (0.881)$ | -0.020 (0.015) | -0.018 (0.013) | -0.016^{*} (0.009) |
| $\begin{array}{l} \text{Observations} \\ R^2 \\ \text{Non-Shutout Mean} \end{array}$ | $7,646 \\ 0.130 \\ 109.000$ | 7,532 0.146 3.252 | $7,646 \\ 0.090 \\ 0.604$ | $7,583 \\ 0.514 \\ 0.693$ |

Table 7: Effect of Shutouts on Student-Level Outcomes

* p < 0.10, ** p < 0.05, *** p < 0.01. Column (1) reports the effects of total first-term shutouts on the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of total first-term shutouts on cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of total first-term shutouts on graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| Panel A: Female | | | | |
|-----------------------|--------------------|----------------|-------------------|------------|
| | Cumulative Credits | Cumulative GPA | 4-Year Graduation | STEM Major |
| | (1) | (2) | (3) | (4) |
| Total Shutouts | -1.778 | -0.048** | -0.051*** | -0.029** |
| | (1.296) | (0.021) | (0.019) | (0.014) |
| Observations | 3,336 | $3,\!315$ | 3,336 | 3,336 |
| R^2 | 0.076 | 0.168 | 0.103 | 0.529 |
| Non-Shutout Mean | 113.000 | 3.330 | 0.669 | 0.579 |
| Panel B: Male | | | | |
| | Cumulative Credits | Cumulative GPA | 4-Year Graduation | STEM Major |
| | (1) | (2) | (3) | (4) |
| Total Shutouts | 1.250 | -0.001 | 0.005 | -0.004 |
| | (1.200) | (0.020) | (0.017) | (0.011) |
| Observations | 4,310 | 4,217 | 4,310 | 4,247 |
| R^2 | 0.164 | 0.132 | 0.084 | 0.459 |
| Non-Shutout Mean | 105.000 | 3.180 | 0.547 | 0.796 |
| Female vs. Male p-val | 0.083 | 0.118 | 0.028 | 0.140 |

Table 8: Effect of Shutouts on Student-Level Outcomes, by Gender

* p < 0.10, ** p < 0.05, *** p < 0.01. Panel A reports results for female students and Panel B reports results for male students. Column (1) reports the effects of total first-term shutouts on the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of total first-term shutouts on cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of total first-term shutouts on graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| Panel A: Female | | | | |
|--|--|-----------------------------|-----------------------------|--|
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| One Shutouts | -1.000 (1.540) | -0.012 (0.024) | -0.051^{**} (0.022) | -0.033^{*} (0.017) |
| Two Plus Shutouts | -3.517 (2.914) | -0.093^{**} (0.046) | -0.111^{***} (0.042) | -0.054^{*} (0.032) |
| Observations R^2 Non-Shutout Mean | $3,336 \\ 0.076 \\ 113.169$ | $3,315 \\ 0.167 \\ 113.793$ | $3,336 \\ 0.103 \\ 113.169$ | $3,336 \\ 0.529 \\ 113.169$ |
| Panel B: Male | | | | |
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| One Shutouts | 1.440 (1.437) | -0.004 (0.024) | $0.008 \\ (0.020)$ | $0.009 \\ (0.013)$ |
| Two Plus Shutouts | 1.388 (2.644) | -0.026 (0.045) | -0.003 (0.037) | -0.014 (0.023) |
| Observations R^2 Non-Shutout Mean | $\begin{array}{c} 4,310 \\ 0.164 \\ 104.709 \end{array}$ | 4,217 0.132 107.135 | $4,310 \\ 0.084 \\ 104.709$ | $\begin{array}{c} 4,247 \\ 0.460 \\ 106.234 \end{array}$ |

Table 9: Exploring Non-Linear Effects of Shutouts on Cumulative Outcomes by Gender

* p < 0.10, ** p < 0.05, *** p < 0.01. A total of 3,478 students experienced a single shutout. Among the 766 students who faced at least two total shutouts, 676 had two shutouts, and 85 experienced three shutouts. Only five students had the experience of encountering four total shutouts. Panel A reports results for female students and Panel B reports results for male students. Column (1) reports the effects of total first-term shutouts on the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of one first-term shutouts and two plus first-term shutouts on cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of one first-term shutouts and two plus first-term shutouts on graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| | Analysis Sample | Female | Male | (2) vs. (3) |
|------------------------|-----------------|-----------|-------|---------------|
| | | | | P-value |
| | (1) | (2) | (3) | (4) |
| Course Characteristics | | | | |
| STEM | 0.53 | 0.52 | 0.54 | 0.002 |
| Upper Level | 0.04 | 0.05 | 0.02 | 0.000 |
| Required | 0.82 | 0.84 | 0.81 | 0.000 |
| Difficult | 0.39 | 0.39 | 0.40 | 0.035 |
| Course College | | | | |
| Agriculture | 0.03 | 0.05 | 0.02 | 0.000 |
| Business | 0.04 | 0.03 | 0.04 | 0.000 |
| Education | 0.03 | 0.05 | 0.01 | 0.000 |
| Engineering | 0.03 | 0.01 | 0.04 | 0.000 |
| Health Sci. | 0.02 | 0.03 | 0.01 | 0.000 |
| Liberal Arts | 0.42 | 0.43 | 0.42 | 0.001 |
| Pharmacy | 0.01 | 0.01 | 0.01 | 0.000 |
| Science | 0.33 | 0.34 | 0.32 | 0.053 |
| Polytechnic | 0.08 | 0.04 | 0.11 | 0.000 |
| Observations | 15,184 | $6,\!566$ | 8,618 | |

Table 10: Course Characteristics of Shutouts

Observations are in the course-by-individual level. Column (1) reports the share of shutouts by different course characteristics and course colleges from our overall analysis sample. Column (2) reports the share of shutouts by different course characteristics and course colleges from our female student sample while Column (3) reports the share of shutouts by different course characteristics and course colleges from our male student sample. Column (4) reports the p-value of the gender difference by course characteristics and course colleges from Column (2) and Column (3).

Appendix I Appendix- Supplementary Figures and Tables

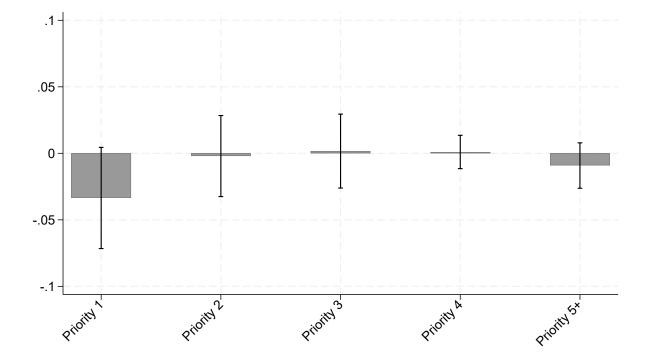


Figure A.1: Effect of Course Shutout on Choosing a Corresponding Major, by Priority

Each bar shows the estimated effect of a shutout on whether a student chooses a major that corresponds to the requested course. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. Majoring in a corresponding subject is defined by choosing a major that shares the same subject code as the course. Observations from courses that do not correspond to a major (e.g. subjects only offered as a minor) or do not fulfill any potential credits in the major (e.g. remedial math courses) are omitted. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects. 95% confidence intervals come from robust standard errors clustered at the individual level.

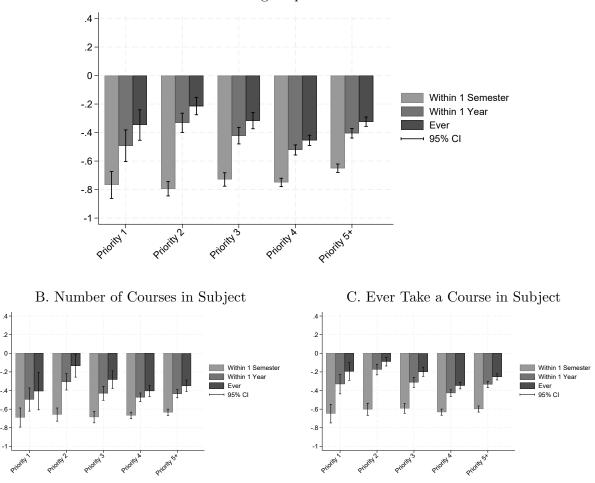


Figure A.2: Effect of Course Shutout on Course-Taking, by Priority

A. Taking Requested Course

Each bar represents the estimated priority-specific effects of a shutout on an outcome. 95% confidence intervals are reported and derived from robust standard errors. Panel A estimates the effects of shutouts in 1st, 2nd, ..., 5th+ priority-ranked requests on taking the corresponding priority-ranked course. Panel B estimates the effects of shutouts in 1st, 2nd, ..., 5th+ priority-ranked requests on the number of courses taken in the corresponding priority-ranked subject. Panel C estimates the effects of shutouts in 1st, 2nd, ..., 5th+ priority-ranked request on whether a student has taken a course in the corresponding priority-ranked subject. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects.

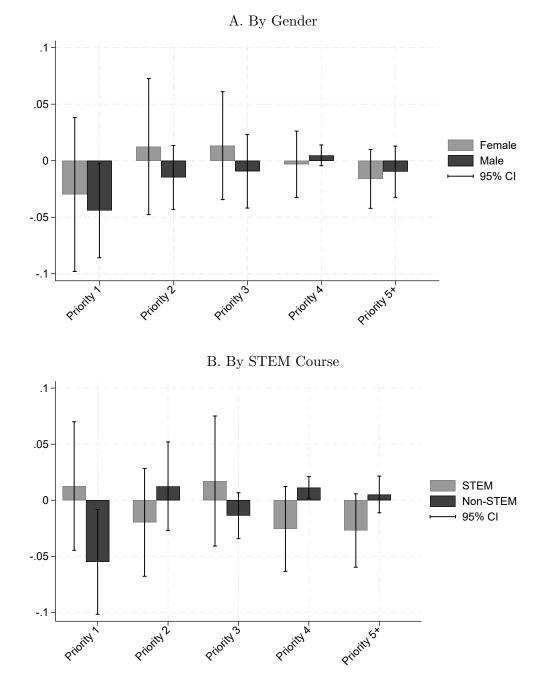


Figure A.3: Heterogeneous Effect of Course Shutout on Choosing a Corresponding Major

Each bar shows the estimated effect of a shutout on whether a student chooses a major that corresponds to the requested course. In Panel A, bars show estimates for each priority-gender pair. In Panel B, bars show estimates for each priority-STEM designation pair. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. Majoring in a corresponding subject is defined by choosing a major that shares the same subject code as the course. Observations from courses that do not correspond to a major (e.g. subjects only offered as a minor) or do not fulfill any potential credits in the major (e.g. remedial math courses) are omitted. Each estimate includes controls for simulated shutout probability, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), an indicator for a reservation course, and course-fixed effects. 95% confidence intervals come from robust standard errors clustered at the individual level.

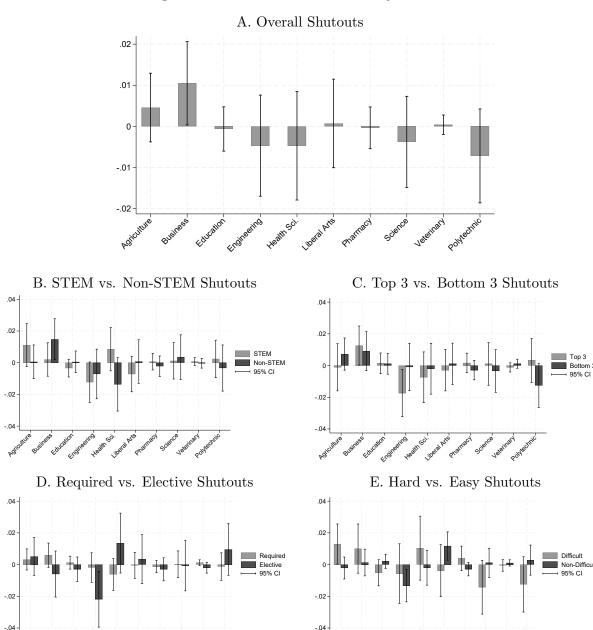


Figure A.4: Effect of Shutouts on Major Choice

This Figure estimates the effect of shutouts on choosing a major from each of the 10 Purdue Colleges. Panel A shows estimates for all students in our sample, Panel B shows estimates separately for shutouts in STEM and Non-STEM courses, and Panel C shows estimates separately for shutouts in top 3 priority and bottom 3 priority courses. Engineering, Health Science, Pharmacy, Science, and Polytechnic Colleges are primarily comprised of STEM majors while Agriculture, Business, Education, Liberal Arts, and Veterinary colleges offer majors that are not STEM designated. Estimates are at the individual level, as outlined in Equation 3. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

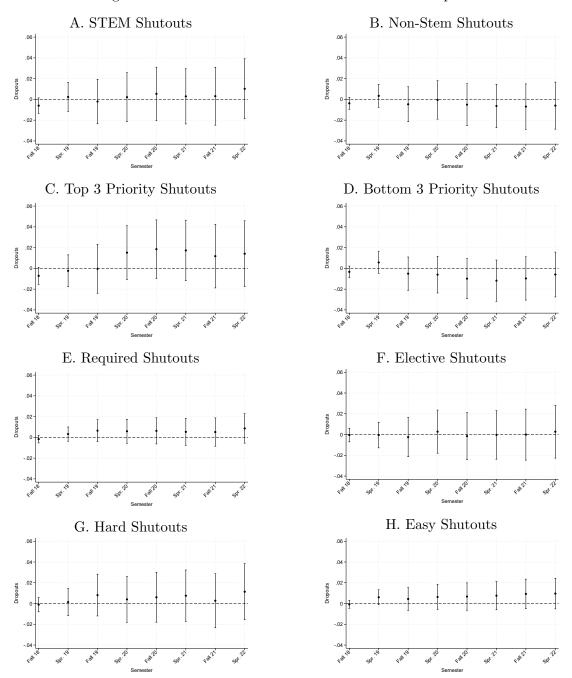


Figure A.5: Effects of First-Semester Shutouts on Dropout

All estimates are at the individual level as outlined in Equation 3. Panels A and B show the effects of STEM and Non-STEM shutouts, respectively. Panels C and D show the effects of shutouts in top 3 priority and bottom three priority courses, respectively. Panels E and F show the effects of shutouts in courses that meet a general education requirement and do not meet a general education requirement, respectively. Finally, Panels G and H show the effects in above-median and below-median difficulty courses, respectively. Difficulty of course is defined as the inverse of residual GPAs in courses, after accounting for student observable characteristics. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent semesters. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

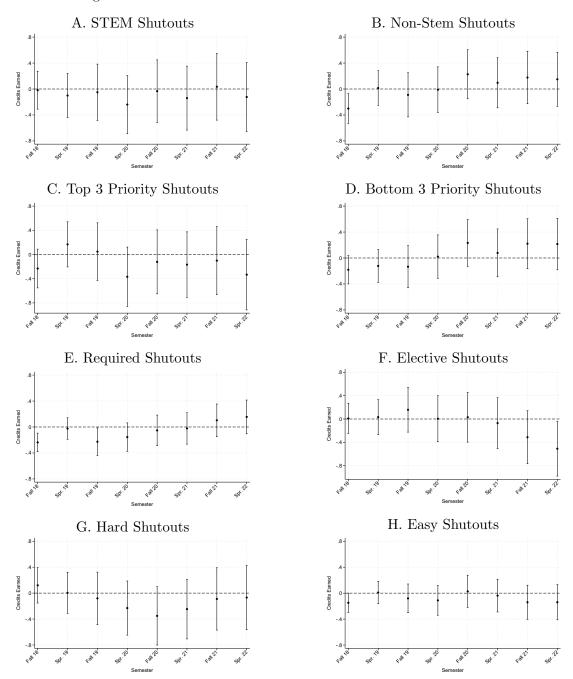


Figure A.6: Effects of First-Semester Shutouts on Credits Earned

All estimates are at the individual level as outlined in Equation 3. Panels A and B show the effects of STEM and Non-STEM shutouts, respectively. Panels C and D show the effects of shutouts in top 3 priority and bottom three priority courses, respectively. Panels E and F show the effects of shutouts in courses that meet a general education requirement and do not meet a general education requirement, respectively. Finally, Panels G and H show the effects in above-median and below-median difficulty courses, respectively. Difficulty of course is defined as the inverse of residual GPAs in courses, after accounting for student observable characteristics. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

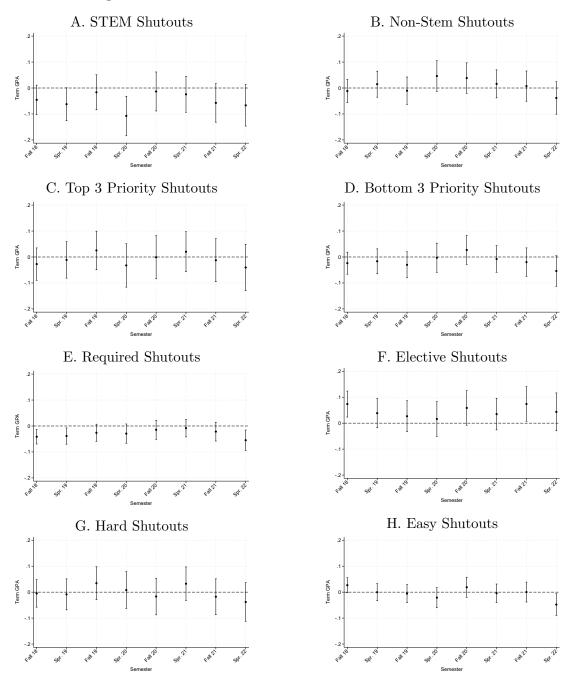


Figure A.7: Effects of First-Semester Shutouts on Term GPA

All estimates are at the individual level as outlined in Equation 3. Panels A and B show the effects of STEM and Non-STEM shutouts, respectively. Panels C and D show the effects of shutouts in top 3 priority and bottom three priority courses, respectively. Panels E and F show the effects of shutouts in courses that meet a general education requirement and do not meet a general education requirement, respectively. Finally, Panels G and H show the effects in above-median and below-median difficulty courses, respectively. Difficulty of course is defined as the inverse of residual GPAs in courses, after accounting for student observable characteristics. GPA is omitted if the student is not enrolled in the term. estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

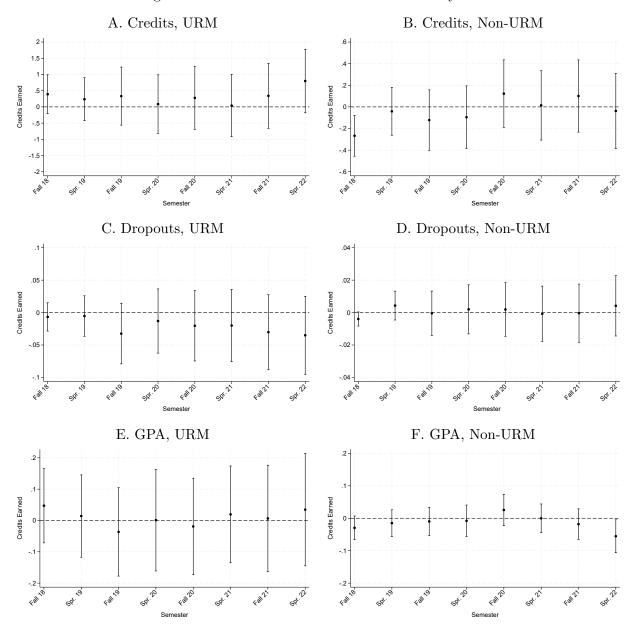


Figure A.8: Effects of First-Semester Shutouts by URM

All estimates are at the individual level as outlined in Equation 3. Panels A, C, and E estimate the effects of shutouts on outcomes for under-represented minority students, including Black, Hispanic, and Pacific Islander students while Panels B, D, and F estimate the effects of shutouts on outcomes for Non-Under-respresented minority students including White and Asian students. Panels A and B report the effects of first-term (Fall 2018) shutouts on credits earned in Fall 2018 and each of the subsequent 7 semesters. Panels C and D report the effects of first-term shutouts on whether individuals have dropped out of Purdue by the referenced semester. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent semesters. Finally, Panels E and F report the effects of first-term shutouts on GPAs earned in Fall 2018 and each of the subsequent is not enrolled in the term. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

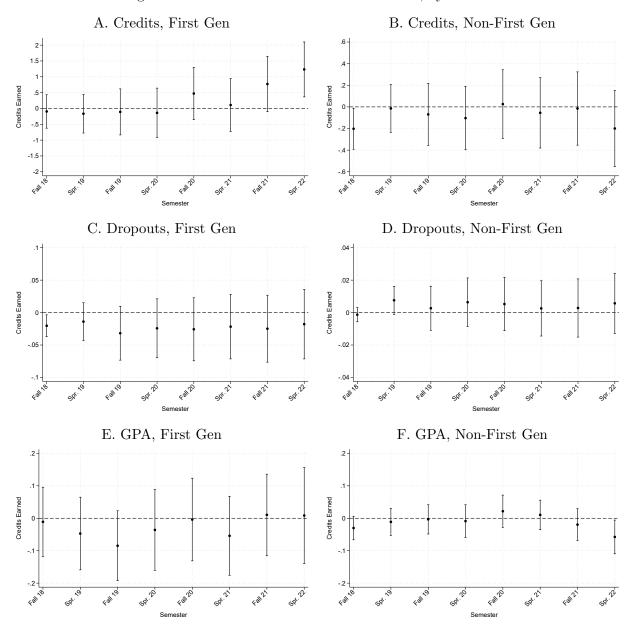


Figure A.9: Effects of First-Semester Shutouts, by First Gen

All estimates are at the individual level as outlined in Equation 3. Panels A, C, and E estimate the effects of shutouts on outcomes for first generation students while Panels B, D, and F estimate the effects of shutouts on outcomes for Non-first generation students. Panels A and B report the effects of first-term (Fall 2018) shutouts on credits earned in Fall 2018 and each of the subsequent 7 semesters. Panels C and D report the effects of first-term shutouts on whether individuals have dropped out of Purdue by the referenced semester. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent semesters. Finally, Panels E and F report the effects of first-term shutouts on GPAs earned in Fall 2018 and each of the subsequent 7 semesters. GPA is omitted if the student is not enrolled in the term. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

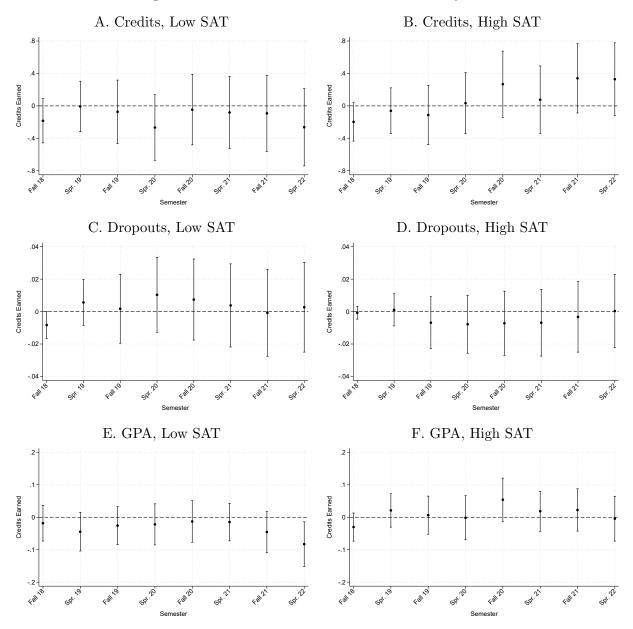


Figure A.10: Effects of First-Semester Shutouts by SAT

All estimates are at the individual level as outlined in Equation 3. Panels A, C, and E estimate the effects of shutouts on outcomes for below-median SAT students while Panels B, D, and F estimate the effects of shutouts on outcomes for above-median SAT students. Panels A and B report the effects of first-term (Fall 2018) shutouts on credits earned in Fall 2018 and each of the subsequent 7 semesters. Panels C and D report the effects of first-term shutouts on whether individuals have dropped out of Purdue by the referenced semester. We define a dropout as having dropped out by a semester if (1) they have not graduated and (2) they do not enroll in any courses in subsequent semesters. Finally, Panels E and F report the effects of first-term shutouts on GPAs earned in Fall 2018 and each of the subsequent 7 semesters is not enrolled in the term. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. 95% intervals from robust standard errors are reported.

| | Fei | male Studer | nts | Male Studuents | | |
|-------------------------------|---------------------------|---|---|---------------------------|-------------------------|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Black | -0.038 (0.023) | -0.052^{**} (0.026) | -0.023 (0.021) | $0.025 \\ (0.024)$ | $0.012 \\ (0.021)$ | $0.021 \\ (0.025)$ |
| Hispanic | -0.021 (0.022) | -0.014 (0.021) | -0.013 (0.017) | -0.003 (0.016) | -0.000 (0.017) | $0.004 \\ (0.015)$ |
| Asian | $0.026 \\ (0.017)$ | $0.019 \\ (0.017)$ | $0.019 \\ (0.014)$ | 0.025^{**} (0.012) | 0.024^{**} (0.012) | $0.011 \\ (0.011)$ |
| Other Race/Ethnicity | 0.028^{*} (0.016) | $\begin{array}{c} 0.011 \\ (0.014) \end{array}$ | $0.019 \\ (0.012)$ | $0.016 \\ (0.012)$ | $0.005 \\ (0.012)$ | $0.003 \\ (0.010)$ |
| First Generation | 0.031^{**} (0.012) | $\begin{array}{c} 0.043^{***} \\ (0.014) \end{array}$ | 0.018^{*} (0.010) | -0.002 (0.013) | $0.002 \\ (0.012)$ | $0.005 \\ (0.010)$ |
| SAT Math | $0.014 \\ (0.010)$ | $0.013 \\ (0.009)$ | -0.002 (0.006) | $0.004 \\ (0.006)$ | $0.001 \\ (0.006)$ | $0.000 \\ (0.005)$ |
| SAT Verbal | -0.001 (0.007) | -0.007 (0.007) | -0.000 (0.006) | -0.000 (0.005) | -0.005 (0.005) | -0.001 (0.006) |
| Course in Pre-enrolled Major | -0.083^{***} (0.030) | -0.027 (0.058) | -0.018 (0.032) | -0.101^{***} (0.030) | -0.014 (0.032) | $0.008 \\ (0.020)$ |
| Simulated Shutout Probability | | | $\begin{array}{c} 0.992^{***} \\ (0.017) \end{array}$ | | | $\begin{array}{c} 0.992^{***} \\ (0.016) \end{array}$ |
| F-stat P-Value | 0 | .004 | .238 | .033 | .633 | .967 |
| Observations | 6559.000 | 6559.000 | 6559.000 | 8562.000 | 8562.000 | 8562.000 |
| R^2 | 0.289 | 0.386 | 0.558 | 0.333 | 0.410 | 0.567 |
| Course FE | Х | _ | Х | Х | — | Х |
| Course-by-Priority FE | — | Х | — | — | Х | — |

Table A.1: Student-by-Course Level Balance Test, by Gender

* p < 0.10, ** p < 0.05, *** p < 0.01. The outcome in this regression is an indicator for being shutout of a potentially oversubscribed course. This sample includes all course-by-individual observations in potentially oversubscribed courses. F-stat p-value comes from a joint test of significance for sex, race/ethnicity, first-generation status, SAT score variables, and whether the course is in the student's pre-enrollment major. Columns (3) and (6) correspond to our primary individual-by-course level specification. Columns (2), (3), (5), and (6) additionally control for an indicator for a potential "reservation course". Certain majors reserve slots in classes for at least some students pre-enrolled in a corresponding major. Individuals are in a "reservation course" if they (1) are pre-enrolled in the major offering the course and (2) the major is one of the majors that reserve slots for some students. if they Standard errors clustered at the individual level are reported in parentheses.

| | Fe | male Stud | ents | Male Studuents | | |
|---|---------|-----------|----------|----------------|---------|----------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Black | -0.091 | -0.067 | -0.058 | 0.077 | 0.057 | 0.058 |
| | (0.068) | (0.075) | (0.042) | (0.075) | (0.078) | (0.048) |
| Hispanic | -0.022 | -0.068 | -0.032 | 0.014 | 0.012 | 0.023 |
| | (0.051) | (0.057) | (0.032) | (0.046) | (0.047) | (0.030) |
| Asian | 0.069 | -0.009 | 0.031 | 0.087^{***} | 0.034 | 0.014 |
| | (0.042) | (0.044) | (0.027) | (0.033) | (0.032) | (0.021) |
| Other Race/Ethnicity | -0.046 | -0.043 | 0.031 | -0.073** | 0.056 | 0.010 |
| | (0.037) | (0.045) | (0.023) | (0.030) | (0.035) | (0.020) |
| First Generation | 0.039 | 0.077** | 0.034* | -0.057* | -0.023 | 0.005 |
| | (0.031) | (0.034) | (0.019) | (0.030) | (0.031) | (0.019) |
| SAT Math | -0.024 | 0.000 | -0.004 | -0.032** | -0.010 | -0.000 |
| | (0.018) | (0.022) | (0.012) | (0.016) | (0.017) | (0.010) |
| SAT Verbal | 0.003 | 0.015 | 0.004 | 0.008 | 0.006 | -0.001 |
| | (0.018) | (0.020) | (0.011) | (0.016) | (0.017) | (0.011) |
| Cumulative Shutout Probability | | | 0.987*** | | | 0.997*** |
| , i i i i i i i i i i i i i i i i i i i | | | (0.015) | | | (0.013) |
| F-stat P-Value | .136 | .286 | .264 | .001 | .714 | .929 |
| Observations | 3341 | 2687 | 3335 | 4311 | 3700 | 4307 |
| R^2 | 0.003 | 0.638 | 0.624 | 0.006 | 0.546 | 0.603 |
| Pre-enrolled Major | _ | Х | Х | | | |
| Number of Reservation Courses | - | Х | Х | - | Х | Х |
| Simulated Shutout Probability | _ | _ | Х | — | — | Х |

Table A.2: Student Level Balance Test, by Gender

* p < 0.10, ** p < 0.05, *** p < 0.01. The outcome in this regression is the number of courses students being shut out. This sample includes all individual observations in potentially oversubscribed courses. F-stat p-value comes from a joint test of significance for sex, race/ethnicity, first-generation status, and SAT score variables. Columns (3) and (6) correspond to our primary individual-level specification. Columns (2), (3), (5), and (6) control for the number of "reservation courses" a student is enrolled in. Certain majors reserve slots in classes for at least some students pre-enrolled in a corresponding major. Individuals are in a "reservation course" if they (1) are pre-enrolled in the major offering the course and (2) the major is one of the majors that reserve slots for some students. Robust standard errors are reported in parentheses

| Panel A: Attendance in the First SemesterFemale URM Minority First Gen Low SAT(1)(2)(3)(4) | | | | | | | |
|--|---------------------------|---------------------------|---|---|--|--|--|
| Shutout of Course | -0.710^{***} (0.010) | -0.719^{***} (0.009) | -0.718^{***} (0.009) | -0.705^{***} (0.010) | | | |
| Interaction | -0.031^{***} (0.011) | -0.048^{***} (0.017) | -0.030^{**} (0.014) | -0.037^{***} (0.011) | | | |
| Demographic | 0.019^{***} (0.006) | 0.025^{**} (0.010) | $0.006 \\ (0.008)$ | $0.002 \\ (0.007)$ | | | |
| Observations R^2 | $15,121 \\ 0.699$ | $15,121 \\ 0.698$ | $15,121 \\ 0.699$ | $15,121 \\ 0.696$ | | | |
| Panel B: Ever Com | aplete | | | | | | |
| Shutout of Course | -0.353^{***} (0.012) | -0.353^{***} (0.011) | -0.354^{***} (0.011) | -0.384^{***} (0.013) | | | |
| Interaction | $0.001 \\ (0.016)$ | -0.006 (0.024) | $\begin{array}{c} 0.010 \\ (0.020) \end{array}$ | 0.063^{***} (0.015) | | | |
| Demographic | 0.023^{***} (0.006) | $0.009 \\ (0.010)$ | -0.005 (0.008) | -0.019^{***} (0.007) | | | |
| Observations R^2 | $15,121 \\ 0.479$ | $15,121 \\ 0.477$ | $15,121 \\ 0.479$ | $15,121 \\ 0.476$ | | | |
| Panel C: Courses 7 | Taken in Sar | ne Subject Ever | | | | | |
| Shutout of Course | -0.265^{***} (0.011) | -0.255^{***} (0.010) | -0.251^{***} (0.010) | -0.285^{***} (0.012) | | | |
| Interaction | 0.026^{*} (0.015) | -0.000 (0.024) | -0.020 (0.020) | $\begin{array}{c} 0.062^{***} \\ (0.015) \end{array}$ | | | |
| Demographic | 0.032^{***} (0.006) | 0.010 (0.009) | -0.003 (0.007) | -0.020^{***} (0.006) | | | |
| Observations R^2 | $15,121 \\ 0.398$ | $15,121 \\ 0.396$ | $15,121 \\ 0.398$ | $15,121 \\ 0.392$ | | | |

Table A.3: How the Effects of a Course Shutout Differ by Student Characteristics

* p < 0.10, ** p < 0.05, *** p < 0.01. Each outcome in this table relates to completion of courses in the same subject area as a course a student was potentially shutout from. "Shutout" is an indicator for being shutout of a requested course. Each regression controls for shutout probability, which is generated from simulations of the algorithm that generated freshman course assignments in the Fall of 2018. Major in subject There are 10 colleges at Purdue including Agriculture, Education, Engineering, Health and Human Sciences, Liberal Arts, Management, Pharmacy, Polytechnic, Science, and Veterinary Medicine. Standard errors clustered at the individual level are reported in parentheses.

| Panel A: Attend Co | Panel A: Attend Course in First Term | | | | | | |
|--------------------|--------------------------------------|-----------|-----------|-----------------|-----------|------------|--|
| | General I | Education | High D | High Difficulty | | references | |
| | Yes | No | Yes | No | Yes | No | |
| | (1) | (2) | (3) | (4) | (5) | (6) | |
| Shutout of Course | -0.762*** | -0.467*** | -0.773*** | -0.826*** | -0.771*** | -0.710*** | |
| | (0.009) | (0.026) | (0.021) | (0.010) | (0.016) | (0.011) | |
| Observations | $12,\!375$ | 2,746 | $5,\!283$ | $8,\!138$ | $5,\!812$ | $9,\!113$ | |
| R^2 | 0.689 | 0.761 | 0.559 | 0.711 | 0.666 | 0.705 | |
| Panel B: Ever Com | <i>aplete</i> Cours | e | | | | | |
| Shutout of Course | -0.380*** | -0.181*** | -0.276*** | -0.433*** | -0.285*** | -0.385*** | |
| | (0.011) | (0.024) | (0.024) | (0.013) | (0.020) | (0.012) | |
| Observations | $12,\!225$ | 2,700 | $5,\!283$ | $8,\!138$ | $5,\!812$ | $9,\!113$ | |
| R^2 | 0.442 | 0.681 | 0.261 | 0.299 | 0.414 | 0.490 | |
| Panel C: Ever Take | e Course in | Subject | | | | | |
| Shutout of Course | -0.277*** | -0.108*** | -0.216*** | -0.301*** | -0.160*** | -0.297*** | |
| | (0.011) | (0.023) | (0.022) | (0.012) | (0.017) | (0.012) | |
| Observations | $12,\!225$ | 2,700 | $5,\!283$ | $8,\!138$ | $5,\!812$ | $9,\!113$ | |
| R^2 | 0.365 | 0.592 | 0.206 | 0.220 | 0.324 | 0.410 | |

Table A.4: Heterogeneous Effects of Course Shutout by Course Characteristics

* p < 0.10, ** p < 0.05, *** p < 0.01. Each outcome in this table relates to completion of courses in the same subject area as a course a student was potentially shutout from. "Shutout" is an indicator for being shutout of a requested course. Each regression controls for shutout probability, which is generated from simulations of the algorithm that generated freshman course assignments in the Fall of 2018. Major in subject There are 10 colleges at Purdue including Agriculture, Education, Engineering, Health and Human Sciences, Liberal Arts, Management, Pharmacy, Polytechnic, Science, and Veterinary Medicine. Standard errors clustered at the individual level are reported in parentheses.

| Panel A: URM | | | | |
|---|---------------------------|---|---|-------------------------|
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | 4.524 (2.933) | $ \begin{array}{c} 0.043 \\ (0.048) \end{array} $ | $ \begin{array}{c} 0.034 \\ (0.037) \end{array} $ | -0.015 (0.029) |
| Observations R^2 | 888 0.395 | 816 0.219 | 888 0.197 | 825 0.489 |
| Non-Shutout Mean | 98.116 | 3.098 | 0.471 | 0.655 |
| Panel B: Non-URM | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | -0.396 (0.922) | -0.025 (0.015) | -0.024* (0.013) | -0.016^{*} (0.009) |
| Observations | 6,758 | 6,716 | 6,758 | 6,758 |
| R^2 | 0.063 110.068 | 0.137 | 0.071 0.621 | 0.521 0.698 |
| Non-Shutout Mean URM vs. Non-URM p-val | 0.100 | 3.270 0.199 | 0.139 | 0.098 |
| Panel C: First Gen | 0.200 | 0.200 | 0.000 | |
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | 2.459 | -0.064 | -0.008 | 0.004 |
| | (2.585) | (0.045) | (0.032) | (0.024) |
| Observations | 1,303 | 1,261 | 1,303 | 1,282 |
| R^2 | 0.161 | 0.122 | 0.125 | 0.496 |
| Non-Shutout Mean | 98.538 | 3.048 | 0.514 | 0.634 |
| Panel D: Non-First Gen | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | -0.593 (0.926) | -0.013 (0.015) | -0.020 (0.014) | -0.019** (0.009) |
| Observations | 6,343 | 6,271 | 6,343 | 6,301 |
| R^2 | 0.114 | 0.142 | 0.088 | 0.522 |
| Non-Shutout Mean | 110.785 | 3.294 | 0.623 | 0.706 |
| 1st Gen vs. non-1st Gen p-val | 0.264 | 0.286 | 0.717 | 0.958 |
| Panel E: Low SAT | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | -0.496 (1.340) | -0.035 (0.022) | -0.028 (0.018) | -0.038*** (0.014) |
| Observations | 3,761 | 3,658 | 3,761 | 3,698 |
| R^2 | 0.174 | 0.141 | 0.118 | 0.462 |
| Non-Shutout Mean | 102.982 | 3.125 | 0.561 | 0.535 |
| Panel F: High SAT | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Total Shutouts | 0.439 (1.148) | -0.000 (0.019) | -0.004 (0.018) | $0.004 \\ (0.010)$ |
| Observations | 3,885 | 3,874 | 3,885 | 3,885 |
| R^2 | 0.052 | 0.118 | 0.071 | 0.459 |
| Non-Shutout Mean | 114.411 | 3.376 | 0.648 | 0.849 |
| Female vs. Male p-val | 0.590 | 0.244 | 0.348 | 0.011 |

Table A.5: Effect of Shutouts on Cumulative Outcomes by Student Characteristics

* p < 0.10, ** p < 0.05, *** p < 0.01. Column (1) reports the effects of total first-term shutouts on the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of total first-term shutouts on cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of total first-term shutouts on graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| Panel A: STEM | | | | |
|---|---------------------------|--------------------------|---|-------------------------|
| Funct A. STEM | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| STEM Shutouts | -0.204 (1.430) | -0.060^{**} (0.024) | -0.035* (0.020) | -0.016 (0.014) |
| Observations R^2 | 7,646 0.130 | 7,532 0.147 | 7,646 0.090 | 7,583 0.514 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| Panel B: Non-STEM | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Non-STEM Shutouts | 0.216 (1.121) | 0.004 (0.019) | -0.007 (0.016) | -0.016 (0.011) |
| Observations P ² | 7,646 | 7,532 | 7,646 | 7,583 |
| R ² Non-Shutout Mean | 0.130 108.668 | 0.146 3.252 | 0.091 0.604 | 0.514 0.693 |
| STEM vs. Non-STEM p-val | 0.816 | 0.041 | 0.277 | 0.093 |
| Panel C: Required | a 1.5 a 5 | a 1 cr: | | 0777777777777 |
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Majo (4) |
| Required Shutouts | -0.339 | -0.017 | -0.028** | -0.018* |
| | (0.947) | (0.016) | (0.013) | (0.009) |
| Observations R^2 | 7,646 | 7,532 | 7,646 | 7,583 |
| Non-Shutout Mean | 0.130 108.668 | 0.146 3.252 | 0.091 0.604 | 0.514 0.693 |
| Panel D: Non-Required | | | | |
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Majo (4) |
| Non-Required Shutouts | 2.502 (2.430) | -0.019 (0.040) | $\begin{array}{c} 0.043 \\ (0.035) \end{array}$ | -0.004 (0.024) |
| Observations | 7,646 | 7,532 | 7,646 | 7,583 |
| R^2 | 0.130 | 0.147 | 0.090 | 0.514 |
| Non-Shutout Mean Required vs. Non-Required p-val | 108.668 0.250 | 3.252 0.952 | 0.604 0.053 | 0.693 0.573 |
| Panel E: Upper Courses | 0.250 | 0.332 | 0.055 | 0.010 |
| Tunci D. Opper Courses | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Majo (4) |
| Upper Shutouts | 4.710 (4.068) | -0.007 (0.068) | 0.063 (0.058) | -0.009 (0.040) |
| Observations | 7,646 | 7,532 | 7,646 | 7,583 |
| R ² Non-Shutout Mean | 0.130 108.668 | 0.146 3.252 | 0.090 0.604 | 0.514 0.693 |
| Panel F: Non-Upper Courses | 100000 | 0.202 | 0.001 | 0.000 |
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Majo (4) |
| Non-Upper Shutouts | -0.196 (0.899) | -0.021 (0.015) | -0.022* (0.013) | -0.016^{*} (0.009) |
| Observations | 7,646 | 7,532 | 7,646 | 7,583 |
| R^2 | 0.130 | 0.146 | 0.091 | 0.514 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| Upper vs. Non-Upper p-val | 0.253 | 0.834 | 0.116 | 0.814 |

Table A.6: Effects of Shutouts on Cumulative Outcomes by Course Characteristics

* p < 0.10, ** p < 0.05, *** p < 0.01. Panel A reports results for the effects of total first-term STEM shutouts on cumulative outcomes, and Panel B reports results for the effects of total first-term non-STEM shutouts on cumulative outcomes. Panel C reports results for the effects of total first-term required course shutouts on cumulative outcomes, and Panel D reports results for the effects of total first-term non-required course shutouts on cumulative outcomes. While Panel E reports results for the effects of total first-term upper-level (300 level or above) course shutouts on cumulative outcomes, Panel F reports results for the effects of total first-term lower-level (100 or 200 level) course shutouts on cumulative outcomes. Column (1) reports the effects of the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| Panel G: Difficult Courses | | | | |
|--|------------------------|-----------------------|--------------------------|-------------------|
| | Cumulative Credits (1) | Cumulative GPA (2) | 4-Year Graduation (3) | STEM Major (4) |
| Difficult Shutouts | 1.158 | -0.048 | -0.062** | 0.006 |
| | (1.895) | (0.032) | (0.027) | (0.019) |
| Observations | 7,646 | 7,532 | 7,646 | $7,\!583$ |
| R^2 | 0.130 | 0.147 | 0.090 | 0.514 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| Panel H: Non-Difficult Courses | | | | |
| | Cumulative Credits | Cumulative GPA | | STEM Major |
| | (1) | (2) | (3) | (4) |
| Non-Difficult Shutouts | -0.284 | -0.013 | -0.006 | -0.022** |
| | (0.994) | (0.016) | (0.014) | (0.010) |
| Observations | 7,646 | 7,532 | 7,646 | 7,583 |
| R^2 | 0.130 | 0.146 | 0.090 | 0.514 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| Difficult vs. Non-Difficult p-val | 0.126 | 0.478 | 0.146 | 0.339 |
| Panel I: Shutouts in Colleges with High Female Enrollm | ent | | | |
| | Cumulative Credits | Cumulative GPA | 4-Year Graduation | STEM Major |
| | (1) | (2) | (3) | (4) |
| Shutouts in High Female Enroll. Colleges | 0.037 | -0.012 | -0.021 | -0.015 |
| 0 | (0.941) | (0.016) | (0.013) | (0.009) |
| Observations | 7,646 | 7,532 | 7,646 | 7.583 |
| R^2 | 0.130 | 0.146 | 0.091 | 0.514 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| Panel J: Shutouts in Colleges with Moderate or Low Fer | nale Enrollment | | | |
| | Cumulative Credits | Cumulative GPA | 4-Year Graduation | STEM Major |
| | (1) | (2) | (3) | (4) |
| Shutouts in Med/low Female Enroll. Colleges | 0.039 | -0.086* | 0.004 | -0.025 |
| i construction of the second sec | (2.627) | (0.044) | (0.037) | (0.026) |
| Observations | 7,646 | 7,532 | 7,646 | 7,583 |
| R^2 | 0.130 | 0.147 | 0.090 | 0.515 |
| Non-Shutout Mean | 108.668 | 3.252 | 0.604 | 0.693 |
| High Female Enroll. vs. Med/low Female Enroll. p-val | 0.999 | 0.087 | 0.508 | 0.715 |

Table A.6: Effects of Shutouts on Cumulative Outcomes by Course Characteristics (Continued)

* p < 0.10, ** p < 0.05, *** p < 0.01. Panel G reports results for the effects of total first-term difficult course shutouts on cumulative outcomes, and Panel H reports results for the effects of total first-term non-difficult course shutouts on cumulative outcomes. While Panel I reports results for the effects of total first-term shutouts in college with high female enrollment on cumulative outcomes, Panel J reports results for the effects of total first-term shutouts in college with moderate or low female enrollment on cumulative outcomes. The colleges with high female enrollment are the College of Agriculture, the College of Education, and the College of Health and Human Sciences. The colleges with moderate or low female enrollment are the College of Engineering, the College of Liberal Arts, the College of Science, the College of Polytechnic Institute, and the College of Business. Column (1) reports the effects of the total credits earned between Fall 2018 and Fall 2022 semesters. Column (2) reports the effects of cumulative GPA between Fall 2018 and Fall 2022 semesters. Students who leave Purdue prior to earning any credits are omitted from this regression. Column (3) reports the effects of graduating from Purdue within 4-years (by the Spring 2022 semester). Column (4) reports the effects of choosing a STEM major. A student's major is defined as their primary graduating major if they have graduated or their most recent primary major if they have not graduated. STEM majors are defined by matching Purdue major CIP codes to the Department of Homeland Security's list of CIP codes that correspond to STEM majors. Each estimate includes controls for summed simulated shutout probabilities, demographic characteristics (sex, race/ethnicity, first-generation status, SAT scores, and pre-enrollment major), and number of reservation courses. Robust standard errors are reported in parentheses.

| Courses |
|----------------|
| Oversubscribed |
| Appendix II |

| ζ | Courses |
|---|-----------|
| - | g |
| _ | e O |
| • | OSCT1 |
| - | versut |
| (| \supset |
| د | ot |
| | List |
| ζ | Complete |
| ŗ | B.I: |
| Ē | Table |

| Accel First-Yr Compos | 696 | 488 | Environmental Scienc & Conserv | 215 | 60 | Intr Pub Pol-Pub Admin | 61 | 43 |
|------------------------------|------|--------|-------------------------------------|------|---------|--------------------------------|-----|-----|
| America In The $1960s$ | 14 | 11 | Essentials Of Nutr | 117 | 95 | Intr To Classical Myth | 75 | 39 |
| Analyte Geom & Calc II | 878 | 565 | Ethics And Animals | 23 | 6 | Intr To Materials Engr | 55 | 16 |
| Ancient Philosophy | ŭ | 4 | Exploring Teaching | 124 | 42 | Intro Acad Prog-Purdue | 124 | 45 |
| Anlytc Geomtry&Calc I | 2099 | 991 | First Year Biology Lab | 242 | 117 | Intro African American Studies | 47 | 19 |
| Applied Calculus I | 1806 | 107 | First-Year Composition | 3169 | 2355 | Intro Anthropology | 127 | 71 |
| Art Appreciation | 39 | 21 | Forensic Investigation | 168 | 128 | Intro Aviation Tech | 180 | 12 |
| Aviation Business | 167 | 101 | Foundation Officership | 33 | 33 | Intro Behvr Neurosci | 24 | × |
| Basic Drawing | 178 | 56 | Foundations Of Org Leadership | 420 | 369 | Intro Costume Des/Tech | က | 2 |
| Bible As Literature | 30 | 22 | French Level II | 81 | 40 | Intro Creative Writing | 14 | ъ |
| Biol Elem Sch Teach | 113 | 7 | French Level V | 10 | 2 | Intro Ed Tech | 153 | 48 |
| Biol Freshman Honr Sem | 09 | 4 | Freshman Seminar In EAS | 49 | 2 | Intro Entr & Innov | 107 | 56 |
| Biol I Divrs Ecol Behv | 341 | IJ | Fundament Of Speech | 3207 | 2103 | Intro Family Processes | 104 | 52 |
| Biology Resource Sem | 220 | ç | Fundamentals Biol I | 1368 | 379 | Intro Health Sciences | 232 | 7 |
| Biotechnology Lab I | 4 | 2 | Fundamentals Biol II | 66 | 11 | Intro Hosp & Tourism Industry | 70 | 2 |
| Black And White Photography | 65 | 52 | Fundamentals Of Music | 13 | 6 | Intro Ling St For Lang | 5 | 1 |
| Bowling | 16 | 15 | Gen Physics | 157 | 120 | Intro Polit Analysis | 16 | 14 |
| Ceramics I | 30 | 29 | Gen Social Psychology | 30 | 27 | Intro Probability Mdls | 21 | 18 |
| Child Psychology | 71 | 28 | General Chemistry | 3956 | 680 | Intro Res Meth In Psy | ° | 1 |
| Chinese Calligraphy | 2 | 1 | General Physics | 169 | 110 | Intro Sci Fields Psy | 122 | 15 |
| Chinese Level I | 20 | 2 | Geosciences Cinema | 78 | 77 | Intro Social Psych | 76 | 40 |
| Chinese Level III | °, | 1 | German Level I | 49 | 20 | Intro To ANSC Programs | 189 | 24 |
| Classical Wrld Civiliz | 38 | 16 | German Level II | 34 | 17 | Intro To Animal Agr | 157 | 10 |
| Collab Leader: Interpersonal | 854 | 684 | German Level V | 12 | 2 | Intro To C Program | 123 | 15 |
| Crit Perspectives Com | 22 | ъ | Global Awareness | 17 | e S | Intro To Com Theory | 43 | 27 |
| Crit Think & Com I | 1535 | 966 | Global Green Politics | 17 | 11 | Intro To Computers | 148 | 95 |
| Crit Think & Com II | 995 | 618 | Global History | 85 | 58 | Intro To Energy Engr | 85 | 28 |
| Crop Production | 43 | x | Global Moral Issues | 117 | 103 | Intro To Int Design | 85 | 29 |
| Customer Relation Mgmt | 42 | 15 | Graphic Communication | 19 | 2 | Intro To Linguistics | 25 | 21 |
| Dance Appreciation | 33 | 20 | Graphics For CE & Const | 83 | ŋ | Intro To Medieval Wrld | 95 | 49 |
| Dscr Astr-Solar System | 154 | 92 | Great American Books | 40 | 35 | Intro To PRMD Programs | 63 | 30 |
| E&M Interactions | 20 | 9 | Honors Multivariate Calculus | 63 | 13 | Intro To Personal Finance | 101 | 56 |
| Earth Through Time | 69 | 50 | Human Antmy & Physiol | 23 | 18 | Intro To Philosophy | 244 | 150 |
| Earthquakes Volcanoes | 103 | 75 | Human Factors Flight Crew | 10 | 9 | Intro To Plant Science | 267 | 139 |
| East Asia & Hist Trad | 28 | 17 | Inclusive Classroom | 49 | 24 | Intro To Retail Mgmt | 26 | 4 |
| Economics | 338 | 142 | Info Tech Architecture | 219 | 157 | Intro To Statistics | 35 | 34 |
| Elem Stat Meth | 164 | 102 | Interdisc Approach To Writing | 83 | 59 | Intro To Tourism Mgmt | 53 | 12 |
| Elementary Psychology | 1838 | 727 | Interpreting America | 31 | 17 | Introduction To Earth Sciences | 30 | 9 |
| Elements Linguistics | 11 | 19 | Intr Actuarial Science | 140 | 45 | Introduction To Management | 332 | 11 |
| | ŗ | ц С | $I_{mkn} = P_{minimum on kn} [P_n]$ | 10 | 66 | T, T T T T | | |

| | anonhour parnoo | | | - | | | | . 1 |
|------------------------------|-----------------|------|--------------------------------|-----|-----|------------------------------|-----|--------|
| Italian Renaissance Impact | 20 | 14 | Prin Of Persuasion | 45 | 19 | The Nuclear Age | 31 | 21 |
| Lab Bio III Cell Strct | 17 | 4 | Private Pilot Lectures | 64 | 11 | The Planets | 187 | 153 |
| Learning & Motivation | 26 | 41 | Prog Appl For Enginrs | 773 | 499 | Theatre Appreciation | 100 | 40 |
| Lodging Management | 32 | 23 | Program With MM Objs | 71 | 15 | Tools | 411 | 6 |
| Macroeconomics | 647 | 477 | Quantitative Reasoning | 290 | 142 | Trans Ideas To Innovation II | 51 | 3 S |
| Magic And Marvels | 24 | 11 | Rac & Ethn Diversity | 63 | 42 | U S Since 1877 | 245 | 134 |
| Materials & Processes I | 184 | 92 | Religions Of The West | 6 | 9 | Visual Programming | 13 | 6 |
| Materials And Processes II | 188 | 89 | Sci & Society In Western Civ I | 13 | 8 | Wildlife In America | 37 | 17 |
| Math For Elm Tchrs I | 110 | 4 | Science Of Food | 54 | 27 | Wind Ensemble I | 14 | 2 |
| Media For Children | 56 | 34 | Science Writing & Presentation | 121 | 101 | Wom Pol And Publ Pol | 27 | 20 |
| Microecon Food & Agbus | 200 | 63 | Screenwriting | 22 | 5 | Women Tech: Expl Possib | 49 | ŝ |
| Microeconomics | 968 | 391 | Second World War | 31 | 27 | World's Forest&Society | 17 | 6 |
| Modern Dance Technique I | 19 | 12 | Social Problems | 321 | 199 | | | |
| Money, Trade & Power | 10 | × | Society & Rock & Roll | 12 | 9 | | | |
| Multicult Leadrshp Sem | 13 | 2 | Spanish Level I | 63 | 26 | | | |
| Multivariate Calculus | 804 | 26 | Spanish Level III | 297 | 16 | | | |
| Music Theory I | 25 | 24 | Spanish Level V | 84 | 24 | | | |
| Native American Cultures | 19 | 10 | Sports & Literature | 18 | 11 | | | |
| Navigating Gender | 51 | 23 | Spreadsht Use Agr Bus | 14 | 2 | | | |
| New Media Culture | 13 | 9 | Statistics & Society | 431 | 196 | | | |
| Org & Mgt Hosp&Tour Ind | 6 | ъ | Std Arabic Level I | 15 | 2 | | | |
| Phil & The Meaning Of Life | 11 | 4 | Strat Success First Yr | 24 | 1 | | | |
| Philos Of Religion | 12 | 9 | Study Skills Seminar | 106 | 27 | | | |
| Philosophy And Law | 71 | 41 | Survey Of Acting | 27 | 41 | | | |
| Physical Geology | 68 | 48 | Symphonic Band | 22 | 6 | | | |
| Pl Anly Geo Calc I | 2007 | 1075 | Technical Graphic Comm | 65 | 44 | | | |
| Pl Anly Geo Calc II | 750 | 335 | Technology And Culture | 76 | 33 | | | |
| Planet Earth | 216 | 146 | Th Minorities In Mgmt | 28 | - | | | |
| Pre-Doctor Of Pharm Orient I | 236 | 2 | The Data Mine I | 84 | 16 | | | |
| Prin Of Economics | 333 | 167 | The Italian Cinema | 23 | 16 | | | |

Appendix III Algorithm and Simulation Exercises

After students submit their preferences, they are assigned a course schedule through Purdue's batch algorithm system (Müller *et al.*, 2010). The algorithm uses each student's preference ranking of courses and preferred class times as inputs in generating schedules for all students.

The objective of the algorithm is to maximize the use of priority course requests while minimizing the use of alternative course requests provided by students (Müller *et al.*, 2010). The algorithm assigned a single weight to each course request based on the following equation:

$$weight(a \in dom(R)) = 0.9^{prior(R)} \times 0.5^{alt(a)},$$
(5)

where dom(R) is the domain of the course request R and prior(R) is the ranking of priority requested course, and alt(a) is the ordering of the alternate requested courses. In our simulations, we assigned different weights based on the main conditions in Equation (5). Whether students were enrolled or not in those courses based on the available spots for each course.

To use the course request template from Figure C.1 as an example, the algorithm assigns $0.9^1 \times 0.5^0 = 0.9$ as the weight in *CNIT18000* while the weight of *ENGL11000* is $0.9^2 \times 0.5^0 = 0.81$. For the third preference, the algorithm assigns $0.9^3 \times 0.5^0 = 0.729$ in *MA16010* while giving the weights of $0.9^3 \times 0.5^1 = 0.3645$ in *PHYS22000* and $0.9^3 \times 0.5^2 = 0.183$ in *CHM11100*. Based on different weights from each course request provided by students, the algorithm solves the problem by implementing Iterative Forward Search (Müller *et al.*, 2004). Equation (5) implies that the algorithm is more costly to reject a course request with only priority listing and without alternative listings than a course request with both priority and alternative listings.

There are four constraints in which the algorithm has to follow while assigning course assignments to students:

- 1. Seats limit
 - Each course section has a seat limit for students to enroll. There is an unlimited number of enrollments in some course sections, such as distance learning sections.
- 2. Overlapping sections
 - Two or more course sections overlapped with each other are not allowed. The algorithm only grants one or none of those courses to students.
- 3. Distance student conflict
 - A distance conflict occurs when locations of two sections are too far apart with little time (10 minutes or less) for students to arrive the later section on time. No distance conflicts are identified if there are more than 20 minutes gap between two sections.
- 4. Course reservations by colleges or departments
 - Colleges or departments reserve spots in certain courses for students who have a declared majors.

The batch registration process has the following steps:

- (i) The algorithm starts by ordering priority students in order of constraints (those with fewest potential of sections for their courses first) and uses a branch and bound technique to evaluate best possible assignments, as defined by their course prioritization
- (ii) Students who are not enrolled in their desired number of courses from step 1 are ordered randomly. In this order, the algorithm looks for assignments that do not conflict with existing assignments or schedule constraints.
- (iii) The algorithm improves the overall schedules by using backtracking technique.
- (iv) The algorithm repeats steps i iii and stores the results.
- (v) After a pre-specified run-time (12 hours in the case of Fall 2018 assignments), the batch assignment with the highest weight score is selected

We run 1,000 simulations of the algorithm used in the Fall of 2018 using Java SE 11 to estimate the shutout probability for each assignment request. Interested readers can learn more from (Müller *et al.*, 2010).

Figure C.1: Course Request Form

| | | | | Stude | nt Cours | e Requests |
|----------------|-------------|-------------------------|-----------------|--------------|--------------|-----------------|
| Student's Nam | ie: | | | PUID: | | |
| Advisor/Email: | | | | PIN #: | | |
| Course Req | uests | | | Term: | | |
| 1. Priority | CNIT18000 - | enrolled | | | | |
| 1. Alter | native | | | | | |
| 2. Alter | native | | | | | |
| 2. Priority | ENGL11000 | - enrolled | | | | |
| 1. Alter | native | | | | | |
| 2. Alter | native | | | | | |
| 3. Priority | MA16010 - e | nrolled | | | | Upper Block |
| 1. Alter | native | PHYS22000 | | | | (Primary = Yes) |
| 2. Alter | native | CHM11100 | | | | |
| 4. Priority | TECH12000 | R - enrolled | | | | |
| 1. Alter | native | CNIT15501 | | | | |
| 2. Alter | native [| | | | | |
| 5. Priority | TLI11200 | | | | | |
| 1. Alter | native | AGEC21700 - enrolled | | | | |
| 2. Alter | native [| AD38300 | | | | |
| 6. Priority | | | | | | |
| 1. Alter | native | | | | | |
| 2. Alter | native | | | | | |
| 7. Priority | - | | | | | |
| 1. Alter | native | | | | | |
| 2. Alter | native | | | | | |
| 8. Priority | | | | | | |
| 1. Alter | native | | | | | |
| 2. Alter | native | | | | | |
| 9. Priority | | | | | | |
| Alternate 0 | Course R | equests (used only if a | course requeste | d above is r | not availabl | e) |
| 1. Priority | ANTH100 | 00 | | | | Lower Block |
| | MUS2500 | | | | | (Primary = No) |
| Student's Sign | | | | Date | | |