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Working Paper

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Working Paper, No. 238/2024

Provided in Cooperation with:

Berlin Institute for International Political Economy (IPE), Berlin School of Economics and Law

Suggested Citation: Farroukh, Arafet; Mazioued, Manel; Pédussel Wu, Jennifer (2024): Revisiting the linkage between remittances inflow and economic growth: A semi-parametric estimation with panel data, Working Paper, No. 238/2024, Hochschule für Wirtschaft und Recht Berlin, Institute for International Political Economy (IPE), Berlin

This Version is available at: https://hdl.handle.net/10419/301859

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Institute for International Political Economy Berlin

Revisiting the linkage between remittances inflow and economic growth: A semi-parametric estimation with panel data

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Working Paper, No. 238/2024

Editors:

Sigrid Betzelt, Eckhard Hein, Martina Metzger, Martina Sproll, Christina Teipen, Markus Wissen, Jennifer Pédussel Wu (lead editor), Reingard Zimmer

Revisiting the linkage between remittances inflow and economic growth:

A semi-parametric estimation with panel data

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Abstract:

We investigate the relationship between remittances inflow and economic growth in a sample of 65 emerging countries over the period 1988-2018 using the semi-parametric panel data model with fixed effects as proposed by Baltagi and Li (2002). Our empirical results show that the effect of remittances inflow on economic growth exhibits a highly nonlinear pattern, which sheds new light on the remittances-growth nexus and provides evidence of a nonlinear relationship.

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Keywords: Remittances; Growth; Nonlinear effects; Semi-parametric regression; Panel data; Emerging countries.

JEL Codes: C14; O15; F43

1. Introduction

Remittances result from migrants sending home a portion of their earnings to support residences in their home countries. Migrant remittances have grown rapidly in the past few years and for many developing countries, represent the largest source of foreign income. While it is difficult to estimate remittance flows since a large portion travel via unofficial methods, remittances are shown to be especially important for low-income countries (Ratha, 2017).

Historically, remittances are considered a very stable financial flow. During the financial crisis in 2009, remittances experienced a downturn of 5 percent, recovering within a year (World Bank Group, 2018), a situation unlikely to result with the COVID-19 pandemic. International remittances are overall the second largest international monetary flow after foreign direct investment (FDI) and it is considered the least volatile and most reliable of all financial flows to developing countries. Economic remittances have therefore become an influential tool for development and economic growth. While remittances can contribute to Dutch Disease if they cause an appreciation in the exchange rate (Fritz et al., 2017), a higher exchange rate makes exports more expensive and imports cheaper, which can also lead to a decline in competitiveness in other sectors of the economy.

The aim of this study is to revisit the linkage between remittances inflow and economic growth by employing the semi-parametric panel fixed effects regression developed by Baltagi and Li (2002). With this approach, a prior parametric shape unassumed in modeling the relationship between remittances and growth. The semi-parametric panel fixed effects regression is supposed to be more flexible than the parametric approach (Wang et al., 2015; Hamadi and Heinen, 2015) and leads to estimating the true shape of the remittances-growth nexus. We supplement our analysis by employing parametric techniques using the generalized method of moments (GMM) estimators.

In the context of the remittances-growth nexus, the term "nonlinear nature" refers to the fact that the relationship between remittances inflow and economic growth rate is not a straight line or a constant rate of change. Instead, it suggests that the impact of remittances on economic growth may vary depending on the level of remittances inflow. In other words, a small increase in remittance inflows may have a different impact on economic growth than a large increase in remittance inflows. The nonlinear

nature of the relationship suggests that there may be a threshold level of remittance inflows beyond which the positive impact on economic growth starts to decrease or even become negative.¹

Understanding the nonlinear nature of the remittances-growth nexus is important because it can help policymakers design more effective policies to maximize the positive impact of remittances on economic growth. For example, policymakers may need to consider different policy interventions depending on the level of remittance inflows to ensure that remittances are channeled towards productive investments or sectors that can contribute to economic growth. Our findings, based on data from 68 countries 1988-2018, provide evidence that the effect of remittances inflow on economic growth exhibits a highly nonlinear pattern and thus our results are inconsistent with the findings provided by previous studies that have assumed a linear parametric functional form or those using a conventional polynomial specification.

The rest of the paper is organized as follows. Section 2 presents the literature review, section 3 presents our econometric models, the data and briefly describes the semi-parametric panel model. Section 4 discusses the empirical results and Section 5 concludes the paper.

2. Literature review

The impact of remittances inflow on economic growth has been extensively examined in recent years although the results, are mixed and inconclusive thus leaving the debate open-ended. A first strand of empirical studies seems to support the evidence of a positive relationship between remittances and economic growth (e.g., Imai et al., 2014; Kumar and Stauvermann, 2014; Nsiah and Fayissa, 2013; Jayaraman et al., 2012; Javid et al., 2012; Nyamongo et al., 2012, Jayaraman et al., 2011; Azam and Khan, 2011; Giuliano and Ruiz-Arranz, 2009; Mundaca, 2009; Pradhan et al., 2008), while other studies suggest a negative or an insignificant relationship (e.g., Feeny et al., 2014; Rao and Hassan, 2011; Barajas et al., 2009; Chami et al., 2003). Another strand of this literature suggested a different view in this debate, and has claimed, by using a threshold approach, that the relationship between remittances inflow and the growth rate is nonlinear (e.g., Eggoh et al., 2019; Bettin and Zazzaro, 2009; Hassan et al., 2012; Giuliano and Ruiz-Arranz, 2009).

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¹ We do not examine the threshold estimation in the context of this paper and recognize that it should be examined in future work.

Table 1 reports the summary of recent literature on the relationship of remittances and growth. Much of this literature examines the importance of financial development in this relationship further contributing to the endogeneity of the question.

Table 1 Tabular Summary of the Literature Review				
Author/Year	Sample	Variables Employed	Methodology	Outcomes
Chami et al. (2003)	113 countries, 1970 to 1998	Remittance /GDP, M2/GDP and GDP per capita	Panel OLS	Negative effects of remittances on growth
Faini (2005)	Europe and USA	Migration figure, Remittances/ GDP and Poverty	Questionnaire	Positive effects of remittances on growth.
Giuliano and Ruiz-Arranz (2005)	100 countries, 1975 to 2002	Remittances/GDP, M2/GDP, private credit to GDP and GDP per capita	Panel OLS	Positive effects of remittances on growth.
Richard et al. (2005)	71 Developing countries 1987 to1996	Gini Coefficient, Per capita GDP and International Migration figure.	Pooled OLS and IV Estimate	International Migration and remittances reduce poverty
Guilano and Ruiz-Arranz (2009)	100 Developing countries	Remittances/GDP Private credit to GDP and Investment	IV Estimate	Positive effects of remittances on Growth.
Karagöz (2009)	Turkey, from 1970 to 2005	Remittances/GDP capital formation/ GDP, FDI, net private flow and FDI.	Cointegration and OLS	Negative effects of remittances on growth.
Siddique et al. (2010)	1976 to 2006 Bangladesh, India and Sri Lanka	Per capita remittances, Export and GDP per capita	VAR Granger Causality test	Two-way causality test existed in Sri Lanka.
Khan and Azam (2011)	1995 to 2010 Azerbaijan and Armenia	GDP and remittance/GDP	Simple Log Linear regression	Positive effects of remittances on growth
Jayaraman, Choong and Kumar (2011)	Samoa and Tonga, 1981 to 2008	Remittance to GDP, M2/GDP, private sector credit to GDP and FDI.	Bound Test	Positive effects of remittances on growth.
Jayaraman, Choong and Kumar (2012)	1970 to 2009, Pacific Island	Remittance/GDP per worker, FDI and private sector credit to GDP	Bound Test	Positive effects of remittances on growth.
Javid et al. (2012)	1973 to 2010 Pakistan	Real GDP, remittance/GDP, investment to GDP, income inequality and trade openess	ARDL	Positive effects of remittances on growth.

Kumar (2012)	Sub-Saharan Africa, 1970 to 2010.	Remittance/GDP, private sector credit to GDP, telephone lines per 100 people, and ODA to GDP.	ARDL Bound Test.	Negative effects of remittances, Financial Development and ICT on growth
Nyamongo et al. (2012)	Africa, 1980 to 2009.	GDP per capita, remittances and financial development.	Pooled OLS fixed effect.	Remittances promote economic growth.
Cooray (2012)	94 Non-OECD, 1990 to 2010	Remittances, financial sector size and government bank ownership.	Pooled OLS and System GMM	Remittances promote financial sector development.
Damnola and Wakili (2013)	Nigeria, 1980 to 2010.	Real GDP per capita, remittance/GDP and M2/GDP	Seemingly unrelated regression (SUR) and ECM	Remittances promotes economic growth.
Akkoyunlu (2013)	Turkey, 1963 to 2009.	Remittances/GDP, bank deposit, private credit/GDP	Toda Yamamoto Non-Granger Causality test.	No causality between financial development and remittances
Koay and Choong (2013)	Malaysia	GDP remittance/GDP and M2/GDP.	ARDL and Granger Causality test.	Remittances promotes growth and causality runs from finance to remittances
Alkhathlan (2013)	Saudi Arabia, 1970 to 2010.	Real GDP, Remittance/GDP and Export.	ARDL	Negative effects of remittances on growth.
Meyer and Shera (2013)	Albania, 1992 to 2012.	GDP, Remittance/ GDP and M2/GDP.	ARDL and ECM	Positive effects of remittances on growth
Goschin (2013)	Romania, 1994 to 2011.	GDP, Remittance/ GDP and M2/GDP.	Multi-factorial regression models	Positive effects of remittances on growth
Kumar (2013)	Guyana, 1982 to 2010	RealGDP, remittances/GDP, ODA/GDP, M2/GDP	ARDL bound Test.	Positive and long run effect of remittances on growth
Kumar and Stauvermann (2014)	Kenya, 1978 to 2010.	Tourism, remittances/GDP and real GDP	Bound Test approach	Positive effects of remittances on growth.
Kumar and Stauvermann (2014)	Lithuania, 1980 to 2012	Real GDP, remittances/GDP and M2/GDP	ARDL	Positive effects of remittances on growth.
Kumar and Stauvermann (2014)	Bangladesh, 1979 to 2012	Real GDP, remittances/GDP and M2/GDP	ARDL Bound procedure	Positive effects of remittances on growth
Imai et al. (2014)	24 Asia and pacific countries, 1980 to 2009.	Real GDP per capita, remittance/GDP, M2/GDP, Volatility of remittance and FDI.	Panel 2 stage Least Square.	Positive effects of remittances on growth
Sarkar and Datta (2014)	Bangladesh, 1975 to 2011	GDP, remittances/GDP, M2/GDP and other control variables	ARDL and Causality	No relationship exists between remittances and GDP
Blouchoutzi and Nikas (2014)	Moldova and Albania, 1990 to 2010.	Remittance/GDP, consumption, GDP, and import.	OLS	Remittances promote growth in both countries.

F (2014)	126	CDD't	CMME	Desirit and Construct
Feeny et al. (2014)	136 states including 25 SIDS,	GDP per capita, remittance/	GMM Estimates	Positive effects of remittances for
	1971 to 2010.	GDP and other control		SIDS countries.
	1571 to 2010.	variables		SIDS Countries.
Sibindi (2014)	Lesotho, 1975 to	Remittances/GDP,	VECM and	Causality running
	2010.	M2/GDP and real	Granger	from remittances
		GDP.	causality.	to growth.
Adarkwa (2015)	Cameroon, Cape	GDP per capita,	OLS	Positive effects of
	Verde, Nigeria and	remittance inflow and		remittances on
	Senegal, 2000 to 2010.	outflow		growth in Nigeria
Coulibaly (2015)	19 sub-Saharan	GDP per capita,	Panel Granger	and Senegal only. No evidence of
Country (2013)	African countries.	remittances, liquidity	causality test	causal relationship
	1980 to 2010.	liability and FDI	causarry test	in SSA
Sobiech (2015)	54 Developing	Remittances/GDP,	GMM Panel	Negative effects of
	countries, 1970 to	M2/GDP, Private	Analysis.	remittances on
	2010.	credit/GDP and GDP		growth
		per capita		
Karikari et al.	50 developing	Remittances/GDP,	Panel VECM	Remittances
(2016)	countries, 1990 to 2011	M2/GDP, Private credit/GDP, FDI and	causality	promotes financial
	2011	GDP per capita	techniques.	development.
Nsiah and Fayissa	1985- 2007 for 64	Remittances, openness	Panel Fully	Positive
(2013)	countries (29 from	of the economy,	Modified OLS	relationship
, ,	Africa, 14 from	capital/labor ratio,	(PFMOLS)	between
	Asia, 21 from	economic freedom		remittances and
	Latin America)			economic growth.
Mundaca (2009)	Latin American	GDP growth per capita,	GMM estimation	Positive
	and Caribbean	investment p.c.,		relationship
	Countries 1970- 2002	remittances/GDP, domestic credit/GDP,		between remittances and
	2002	FDI/GDP		economic growth.
Pradhan et al.	39 developing	Export-GDP-ratio	Standard growth	Positive
(2008)	countries (1980 –	(trade openness), GDP _t .	model with	relationship
	2004)	1, investment, polity	fixed-effects and	between
		(political regime and	random-effects	remittances and
		authority		economic growth.
Earth et al. 2010	40 41	characteristics)	Dan al Cara a stla	Nonlinear
Eggoh et al., 2019	49 developing countries, 2001 -	GDP p.c. (log), remittances/GDP,	Panel Smooth Transition	relationship
	2013	economic growth	Regression	between
	2013	factors,	(PSTR)	remittances inflow
		,		and economic
				growth rate.
Bettin and	66 developing	Remittances/GDP,	Growth	Nonlinear
Zazzaro, 2009	countries 1970 -	financial development,	regression	relationship
	2005	interaction term,		between remittances inflow
		financial control variables		and economic
		variables		growth rate.
Hassan et al., 2012	Bangladesh 1974 -	Remittances/GDP,	OLS,	Nonlinear
, ,	2006	gross capital	Instrumental	relationship
		formation/GDP,	Variable-Two	between
		populations growth	Stage Least	remittances inflow
		government	Square (IV-	and economic
		consumption/GDP,	2SLS), and IV-	growth rate.
		M2/GDP, inflation rate	GMM estimators	
		(all log)		

Source: Author compilation.

Remittances do not automatically lead to higher long-term growth rates for several reasons. First, endogeneity is an issue and, second, while remittances have important spillover effects, other contributions to growth are equally important. The existing studies in this literature suffer from important limitations. Most of the literature imposes a specific functional form (i.e., a linear relationship) between remittances and economic growth, while other studies, which have supported the existence of nonlinear effect, may suffer from important endogeneity issues related to the choice of the threshold variable. The threshold in the context of the remittances-growth nexus refers to the point at which the positive impact of remittance inflows on economic growth starts to decrease or become negative. In other words, there is a level of remittance inflows beyond which additional remittances may not have a significant, or a positive impact, on economic growth. The exact threshold level may vary depending on a variety of factors such as the economic conditions of the recipient country, the source and type of remittances, and the policy environment. Therefore, it is difficult to determine a universal threshold level for all countries or contexts.

3. Econometric models, data, and estimation strategy

As a starting exercise and following Mundaca (2009), Bettin and Zazzaro (2012), Fenny et al. (2014) and Giuliano and Ruiz-Arranz (2009), we estimate the following econometric models. The included control variables are standard in the literature and we follow previous protocol. To test the relationship between remittances inflow and economic growth based on prior studies, our baseline parametric model (in a dynamic form) can be illustrated by the following equation:

$$\begin{split} GDPG_{i,t} &= \beta_0 + \beta_1 GDPG_{i,t-1} + \beta_2 REMIT_{i,t} + \beta_3 TRADE_{i,t} + \beta_4 INVEST_{i,t} + \\ & \beta_5 \ EXPEND_{i,t} + \beta_6 INFL_{i,t} + \beta_7 POPG_{i,t} + \mu_t + \gamma_i + \\ & \varepsilon_{i,t} \end{split} \tag{1}$$

where GDPG is the annual growth rate of real GDP per capita for country i at time t, REMIT refers to remittances received as a percentage of GDP, TRADE is international trade, defined as the ratio of the sum of exports plus imports of goods to total output, INVEST is the gross fixed capital formation as a percent of real GDP used as a proxy for investment in physical capita, EXPEND is Government spending, measured as the ratio of government consumption to GDP, INFL is inflation measured as the annual percentage change in the consumption price index, POPG measures the average annual

population growth rate, μ_t is a time specific effect, γ_i is an unobserved country-specific fixed effect and $\varepsilon_{i,t}$ is the error term.

As for previous studies, to check if the relationship between remittances and growth is nonlinear, we modify to parametric specification in Equation (1) into different polynomial forms (i.e, quadratic and cubed) on remittances inflow variable. The resulting parametric nonlinear models where the variable is nonlinear and we modify the parametric specification from Equation 1, are given as follows:

$$GDPG_{i,t} = \beta_0 + \beta_1 GDPG_{i,t-1} + \beta_2 REMIT_{i,t} + \beta_3 REMIT_{i,t}^2 + \beta_4 TRADE_{i,t}$$

$$+ \beta_5 INVEST_{i,t} + \beta_6 EXPEND_{i,t} + \beta_7 INFL_{i,t} + \beta_8 POPG_{i,t} + \mu_t$$

$$+ \gamma_i \varepsilon_{i,t}$$
 (2)

and

$$GDPG_{i,t} = \beta_0 + \beta_1 GDPG_{i,t-1} + \beta_2 REMIT_{i,t}$$

$$+ \beta_3 REMIT_{i,t}^2 + \beta_4 REMIT_{i,t}^3 + \beta_5 TRADE_{i,t} + \beta_6 INVEST_{i,t}$$

$$+ \beta_7 EXPEND_{i,t} + \beta_8 INFL_{i,t} + \beta_9 POPG_{i,t} + \mu_t + \gamma_i$$

$$+ \varepsilon_{i,t}$$
(3)

As given in Equations (1), (2) and (3), we use the system generalized method of moments (SYS-GMM) of Arellano and Bond (1991) and Arellano and Bover (1995). This methodology is more valuable than OLS for dealing with endogeneity problems that might occur because of omitted variable bias, reverse causality, as well as measurement error.

In order to estimate the aforementioned parametric models, we rely on a panel data set, which consists of cross-country observations during the period 1988-2018 for 68 countries². The countries included in this study and the sample period are due to data

Philippines, Romania, Rwanda, Senegal, Serbia, South Africa, Sri Lanka, Sudan, Syrian Arab, Tanzania, Thailand, Togo, Tunisia, Turkey, Uganda, Ukraine, Vietnam, Yemen Rep., Zambia, Zimbabwe.

² The countries selected for this study are: Algeria, Armenia, Azerbaijan, Bangladesh, Belarus, Benin, Bolivia, Botswana, Brazil, Bulgaria, Burkina Faso, Cameroon, Central African Republic, Colombia, Congo, Rep., Costa Rica, Cote d'Ivoire, Dominican Republic, Egypt Arab Rep., El Salvador, Eswatini (Swaziland), Fiji, Gabon, Gambia, Ghana, Guatemala, India, Indonesia, Iran, Jamaica, Jordan, Kazakhstan, Kenya, Lebanon, Madagascar, Malaysia, Mali, Mauritania, Mauritius, Mexico, Morocco, Mozambique, Niger, Nigeria, Pakistan, Panama, Paraguay, Peru,

availability. All the variables, as defined above, have been extracted from World Bank³ World Development Indicators (WDI) and International Monetary Fund (IMF) datasets. The descriptive statistics of our variables of interest are given in Table 2.

Table n°2: Summary statistics

Variables	Mean	Stand. Dev.	Minimum	Maximum
Real GDP Growth (GDPG)	0.042	0.032	-0.046	0.114
Remittances (REMIT)	0.027	0.029	0.000	0.130
TRADE	0.623	0.271	0.190	1.287
Investment (INVEST)	0.206	0.054	0.101	0.340
Expenditures (EXPEND)	0.137	0.037	0.068	0.232
Inflation (INFL)	0.094	0.088	-0.002	0.527
Population Growth (POPG)	0.019	0.019	0.000	0.071

Source: Author compilation using IMF (2022) and World Bank (2021).

As indicated in the introduction, many of the previous studies have assumed a specific functional form (i.e., linear) to describe the relationship between remittances inflow and economic growth. However, as argued by Yatchew (1998), the economic theories never imply a specific functional form when relating a response variable to a set of exogenous covariates.

Therefore, to find out the "true" functional form of the remittances-growth nexus and unlike pre-existing studies, we employ the semi-parametric panel model with fixed effects as developed by Baltagi and Li (2002). This approach allows the effect of remittances to be modeled by an unrestricted and more flexible functional form. Formally, the semi-parametric panel model with fixed effects can be described by the following equation:

$$GDPG_{i,t} = \alpha_i + \gamma \mathbf{X}_{i,t}^T + f(REMIT_{i,t}) + \varepsilon_{i,t}, \quad i = 1, \dots, N; t = 1, \dots, T$$

$$\tag{4}$$

Where f(.) is an unspecified function of $REMIT_{i,t}$, α_i refers to the country-specific fixed effects which are correlated with $REMIT_{i,t}$, $\boldsymbol{X}_{i,t}^T$ is a k-dimensional vector of control variables (TRADE, INVEST, EXPEND, INFL, POPG), , and $\varepsilon_{i,t}$ are error terms

³ The World Bank provides data on remittances through its Migration and Remittances Data portal. This database includes information on remittance flows, the cost of remittances, and migration trends.

assumed to be *i.i.d* with finite variance σ_{ε}^2 and mean-independent of $REMIT_{i,t}$, namely, $E(\varepsilon_{i,t}|REMIT_{i,t}) = 0$.

The non-parametric component, f(REMIT), is estimated using splines with optimal basis functions, as discussed by Keele (2008). To estimate the model given in equation (4) with panel data, the common procedure used is to eliminate the fixed effects, α_i , by differentiating equation (4) over time which gives the following expression:

$$\Delta(GDPG_{i,t}) = \Delta(\mathbf{X}_{i,t})\boldsymbol{\gamma} + f[\Delta(REMIT_{i,t})] + \Delta(\varepsilon_{i,t})$$

(5)

Or

$$GDPG_{it} = X_{it}\gamma + F(REMIT_{i,t}, REMIT_{i,t-1}) + U_{i,t}$$
(6)

where
$$\mathbf{GDPG}_{it} = GDPG_{i,t} - GDP_{i,t-1}, \mathbf{X}_{it} = \mathbf{X}_{i,t} - \mathbf{X}_{i,t-1}, \ \mathbf{F}(REMIT_{i,t}, REMIT_{i,t-1})$$

= $f(REMIT_{i,t}) - f(REMIT_{i,t-1})$ and $\mathbf{U}_{i,t} = \varepsilon_{i,t} - \varepsilon_{i,t-1}$.

The main problem here is to estimate consistently the unknown function of remittances inflow variable, REMIT $\equiv \mathbf{F}(REMIT_{i,t}, REMIT_{i,t-1}) = [f(REMIT_{i,t}) - f(REMIT_{i,t-1})].$

Following Baltagi and Li (2002), we approximate f(.) by series differences $p^k(.)$ and therefore approximate $\mathbf{F}(REMIT_{i,t}, REMIT_{i,t-1}) = [f(REMIT_{i,t}) - f(REMIT_{i,t-1})]$ by $\mathbf{p}^k(REMIT_{i,t}, REMIT_{i,t-1}) = [p^k(REMIT_{i,t}) - p^k(REMIT_{i,t-1})]$ where $p^k(.)$ are the first k terms of a sequence of functions $(p_1(REMITD_{i,t}), p_2(REMIT_{i,t}), ...)$. Hence equation (5) becomes:

$$GDPG_{i,t} - GDPG_{i,t-1} = (X_{i,t} - X_{i,t-1})\gamma + [p^k(REMIT_{i,t}) - p^k(REMIT_{i,t-1})]\beta + (\varepsilon_{i,t} - \varepsilon_{i,t-1})$$
(7)

Typically, the $p^k(REMIT_{i,t})$ series is a spline, which is a segmented polynomial with components characterized by a sequence of smooth knots. A simple method is to solve using a linear spline of degree m with k knots, which can be represented by:

$$S(REMIT) = \sum_{j=0}^{m} \theta_{j} REMIT^{j} + \sum_{j=1}^{k} \varphi_{j} (REMIT - c_{j})_{+}^{m} \text{ where } (REMIT - c_{j})_{+}^{m} = \begin{cases} REMIT - c_{j} & \text{if } REMIT > c_{j} \\ 0 & \text{otherwise} \end{cases}$$
 and c_{j} a sequence of k knots where they join smoothly.

The better method is to use splines as a linear combination of a set of basic splines called B-splines which are defined as:

$$B(REMIT, c_1, c_2, ..., c_{k+1})$$

$$= (k+1) \sum_{j=1}^{k+2} \left[\prod_{1 \le h \le k+2, h \ne j} (c_h - c_j) \right]^{-1} (REMIT - c_j)_+^k$$

The model in equation (7) can be estimated using OLS. Once estimated, $\hat{\gamma}$ and $\hat{\beta}$ are obtained, it is straightforward to fit the fixed effects α_i and use equation (4) the error component as follows:

$$\hat{\varepsilon}_{i,t} = GDPG_{i,t} - X_{i,t}\hat{\beta} - \hat{u}_i = f(REMITT_{i,t}) + \varepsilon_{i,t}$$

(8)

The curve of the function $f(REMIT_{i,t})$ can be estimated by spline regressing $\hat{\varepsilon}_{i,t}$ against $REMIT_{i,t}$ using standard non-parametric regression estimator (Baltagi and Li, 2000; Libois and Verardi, 2012).

4. Results and discussion

4.1.Parametric estimates

We start by estimating the parametric models described in Equation (1), (2) and (3) expressed in linear, quadratic and cubic form, respectively. The results of such estimation are reported in Table 3. The regressions satisfy the specification tests (AR1, AR2 and Hansen test). There is strong evidence of a first serial correlation, but there is no evidence of a second serial correlation. Furthermore, the regressions pass the Hansen test and confirm the validity of the chosen instruments. Regarding the linear

specification (column 1 of Table 3), we note that the effect of remittances on economic growth is negative and statistically significant.

This result is in line with the findings of Feeny et al., 2014, Rao and Hassan, 2011, Barajas et al., 2009; Chami et al., 2003. About the quadratic model (column 2 of Table 3), we observe that the remittances variable (REMIT) and its quadratic term (REMIT²) are both significant at 1% level and with opposite sign. This result reveals the existence of an inverted U-shaped relationship between remittances inflow and economic growth. Lastly, the estimation of the cubic polynomial model (column 3 of Table 3) reveals an S-shaped curve between remittances and per capita GDP growth. Overall, these findings are in line with the idea that the effect of remittances on growth rate is not monotonic.

For the control variables and consistent with previous studies, we find strong evidence of a positive and significant effect of trade and gross fixed capital formation variables on economic growth, while we find negative and significant effect of government expenditure, inflation, and population growth.

Table 3: The effect of remittances inflow on economic growth: parametric estimation

	Linear model	Quadratic polynomial model	Cubic polynomial model
	(1)	(2).	(3).
L.GDPG	0.2750***	0.3013***	0.2546***
	(20.97)	(24.48)	(9.89)
REMIT	-0.0006***	0.0387***	-1.2870***
	(-2.86)	(6.50)	(-12.82)
$REMIT^2$		-0.0011***	0.0691***
		(-6.45)	(13.37)
$REMIT^3$			-0.0012***
			(-13.91)
TRADE	0.0228***	0.0213***	0.0297***
	(7.30)	(5.59)	(5.99)
INVEST	0.0511***	0.0414***	0.0814***
	(5.88)	(3.73)	(8.02)
EXPEND	-0.1095*	-0.0975	-0.1191***
	(-1.92)	(-1.66)	(-5.04)
INFL	-0.0291***	-0.0287***	-0.0438***
	(-5.04)	(-3.43)	(-6.15)
POPG	-0.1998***	-0.2632***	-0.2940***
	(-4.07)	(-5.75)	(-4.73)
No. of Obs.	476	476	476
F Statistic	515.064***	372.306***	231.198***
AR1	0.000	0.000	0.000
AR2	0.527	0.655	0.532
P-value Hansen test	0.816	0.794	0.934

Note: (*) and (***) denotes the rejection of null hypothesis at 1% and 10% significance levels, respectively. The values in parentheses represent t statistics

Source: Author compilation of analysis performed using IMF (2022) and World Bank (2021).

4.2. Semi-parametric estimates

In the next stage, we proceed to estimate our baseline model using the semi-parametric approach by allowing the remittances variable (REMIT) to enter non-parametrically into the model as described by the Equation (4). The results are reported in Table 4. We use two estimators: (i) the Robinson (1988) double residual estimator (column 1 of Table 4) and (ii) Baltagi and Li (2002) fixed-effects estimator (column 2 of Table 4). The coefficients estimation of control variables are close to those obtained under the parametric model and they are statistically significant at conventional levels (1%, 5% and 10%).

Table 4: The effect of remittances on economic growth: Semi-parametric estimation

Independent variables	Robinson's (1988) Semi- parametric estimator	Baltagi and Li's (2002) Semi- parametric panel fixed effects estimator	
	(1)	(2)	
REMIT	See Fig. 1	See Fig. 2	
TRADE	-0.0038	0.0546**	
	(-0.65)	(2.10)	
INVEST	0.0587**	0.1660**	
	(2.47)	(2.33)	
EXPEND	-0.1183***	-0.4394***	
	(-3.61)	(-3.31)	
INFL	-0.0401***	-0.0874***	
	(-2.98)	(-4.09)	
POPG	0.0028**	-0.0062*	
	(2.19)	(-1.70)	
N	582	325	
$R^2(\%)$	4.90	15.73	

Note: (*), (**) and (***) denotes the rejection of null hypothesis at 10%, 5% and 1% significance levels, respectively. The values in parentheses represent t statistics.

Figure 1 and Figure 2 display the impact of remittances (horizontal axis) on economic growth (vertical axis). More precisely, they illustrate the semi-parametric estimation of f(.) along with 95% confidence bands (grey shaded area) obtained by using the Robinson's (1988) double residual estimator and Baltagi and Li's (2002) fixed-effects estimator, respectively. The circles in the graphs are partial residual for economic growth (GDPG) in the semi-parametric specification. They are obtained from equation 8. The maroon curve represents the fitted value for adjusted effects of the other covariates in the semi-parametric model. The result for both estimators, in accordance with previous literature, show that remittances has a nonlinear effect on economic growth.

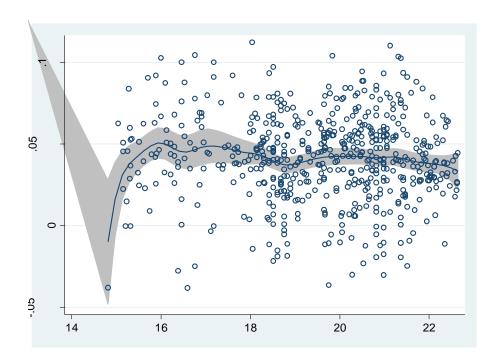


Figure 1. Partial fit of remittances and economic growth nexus. Note: the circles in the graph are estimated partial residuals for economic growth in the semi-parametric model using the Robinsson's (1988) estimator. The blue curve represents the semi-parametric estimation of f(REMIT).95% confidence bands are indicated by shaded areas.

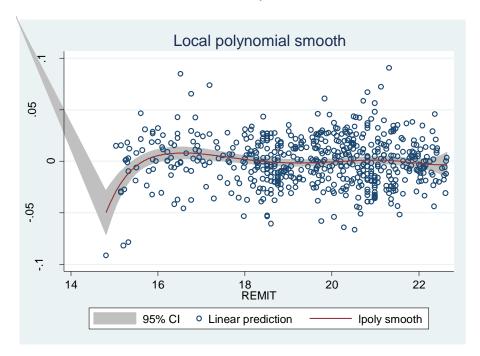


Figure 2. Partial fit of remittances and economic growth nexus. Note: the circles in the graph are estimated partial residuals for economic growth in the semi-parametric model using the Baltagi and Li (2002) estimator. The red curve represents the semi-parametric estimation of f(REMIT). 95%confidence bands are indicated by shaded areas.

5. Conclusion

This paper re-examines the remittances-growth nexus by employing the semiparametric panel model with fixed effects as developed by Baltagi and Li (2002). We investigate a sample of 68 countries over the period 1988-2018. We find that there is a nonlinear relationship between remittances inflow and economic growth rate, in contrast to most of the literature.

One of the strengths of this paper is that it assumes no prior functional form on the regression model, which helps to reduce the misspecification error that may arise from imposing specific functional forms. This approach increases the robustness of the findings and contributes to a better understanding of the relationship between remittances inflow and economic growth rate. We shed light on the nonlinear nature of the remittances-growth nexus. The results may have important implications for financial and economic regulation, as policymakers can better understand the relationship between remittances inflow and economic growth rate, which can inform policy decisions related to the regulation of remittance flows and their impact on economic growth.

For policymakers, the findings of this paper have several implications. First, the nonlinear relationship between remittances inflow and economic growth rate suggests that the impact of remittances on economic growth is not straightforward and depends on the level of remittances inflow. Policymakers should consider the nonlinear nature of this relationship when designing policies related to remittances inflow.

Second, the results of the study highlight the importance of regulating remittance flows to maximize their positive impact on economic growth. Policymakers should consider implementing policies that encourage remittance inflows to flow into productive investments or sectors, which can contribute to economic growth. Additionally, policies aimed at reducing the costs associated with remittance transfers could help to increase remittance inflows, which may in turn boost economic growth.

Overall, the findings of this study suggest that policymakers should take a nuanced approach when designing policies related to remittance inflows and their impact on economic growth. Understanding the nonlinear nature of the remittancesgrowth nexus can help policymakers make more informed decisions that can contribute to economic growth and development.

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Imprint
Editors: Sigrid Betzelt, Eckhard Hein, Martina Metzger, Martina Sproll, Christina Teipen, Markus Wissen, Jennifer Pédussel Wu (lead editor), Reingard Zimmer
ISSN 1869-6406
Printed by HWR Berlin
Berlin, August 2024